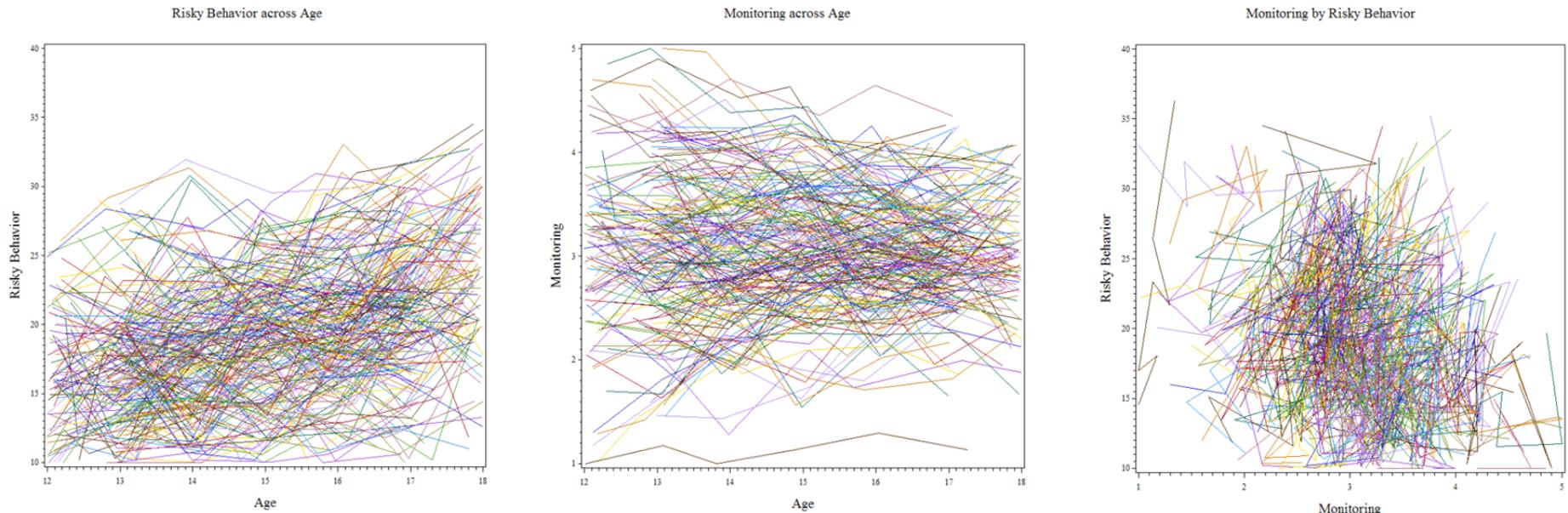


Revised Example 4: Three Ways of Estimating Multivariate Change: in SAS PROC MIXED, as well as in Multivariate MLM (“Multilevel SEM”) and Single-Level SEM in Mplus v. 8.1 (complete syntax and output available electronically)

These simulated data are from Hoffman (2015) chapter 9, and include 200 girls measured approximately annually from ages 12–18 (time 0 = age 18) on their risky behavior (the outcome, a sum ranging from 10 to 50) and the extent to which their mothers monitored their activities (the time-varying predictor, a mean ranging from 1 to 5, centered at 3). A time-invariant predictor of the conservativeness of mothers' attitudes about the smoking and drinking (a mean ranging from 1 to 5, centered at 4) was also collected at the age 12 occasion. Here are the individual growth trajectories for risky behavior and monitoring:



Level 1:

Multivariate Multilevel Model

$$\begin{aligned} y_{tiR} &= dvR \left[\beta_{0iR} + \beta_{1iR} (\text{Age}_{tiR} - 18) + \beta_{2iR} (\text{Age}_{tiR} - 18)^2 + e_{tiR} \right] + \\ dvM &\left[\beta_{0iM} + \beta_{1iM} (\text{Age}_{tiM} - 18) + e_{tiM} \right] \end{aligned}$$

Level 2:

$$\text{Risky Intercept: } \beta_{0iR} = \gamma_{00R} + \gamma_{01R} (\text{Attitudes12}_i - 4) + U_{0iR}$$

$$\text{Risky Age: } \beta_{1iR} = \gamma_{10R} + \gamma_{11R} (\text{Attitudes12}_i - 4) + U_{1iR}$$

$$\text{Risky Age}^2: \quad \beta_{2iR} = \gamma_{20R}$$

$$\text{Monitor Intercept: } \beta_{0iM} = \gamma_{00M} + \gamma_{01M} (\text{Attitudes12}_i - 4) + U_{0iM}$$

$$\text{Monitor Age: } \beta_{1iM} = \gamma_{10M} + \gamma_{11M} (\text{Attitudes12}_i - 4) + U_{1iM}$$

The best-fitting unconditional longitudinal models included fixed quadratic and random linear effects of age for risky behavior, but a random linear effect of age for monitoring (although the fixed linear age slope was nonsignificant). In addition, mother's attitudes significantly predicted the intercept and linear age slope for risky behavior. Although they did not significantly predict monitoring, I have added them here to illustrate computing indirect effects.

Chapter 9 began with person-mean-centering and baseline-centering of monitoring of a time-varying predictor of risky behavior. Both were shown to be inadequate because they do not properly distinguish the intercept, linear age slope, and residual variance contained in the monitoring predictor, each of which could potentially relate to those of risky behavior. So the purpose of this example is to demonstrate alternative software methods of estimating models of multivariate change so that you can decide what approach (software and syntax combination) will be most optimal for your own data. See chapter 9 for the results from a directed path model very similar to 2c.

Undirected Multivariate Growth Model for Risky Behavior and Monitoring in SAS PROC MIXED, controlling risky behavior for time-invariant attitudes (Model 1):

```
* Stack longitudinal data into multivariate longitudinal;
DATA RiskyStacked2; SET RiskyStacked;
DV="1risky"; dvR=1; dvM=0; outcome=risky; OUTPUT;
DV="2monitor"; dvR=0; dvM=1; outcome=mon3; OUTPUT;
RUN;

TITLE1 "Multivariate Model at Age 18 = Time 0";
PROC MIXED DATA=work.Chapter9 NOCLPRINT COVTEST IC
    NAMELEN=100 METHOD=ML;
CLASS FamilyID occasion DV;

MODEL outcome = dvR dvM dvR*agec18 dvM*agec18 dvR*agec18*agec18
    dvR*att4 dvR*agec18*att4 dvM*att4 dvM*agec18*att4
    / NOINT SOLUTION DDFM=Satterthwaite;
RANDOM dvR dvM dvR*agec18 dvM*agec18
    / G GCORR TYPE=UN SUBJECT=FamilyID;
REPEATED DV / R RCORR TYPE=UN SUBJECT=occasion*FamilyID;
RUN; TITLE1;
```

Results start here: This is the same model as in SAS...

SAS:

Fit Statistics	
-2 Log Likelihood	8783.8
AIC (Smaller is Better)	8827.8
AICC (Smaller is Better)	8828.2
BIC (Smaller is Better)	8900.4

MPLUS:

Number of Free Parameters	22
Loglikelihood	
H0 Value	-4391.885
Information Criteria	
Akaike (AIC)	8827.771
Bayesian (BIC)	8943.144
Sample-Size Adjusted BIC	8873.258
(n* = (n + 2) / 24)	

In Mplus, the same Model 1 as an undirected multivariate MLM:

```
TITLE: Model 1: Undirected Multivariate Growth Model as MLM
DATA: FILE = Example4.csv; ! Syntax in same folder as data
VARIABLE:
! List of variables in data file
NAMES = PersonID occasion risky age18 att4 agesq mon3;
! Variables to be analyzed in this model
USEVARIABLE = age18 agesq att4 risky mon3;
MISSING ARE ALL (-999); ! Missing data identifier
! MLM options
CLUSTER = PersonID; ! Level-2 ID
BETWEEN = att4; ! Observed ONLY level-2 predictors
WITHIN = age18 agesq; ! Observed ONLY level-1 predictors

ANALYSIS: TYPE = TWOLEVEL RANDOM; ESTIMATOR = ML;

MODEL: ! R = risky behavior, M = monitoring
%WITHIN%
risky mon3 (Rresvar Mresvar); ! L1 R: Residual variances (labels)
Rslp | risky ON age18; ! Placeholder for R linear age slope
Rquad | risky ON agesq; ! Placeholder for R quadratic age slope
Msdp | mon3 ON age18; ! Placeholder for M linear age slope
risky WITH mon3 (ResCov); ! L1 R: Residual covariance

%BETWEEN%
[risky mon3]; ! Fixed intercepts
risky mon3 (Rintvar Mintvar); ! L2 G: Random intercept variances (labels)
[Rslp Rquad Msdp];
Rslp Msdp (Rslpvar Msdpvar); ! Fixed age slopes (as defined earlier)
Rquad@0; ! No quadratic age slope variance

risky Rslp ON att4; ! Att-> R int, linear age slope
mon3 Msdp ON att4; ! Att-> M int, linear age slope
risky WITH Rslp (RIntSlp); ! R Int-slope covariance (label)
mon3 WITH Msdp (MIntSlp); ! M Int-slope covariance (label)

risky WITH mon3 (IntCov); ! L2 G: Random intercept covariance
Rslp WITH Msdp (SlpCov); ! L2 G: Random linear age slope covariance
mon3 WITH Rslp (Int2Slp); ! L2 G: M int, R slope covariance
Msdp WITH risky (Slp2Int); ! L2 G: M slope, R int covariance

MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter
! Need to name each new created effect
NEW(ResCor IntCor SlpCor RISCor MISCor I2SCor S2ICor);

! Estimating correlations found in SAS RCORR and GCORR
! Corr = Cov / (SQRT(Xvar)*SQRT(Xvar))
ResCor = ResCov / (SQRT(Rresvar)*SQRT(Mresvar));
IntCor = IntCov / (SQRT(Rintvar)*SQRT(Mintvar));
SlpCor = SlpCov / (SQRT(Rslpvar)*SQRT(Mslpvar));
RISCor = RIntSlp / (SQRT(Rintvar)*SQRT(Rslpvar));
MISCor = MIntSlp / (SQRT(Mintvar)*SQRT(Mslpvar));
I2SCor = Int2Slp / (SQRT(Mintvar)*SQRT(Rslpvar));
S2ICor = Slp2Int / (SQRT(Mslpvar)*SQRT(Rintvar));
```

SAS Undirected Multivariate MLM Results:

