

## Example 2: Unconditional Polynomial Models for Change in Number Match 3 Response Time (complete data, syntax, and output available for SAS, SPSS, and STATA electronically)

These data (in "Example23" data files) come from a short-term longitudinal study of 6 observations over 2 weeks for 101 adults age 65–80. The goal is to see how performance on this processing speed task ("number match 3"), as measured by response time in milliseconds, declines over the 6 practice sessions.

### SAS Code for Data Manipulation:

```
* SAS code to import data, center time for polynomial models;
DATA work.example23; SET filepath.example23;
    clsess = session - 1; LABEL clsess = "clsess: Session Centered at 1";
RUN;
```

### SPSS Code for Data Manipulation:

```
* SPSS code to import data, center time for polynomial models.
GET FILE = "example/Example23.sav".
DATASET NAME example23 WINDOW=FRONT.
COMPUTE clsess = session - 1.
VARIABLE LABELS clsess "clsess: Session Centered at 1".
```

### STATA Code for Data Manipulation:

```
* STATA code to center time for polynomial models (and make quadratic version)
gen clsess = session - 1
gen clsess2 = clsess * clsess
label variable clsess "clsess: Session Centered at 1"
label variable clsess2 "clsess2: Quadratic Session Centered at 1"
```

### Model 1a. Most Conservative Baseline: Empty Means, Random Intercept

$$\text{Level 1: } y_{ti} = \beta_{0i} + e_{ti}$$

$$\text{Level 2: Intercept: } \beta_{0i} = \gamma_{00} + U_{0i}$$

TITLE1 "SAS Model 1a: Empty Means, Random Intercept Only";

```
PROC MIXED DATA=work.example23 NOCLPRINT NOITPRINT
    COVTEST NAMELEN=100 METHOD=REML;
    CLASS ID session;
    MODEL nm3rt = / SOLUTION DDFM=Satterthwaite;
    RANDOM INTERCEPT / G V V CORR TYPE=UN SUBJECT=ID;
    REPEATED session / R TYPE=VC SUBJECT=ID; RUN;
```

METHOD = ML or REML (default)  
CLASS = categorical predictors, nesting  
MODEL dv = fixed effects / print solution  
RANDOM = person variances in **G**  
REPEATED = residuals in **R** matrix

TITLE "SPSS Model 1a: Empty Means, Random Intercept".

```
MIXED nm3rt BY ID session
    /METHOD = REML
    /PRINT = SOLUTION TESTCOV G R
    /FIXED =
    /RANDOM = INTERCEPT | SUBJECT(ID) COVTYPE(UN)
    /REPEATED = session | SUBJECT(ID) COVTYPE(ID).
```

MIXED dv BY categorical predictors  
WITH continuous predictors  
/METHOD = REML or ML  
/PRINT = regression solution  
/FIXED = predictors for means model  
/RANDOM = person variances in **G**

\* STATA Model 1a: Empty Means, Random Intercept

```
xtmixed nm3rt , || id: , ///
    variance reml covariance(unstructured) residuals(independent,t(session)),
    estat ic, n(101)
    estat recovariance, level(id)
```

DV = nm3rt, random part after ||  
Level 2 ID is PersonID, random intercept by default  
Print variances instead of SD, use reml  
covariance(unstructured) refers to **G** matrix  
residuals(independent) → refers to **R** matrix by session  
estat ic → Print IC given N = 101 persons

**STATA output:**

```

Mixed-effects REML regression              Number of obs   =       606
Group variable: id                        Number of groups  =       101

                                         Obs per group: min =        6
                                         avg   =       6.0
                                         max   =        6

                                         Wald chi2(0)      =        .
                                         Prob > chi2       =        .

```

**NOTE: LL is given rather than -2LL**

Log restricted-likelihood = -4268.4304

```

-----+-----
      nm3rt |      Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      _cons |    1770.701   45.42063    38.98  0.000    1681.679    1859.724
-----+-----

```

This is the fixed intercept  
(just grand mean so far).

```

-----+-----
Random-effects Parameters |   Estimate    Std. Err.      [95% Conf. Interval]
-----+-----
id: Identity
      var(_cons) |       200883   29471.23     150683.2     267806.8
-----+-----
      var(Residual) |   44899.96   2825.63     39689.76     50794.13
-----+-----

```

LR test vs. linear regression: chibar2(01) = 691.74 Prob >= chibar2 = 0.0000

. estat ic, n(101)

```

-----+-----
      Model |   Obs   ll(null)   ll(model)    df       AIC       BIC
-----+-----
      .    |   101         .   -4268.43     3   8542.861   8550.706
-----+-----

```

Note: N=101 used in calculating BIC

**Calculate the ICC for the  
Number Match 3 outcome:**

$$ICC = \frac{200883}{200883 + 44900} = .82$$

This LR test tells us that the  
random intercept variance is  
significantly greater than 0,  
and thus so is the ICC.

REML-based AIC and BIC are  
calculated differently in  
STATA (they count fixed  
effects), so they won't match  
the values in other programs.

```

.      estat recovariance, level(id)
Random-effects covariance matrix for level id

```

```

-----+-----
      |      _cons
-----+-----
      |
      _cons |      200883
-----+-----

```

This is the level-2 **G** matrix, just a  
random intercept variance so far.

**Extra SAS output not provided by STATA:****Estimated R Matrix for ID 101**

Row	Col1	Col2	Col3	Col4	Col5	Col6
1	44900					
2		44900				
3			44900			
4				44900		
5					44900	
6						44900

This level-1 **R** matrix (with  
equal variance over time, no  
covariance of any kind, known  
as VC or independence) will  
be used repeatedly as we add  
fixed and random effects.

**Estimated G Matrix**

Row	Effect	Participant ID	Col1
1	Intercept	101	200883

This is the level-2 **G** matrix, just a  
random intercept variance so far.

**Estimated V Matrix for ID 101**

Row	Col1	Col2	Col3	Col4	Col5	Col6
1	<b>245783</b>	200883	200883	200883	200883	200883
2	200883	<b>245783</b>	200883	200883	200883	200883
3	200883	200883	<b>245783</b>	200883	200883	200883
4	200883	200883	200883	<b>245783</b>	200883	200883
5	200883	200883	200883	200883	<b>245783</b>	200883
6	200883	200883	200883	200883	200883	<b>245783</b>

The **V** matrix is the total  
variance-covariance matrix  
after combining the level-2  
**G** and level-1 **R** matrices.

**Model 1b. Most Liberal Baseline – Saturated Means, Unstructured Variances (Model Answer Key)**

```

TITLE1 "SAS Model 1b: Saturated Means, Unstructured Variances";
PROC MIXED DATA=work.example23 NOCLPRINT NOITPRINT COVTEST NAMELEN=100 METHOD=REML;
  CLASS ID session;
  MODEL nm3rt = session / SOLUTION DDFM=Satterthwaite;
  REPEATED session / R RCORR TYPE=UN SUBJECT=ID;
  LSMEANS session /; RUN;

```

Placing *session* on the CLASS/BY statements and in the FIXED/MODEL statements treats it as a categorical predictor. So this is an ANOVA means model. No RANDOM statements mean no random effects.

```

TITLE "SPSS Model 1b: Saturated Means, Unstructured Variances".
MIXED nm3rt BY ID session
  /METHOD = REML
  /PRINT = SOLUTION TESTCOV R
  /FIXED = session
  /REPEATED = session | SUBJECT(ID) COVTYPE(UN)
  /EMMEANS = TABLES(session).

```

```

* STATA Model 1b: Saturated Means, Unstructured Variances
xtmixed nm3rt ib(last).session, || id: , noconstant ///
  variance reml residuals(unstructured, t(session)),
  estat ic, n(101),
  contrast session,          // omnibus test of mean differences
  margins i.session,         // observed means per session
  marginsplot name(observed_means, replace) // plot observed means

```

i. indicates categorical predictor of *session* (ref=last to match others) noconstant = no random intercept (just **R** matrix)

**STATA output:**

```

Mixed-effects REML regression
Group variable: id

```

Number of obs	=	606
Number of groups	=	101
Obs per group: min	=	6
avg	=	6.0
max	=	6
Wald chi2(5)	=	83.60
Prob > chi2	=	0.0000

Log restricted-likelihood = -4114.8942

This is the multivariate Wald test for all the fixed effects simultaneously (5 mean differences from the fixed intercept here).

nm3rt	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
session					
1	289.7574	32.69997	8.86	0.000	225.6666 353.8481
2	143.0364	26.20308	5.46	0.000	91.67927 194.3935
3	77.89864	22.8842	3.40	0.001	33.04642 122.7509
4	45.66045	20.78533	2.20	0.028	4.921952 86.39894
5	35.03972	18.11681	1.93	0.053	-.468579 70.54802
_cons	1672.136	44.13439	37.89	0.000	1585.634 1758.638

Mean diffs relative to session 6

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]
id:	(empty)		
Residual: Unstructured			
var(e1)	301983.1	42696.65	228893.4 398411.6
var(e2)	259148.8	36635.7	196433.4 341887.4
var(e3)	233366.9	32990.48	176891.6 307872.9
var(e4)	217542.8	30753.82	164896.4 286997.6
var(e5)	212096.8	29984.63	160767.3 279814.7
var(e6)	196732.3	27812.21	149121.6 259543.9
cov(e1,e2)	235657.1	36563.79	163993.4 307320.8

These are the total variances at each occasion...

```

cov(e1,e3) | 217992.5 34336.3 150694.6 285290.4
cov(e1,e4) | 202605.4 32657.69 138597.5 266613.2
cov(e1,e5) | 192152.4 31762.13 129899.7 254405
cov(e1,e6) | 195358.7 31224.07 134160.6 256556.7
cov(e2,e3) | 230215.6 33672.52 164218.7 296212.5
cov(e2,e4) | 213230.6 31899.27 150709.2 275752
cov(e2,e5) | 202091 30938.65 141452.3 262729.6
cov(e2,e6) | 193267.2 29707.66 135041.2 251493.1
cov(e3,e4) | 205208 30462.92 145501.8 264914.2
cov(e3,e5) | 196917.7 29697.89 138710.9 255124.5
cov(e3,e6) | 188603.5 28532.39 132681 244525.9
cov(e4,e5) | 193674.7 28910.48 137011.2 250338.2
cov(e4,e6) | 185320 27762.64 130906.2 239733.8
cov(e5,e6) | 187839.5 27739.82 133470.5 242208.6

```

And these are  
the total  
covariances  
across  
occasions...

LR test vs. linear regression:      chi2(20) = 925.64      Prob > chi2 = 0.0000

Note: The reported degrees of freedom assumes the null hypothesis is not on the boundary of the parameter space. If this is not true, then the reported test is conservative.

This is the LRT of whether the  
unstructured R model fits better  
than the e-only R model...

```
.      estat ic, n(101),
```

```

-----+-----
Model |    Obs    ll(null)    ll(model)    df           AIC           BIC
-----+-----
. |    101           .    -4114.894    27       8283.788       8354.397
-----+-----

```

Note: N=101 used in calculating BIC

```
.      contrast session,    // omnibus test of mean differences
```

Contrasts of marginal linear predictions

Margins      : asbalanced

```

-----+-----
                 |           df           chi2           P>chi2
-----+-----
nm3rt           |
session |           5           83.60           0.0000
-----+-----

```

This is the omnibus test of mean  
differences across 6 sessions.

```
.      margins i.session,    // observed means per session
```

Adjusted predictions                                  Number of obs    =           606

Expression    : Linear prediction, fixed portion, predict()

**THESE ARE THE SATURATED MEANS THE FIXED EFFECTS WILL BE TRYING TO REPRODUCE.**

```

-----+-----
                 |           Delta-method
                 |    Margin    Std. Err.           z    P>|z|           [95% Conf. Interval]
-----+-----
session |
1 |    1961.893    54.68027       35.88    0.000       1854.722       2069.065
2 |    1815.172    50.65402       35.83    0.000       1715.892       1914.452
3 |    1750.035    48.06832       36.41    0.000       1655.822       1844.247
4 |    1717.796    46.41001       37.01    0.000       1626.835       1808.758
5 |    1707.176    45.82541       37.25    0.000       1617.36       1796.992
6 |    1672.136    44.13439       37.89    0.000       1585.634       1758.638
-----+-----