Estimated R Matrix for PersonID*occasion 1 12			Estimated R Correlation Matrix for PersonID*occasion 1 12		
Row	Col1	Col2	Row	Col1	Col2
1	8.3538	0.2874	1	1.0000	0.3499
2	0.2874	0.08077	2	0.3499	1.0000
Estimated G Matrix					
Row	Effect	PersonID	Col1	Col2	Col3
1	dvR	1	18.0535	-0.8527	1.8821
2	dvM	1	-0.8527	0.1946	-0.1062
3	dvR*agec18	1	1.8821	-0.1062	0.4882
4	dvM*agec18	1	0.04051	-0.00042	-0.01817
Estimated G Correlation Matrix					
Row	Effect	PersonID	Col1	Col2	Col3
1	dvR	1	1.0000	-0.4549	0.6339
2	dvM	1	-0.4549	1.0000	-0.3446
3	dvR*agec18	1	0.6339	-0.3446	1.0000
4	dvM*agec18	1	0.0931	-0.0093	-0.2539
Covariance Parameter Estimates (covariances only to save space)					
Cov Parm	Subject	Estimate	Standard	Z	
UN(2,1)	PersonID	-0.8527	0.1680	-5.08	<.0001
UN(3,1)	PersonID	1.8821	0.3562	5.28	<.0001
UN(3,2)	PersonID	-0.1062	0.03077	-3.45	0.0006
UN(4,1)	PersonID	0.04051	0.03877	1.04	0.2961
UN(4,2)	PersonID	-0.00042	0.004000	-0.11	0.9164
UN(4,3)	PersonID	-0.01817	0.007341	-2.47	0.0133
UN(2,1)	PersonID*occasion	0.2874	0.02753	10.44	<.0001
Solution for Fixed Effects					
Effect	Estimate	Error	DF	t Value	Pr > t
dvR	23.3224	0.3477	239	67.07	<.0001
dvm	0.06287	0.03418	200	1.84	0.0674
dvR*agec18	1.9749	0.1386	1185	14.25	<.0001
dvm*agec18	-0.00312	0.008202	200	-0.38	0.7040
dvR*agec18*agec18	0.1466	0.02058	1010	7.12	<.0001
dvR*att4	-3.1601	0.5509	200	-5.74	<.0001
dvR*agec18*att4	-0.5173	0.1043	199	-4.96	<.0001
dvm*att4	-0.04418	0.05668	200	-0.78	0.4366
dvm*agec18*att4	0.003269	0.01360	200	0.24	0.8103

Mplus results continue: This is the same model as in SAS...

	Within Level	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
RISKY	WITH				
MON3		0.287	0.028	10.441	0.000
Residual Variances					
RISKY		8.352	0.374	22.351	0.000
MON3		0.081	0.004	22.354	0.000
Between Level					
RSLP	ON				
ATT4		-0.518	0.104	-4.963	0.000
MSLP	ON				
ATT4		0.003	0.014	0.240	0.810
RISKY	ON				
ATT4		-3.160	0.551	-5.737	0.000
MON3	ON				
ATT4		-0.044	0.057	-0.779	0.436
RISKY	WITH				
RSLP		1.878	0.356	5.273	0.000
MSLP		0.040	0.039	1.044	0.296
MON3	WITH				
MSLP		0.000	0.004	-0.105	0.916
RSLP		-0.106	0.031	-3.449	0.001
RSLP	WITH				
MSLP		-0.018	0.007	-2.478	0.013
RISKY	WITH				
MON3		-0.853	0.168	-5.076	0.000
Means					
RQUAD		0.147	0.021	7.117	0.000
Intercepts					
RISKY		23.322	0.348	67.075	0.000
MON3		0.063	0.034	1.839	0.066
RSLP		1.975	0.138	14.259	0.000
MSLP		-0.003	0.008	-0.380	0.704
Variances					
RQUAD		0.000	0.000	999.000	999.000
Residual Variances					
RISKY		18.049	2.202	8.198	0.000
MON3		0.195	0.023	8.371	0.000
RSLP		0.484	0.080	6.071	0.000
MSLP		0.010	0.001	7.802	0.000
New/Additional Parameters - look for the first 3 across outputs					
RESCOR		0.350	0.028	12.607	0.000
INTCOR		-0.455	0.074	-6.119	0.000
SLPCOR		-0.255	0.103	-2.483	0.013
RISCOR		0.635	0.057	11.088	0.000
MISCOR		-0.009	0.089	-0.105	0.917
I2SCOR		-0.346	0.095	-3.646	0.000
S2ICOR		0.093	0.087	1.066	0.286

In Mplus, the same Model 1 as an undirected single-level SEM:

```

TITLE: Model 1: Undirected Multivariate Growth Model as Single-Level SEM
DATA: FILE = Example4.csv; ! Syntax in same folder as data
! Unstacking to multivariate format
DATA LONGTOWIDE:
! Names of old stacked former variables (without numbers)
LONG = risky|mon|age;
! Names of new multivariate variables (that use numbers)
WIDE = risky12-risky18|mon12-mon18|age12-age18;
! Variable with level-2 ID info
IDVARIABLE = PersonID;
! Old level-1 identifier
REPETITION = age (12 13 14 15 16 17 18);
VARIABLE:
! List of variables in original data file
NAMES = PersonID occasion risky age18 att4 mon3 agesq;
! Variables to be analyzed in this model
USEVARIABLE = att4 age12-age18 mon12-mon18 risky12-risky18;
MISSING ARE ALL (-999); ! Missing data identifier
TSCORES = age12-age18; ! Exact time indicator

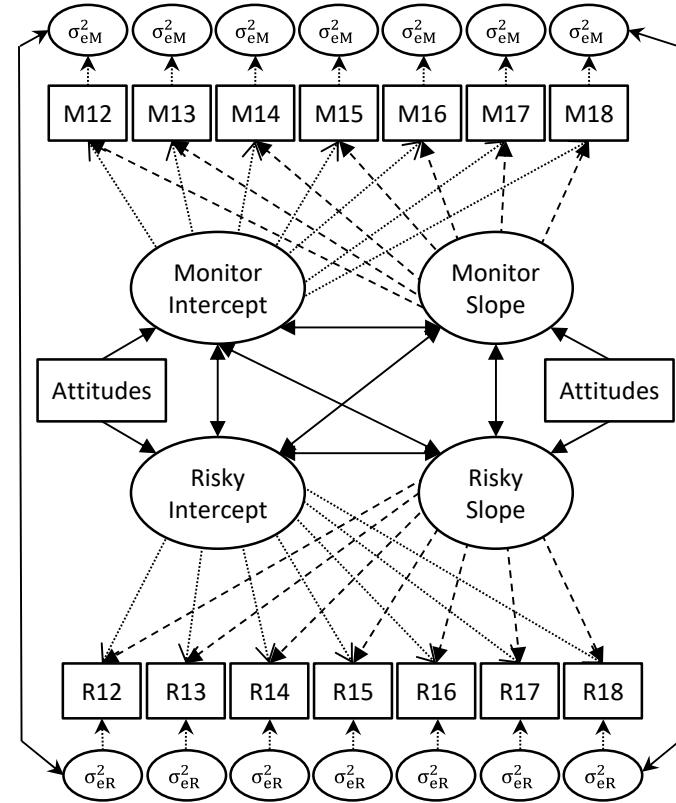
ANALYSIS: TYPE = RANDOM; ESTIMATOR = ML; MODEL = NOCOVARIANCES;
MODEL: ! R = risky behavior, M = monitoring
[risky12-risky18@0 mon12-mon18@0]; ! All variable intercepts fixed to 0
risky12-risky18 (Rresvar); ! L1 R: R residual variances held equal
mon12-mon18 (Mresvar); ! L1 R: M residual variances held equal

! Risky behavior quadratic growth model using exact age as loadings
Rint Rslp Rquad | risky12-risky18 AT age12-age18;
! Monitoring linear growth model using exact age as loadings
Mint Mslp | mon12-mon18 AT age12-age18;
! Fixed growth effects for R and M
[Rint Rslp Rquad Mint Mslp];
! L2 G: Random int and linear age slope variances, no quad age variance
Rint Rslp Mint Mslp (Rintvar Rslpvar Mintvar Mslpvar); Rquad@0;
! L2 G: Within-variable random int-slope covariances for R, M
Rint WITH Rslp (Rintslp); Mint WITH Mslp (Mintsdp);
! Attitudes --> R int, R linear slope, M int, M linear slope
Rint Rslp Mint Mslp ON att4;

! Covariances between outcomes
Rint WITH Mint (IntCov); ! L2 G: Random intercept covariance
Rslp WITH Mslp (SlpCov); ! L2 G: Random linear age slope covariance
Mint WITH Rslp (Int2Slp); ! L2 G: M int, R slope covariance
Mslp WITH Rint (Slp2Int); ! L2 G: M slope, R int covariance
! Residual WP covariance between same ages, held equal across age
risky12-risky18 PWITH mon12-mon18 (ResCov);

MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter
! Need to name each new created effect
NEW(ResCor IntCor SlpCor);
! Estimating correlations found in SAS RCORR and GCORR
! Corr = Cov / (SQRT(Yvar)*SQRT(Xvar))
ResCor = ResCov / (SQRT(Rresvar)*SQRT(Mresvar));
IntCor = IntCov / (SQRT(Rintvar)*SQRT(Mintvar));
SlpCor = SlpCov / (SQRT(Rslpvar)*SQRT(Mslpvar));

```



-→ Indicates paths fixed = 1
- - - → Indicates paths fixed = time values
- ↔ Indicates paths freely estimated
- ↔ Indicates paths freely estimated between residuals at the same occasion but held equal over time

For balanced time, a linear growth model would look like this instead (add quad as third place before |):

```
Mint Mslp | mon12@-6 mon13@-5 mon14@-4 mon15@-3
mon16@-2 mon17@-1 mon18@0;
```