```

**Extra SAS output not provided by STATA:**

Estimated R Matrix for ID 101

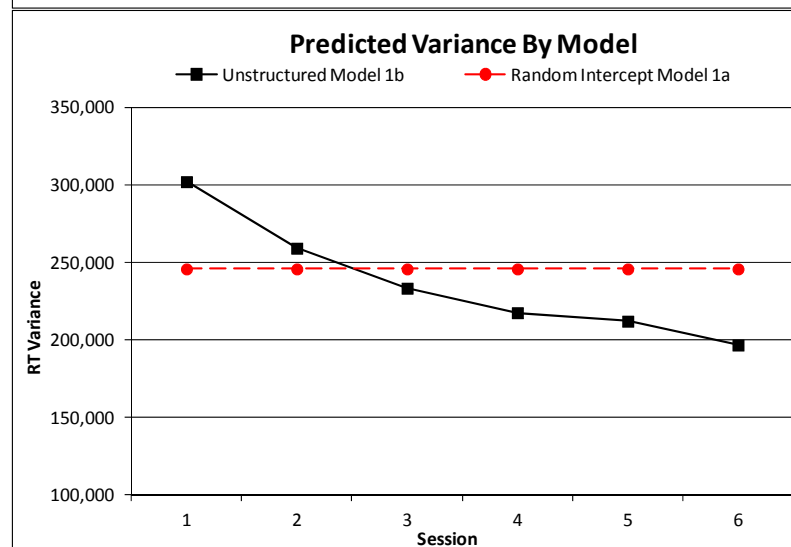
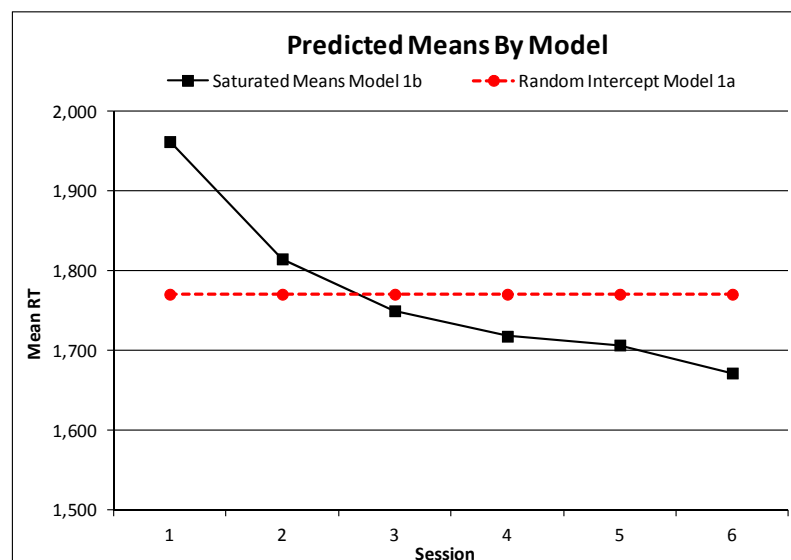
Row	Col1	Col2	Col3	Col4	Col5	Col6
1	<b>301985</b>	235659	217994	202607	192154	195360
2	235659	<b>259150</b>	230217	213232	202092	193268
3	217994	230217	<b>233368</b>	205209	196919	188604
4	202607	213232	205209	<b>217544</b>	193676	185321
5	192154	202092	196919	193676	<b>212098</b>	187840
6	195360	193268	188604	185321	187840	<b>196733</b>

This Unstructured **R matrix** estimates all variances and covariances separately. THIS IS THE DATA we are trying to duplicate with our model for the variances.

Estimated R Correlation Matrix for ID 101

Row	Col1	Col2	Col3	Col4	Col5	Col6
1	1.0000	0.8424	0.8212	0.7905	0.7593	0.8015
2	0.8424	1.0000	0.9361	0.8981	0.8620	0.8559
3	0.8212	0.9361	1.0000	0.9108	0.8851	0.8802
4	0.7905	0.8981	0.9108	1.0000	0.9016	0.8958
5	0.7593	0.8620	0.8851	0.9016	1.0000	0.9196
6	0.8015	0.8559	0.8802	0.8958	0.9196	1.0000

So here is what are we trying to model—means and variances, where model 1b is the data:



**Model 2a. Fixed Linear Time, Random Intercept**

$$\text{Level 1: } y_{ti} = \beta_{0i} + \beta_{1i}(\text{Session}_{ti} - 1) + e_{ti}$$

$$\text{Level 2: Intercept: } \beta_{0i} = \gamma_{00} + U_{0i}$$

$$\text{Linear Session: } \beta_{1i} = \gamma_{10}$$

```
TITLE1 "SAS Model 2a: Fixed Linear Time, Random Intercept";
PROC MIXED DATA=work.example23 NOCLPRINT NOITPRINT COVTEST NAMELEN=100 METHOD=REML;
  CLASS ID session;
  MODEL nm3rt = c1sess / SOLUTION DDFM=Satterthwaite;
  RANDOM INTERCEPT / G V VCORR TYPE=UN SUBJECT=ID;
  REPEATED session / R TYPE=VC SUBJECT=ID; RUN;
```

The predictor of *c1sess* will be treated as continuous given that it is not on the CLASS statement (SAS) and it is on WITH (SPSS).

```
TITLE "SPSS Model 2a: Fixed Linear Time, Random Intercept".
MIXED nm3rt BY ID session WITH c1sess
  /METHOD = REML
  /PRINT = SOLUTION TESTCOV G R
  /FIXED = c1sess
  /RANDOM = INTERCEPT | SUBJECT(ID) COVTYPE(UN)
  /REPEATED = session | SUBJECT(ID) COVTYPE(ID).
```

```
* STATA Model 2a: Fixed Linear Time, Random Intercept
xtmixed nm3rt c.c1sess, || id: , ///
  variance reml covariance(un) residuals(independent,t(session)),
  estat ic, n(101),
  estimates store FixLin
```