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MODEL RESULTS						Means				
		Estimate	S.E.	Est./S.E.	P-Value	RQUAD	0.147	0.021	7.117	0.000
RINT	ON	-3.160	0.551	-5.737	0.000	Intercepts				
ATT4						RISKY12	0.000	0.000	999.000	999.000
RSLP	ON	-0.518	0.104	-4.963	0.000	RISKY13	0.000	0.000	999.000	999.000
ATT4						RISKY14	0.000	0.000	999.000	999.000
MINT	ON	-0.044	0.057	-0.779	0.436	RISKY15	0.000	0.000	999.000	999.000
ATT4						RISKY16	0.000	0.000	999.000	999.000
MSLP	ON	0.003	0.014	0.240	0.810	RISKY17	0.000	0.000	999.000	999.000
ATT4						RISKY18	0.000	0.000	999.000	999.000
RINT	WITH					MON12	0.000	0.000	999.000	999.000
RSLP		1.878	0.356	5.273	0.000	MON13	0.000	0.000	999.000	999.000
MINT		-0.853	0.168	-5.076	0.000	MON14	0.000	0.000	999.000	999.000
MSLP		0.040	0.039	1.044	0.296	MON15	0.000	0.000	999.000	999.000
						MON16	0.000	0.000	999.000	999.000
MINT	WITH					MON17	0.000	0.000	999.000	999.000
MSLP		0.000	0.004	-0.105	0.916	MON18	0.000	0.000	999.000	999.000
RSLP		-0.106	0.031	-3.449	0.001	RINT	23.322	0.348	67.074	0.000
						RSLP	1.975	0.138	14.259	0.000
						MINT	0.063	0.034	1.839	0.066
						MSLP	-0.003	0.008	-0.380	0.704
RSLP	WITH					Variances				
MSLP		-0.018	0.007	-2.478	0.013	RQUAD	0.000	0.000	999.000	999.000
RISKY12	WITH					Residual Variances				
MON12		0.287	0.028	10.441	0.000	RISKY12	8.352	0.374	22.351	0.000
RISKY13	WITH					RISKY13	8.352	0.374	22.351	0.000
MON13		0.287	0.028	10.441	0.000	RISKY14	8.352	0.374	22.351	0.000
RISKY14	WITH					RISKY15	8.352	0.374	22.351	0.000
MON14		0.287	0.028	10.441	0.000	RISKY16	8.352	0.374	22.351	0.000
RISKY15	WITH					RISKY17	8.352	0.374	22.351	0.000
MON15		0.287	0.028	10.441	0.000	RISKY18	8.352	0.374	22.351	0.000
RISKY16	WITH					MON12	0.081	0.004	22.354	0.000
MON16		0.287	0.028	10.441	0.000	MON13	0.081	0.004	22.354	0.000
RISKY17	WITH					MON14	0.081	0.004	22.354	0.000
MON17		0.287	0.028	10.441	0.000	MON15	0.081	0.004	22.354	0.000
RISKY18	WITH					MON16	0.081	0.004	22.354	0.000
MON18		0.287	0.028	10.441	0.000	MON17	0.081	0.004	22.354	0.000
						MON18	0.081	0.004	22.354	0.000
						RINT	18.049	2.202	8.198	0.000
						RSLP	0.484	0.080	6.071	0.000
						MINT	0.195	0.023	8.371	0.000
						MSLP	0.010	0.001	7.802	0.000
						New/Additional Parameters				
						RESCOR	0.350	0.028	12.607	0.000
						INTCOR	-0.455	0.074	-6.119	0.000
						SLPCOR	-0.255	0.103	-2.483	0.013

Model 2a: Partially Directed Path Multivariate MLM in Mplus: Monitor → Risky for intercepts and slopes, but residuals covary

```

TITLE: Model 2a: Partially Directed Multivariate Growth Model as MLM
L1 WP effect as residual covariance

( DATA, VARIABLE, and ANALYSIS are the same as for Model 1 )

MODEL: ! R = risky behavior, M = monitoring
%WITHIN%
risky mon3 (Rresvar Mresvar); ! L1 R: Residual variances (labels)
Rslp | risky ON age18; ! Placeholder for R linear age slope
Rquad | risky ON agesq; ! Placeholder for R quadratic age slope
Mslp | mon3 ON age18; ! Placeholder for M linear age slope
risky WITH mon3 (ResCov); ! L1 R: Still residual covariance

%BETWEEN%
[risky mon3]; ! Fixed intercepts
risky mon3 (Rintvar Mintvar); ! L2 G: Random intercept variances (labels)
[Rquad Rslp Mslp];
Rslp Mslp (Rslpvar Mslpvar); ! L2 G: Random linear age slope variances
Rquad@0;
risky Rslp ON att4; ! Att-> R int, linear age slope
mon3 Mslp ON att4; ! Att-> M int, linear age slope
risky WITH Rslp (RIntSlp); ! R Int-slope covariance (label)
mon3 WITH Mslp (MIntSlp); ! M Int-slope covariance (label)

! Although we have changed the int-int and slope-slope relations to direct
! paths from M -> R instead of covariances, they still represent total BP
! relationships because the L1 relationship is still a covariance

risky ON mon3 (BPIntEff); ! Total BP intercept effect
Rslp ON Mslp (BPSlpEff); ! Total BP age slope effect

mon3 WITH Rslp (Int2Slp); ! L2 G: M int, R slope covariance
Mslp WITH risky (Slp2Int); ! L2 G: M slope, R int covariance

MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter
! Need to name each new created effect -- all values from undirected model
NEW(ResCor IntStd SlpStd);

! Corr = Cov / (SQRT(Yvar)*SQRT(Xvar))
ResCor = ResCov / (SQRT(8.3538)*SQRT(0.08077)); ! WP Res corr

! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP Int effect
SlpStd = BPSlpEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP Slope effect

```

This is the same model, just with a different way of specifying the level-2 intercept to intercept and slope to slope relationships.

		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Within Level					
RISKY	WITH				
MON3		0.287	0.028	10.441	0.000
Residual Variances					
RISKY		8.352	0.374	22.351	0.000
MON3		0.081	0.004	22.354	0.000
Between Level - new parameters are in <u>BOLD</u> underline					
<u>RSLP</u>	<u>ON</u>				
<u>MSLP</u>		-1.736	0.713	-2.434	0.015
RSLP	ON				
ATT4		-0.512	0.105	-4.869	0.000
MSLP	ON				
ATT4		0.003	0.014	0.240	0.810
RISKY	ON				
ATT4		-3.354	0.528	-6.348	0.000
MON3		-4.380	0.797	-5.497	0.000
MON3	ON				
ATT4		-0.044	0.057	-0.779	0.436
RISKY	WITH				
RSLP		1.480	0.345	4.285	0.000
MSLP		0.039	0.038	1.023	0.306
MON3	WITH				
MSLP		0.000	0.004	-0.105	0.916
RSLP		-0.107	0.031	-3.454	0.001
Means					
RQUAD		0.147	0.021	7.117	0.000
Intercepts					
RISKY		23.598	0.338	69.837	0.000
MON3		0.063	0.034	1.839	0.066
RSLP		1.969	0.139	14.195	0.000
MSLP		-0.003	0.008	-0.380	0.704
Variances					
RQUAD		0.000	0.000	999.000	999.000
Residual Variances					
RISKY		14.315	2.030	7.053	0.000
MON3		0.195	0.023	8.371	0.000
RSLP		0.453	0.081	5.564	0.000
MSLP		0.010	0.001	7.802	0.000
New/Additional Parameters					
RESCOR		0.350	0.034	10.441	0.000
INTSTD		-0.455	0.083	-5.497	0.000
SLPSTD		-0.254	0.104	-2.434	0.015

**In Mplus, Model 2a as a partially directed single-level SEM:
Monitor → Risky for intercepts and slopes, but residuals covary**

```

TITLE: Model 2a: Partially Directed Multivariate Growth Model as
       Single-Level SEM, L1 WP effect as residual covariance

( DATA, VARIABLE, and ANALYSIS are the same as for Model 1 )

MODEL: ! R = risky behavior, M = monitoring
[risky12-risky18@0 mon12-mon18@0]; ! All variable intercepts fixed to 0
risky12-risky18 (Rresvar);           ! L1 R: R residual variances held equal
mon12-mon18 (Mresvar);             ! L1 R: M residual variances held equal

! Risky behavior quadratic growth model using exact age as loadings
Rint Rsdp Rquad | risky12-risky18 AT age12-age18;
! Monitoring linear growth model using exact age as loadings
Mint Msdp | mon12-mon18 AT age12-age18;
! Fixed growth effects for R and M
[Rint Rsdp Rquad Mint Msdp];
! L2 G: Random int and linear age slope variances, no quad age variance
Rint Rsdp Mint Msdp (Rintvar Rsdpvar Mintvar Msdpvar); Rquad@0;
! L2 G: Within-variable random int-slope covariances for R, M
Rint WITH Rsdp (Rintslp); Mint WITH Msdp (Mintslp);
! Attitudes --> R int, R linear slope, M int, M linear slope
Rint Rsdp Mint Msdp ON att4;

! Although we have changed the int-int and slope-slope relations to direct
! paths from M -> R instead of covariances, they still represent total BP
! relationships because the L1 relationship is still a covariance

Rint ON Mint (BPIntEff);          ! Total BP intercept effect
Rsdp ON Msdp (BPSdpEff);          ! Total BP age slope effect

Mint WITH Rsdp (Int2Sdp);          ! L2 G: M int, R slope covariance
Msdp WITH Rint (Sdp2Int);          ! L2 G: M slope, R int covariance

! Residual WP covariance between same ages, held equal across age
risky12-risky18 PWITH mon12-mon18 (ResCov);

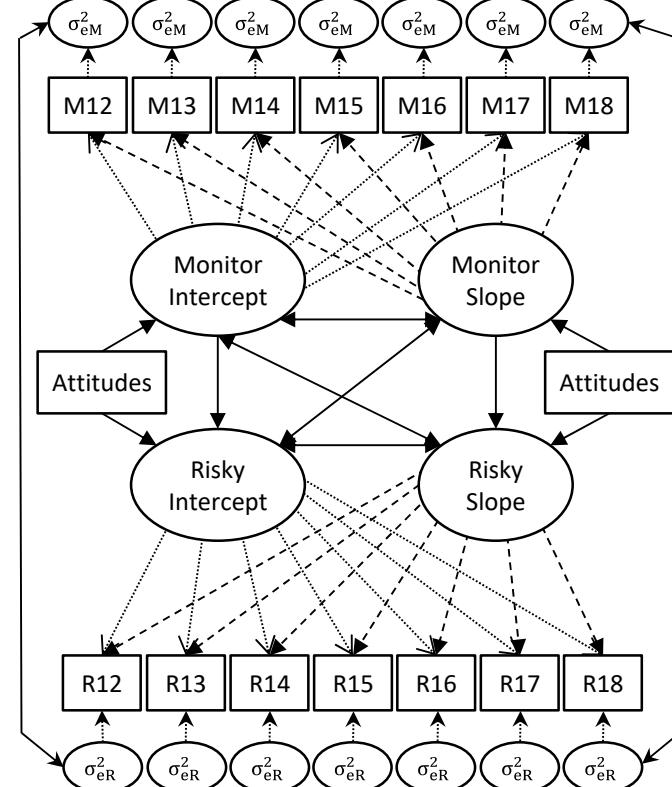
MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter
! Need to name each new created effect -- all values from undirected model
NEW(ResCor IntStd SdpStd);

! Corr = Cov / (SQRT(Yvar)*SQRT(Xvar))
ResCor = ResCov / (SQRT(8.3538)*SQRT(0.08077));      ! WP Res corr

! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP Int effect
SdpStd = BPSdpEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP Slope effect

```

This is the same model, just with a different way of specifying the level-2 intercept to intercept and slope to slope relationships.