DV = nm3rt, c. means continuous fixed slope for *c1sess*  
 Level 2 ID is id, random intercept by default  
 estimates → save results as "FixLin" for next LRT

**STATA output:**

```
Mixed-effects REML regression
Group variable: id
Number of obs      =      606
Number of groups   =      101
Obs per group: min =         6
                  avg =        6.0
                  max =         6
Wald chi2(1)       =     131.82
Prob > chi2        =      0.0000
Log restricted-likelihood = -4207.344
```

nm3rt	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
c1sess	-51.57185	4.491815	-11.48	0.000	-60.37565 -42.76806
_cons	1899.631	46.7882	40.60	0.000	1807.928 1991.334

The fixed linear effect of *c1sess* is significant according to the Wald test (*p*-value for fixed effect).

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]
id: Identity			
var(_cons)	202422.7	29469.85	152172.6 269266.3
var(Residual)	35661.79	2246.481	31519.73 40348.16

Relative to the empty means, random intercept model 1a, the fixed linear effect of session explained ~21% of the residual variance (which made the random intercept variance increase due to its residual variance correction factor).

```
LR test vs. linear regression: chibar2(01) = 787.61 Prob >= chibar2 = 0.0000
```

```
. estat ic, n(101),
-----+-----
Model | Obs   ll(null)   ll(model)   df       AIC       BIC
-----+-----
. | 101      . -4207.344      4   8422.688   8433.149
```

Note: N=101 used in calculating BIC

**Model 2b. Random Linear Time**

```

TITLE1 "SAS Model 2b: Random Linear Time";
PROC MIXED DATA=work.example23 NOCLPRINT
NOITPRINT COVTEST NAMELEN=100 METHOD=REML;
    CLASS ID session;
    MODEL nm3rt = c1sess / SOLUTION DDFM=Satterthwaite;
    RANDOM INTERCEPT c1sess / G V VCORR TYPE=UN SUBJECT=ID;
    REPEATED session / R TYPE=VC SUBJECT=ID; RUN;

```

$$\text{Level 1: } y_{ti} = \beta_{0i} + \beta_{1i} (\text{Session}_{ti} - 1) + e_{ti}$$

$$\text{Level 2: Intercept: } \beta_{0i} = \gamma_{00} + U_{0i}$$

$$\text{Linear Session: } \beta_{1i} = \gamma_{10} + U_{1i}$$

```

TITLE "SPSS Model 2b: Random Linear Time".
MIXED nm3rt BY ID session WITH c1sess
    /METHOD = REML
    /PRINT = SOLUTION TESTCOV G R
    /FIXED = c1sess
    /RANDOM = INTERCEPT c1sess | SUBJECT(ID) COVTYPE(UN)
    /REPEATED = session | SUBJECT(ID) COVTYPE(ID).

```

Now there are 2 random effects: intercept and linear slope, given by c1sess on the RANDOM statements.

```

* STATA Model 2b: Random Linear Time
xtmixed nm3rt c.c1sess, || id: c1sess, ///
    variance reml covariance(un) residuals(independent,t(session)),
    estat ic, n(101),
    estat recovariance, level(id),
    estimates store RandLin,
    lrtest RandLin FixLin

```

DV = nm3rt, c. means continuous fixed slope for c1sess  
 Level 2 ID is id, random intercept and c1sess now  
 estimates → save results as "RandLin" for LRT

**STATA output:**

```

Mixed-effects REML regression
Group variable: id
Number of obs      =      606
Number of groups   =      101
Obs per group: min =        6
                  avg =       6.0
                  max =        6

```

```

Log restricted-likelihood = -4186.0512
Wald chi2(1)           =      70.17
Prob > chi2            =      0.0000

```

nm3rt	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
c1sess	-51.57185	6.156722	-8.38	0.000	-63.63881 -39.5049
_cons	1899.631	51.4998	36.89	0.000	1798.693 2000.569

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]
id: Unstructured			
var(c1sess)	2233.833	552.9239	1375.178 3628.626
var(_cons)	253258	37897.26	188881.9 339575.3
cov(c1sess,_cons)	-12700.79	3621.977	-19799.74 -5601.848
var(Residual)	27905.42	1963.419	24310.74 32031.62

```

LR test vs. linear regression:      chi2(3) = 830.20 Prob > chi2 = 0.0000
Note: LR test is conservative and provided only for reference.

```

```

.      estat ic, n(101),
-----+-----
Model | Obs   ll(null)   ll(model)   df       AIC       BIC
-----+-----
. | 101      . -4186.051      6   8384.102   8399.793

```

Note: N=101 used in calculating BIC

```
.      estat recovariance, level(id),

Random-effects covariance matrix for level id
      |      c1sess      _cons
-----+-----
c1sess | 2233.833
_cons  | -12700.79    253258

.      estimates store RandLin,
.      lrtest RandLin FixLin
```

**Is the random linear time model (2b) better than the fixed linear time, random intercept model (2a)?**

Yep,  $-2\Delta LL = 43$ , which is bigger than the critical value of 5.99ish on  $df \approx 2$ ish

```
Likelihood-ratio test      LR chi2(2) =      42.59
(Assumption: FixLin nested in RandLin)  Prob > chi2 =      0.0000
```

Note: The reported degrees of freedom assumes the null hypothesis is not on the boundary of the parameter space. If this is not true, then the reported test is conservative. Note: LR tests based on REML are valid only when the fixed-effects specification is identical for both models.