-> Indicates paths fixed = 1
- - -> Indicates paths fixed = time values
- ←→ Indicates paths freely estimated
- ←→ Indicates paths freely estimated between residuals at the same occasion but held equal over time

For balanced time, a linear growth model would look like this instead (add quad as third place before |):

```
Mint Msdp | mon12@-6 mon13@-5 mon14@-4 mon15@-3
           mon16@-2 mon17@-1 mon18@0;
```

MODEL RESULTS - <u>changed parameters are in BOLD underline</u>					Means					
					RQUAD	0.147	0.021	7.117	0.000	
					Intercepts					
<u>RINT</u>	<u>ON</u>				RISKY12	0.000	0.000	999.000	999.000	
MINT		<u>-4.380</u>	<u>0.797</u>	<u>-5.496</u>	RISKY13	0.000	0.000	999.000	999.000	
<u>RSLP</u>	<u>ON</u>				RISKY14	0.000	0.000	999.000	999.000	
MSLP		<u>-1.736</u>	<u>0.713</u>	<u>-2.434</u>	RISKY15	0.000	0.000	999.000	999.000	
RINT	ON				RISKY16	0.000	0.000	999.000	999.000	
ATT4		-3.354	0.528	-6.348	RISKY17	0.000	0.000	999.000	999.000	
RSLP	ON			RISKY18	0.000	0.000	999.000	999.000		
ATT4		-0.512	0.105	-4.869	MON12	0.000	0.000	999.000	999.000	
MINT	ON			MON13	0.000	0.000	999.000	999.000		
ATT4		-0.044	0.057	-0.779	MON14	0.000	0.000	999.000	999.000	
MSLP	ON			MON15	0.000	0.000	999.000	999.000		
ATT4		0.003	0.014	0.240	MON16	0.000	0.000	999.000	999.000	
RINT	WITH			MON17	0.000	0.000	999.000	999.000		
RSLP		1.480	0.345	4.285	MON18	0.000	0.000	999.000	999.000	
MSLP		0.039	0.038	1.023	RINT	23.598	0.338	69.836	0.000	
RINT	WITH			RSLP	1.969	0.139	14.195	0.000		
MINT				MINT	0.063	0.034	1.839	0.066		
MSLP				MSLP	-0.003	0.008	-0.380	0.704		
MINT	WITH			Variances						
MSLP		0.000	0.004	-0.105	RQUAD	0.000	0.000	999.000	999.000	
RSLP		-0.107	0.031	-3.454	Residual Variances					
RISKY12	WITH			RISKY12	8.352	0.374	22.351	0.000		
MON12		0.287	0.028	10.441	RISKY13	8.352	0.374	22.351	0.000	
RISKY13	WITH			RISKY14	8.352	0.374	22.351	0.000		
MON13		0.287	0.028	10.441	RISKY15	8.352	0.374	22.351	0.000	
RISKY14	WITH			RISKY16	8.352	0.374	22.351	0.000		
MON14		0.287	0.028	10.441	RISKY17	8.352	0.374	22.351	0.000	
RISKY15	WITH			RISKY18	8.352	0.374	22.351	0.000		
MON15		0.287	0.028	10.441	MON12	0.081	0.004	22.354	0.000	
RISKY16	WITH			MON13	0.081	0.004	22.354	0.000		
MON16		0.287	0.028	10.441	MON14	0.081	0.004	22.354	0.000	
RISKY17	WITH			MON15	0.081	0.004	22.354	0.000		
MON17		0.287	0.028	10.441	MON16	0.081	0.004	22.354	0.000	
RISKY18	WITH			MON17	0.081	0.004	22.354	0.000		
MON18		0.287	0.028	10.441	MON18	0.081	0.004	22.354	0.000	
					RINT	14.315	2.030	7.053	0.000	
					RSLP	0.453	0.081	5.564	0.000	
					MINT	0.195	0.023	8.371	0.000	
					MSLP	0.010	0.001	7.802	0.000	
New/Additional Parameters										
					RESCOR	0.350	0.034	10.441	0.000	
					INTSTD	-0.455	0.083	-5.496	0.000	
					SLPSTD	-0.254	0.104	-2.434	0.015	

Model 2b: Partially Directed Path Multivariate MLM in Mplus: Monitor → Risky for WP residuals within L1 model
 Also demonstrating how to request BP indirect effects

```

TITLE: Model 2b: Partially Directed Multivariate Growth Model as MLM
L1 WP effect as direct path specified in L1 WITHIN

( DATA, VARIABLE, and ANALYSIS are the same as for Model 1 )

MODEL: ! R = risky behavior, M = monitoring
%WITHIN%
risky mon3 (Rresvar Mresvar); ! L1 R: Residual variances (labels)
Rslp | risky ON age18; ! Placeholder for R linear age slope
Rquad | risky ON agesq; ! Placeholder for R quadratic age slope
Mslp | mon3 ON age18; ! Placeholder for M linear age slope
risky ON mon3 (ResEff); ! L1 WP fixed effect M->R here (label)

%BETWEEN%
[risky mon3]; ! Fixed intercepts
risky mon3 (Rintvar Mintvar); ! L2 G: Random intercept variances (labels)
[Rquad Rslp Mslp];
! Fixed age slopes (as defined earlier)
Rslp Mslp (Rslpvar Mslpvar); ! L2 G: Random linear age slope variances
Rquad@0; ! No quadratic age slope variance
risky Rslp ON att4 (XtoYint XtoYslp); ! Att-> R int, linear age slope
mon3 Mslp ON att4 (XtoMint XtoMsdp); ! Att-> M int, linear age slope
risky WITH Rslp (RIntSlp); ! R Int-slope covariance (label)
mon3 WITH Mslp (MIntSlp); ! M Int-slope covariance (label)

! Although the intercept -> intercept path remains the total BP effect,
! now the slope -> slope path becomes the contextual BP effect instead

risky ON mon3 (BPIntEff); ! STILL total BP intercept effect
Rslp ON Mslp (SlpCont); ! NOW contextual BP age slope effect

mon3 WITH Rslp (Int2Slp); ! L2 G: M int, R slope covariance
Mslp WITH risky (Slp2Int); ! L2 G: M slope, R int covariance

MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter
! Need to name each new created effect -- all values from undirected model
NEW(ResStd IntStd BPSlpEff SlpStd indBPint indBPsdp);

! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
ResStd = ResEff * SQRT(0.08077) / SQRT(8.3538); ! STD WP Res effect
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP Int effect
BPSlpEff = ResEff + SlpCont; ! WP + Context = BP slp
SlpStd = BPSlpEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP Slope effect
indBPint = XtoMint * BPIntEff; ! BP intercept indirect effect
indBPsdp = XtoMsdp * BPSlpEff; ! BP age slope indirect effect

This is still the same model, just with a different way of specifying the
level-1 residual to residual relationship. This method will only work for
level-1 effects that are fixed, though.

```

		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
<u>Within Level - new parameters are in BOLD underline</u>					
<u>RISKY</u>	<u>ON</u>	<u>MON3</u>	<u>3.559</u>	<u>0.301</u>	<u>11.809</u>
<u>Residual Variances</u>					
RISKY			7.329	0.328	22.353
MON3			0.081	0.004	22.354
<u>Between Level - changed parameters are in BOLD underline</u>					
<u>RSLP</u>	<u>ON</u>	<u>MSLP</u>	<u>-5.294</u>	<u>0.806</u>	<u>-6.569</u>
RSLP	ON	ATT4	-0.512	0.105	-4.869
MSLP	ON	ATT4	0.003	0.014	0.240
<u>RISKY</u>	<u>ON</u>	<u>ATT4</u>	<u>-3.354</u>	<u>0.528</u>	<u>-6.348</u>
MON3	ON	MON3	-4.380	0.797	-5.497
MON3	ON	ATT4	-0.044	0.057	-0.779
RISKY	WITH	RSLP	1.480	0.345	4.285
		MSLP	0.039	0.038	1.023
MON3	WITH	MSLP	0.000	0.004	-0.105
		RSLP	-0.107	0.031	-3.454
<u>Means</u>					
RQUAD		0.147	0.021	7.117	0.000
<u>Intercepts</u>					
RISKY		23.598	0.338	69.837	0.000
MON3		0.063	0.034	1.839	0.066
RSLP		1.969	0.139	14.195	0.000
MSLP		-0.003	0.008	-0.380	0.704
<u>Variances</u>					
RQUAD		0.000	0.000	999.000	999.000
<u>Residual Variances</u>					
RISKY		14.315	2.030	7.053	0.000
MON3		0.195	0.023	8.371	0.000
RSLP		0.453	0.081	5.564	0.000
MSLP		0.010	0.001	7.802	0.000
<u>New/Additional Parameters</u>					
RESCOR		0.350	0.034	10.441	0.000
INTSTD		-0.455	0.083	-5.497	0.000
BPSLPEFF		-1.736	0.713	-2.434	0.015
SLPSTD		-0.254	0.104	-2.434	0.015
INDBPINT		0.194	0.251	0.772	0.440
INDBPSLP		-0.006	0.024	-0.239	0.811