### Extra SAS output not provided by STATA:

```
Estimated R Matrix for ID 101
Row      Col1      Col2      Col3      Col4      Col5      Col6
1      27905
2
3
4
5
6
```

```
Estimated G Matrix
Participant
Row      Effect      ID      Col1      Col2
1      Intercept      101      253258      -12701
2      C1sess      101      -12701      2233.83
```

```
Estimated V Matrix for ID 101
Row      Col1      Col2      Col3      Col4      Col5      Col6
1      281163      240557      227856      215155      202455      189754
2      240557      257995      219623      209156      198689      188222
3      227856      219623      239295      203157      194924      186691
4      215155      209156      203157      225063      191158      185159
5      202455      198689      194924      191158      215298      183627
6      189754      188222      186691      185159      183627      210001
```

The **V** matrix is the total variance-covariance matrix after combining the level-2 **G** and level-1 **R** matrices. Now the variances and covariances are predicted to change based on time.

```
Estimated V Correlation Matrix for ID 101
Row      Col1      Col2      Col3      Col4      Col5      Col6
1      1.0000      0.8932      0.8784      0.8553      0.8229      0.7809
2      0.8932      1.0000      0.8839      0.8680      0.8430      0.8086
3      0.8784      0.8839      1.0000      0.8754      0.8588      0.8328
4      0.8553      0.8680      0.8754      1.0000      0.8684      0.8517
5      0.8229      0.8430      0.8588      0.8684      1.0000      0.8636
6      0.7809      0.8086      0.8328      0.8517      0.8636      1.0000
```

The **VCORR** matrix is the correlation version. The ICC is now predicted to change over time, too (and conditional on linear time).

### How the V matrix variances and covariances get calculated in a random linear time model:

$$\mathbf{V}_i \text{ matrix: Variance}[y_{\text{time}}] = \tau_{U_0}^2 + \left[ (\text{Session} - 1)^2 \tau_{U_1}^2 \right] + \left[ 2(\text{Session} - 1) \tau_{U_{01}} \right] + \sigma_e^2$$

$$\mathbf{V}_i \text{ matrix: Covariance}[y_A, y_B] = \tau_{U_0}^2 + \left[ (A + B) \tau_{U_{01}} \right] + \left[ (AB) \tau_{U_1}^2 \right]$$



**Model 3a. Fixed Quadratic, Random Linear Time**

Level 1:  $y_{ti} = \beta_{0i} + \beta_{1i}(\text{Session}_{ti} - 1) + \beta_{2i}(\text{Session}_{ti} - 1)^2 + e_{ti}$

Level 2: Intercept:  $\beta_{0i} = \gamma_{00} + U_{0i}$

Linear Session:  $\beta_{1i} = \gamma_{10} + U_{1i}$

Quadratic Session:  $\beta_{2i} = \gamma_{20}$

```
TITLE1 "SAS Model 3a: Fixed Quadratic, Random Linear Time";
PROC MIXED DATA=work.example23 NOCLPRINT NOITPRINT COVTEST NAMELEN=100 METHOD=REML;
  CLASS ID session;
  MODEL nm3rt = c1sess c1sess*c1sess / SOLUTION DDFM=Satterthwaite;
  RANDOM INTERCEPT c1sess / G V V CORR TYPE=UN SUBJECT=ID;
  REPEATED session / R TYPE=VC SUBJECT=ID; RUN;
```

```
TITLE "SPSS Model 3a: Fixed Quadratic, Random Linear Time".
```

```
MIXED nm3rt BY ID session WITH c1sess
  /METHOD = REML
  /PRINT = SOLUTION TESTCOV G R
  /FIXED = c1sess c1sess*c1sess
  /RANDOM = INTERCEPT c1sess | SUBJECT(ID) COVTYPE(UN)
  /REPEATED = session | SUBJECT(ID) COVTYPE(ID).
```

Interactions can be defined on the fly in SAS and SPSS using \*, or in STATA using # (but only for fixed effects in STATA).

```
* STATA Model 3a: Fixed Quadratic, Random Linear Time
xtmixed nm3rt c.c1sess c.c1sess#c.c1sess, || id: c1sess, ///
  variance reml covariance(un) residuals(independent,t(session)),
  estat ic, n(101),
  estat recovariance, level(id),
  estimates store FixQuad
```

**STATA output:**

```
Mixed-effects REML regression
Group variable: id
Number of obs      =      606
Number of groups   =      101
Obs per group: min =         6
                  avg =        6.0
                  max =         6
Wald chi2(2)       =      97.86
Prob > chi2        =      0.0000
Log restricted-likelihood = -4170.7386
```

nm3rt	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
c1sess	-120.8999	14.54147	-8.31	0.000	-149.4007	-92.39917
c.c1sess#c.c1sess	13.86561	2.634761	5.26	0.000	8.701578	19.02965
_cons	1945.85	52.2433	37.25	0.000	1843.455	2048.245

The fixed quadratic effect of c1sess is significant according to the Wald test ( $p$ -value for fixed effect).

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]	
id: Unstructured				
var(c1sess)	2332.667	551.5799	1467.501	3707.891
var(_cons)	254164	37895.62	189758.3	340429.7
cov(c1sess,_cons)	-12947.88	3620.697	-20044.31	-5851.442
var(Residual)	26175.83	1844.008	22800.05	30051.42

Relative to the random linear time model 2b, the fixed quadratic effect of session explained another ~6% of the residual variance (which made the random intercept variance increase due to its residual variance correction factor).

```
LR test vs. linear regression:      chi2(3) =    851.78    Prob > chi2 = 0.0000
```

Note: LR test is conservative and provided only for reference.

```
.      estat ic, n(101),
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	101	.	-4170.739	7	8355.477	8373.783

Note: N=101 used in calculating BIC

### Model 3b. Random Quadratic Time (and an example of ESTIMATE/TEST/MARGINS statements)

$$\text{Level 1: } y_{ti} = \beta_{0i} + \beta_{1i}(\text{Session}_{ti} - 1) + \beta_{2i}(\text{Session}_{ti} - 1)^2 + e_{ti}$$

$$\text{Level 2: Intercept: } \beta_{0i} = \gamma_{00} + U_{0i}$$

$$\text{Linear Session: } \beta_{1i} = \gamma_{10} + U_{1i}$$

$$\text{Quadratic Session: } \beta_{2i} = \gamma_{20} + U_{2i}$$

```
TITLE1 "SAS Model 3b: Random Quadratic Time";
PROC MIXED DATA=work.example23 NOCLPRINT NOITPRINT COVTEST NAMELEN=100 METHOD=REML;
  CLASS ID session;
  MODEL nm3rt = clsess clsess*clsess / SOLUTION DDFM=Satterthwaite;
  RANDOM INTERCEPT clsess clsess*clsess / G V V CORR TYPE=UN SUBJECT=ID;
  REPEATED session / R TYPE=VC SUBJECT=ID;
  ESTIMATE "Intercept at Session 1"    intercept 1 clsess 0    clsess*clsess 0;
  ESTIMATE "Intercept at Session 2"    intercept 1 clsess 1    clsess*clsess 1;
  ESTIMATE "Intercept at Session 3"    intercept 1 clsess 2    clsess*clsess 4;
  ESTIMATE "Intercept at Session 4"    intercept 1 clsess 3    clsess*clsess 9;
  ESTIMATE "Intercept at Session 5"    intercept 1 clsess 4    clsess*clsess 16;
  ESTIMATE "Intercept at Session 6"    intercept 1 clsess 5    clsess*clsess 25;
  * Predicting linear rate of change at each session (linear changes by 2*quad);
  ESTIMATE "Linear Slope at Session 1"  clsess 1    clsess*clsess 0;
  ESTIMATE "Linear Slope at Session 2"  clsess 1    clsess*clsess 2;
  ESTIMATE "Linear Slope at Session 3"  clsess 1    clsess*clsess 4;
  ESTIMATE "Linear Slope at Session 4"  clsess 1    clsess*clsess 6;
  ESTIMATE "Linear Slope at Session 5"  clsess 1    clsess*clsess 8;
  ESTIMATE "Linear Slope at Session 6"  clsess 1    clsess*clsess 10; RUN;

TITLE "SPSS Model 3b: Random Quadratic Time".
MIXED nm3rt BY ID session WITH clsess
  /METHOD = REML
  /PRINT  = SOLUTION TESTCOV G R
  /FIXED  = clsess clsess*clsess
  /RANDOM  = INTERCEPT clsess clsess*clsess | SUBJECT(ID) COVTYPE(UN)
  /REPEATED = session | SUBJECT(ID) COVTYPE(ID)
  /TEST   = "Intercept at Session 1"    intercept 1 clsess 0    clsess*clsess 0
  /TEST   = "Intercept at Session 2"    intercept 1 clsess 1    clsess*clsess 1
  /TEST   = "Intercept at Session 3"    intercept 1 clsess 2    clsess*clsess 4
  /TEST   = "Intercept at Session 4"    intercept 1 clsess 3    clsess*clsess 9
  /TEST   = "Intercept at Session 5"    intercept 1 clsess 4    clsess*clsess 16
  /TEST   = "Intercept at Session 6"    intercept 1 clsess 5    clsess*clsess 25
  /TEST   = "Linear Slope at Session 1"  clsess 1    clsess*clsess 0
  /TEST   = "Linear Slope at Session 2"  clsess 1    clsess*clsess 2
  /TEST   = "Linear Slope at Session 3"  clsess 1    clsess*clsess 4
  /TEST   = "Linear Slope at Session 4"  clsess 1    clsess*clsess 6
  /TEST   = "Linear Slope at Session 5"  clsess 1    clsess*clsess 8
  /TEST   = "Linear Slope at Session 6"  clsess 1    clsess*clsess 10.
```

Because twice the quadratic slope is how the linear slope changes per unit time, the value for *clsess* used in estimating the linear slope per session gets multiplied by 2.

**\* STATA Model 3b: Random Quadratic Time**

```

xtmixed nm3rt c.c1sess c.c1sess#c.c1sess, || id: c1sess c1sess2, ///
    variance reml covariance(un) residuals(independent,t(session)),
    estat ic, n(101),
    estat recovariance, level(id),
    estimates store RandQuad,
    lrtest RandQuad FixQuad,
    margins, at(c.c1sess=(0(1)5)) vsquish // intercepts per session
    marginsplot, name(predicted_means, replace) // plot intercepts
    margins, at(c.c1sess=(0(1)5)) dydx(c.c1sess) vsquish // linear slope per session
    marginsplot, name(change_in_linear_slope, replace) // plot quadratic effect

```

The random statement will not accept interaction terms, so we are using the c1sess2 created manually before.

**STATA output:**

```

Mixed-effects REML regression
Group variable: id
Number of obs      =      606
Number of groups   =      101
Obs per group: min =         6
                  avg =        6.0
                  max =         6
Wald chi2(2)       =      71.74
Prob > chi2        =      0.0000
Log restricted-likelihood = -4151.3728

```

	nm3rt	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
c1sess		-120.8999	20.04752	-6.03	0.000	-160.1923 -81.6075
c.c1sess#c.c1sess		13.86561	3.41541	4.06	0.000	7.171534 20.55969
_cons		1945.85	53.84993	36.13	0.000	1840.306 2051.394

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]
id: Unstructured			
var(c1sess)	25839.79	5864.685	16561.42 40316.29
var(c1sess2)	634.4659	172.375	372.5198 1080.605
var(_cons)	276207.8	41445.59	205831.2 370647.1
cov(c1sess,c1sess2)	-3903.291	982.6248	-5829.2 -1977.381
cov(c1sess,_cons)	-35734.05	11947.96	-59151.62 -12316.48
cov(c1sess2,_cons)	3901.974	1950.304	79.44722 7724.5
var(Residual)	20298.19	1649.117	17310.19 23801.96

```

LR test vs. linear regression:      chi2(6) = 890.51  Prob > chi2 = 0.0000
Note: LR test is conservative and provided only for reference.

```

```

.      estat ic, n(101),

```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	101	.	-4151.373	10	8322.746	8348.897

Note: N=101 used in calculating BIC

```

.      estimates store RandQuad,

```

```

.      lrtest RandQuad FixQuad,

```

```

Likelihood-ratio test
(Assumption: FixQuad nested in RandQuad)

```

**Is the random quadratic model (3b) better than the fixed quadratic, random linear model (3a)?**

Yep,  $-2\Delta LL = 39$ , which is bigger than the critical value of 7.82ish on  $df \sim 3$ ish

```

LR chi2(3) = 38.73
Prob > chi2 = 0.0000

```

Note: The reported degrees of freedom assumes the null hypothesis is not on the boundary of the parameter space. If this is not true, then the reported test is conservative. Note: LR tests based on REML are valid only when the fixed-effects specification is identical for both models.