**In Mplus, Model 2b as a partially directed single-level SEM:
Monitor → Risky for WP residuals using structured residuals**

```

TITLE: Model 2b: Partially Directed Multivariate Growth Model as Single-Level
      SEM, L1 WP effect as direct path using structured residuals
      ( DATA, VARIABLE, and ANALYSIS are the same as for Model 1 )
MODEL: ! R = risky behavior, M = monitoring
[risky12-risky18@0 mon12-mon18@0]; ! All variable intercepts fixed to 0

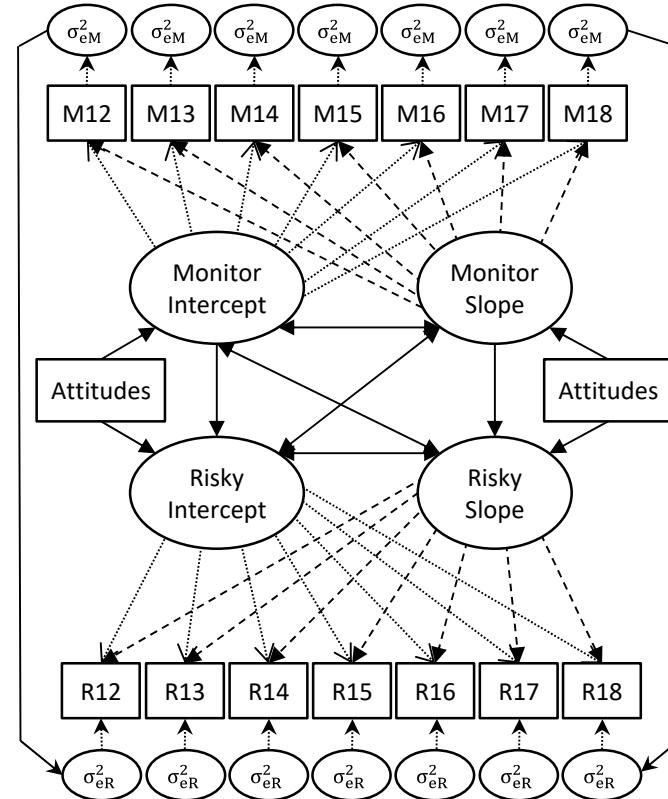
! Risky behavior quadratic growth model using exact age as loadings
Rint Rslp Rquad | risky12-risky18 AT age12-age18;
! Monitoring linear growth model using exact age as loadings
Mint Mslp | mon12-mon18 AT age12-age18;
! Fixed growth effects for R and M
[Rint Rslp Rquad Mint Mslp];
! L2 G: Random int and linear age slope variances, no quad age variance
Rint Rslp Mint Mslp (Rintvar Rslpvar Mintvar Mslpvar); Rquad@0;
! L2 G: Within-variable random int-slope covariances for R, M
Rint WITH Rslp (Rintsdp); Mint WITH Mslp (Mintsdp);
! Attitudes --> R int, R linear slope, M int, M linear slope
Rint Rslp Mint Mslp ON att4 (XtoYint XtoYslp XtoMint XtoMslp);
Rint ON Mint (BPIntEff);      ! Total BP intercept effect
Rslp ON Mslp (BPSlpEff);     ! Total BP age slope effect
Mint WITH Rslp (Int2Slp);    ! L2 G: M int, R slope covariance
Mslp WITH Rint (Slp2Int);    ! L2 G: M slope, R int covariance

! Define new latent factors for residuals at each occasion
Frisky12 BY risky12@1; Frisky13 BY risky13@1; Frisky14 BY risky14@1;
Frisky15 BY risky15@1; Frisky16 BY risky16@1; Frisky17 BY risky17@1;
Frisky18 BY risky18@1; Fmon12 BY mon12@1; Fmon13 BY mon13@1;
Fmon14 BY mon14@1; Fmon15 BY mon15@1; Fmon16 BY mon16@1;
Fmon17 BY mon17@1; Fmon18 BY mon18@1;
! All factor means fixed to 0
[Frisky12-Friskyl18@0 Fmon12-Fmon18@0];
! Shut off old residual variances
risky12-risky18@0 mon12-mon18@0;
! Hold new residual variances equal over time
Frisky12-Friskyl18 (Rresvar); ! L1 R: R residual variances held equal
Fmon12-Fmon18 (Mresvar); ! L1 R: M residual variances held equal
! Factor residual WP effect between same ages, held equal across age
Frisky12-Friskyl18 PON Fmon12-Fmon18 (ResEff);

MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter
! Need to name each new created effect -- all values from undirected model
NEW(ResStd IntStd SlpStd indBpint indBpslp);
! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
ResStd = ResEff * SQRT(0.08077) / SQRT(8.3538); ! STD WP Res effect
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP Int effect
SlpStd = BPSlpEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP Slope effect
indBpint = XtoMint * BPIntEff;                      ! BP intercept indirect effect
indBpslp = XtoMslp * BPSlpEff;                      ! BP age slope indirect effect

```

This is the same model, just with a different way of specifying the level-2 intercept to intercept and slope to slope relationships.



-→ Indicates paths fixed = 1
- - - - - → Indicates paths fixed = time values
- ← - - - - - Indicates paths freely estimated
- → Indicates paths freely estimated
- — → Indicates paths freely estimated between residuals at the same occasion but held equal over time

For balanced time, a linear growth model would look like this instead (add quad as third place before |):

```
Mint Mslp | mon12@-6 mon13@-5 mon14@-4 mon15@-3
           mon16@-2 mon17@-1 mon18@0;
```

Revised OSU Workshop 2018 Example 4 page 11

MODEL RESULTS - <u>changed parameters are in BOLD underline</u>					Revised CCC Workshop 2010 Example 1 page 11							
					Two-Tailed							
		Estimate	S.E.	Est./S.E.	P-Value			RQUAD	0.147	0.021	7.117	0.000
Factor loadings set to 1 omitted												
RINT	ON					Intercepts						
MINT	ON	-4.380	0.797	-5.496	0.000	Intercepts fixed to 0 omitted						
RSLP	ON					RINT	23.598	0.338	69.836	0.000		
MSLP	ON	-1.736	0.713	-2.434	0.015	RSLP	1.969	0.139	14.195	0.000		
<u>FRISKY12</u>	<u>ON</u>					MINT	0.063	0.034	1.839	0.066		
<u>FMON12</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	MSLP	-0.003	0.008	-0.380	0.704		
<u>FRISKY13</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	Variances						
<u>FMON13</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	FMON12	0.081	0.004	22.327	0.000		
<u>FRISKY14</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	FMON13	0.081	0.004	22.327	0.000		
<u>FMON14</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	FMON14	0.081	0.004	22.327	0.000		
<u>FRISKY15</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	FMON15	0.081	0.004	22.327	0.000		
<u>FMON15</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	FMON16	0.081	0.004	22.327	0.000		
<u>FRISKY16</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	FMON17	0.081	0.004	22.327	0.000		
<u>FMON16</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	FMON18	0.081	0.004	22.327	0.000		
<u>FRISKY17</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	RQUAD	0.000	0.000	999.000	999.000		
<u>FRISKY18</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	Residual Variances						
<u>FMON18</u>	<u>ON</u>	<u>3.563</u>	<u>0.302</u>	<u>11.810</u>	<u>0.000</u>	Residual variances fixed to 0 omitted						
RINT	ON					FRISKY12	7.328	0.328	22.349	0.000		
ATT4	ON	-3.354	0.528	-6.348	0.000	FRISKY13	7.328	0.328	22.349	0.000		
RSLP	ON					FRISKY14	7.328	0.328	22.349	0.000		
ATT4	ON	-0.512	0.105	-4.869	0.000	FRISKY15	7.328	0.328	22.349	0.000		
MINT	ON					FRISKY16	7.328	0.328	22.349	0.000		
ATT4	ON	-0.044	0.057	-0.779	0.436	FRISKY17	7.328	0.328	22.349	0.000		
MSLP	ON					FRISKY18	7.328	0.328	22.349	0.000		
ATT4	ON	0.003	0.014	0.240	0.810	RINT	14.315	2.030	7.053	0.000		
RINT	WITH					RSLP	0.453	0.081	5.564	0.000		
RSLP	WITH	1.480	0.345	4.285	0.000	MINT	0.195	0.023	8.371	0.000		
MSLP	WITH	0.039	0.038	1.023	0.306	MSLP	0.010	0.001	7.802	0.000		
MINT	WITH					New/Additional Parameters						
MSLP	WITH	0.000	0.004	-0.105	0.916	RESCOR	0.350	0.034	10.441	0.000		
RSLP	WITH	-0.107	0.031	-3.454	0.001	INTSTD	-0.455	0.083	-5.496	0.000		
					SLPSTD	-0.254	0.104	-2.434	0.015			
					INDBPINT	0.194	0.251	0.772	0.440			
					INDBPSLP	-0.006	0.024	-0.239	0.811			

Model 2c: Partially Directed Path Multivariate MLM in Mplus: Monitor → Risky for WP residuals within L2 model via placeholder syntax
Also demonstrating how to request BP indirect effects

```

TITLE: Model 2c: Partially Directed Multivariate Growth Model as MLM
L1 WP effect as direct path specified in L2 BETWEEN
( DATA, VARIABLE, and ANALYSIS are the same as for Model 1 )

MODEL: ! R = risky behavior, M = monitoring
%WITHIN%
risky mon3 (Rresvar Mresvar); ! L1 R: Residual variances (labels)
Rslp | risky ON age18; ! Placeholder for R linear age slope
Rquad | risky ON agesq; ! Placeholder for R quadratic age slope
Mslp | mon3 ON age18; ! Placeholder for M linear age slope
WPres | risky ON mon3; ! NOW placeholder for L1 WP effect M->R

%BETWEEN%
[risky mon3]; ! Fixed intercepts
risky mon3 (Rintvar Mintvar); ! L2 G: Random intercept variances (labels)
[Rquad Rslp Mslp];
! Fixed age slopes (as defined earlier)
Rslp Mslp (Rslpvar Mslpvar); ! L2 G: Random linear age slope variances
Rquad@0;
! No quadratic age slope variance
risky Rslp ON att4 (XtoYint XtoYslp); ! Att-> R int, linear age slope
mon3 Mslp ON att4 (XtoMint XtoMslp); ! Att-> M int, linear age slope
risky WITH Rslp (RIntSlp); ! R Int-slope covariance (label)
mon3 WITH Mslp (MIntSlp); ! M Int-slope covariance (label)

! And now both the intercept -> intercept path and the slope -> slope path
! are contextual BP effects given the L1 placeholder for WP residual effect

risky ON mon3 (IntCont); ! NOW contextual BP intercept effect
Rslp ON Mslp (SlpCont); ! NOW contextual BP slope effect
mon3 WITH Rslp (Int2Slp); ! L2 G: M int, R slope covariance
Mslp WITH risky (Slp2Int); ! L2 G: M slope, R int covariance

[WPres] (ResEff); ! Fixed effect for L1 WP M->R (as defined earlier)
WPres@0; ! No random L1 WP M->R effect variance

MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter
! Need to name each new created effect -- all values from undirected model
NEW(ResStd BPIntEff IntStd BPSlpEff SlpStd indBPint indBPsdp);

! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
ResStd = ResEff * SQRT(0.08077) / SQRT(8.3538); ! STD WP Res effect
BPIntEff = ResEff + IntCont; ! WP + Context = BP int
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP Int effect
BPSlpEff = ResEff + SlpCont; ! WP + Context = BP slp
SlpStd = BPSlpEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP Slope effect
indBPint = XtoMint * BPIntEff; ! BP intercept indirect effect
indBPsdp = XtoMslp * BPSlpEff; ! BP age slope indirect effect

This is still the same model, just with a different syntax for specifying
the same level-1 residual to residual directed relationship. This is the
version that is necessary in order to have the level-1 effect become
random or systematically varying.