```
.      margins, at(c1sess=(0(1)5))      vsquish      // intercepts per session
Adjusted predictions      Number of obs      =      606
Expression : Linear prediction, fixed portion, predict()
1._at      : c1sess      =      0
2._at      : c1sess      =      1
3._at      : c1sess      =      2
4._at      : c1sess      =      3
5._at      : c1sess      =      4
6._at      : c1sess      =      5
```

These are the quadratic-model-predicted means per session.

		Delta-method					
		Margin	Std. Err.	z	P> z	[95% Conf. Interval]	
_at							
1		1945.85	53.84993	36.13	0.000	1840.306	2051.394
2		1838.815	48.48658	37.92	0.000	1743.784	1933.847
3		1759.512	46.99744	37.44	0.000	1667.399	1851.626
4		1707.941	45.89598	37.21	0.000	1617.986	1797.895
5		1684.1	44.23964	38.07	0.000	1597.392	1770.808
6		1687.991	44.20394	38.19	0.000	1601.352	1774.629

```
.      marginsplot, name(predicted_means, replace)      // plot intercepts
Variables that uniquely identify margins: c1sess
.      margins, at(c1sess=(0(1)5)) dydx(c1sess) vsquish      // linear slope per session
Conditional marginal effects      Number of obs      =      606
```

```
Expression : Linear prediction, fixed portion, predict()
dy/dx w.r.t. : c1sess
```