```

		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Within Level					
Residual Variances					
RISKY	ON	7.329	0.328	22.353	0.000
MON3	ON	0.081	0.004	22.354	0.000
Between Level - <u>changed parameters are in BOLD underline</u>					
RSLP	ON	-5.294	0.806	-6.569	0.000
MSLP	ON	-0.512	0.105	-4.869	0.000
ATT4	ON	0.003	0.014	0.240	0.810
RISKY	ON	-3.354	0.528	-6.348	0.000
ATT4	ON	-7.938	0.872	-9.099	0.000
MON3	ON	-0.044	0.057	-0.779	0.436
RISKY	WITH				
RSLP	WITH	1.480	0.345	4.285	0.000
MSLP	WITH	0.039	0.038	1.023	0.306
MON3	WITH				
MSLP	WITH	0.000	0.004	-0.105	0.916
RSLP	WITH	-0.107	0.031	-3.454	0.001
Means					
RQUAD	MEAN	0.147	0.021	7.117	0.000
WPRES (here now)	MEAN	3.559	0.301	11.810	0.000
Intercepts					
RISKY	INTERCEPT	23.598	0.338	69.837	0.000
MON3	INTERCEPT	0.063	0.034	1.839	0.066
RSLP	INTERCEPT	1.969	0.139	14.195	0.000
MSLP	INTERCEPT	-0.003	0.008	-0.380	0.704
Variances					
RQUAD	VARIANCE	0.000	0.000	999.000	999.000
WPRES	VARIANCE	0.000	0.000	999.000	999.000
Residual Variances					
RISKY	RESIDUAL	14.315	2.030	7.053	0.000
MON3	RESIDUAL	0.195	0.023	8.371	0.000
RSLP	RESIDUAL	0.453	0.081	5.564	0.000
MSLP	RESIDUAL	0.010	0.001	7.802	0.000
New/Additional Parameters					
RESSTD	PARAMETER	0.350	0.030	11.810	0.000
BPINTEFF	PARAMETER	-4.380	0.797	-5.496	0.000
INTSTD	PARAMETER	-0.455	0.083	-5.496	0.000
BPSLPEFF	PARAMETER	-1.735	0.713	-2.434	0.015
SLPSTD	PARAMETER	-0.254	0.104	-2.434	0.015
INDBPINT	PARAMETER	0.194	0.251	0.772	0.440
INDBPSLP	PARAMETER	-0.006	0.024	-0.239	0.811

**In Mplus, Model 2c as a partially directed single-level SEM:
Monitor → Risky for WP residuals using original residuals**

```

TITLE: Model 2c: Partially Directed Multivariate Growth Model as Single-
       Level SEM, L1 WP effect as direct path using original residuals

( DATA, VARIABLE, and ANALYSIS are the same as for Model 1 )
MODEL: ! R = risky behavior, M = monitoring
[risky12-risky18@0 mon12-mon18@0]; ! All variable intercepts fixed to 0
risky12-risky18 (Rresvar);           ! L1 R: R residual variances held equal
mon12-mon18 (Mresvar);             ! L1 R: M residual variances held equal

! Risky behavior quadratic growth model using exact age as loadings
Rint Rslp Rquad | risky12-risky18 AT age12-age18;
! Monitoring linear growth model using exact age as loadings
Mint Mslp | mon12-mon18 AT age12-age18;
! Fixed growth effects for R and M
[Rint Rslp Rquad Mint Mslp];
! L2 G: Random int and linear age slope variances, no quad age variance
Rint Rslp Mint Mslp (Rintvar Rslpvar Mintvar Mslpvar); Rquad@0;
! L2 G: Within-variable random int-slope covariances for R, M
Rint WITH Rslp (Rintsdp); Mint WITH Mslp (Mintsdp);
! Attitudes --> R int, R linear slope, M int, M linear slope
Rint Rslp Mint Mslp ON att4 (XtoYint XtoYslp XtoMint XtoMslp);

Rint ON Mint (IntCont);      ! NOW contextual BP intercept effect
Rslp ON Mslp (SlpCont);    ! NOW contextual BP age slope effect
Mint WITH Rslp (Int2Slp);   ! L2 G: M int, R slope covariance
Mslp WITH Rint (Slp2Int);   ! L2 G: M slope, R int covariance

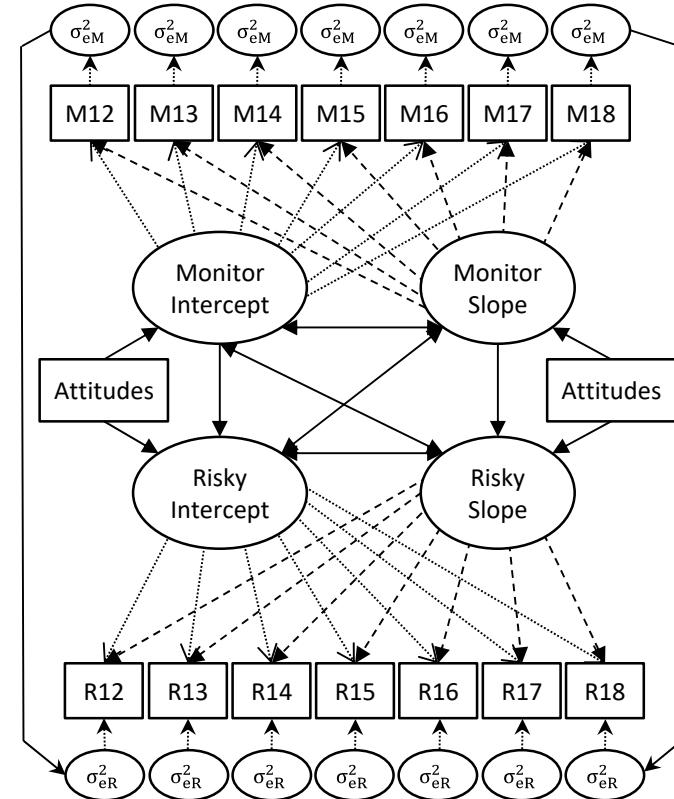
! Residual WP covariance between same ages, held equal across age
risky12-risky18 PON mon12-mon18 (ResEff);

MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter
! Need to name each new created effect -- all values from undirected model
NEW(ResStd BPIntEff IntStd BPSlpEff SlpStd indBPint indBPsdp);

! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
ResStd = ResEff * SQRT(0.08077) / SQRT(8.3538); ! STD WP Res effect
BPIntEff = ResEff + IntCont;                      ! WP + Context = BP int
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP Int effect
BPSlpEff = ResEff + SlpCont;                      ! WP + Context = BP slp
SlpStd = BPSlpEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP Slope effect
indBPint = XtoMint * BPIntEff;                    ! BP intercept indirect effect
indBPsdp = XtoMslp * BPSlpEff;                   ! BP age slope indirect effect

```

This is still the same model, just with a different syntax for specifying the same level-1 residual to residual directed relationship. The consequence is that the intercept to intercept and slope to slope relationships become contextual BP effects. Oddly, if we were to switch to ON for the int-slope cross-variable relationships, those stay total BP (see chapter 9 for an example using this version of the model).



-> Indicates paths fixed = 1
- - -> Indicates paths fixed = time values
- < - -> Indicates paths freely estimated
- > Indicates paths freely estimated
- > Indicates paths freely estimated between residuals at the same occasion but held equal over time

For balanced time, a linear growth model would look like this instead (add quad as third place before |):

```
Mint Mslp | mon12@-6 mon13@-5 mon14@-4 mon15@-3
          mon16@-2 mon17@-1 mon18@0;
```

MODEL RESULTS - <u>changed parameters are in BOLD underline</u>					Means					
					RQUAD	0.147	0.021	7.117	0.000	
Estimate S.E. Est./S.E. P-Value										
Factor loadings set to 1 omitted										
RINT	ON				Intercepts					
MINT		-7.939	0.872	-9.099	0.000	Intercepts fixed to 0 omitted				
RSLP	ON					RINT	23.598	0.338	69.836	0.000
MSLP		-5.294	0.806	-6.569	0.000	RSLP	1.969	0.139	14.195	0.000
RINT	ON					MINT	0.063	0.034	1.839	0.066
ATT4		-3.354	0.528	-6.348	0.000	MSLP	-0.003	0.008	-0.380	0.704
RSLP	ON					Variances				
ATT4		-0.512	0.105	-4.869	0.000	RQUAD	0.000	0.000	999.000	999.000
MINT	ON					Residual Variances				
ATT4		-0.044	0.057	-0.779	0.436	RISKY12	7.329	0.328	22.353	0.000
MSLP	ON					RISKY13	7.329	0.328	22.353	0.000
ATT4		0.003	0.014	0.240	0.810	RISKY14	7.329	0.328	22.353	0.000
RISKY12	ON					RISKY15	7.329	0.328	22.353	0.000
MON12		3.559	0.301	11.810	0.000	RISKY16	7.329	0.328	22.353	0.000
RISKY13	ON					RISKY17	7.329	0.328	22.353	0.000
MON13		3.559	0.301	11.810	0.000	RISKY18	7.329	0.328	22.353	0.000
RISKY14	ON					MON12	0.081	0.004	22.354	0.000
MON14		3.559	0.301	11.810	0.000	MON13	0.081	0.004	22.354	0.000
RISKY15	ON					MON14	0.081	0.004	22.354	0.000
MON15		3.559	0.301	11.810	0.000	MON15	0.081	0.004	22.354	0.000
RISKY16	ON					MON16	0.081	0.004	22.354	0.000
MON16		3.559	0.301	11.810	0.000	MON17	0.081	0.004	22.354	0.000
RISKY17	ON					MON18	0.081	0.004	22.354	0.000
MON17		3.559	0.301	11.810	0.000	RINT	14.315	2.030	7.053	0.000
RISKY18	ON					RSLP	0.453	0.081	5.564	0.000
MON18		3.559	0.301	11.810	0.000	MINT	0.195	0.023	8.371	0.000
RINT	WITH					MSLP	0.010	0.001	7.802	0.000
RSLP		1.480	0.345	4.285	0.000	New/Additional Parameters				
MSLP		0.039	0.038	1.023	0.306	RESSTD	0.350	0.030	11.810	0.000
MINT	WITH					BPINTEFF	-4.380	0.797	-5.496	0.000
MSLP		0.000	0.004	-0.105	0.916	INTSTD	-0.455	0.083	-5.496	0.000
RSLP		-0.107	0.031	-3.454	0.001	BPSLPEFF	-1.735	0.713	-2.434	0.015
						SLPSTD	-0.254	0.104	-2.434	0.015
						INDBPINT	0.194	0.251	0.772	0.440
						INDBPSLP	-0.006	0.024	-0.239	0.811