```
1._at      : c1sess      =      0
2._at      : c1sess      =      1
3._at      : c1sess      =      2
4._at      : c1sess      =      3
5._at      : c1sess      =      4
6._at      : c1sess      =      5
```

These are the instantaneous linear slopes at each session. Note how the SEs narrow towards the middle of the data.

		Delta-method					
		dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]	
c1sess	_at						
	1	-120.8999	20.04752	-6.03	0.000	-160.1923	-81.6075
	2	-93.1687	13.64968	-6.83	0.000	-119.9216	-66.4158
	3	-65.43747	8.002796	-8.18	0.000	-81.12266	-49.75228
	4	-37.70624	5.92417	-6.36	0.000	-49.3174	-26.09508
	5	-9.975015	9.973315	-1.00	0.317	-29.52235	9.572324
	6	17.75621	16.03616	1.11	0.268	-13.67408	49.18651

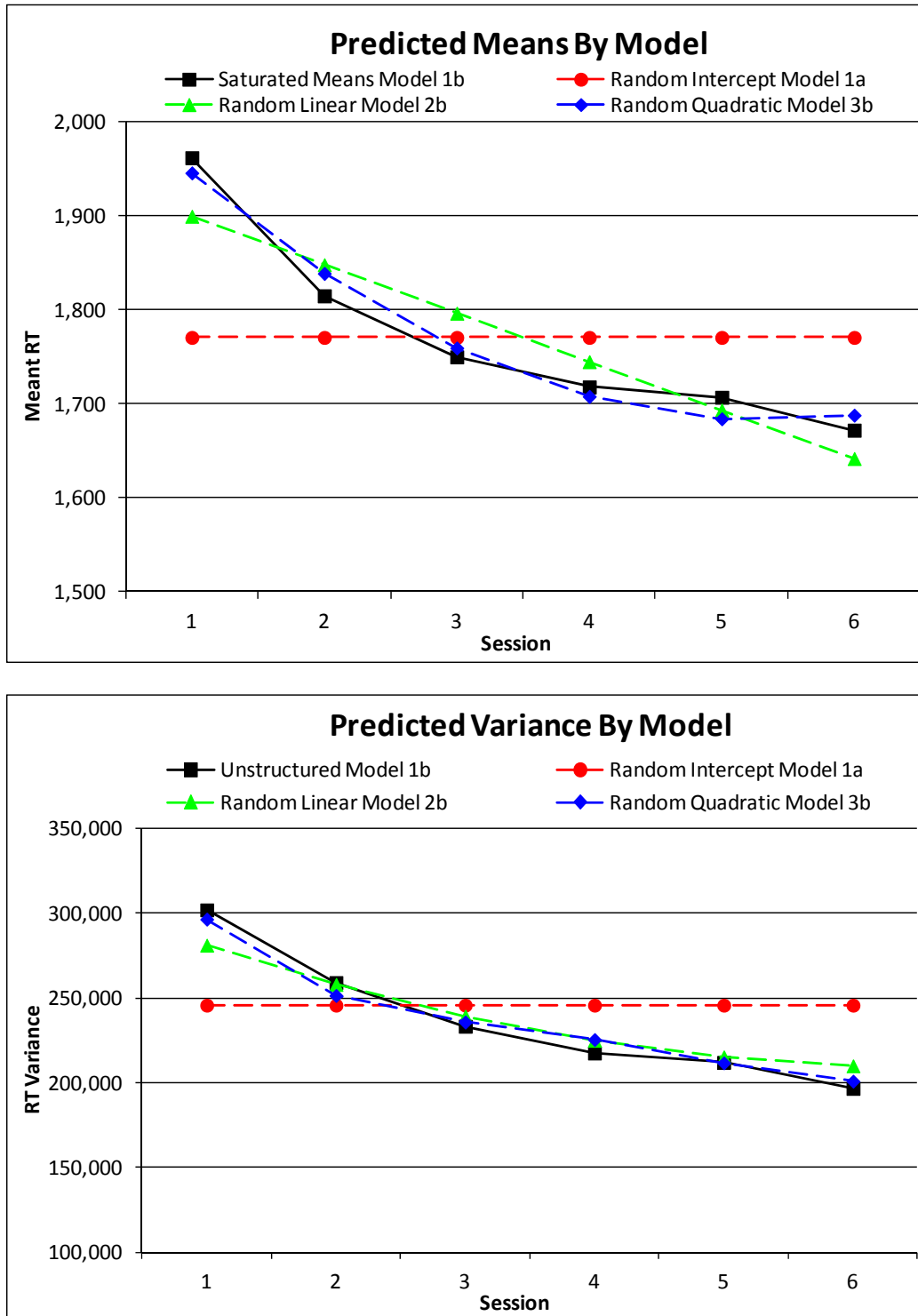
**How well do the predicted means, variances, and covariances from the random quadratic model (3b) match the original means, variances, and covariances from the saturated means model (1b)?**

**Extra SAS output not provided by STATA:**

Estimated V Matrix for ID 101						
Row	Col1	Col2	Col3	Col4	Col5	Col6
1	<b>296504</b>	244374	220346	204122	195702	195085
2	244374	<b>251508</b>	219312	208680	199315	191215
3	220346	219312	<b>235842</b>	209043	199808	187840
4	204122	208680	209043	<b>225508</b>	197182	184958
5	195702	199315	199808	197182	<b>211735</b>	182571
6	195085	191215	187840	184958	182571	<b>200977</b>

The **V** matrix is the total variance-covariance matrix after combining the level-2 **G** and level-1 **R** matrices. The variances and covariances are predicted to change based on time, but differently.

Figure 1



**How the V matrix variances and covariances get calculated in a random quadratic time model:**

Predicted Variance at Time  $T$ :

$$\text{Var}(y_T) = \sigma_e^2 + \tau_{U_0}^2 + 2*T*\tau_{U_{01}} + T^2*\tau_{U_1}^2 + 2*T^2*\tau_{U_{02}} + 2*T^3*\tau_{U_{12}} + T^4*\tau_{U_2}^2$$

Predicted Covariance between Time  $A$  and  $B$ :

$$\text{Cov}(y_A, y_B) = \tau_{U_0}^2 + (A+B)*\tau_{U_{01}} + (AB)*\tau_{U_1}^2 + (A^2+B^2)*\tau_{U_{02}} + (AB^2)+(A^2B)*\tau_{U_{12}} + (A^2B^2)*\tau_{U_2}^2$$

## Simple Processing Speed – Example Unconditional Models of Change Results

### Model Specification

Linear mixed models were estimated using restricted maximum likelihood (REML) in order to examine the overall pattern of and individual differences in response time over six sessions for a simple processing speed test (number match three). The significance of new fixed effects were evaluated using Wald tests, whereas the significance of new random effects was evaluated using likelihood ratio tests (i.e.,  $-2\Delta LL$ ), with degrees of freedom equal to the number of new random effects variances and covariances. The 95% confidence interval (CI) for random variation around each fixed effect was calculated as  $\pm 1.96$  standard deviations of its accompanying random variance term.

Although the six sessions were held over a period of 6–10 days, given that experience to the test (and not *time* per se) was the most likely reason for changes in response time, session was used as the metric of time (i.e., as opposed to age or day). Session was centered at the first occasion, such that the intercept represented initial status in all models. Observed mean response times (in milliseconds) estimated from a saturated means model (i.e., multivariate analysis of variance) are shown in Figure 1. The intraclass correlation from the unconditional means model (i.e., empty model; random intercept only) was calculated as .82, indicating that over 80% of the variance in number match 3 across sessions occurred between persons in mean RT. Polynomial models were then estimated to approximate the effects of practice across the six sessions, as presented below.

### Polynomial Models

Polynomial models were first specified with a random intercept only. A fixed linear effect of session was significant ( $p < .001$ ), such that average response time declined across sessions. The addition of a random linear slope (as well as a covariance between the random intercept and random linear slope) resulted in a significant improvement to the model,  $-2\Delta LL(2) = 43$ ,  $p < .001$ . However, the magnitude of this linear decline was reduced in later sessions, as indicated by a significant fixed quadratic effect of session (i.e., a decelerating negative trend;  $p < .001$ ). The addition of a random quadratic slope (and its two accompanying covariances with the random intercept and random linear slope) also resulted in a significant improvement in model fit,  $-2\Delta LL(3) = 39$ ,  $p < .001$ .

The predicted means from the unconditional random quadratic polynomial model for session (i.e., without predictors) are shown in Figure 1, and model parameters using REML estimation are given in Table 1. As shown, the mean predicted response time at session 1 was 1946 ms, with a 95% CI of 916 to 2976 ms. The mean instantaneous linear rate of change at session 1 was  $-121$  ms per session, with a 95% CI of  $-436$  to  $194$  ms, indicating that not all participants were predicted to improve as evaluated at session 1. Half the mean deceleration in linear rate of change was  $14$  ms per session, such that the linear rate of change became less negative by  $28$  ms with each session. The 95% CI for the quadratic effect was of  $-36$  to  $63$  ms, indicating that not all participants were predicted to decelerate in their rate of improvement across sessions.

### Computing random effects confidence intervals for each random effect:

$$\text{Random Effect 95\% CI} = \text{fixed effect} \pm \left(1.96 * \sqrt{\text{Random Variance}}\right)$$

$$\text{Intercept 95\% CI} = \gamma_{00} \pm \left(1.96 * \sqrt{\tau_{U_0}^2}\right) \rightarrow 1,945.9 \pm \left(1.96 * \sqrt{276,209}\right) = 916 \text{ to } 2,976$$

$$\text{Linear Time Slope 95\% CI} = \gamma_{10} \pm \left(1.96 * \sqrt{\tau_{U_1}^2}\right) \rightarrow -120.9 \pm \left(1.96 * \sqrt{25,840}\right) = -436 \text{ to } 194$$

$$\text{Quadratic Time Slope 95\% CI} = \gamma_{20} \pm \left(1.96 * \sqrt{\tau_{U_2}^2}\right) \rightarrow 13.9 \pm \left(1.96 * \sqrt{634}\right) = -36 \text{ to } 63$$