By popular demand, here is an example of how to use “structured residuals” to fit two cross-lag effects at level 1: Model 3a, which switches to covariances at level 2 when fitting these models (per convention, to be agnostic as to which comes first)

TITLE: Model 3a: SEM Structured Residuals to Fit 2 Cross-Lag Paths (DATA, VARIABLE, and ANALYSIS are the same as for Model 1)		MODEL RESULTS - Parameters fixed to 0 or 1 are omitted for brevity				
		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value	
MODEL: ! R = risky behavior, M = monitoring	FRISKY13 ON	-0.255	0.373	-0.682	0.495	
[risky12-risky18@0 mon12-mon18@0]; ! All variable intercepts fixed to 0	FRISKY14 ON	-0.255	0.373	-0.682	0.495	
! Risky behavior quadratic growth model using exact age as loadings	FRISKY15 ON	-0.255	0.373	-0.682	0.495	
Rint Rslp Rquad risky12-risky18 AT age12-age18;	FRISKY16 ON	-0.255	0.373	-0.682	0.495	
! Monitoring linear growth model using exact age as loadings	FRISKY17 ON	-0.255	0.373	-0.682	0.495	
Mint Mslp mon12-mon18 AT age12-age18;	FRISKY18 ON	-0.255	0.373	-0.682	0.495	
! Fixed growth effects for R and M	FMON13 ON	0.008	0.004	2.082	0.037	
[Rint Rslp Rquad Mint Mslp];	FMON14 ON	0.008	0.004	2.082	0.037	
! L2 G: Random int and linear age slope variances, no quad age variance	FMON15 ON	0.008	0.004	2.082	0.037	
Rint Rslp Mint Mslp (Rintvar Rslpvar Mintvar Mslpvar); Rquad@0;	FMON16 ON	0.008	0.004	2.082	0.037	
! Attitudes --> R int, R linear slope, M int, M linear slope	FMON17 ON	0.008	0.004	2.082	0.037	
Rint Rslp Mint Mslp ON att4 (XtoYint XtoYslp XtoMint XtoMslp);	FRISKY12	0.008	0.004	2.082	0.037	
! L2 G: covariances for random intercepts and slopes across outcomes	FRISKY13	0.008	0.004	2.082	0.037	
Rint Rslp Mint Mslp WITH Rint Rslp Mint Mslp;	FRISKY14	0.008	0.004	2.082	0.037	
! Define new latent factors for residuals at each occasion	FRISKY15	0.008	0.004	2.082	0.037	
Frisky12 BY risky12@1; Frisky13 BY risky13@1; Frisky14 BY risky14@1;	FRISKY16	0.008	0.004	2.082	0.037	
Frisky15 BY risky15@1; Frisky16 BY risky16@1; Frisky17 BY risky17@1;	FRISKY17	0.008	0.004	2.082	0.037	
Frisky18 BY risky18@1; Fmon12 BY mon12@1; Fmon13 BY mon13@1;	FMON13	0.008	0.004	2.082	0.037	
Fmon14 BY mon14@1; Fmon15 BY mon15@1; Fmon16 BY mon16@1;	FMON14	0.008	0.004	2.082	0.037	
Fmon17 BY mon17@1; Fmon18 BY mon18@1;	FMON15	0.008	0.004	2.082	0.037	
! All factor means fixed to 0	FMON16	0.008	0.004	2.082	0.037	
[Frisky12-Friskyl18@0 Fmon12-Fmon18@0];	FMON17	0.008	0.004	2.082	0.037	
! Shut off old residual variances	FMON18	0.008	0.004	2.082	0.037	
risky12-risky18@0 mon12-mon18@0;	FRISKY12	0.008	0.004	2.082	0.037	
! Hold new residual variances equal over time	FRISKY13	0.008	0.004	2.082	0.037	
Frisky12-Friskyl18 (Rresvar); ! L1 R: R residual variances held equal	FRISKY14	0.008	0.004	2.082	0.037	
Fmon12-Fmon18 (Mresvar); ! L1 R: M residual variances held equal	FRISKY15	0.008	0.004	2.082	0.037	
! Factor residual WP effect between same ages, held equal across age	FRISKY16	0.008	0.004	2.082	0.037	
Frisky12-Friskyl18 PWITH Fmon12-Fmon18 (ResCov);	FRISKY17	0.008	0.004	2.082	0.037	
! Cross-lag WP effects predicting next age, held equal across age	RINT ON	-3.156	0.551	-5.725	0.000	
Frisky13-Friskyl18 PON Fmon12-Fmon17 (MR2RR);	RSLP ON	-0.516	0.104	-4.945	0.000	
Fmon13-Fmon18 PON Frisky12-Friskyl17 (RR2MR);	MINT ON	-0.045	0.057	-0.796	0.426	
MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter	ATT4	0.003	0.014	0.221	0.825	
! Need to name each new created effect - using actual values	MSLP ON	1.901	0.357	5.318	0.000	
NEW(ResCor MR2RRsd RR2MRsd);	ATT4	-0.880	0.170	-5.172	0.000	
! Corr = Cov / (SQRT(Yvar)*SQRT(Xvar))	ATT4	0.033	0.039	0.847	0.397	
ResCor = ResCov / (SQRT(8.3538)*SQRT(0.08077));	MINT WITH	-0.001	0.004	-0.186	0.852	
! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)	MSLP	-0.110	0.031	-3.529	0.000	
MR2RRsd = MR2RR * SQRT(0.08077) / SQRT(8.3538); ! STD M->R lag effect	RSLP WITH	-0.020	0.008	-2.642	0.008	
RR2MRsd = RR2MR * SQRT(8.3538) / SQRT(0.08077); ! STD R->M lag effect	MSLP					

	What if we controlled for the contemporaneous effect of M → before examining the lagged effect of M → R (my own preference)?								
FRISKY12 WITH FMON12	0.298	0.031	9.607	0.000	TITLE: Model 3b: Example of Structured Residuals to Fit M→R Cross-Lag Path Controls for contemporaneous effect before fitting the lagged effect				
FRISKY13 WITH FMON13	0.298	0.031	9.607	0.000	All else is the same until here...				
FRISKY14 WITH FMON14	0.298	0.031	9.607	0.000	! Factor residual WP effect between same ages, held equal across age Frisky12-Friskyl8 PON Fmon12-Fmon18 (ResEff);				
FRISKY15 WITH FMON15	0.298	0.031	9.607	0.000	! Cross-lag WP effects predicting next age, held equal across age Frisky13-Friskyl8 PON Fmon12-Fmon17 (MR2RR);				
FRISKY16 WITH FMON16	0.298	0.031	9.607	0.000	MODEL CONSTRAINT:				
FRISKY17 WITH FMON17	0.298	0.031	9.607	0.000	NEW(ResStd MR2RRsd);				
FRISKY18 WITH FMON18	0.298	0.031	9.607	0.000	! STD = Unstd * SQRT(Xvar) / SQRT(Yvar) ResStd = ResEff * SQRT(Mresvar) / SQRT(Rresvar); ! STD M→R contemporaneous MR2RRsd = MR2RR * SQRT(Mresvar) / SQRT(Rresvar); ! STD M→R lag effect				
Means									
RQUAD	0.146	0.020	7.195	0.000					
Intercepts									
RINT	23.320	0.347	67.184	0.000					
RSLP	1.971	0.136	14.464	0.000					
MINT	0.063	0.034	1.844	0.065					
MSLP	-0.003	0.008	-0.369	0.712					
Variances									
FRISKY12	8.301	0.379	21.890	0.000					
FMON12	0.081	0.004	22.125	0.000					
RQUAD	0.000	0.000	999.000	999.000					
Residual Variances									
FRISKY13	8.301	0.379	21.890	0.000					
FRISKY14	8.301	0.379	21.890	0.000					
FRISKY15	8.301	0.379	21.890	0.000					
FRISKY16	8.301	0.379	21.890	0.000					
FRISKY17	8.301	0.379	21.890	0.000					
FRISKY18	8.301	0.379	21.890	0.000					
FMON13	0.081	0.004	22.126	0.000					
FMON14	0.081	0.004	22.126	0.000					
FMON15	0.081	0.004	22.126	0.000					
FMON16	0.081	0.004	22.126	0.000					
FMON17	0.081	0.004	22.126	0.000					
FMON18	0.081	0.004	22.126	0.000					
RINT	18.153	2.208	8.223	0.000					
RSLP	0.492	0.080	6.132	0.000					
MINT	0.194	0.023	8.300	0.000					
MSLP	0.010	0.001	7.676	0.000					
New/Additional Parameters									
RESCOR	0.363	0.031	11.624	0.000					
MR2RRSD	-0.025	0.037	-0.681	0.496					
RR2MRSD	0.077	0.037	2.083	0.037					

It looks like evidence for a lagged R → M effect is even stronger after controlling for the contemporaneous effect (and vice-versa).

Here is a comparison of the SEM cross-lagged effects to those from MLM using either manually lagged predictors (in WITHIN only, using ML or BAYES), and to those from MLM using the new LAGGED option (BAYES only). To do so, I had to move all fixed-only L1 effects to the WITHIN model (not use L1 placeholders).

Here is Model 3a using new LAGGED and BAYES: code is same when using my own lagged versions (in which occasion12 = 0), except that lagged variables get added to USEVARIABLE and WITHIN lines, and are used in %WITHIN% model in place of mon3&1 or risky&1 (or ^mon3 or ^risky in Mplus 8.1).

```

DATA: FILE = Example4Lag.csv; ! Different data
VARIABLE: ! List of variables in data file
NAMES = PersonID occasion risky age18 att4 mon3 agesq LagRisky LagMon3 Use12;
! Variables to be analyzed in this model
USEVARIABLE = age18 agesq att4 risky mon3;
MISSING ARE ALL (-999); ! Missing data identifier
! MLM options
CLUSTER = PersonID; ! Level-2 ID
BETWEEN = att4; ! Observed ONLY level-2 predictors
WITHIN = age18 agesq; ! Observed ONLY level-1 predictors
! Was removed when I used my own lagged variables
LAGGED = risky(1) mon3(1); ! Create Mplus lag-1 versions

ANALYSIS: TYPE = TWOLEVEL RANDOM; ESTIMATOR = BAYES; ! OR ML
MODEL: ! R = risky behavior, M = monitoring
%WITHIN%
risky mon3 (Rresvar Mresvar); ! L1 R: Residual variances (labels)
Rsdp | risky ON age18; ! Placeholder for R linear age slope
risky ON agesq; ! R fixed quadratic age slope
Msdp | mon3 ON age18; ! Placeholder for M linear age slope
risky WITH mon3 (ResCov); ! L1 WP covariance for contemp M->R
risky ON mon3&1 (MRLagEff); ! L1 WP fixed effect of lagged M->R
mon3 ON risky&1 (RMLagEff); ! L1 WP fixed effect of lagged R->M
! New in 8.1 to create lagged residual effects instead (remove other lag lines)
risky^ ON mon3^1 (MRLagEff); ! L1 WP fixed effect of ^res lagged M->R
mon3^ ON risky^1 (RMLagEff); ! L1 WP fixed effect of ^res lagged R->M

%BETWEEN%
[risky mon3]; ! Fixed intercepts
risky mon3 (Rintvar Mintvar); ! L2 G: Random intercept variances (labels)
[Rsdp Msdp]; ! Fixed age slopes (as defined earlier)
Rsdp Msdp (Rsdpvar Msdpvar); ! L2 G: Random linear age slope variances
risky Rsdp ON att4 (XtoYint XtoYslp); ! Att-> R int, linear age slope
mon3 Msdp ON att4 (XtoMint XtoMsdp); ! Att-> M int, linear age slope
! L2 G: covariances for random intercepts and slopes across outcomes
risky Rsdp mon3 Msdp WITH risky Rsdp mon3 Msdp;

MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter
NEW(ResCor MR2RRsd RR2MRsd);
! Corr = Cov / (SQRT(Yvar)*SQRT(Xvar))
ResCor = ResCov / (SQRT(8.3538)*SQRT(0.08077));
! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
MR2RRsd = MRLagEff * SQRT(0.08077) / SQRT(8.3538); ! STD M->R lag effect
RR2MRsd = RMLagEff * SQRT(8.3538) / SQRT(0.08077); ! STD R->M lag effect

```

SEs for the lagged effects are noticeably different when using a non-model-provided lagged variable. Tread carefully in this case—it appears to be better to fit cross-lagged effects one of the other two ways.

However: we cannot discern the reason for the differences between the SEM lagged version and the MLM LAGGED version because they must use different estimators (ML vs Bayes; LAGGED is not available in ML and SEM cannot use BAYES with TSCORES).

3a SEM via ML and its lagged effects

	Two-Tailed			
	Estimate	S.E.	Est./S.E.	P-Value
RESCOR	0.363	0.031	11.624	0.000 *
MR2RRSD	-0.025	0.037	-0.681	0.496
RR2MRSD	0.077	0.037	2.083	0.037 *

3a MLM via ML and my lagged effects

	Two-Tailed			
	Estimate	S.E.	Est./S.E.	P-Value
RESCOR	0.341	0.034	10.036	0.000 *
MR2RRSD	-0.051	0.022	-2.310	0.021 *
RR2MRSD	0.017	0.015	1.115	0.265

3a MLM via Bayes and my lagged effects

	Posterior 95% C.I.			
	Estimate	S.D.	Lower 2.5%	Upper 2.5%
RESCOR	0.336	0.033	0.276	0.405 *
MR2RRSD	-0.054	0.023	-0.092	-0.004 *
RR2MRSD	0.013	0.016	-0.014	0.045

3a MLM via Bayes and their lagged effects

	Posterior 95% C.I.			
	Estimate	S.D.	Lower 2.5%	Upper 2.5%
RESCOR	0.349	0.034	0.288	0.419 *
MR2RRSD	-0.048	0.025	-0.101	-0.003 *
RR2MRSD	0.028	0.024	-0.018	0.076

The new version of residual lagged effects below appears to match the SEM version most closely:

New in 8.1: 3a MLM via Bayes and their ^residual lagged effects				
RESCOR	0.366	0.039	0.300	0.449 *
MR2RRSD	-0.021	0.035	-0.085	0.050
RR2MRSD	0.077	0.033	0.012	0.144 *

Here are Models 3b and 3c using new LAGGED and BAYES: code is same when using my own lagged versions (in which occasion12 = 0), except that lagged variables get added to USEVARIABLE and WITHIN lines, and are used in %WITHIN% model in place of mon3&1 or risky&1 (or ^mon3 or ^risky in 8.1).

```

DATA: FILE = Example4Lag.csv; ! Different data
VARIABLE: ! List of variables in data file
  NAMES = PersonID occasion risky age18 att4 mon3 agesq LagRisky LagMon3 Use12;
! Variables to be analyzed in this model
  USEVARIABLE = age18 agesq att4 risky mon3;
  MISSING ALL (-999); ! Missing data identifier
! MLM options
  CLUSTER = PersonID; ! Level-2 ID
  BETWEEN = att4; ! Observed ONLY level-2 predictors
  WITHIN = age18 agesq; ! Observed ONLY level-1 predictors
! LAGGED was removed when I used my own lagged variables
  LAGGED = mon3(1); ! For Model 3b: Create Mplus lag-1 version
  LAGGED = risky(1); ! For Model 3c: Create Mplus lag-1 version
  LAGGED = mon3(1) risky(1); ! For 8.1 res lag-1 versions

ANALYSIS: TYPE = TWOLEVEL RANDOM; ESTIMATOR = BAYES; ! OR ML
MODEL: ! R = risky behavior, M = monitoring
%WITHIN%
  risky mon3 (Rresvar Mresvar); ! L1 R: Residual variances (labels)
  Rsdp | risky ON age18; ! Placeholder for R linear age slope
  risky ON agesq; ! R fixed quadratic age slope
  Msdp | mon3 ON age18; ! Placeholder for M linear age slope
  risky ON mon3 (ResEff); ! Model 3b: L1 WP covariance for contemp M->R
  mon3 ON risky (ResEff); ! Model 3c: L1 WP covariance for contemp R->M
  risky ON mon3&1 (MLlagEff); ! Model 3b: L1 WP fixed effect of lagged M->R
  mon3 ON risky&1 (MLlagEff); ! Model 3c: L1 WP fixed effect of lagged R->M
! New in 8.1 to create lagged residual effects instead (remove other lag lines)
  risky^ ON mon3^1 (MLlagEff); ! Model 3b: L1 WP fixed eff of res lagged M->R
  mon3^ ON risky^1 (MLlagEff); ! Model 3c: L1 WP fixed eff of res lagged R->M

%BETWEEN%
  [risky mon3]; ! Fixed intercepts
  risky mon3 (Rintvar Mintvar); ! L2 G: Random intercept variances (labels)
  [Rsdp Msdp]; ! Fixed age slopes (as defined earlier)
  Rsdp Msdp (Rsdpvar Msdpvar); ! L2 G: Random linear age slope variances
  risky Rsdp ON att4 (XtoYint XtoYsdp); ! Att-> R int, linear age slope
  mon3 Msdp ON att4 (XtoMint XtoMsdp); ! Att-> M int, linear age slope
! L2 G: covariances for random intercepts and slopes across outcomes
  risky Rsdp mon3 Msdp WITH risky Rsdp mon3 Msdp;

MODEL CONSTRAINT: ! Like ESTIMATE in SAS, but can refer to any parameter
! Need to name each new created effect -- using actual values
NEW(ResStd MR2RRsd RR2MRsd);
! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
  ResStd = ResEff * SQRT(0.08077) / SQRT(8.3538); ! 3b: STD M->R contemp
  MR2RRsd = MLlagEff * SQRT(0.08077) / SQRT(8.3538); ! 3b: STD M->R lag effect
  ResStd = ResEff * SQRT(8.3538) / SQRT(0.08077); ! 3c: STD R->M contemp
  RR2MRsd = MLlagEff * SQRT(8.3538) / SQRT(0.08077); ! 3c: STD R->M lag effect

```

3b SEM via ML and its lagged effects					
	Estimate	S.E.	Est./S.E.	2-T	P-Value
RESSTD	0.335	0.031	10.789	0.000	*
MR2RRSD	-0.056	0.034	-1.654	0.098	
3b MLM via ML and my lagged effects					
	Estimate	S.E.	Est./S.E.	2-T	P-Value
RESSTD	0.336	0.030	11.185	0.000	*
MR2RRSD	-0.053	0.022	-2.438	0.015	*
3b MLM via Bayes and my lagged effects					
	Posterior	95% C.I.			
	Estimate	S.D.	Lower 2.5%	Upper 2.5%	
RESSTD	0.341	0.027	0.288	0.391	*
MR2RRSD	-0.047	0.022	-0.092	-0.006	*
3b MLM via Bayes and their lagged effects					
	Posterior	95% C.I.			
	Estimate	S.D.	Lower 2.5%	Upper 2.5%	
RESSTD	0.330	0.032	0.265	0.391	*
MR2RRSD	-0.059	0.022	-0.097	-0.013	*
New in 8.1: 3b MLM via Bayes and their ^residual lagged effects					
	Posterior	95% C.I.			
	Estimate	S.D.	Lower 2.5%	Upper 2.5%	
RESSTD	0.330	0.028	0.276	0.377	*
MR2RRSD	-0.058	0.033	-0.123	0.005	
3c SEM via ML and its lagged effects					
	Estimate	S.E.	Est./S.E.	2-T	P-Value
RESSTD	0.373	0.031	12.097	0.000	*
RR2MRSD	0.087	0.034	2.570	0.010	*
3c MLM via ML and my lagged effects					
	Estimate	S.E.	Est./S.E.	2-T	P-Value
RESSTD	0.353	0.030	11.858	0.000	*
RR2MRSD	0.041	0.015	2.755	0.006	*
3c MLM via Bayes and my lagged effects					
	Posterior	95% C.I.			
	Estimate	S.D.	Lower 2.5%	Upper 2.5%	
RESCOR	0.353	0.027	0.301	0.404	*
RR2MRSD	0.045	0.014	0.017	0.069	*
3c MLM via Bayes and their lagged effects					
	Posterior	95% C.I.			
	Estimate	S.D.	Lower 2.5%	Upper 2.5%	
RESCOR	0.327	0.027	0.271	0.381	*
RR2MRSD	0.005	0.023	-0.038	0.051	
New in 8.1: 3c MLM via Bayes and their ^residual lagged effects					
	Posterior	95% C.I.			
	Estimate	S.D.	Lower 2.5%	Upper 2.5%	
RESCOR	0.361	0.031	0.289	0.414	*
RR2MRSD	0.084	0.036	0.013	0.155	*