

# Interactions among Continuous Predictors

- Today's Class:
  - Simple main effects within two-way interactions
  - Conquering TEST/ESTIMATE/LINCOM statements
  - Regions of significance
  - Three-way interactions (and beyond...)

# Representing the Effects of Predictors

- From now on, we will think carefully about exactly **how** the **predictor variables** are entered into the **model for the means** (i.e., by which a predicted outcome is created for each person)
- Why don't people always care? Because the scale of predictors:
  - Does NOT affect the amount of outcome variance accounted for ( $R^2$ )
  - Does NOT affect the outcomes values predicted by the model for the means (so long as the same predictor fixed effects are included)
- Why should this matter to us?
  - ***Because the Intercept = expected outcome value when  $X = 0$***
  - Can end up with nonsense values for intercept if  $X = 0$  isn't in the data, so we need to change the scale of the predictors to include 0
  - Scaling becomes more important **once interactions are included** or once random intercepts are included (i.e., variability around fixed intercept)

# Adjusting the Scale of Predictors

- For **continuous** (quantitative) predictors, we will make the intercept interpretable by **centering**:
  - **Centering** = subtract a constant from each variable value so that **the 0 value** falls within the range of the new centered predictor variable
  - Typical → Center around predictor's mean: Centered  $X_1 = X_1 - \bar{X}_1$ 
    - Intercept is then expected outcome for "average  $X_1$  person"
  - Better → Center around meaningful constant  $C$ : Centered  $X_1 = X_1 - C$ 
    - Intercept is then expected outcome for person with that constant (even 0 may be ok)
- For **binary** predictors, it can be more convenient to treat them as "continuous" than as "categorical" by doing manual coding:
  - Accomplished via "dummy coding" (aka, "reference group coding")
    - Two-group example using *Gender*: 0 = Men, 1 = Women  
(or 0 = Women, 1 = Men)

# Interactions: $y_i = \beta_0 + \beta_1 X_i + \beta_2 Z_i + \beta_3 X_i Z_i + e_i$

- **Interaction = Moderation:** the effect of a predictor depends on the value of the interacting predictor
  - Either predictor can be “the moderator” (interpretive distinction only)
- Interactions can always be evaluated for any combination of categorical and continuous predictors, although traditionally...
  - In “ANOVA”: By default, all possible interactions are estimated
    - Software does this for you; oddly enough, nonsignificant interactions usually still are kept in the model (even if only significant interactions are interpreted)
  - In “ANCOVA”: Continuous predictors (“covariates”) do not get to be part of interaction terms → “homogeneity of regression assumption”
    - There is no reason to assume this – it is a testable hypothesis!
  - In “Regression”: No default – the effects of predictors are as you specify
    - Requires most thought, but gets annoying because in regression programs you usually have to manually create the interaction as an observed variable:
    - e.g.,  $XZ_{interaction} = centeredX * centeredZ$

Interaction variables are created on the fly in MIXED instead! 😊

# Main Effects of Predictors within Interactions

- Main effects of predictors within interactions should remain in the model regardless of whether or not they are significant
  - An interaction is an over-additive (enhancing) or under-additive (dampening) effect, so *what it is additive to* must be included
- **The role of a two-way interaction is to adjust its main effects...**
- However, the idea of a “main effect” no longer applies... each main effect is ***conditional*** on the interacting predictor = 0
- e.g., Model of  $Y = W, X, Z, X*Z$ :
  - The effect of W is still a “main effect” because it is not part of an interaction
  - The effect of X is now the conditional main effect of X *specifically when Z=0*
  - The effect of Z is now the conditional main effect of Z *specifically when X=0*
- Note that this is a different type of conditionality than just “holding the other predictors constant” (which means constant at any value)
  - Constant at the 0 value of the interacting predictor(s), specifically

# Interactions: Why 0 Matters

- Y = Student achievement (GPA as percentage out of 100)  
X = Parent attitudes about education (measured on 1-5 scale)  
Z = Father's education level (measured in years of education)
- $\text{GPA}_i = \beta_0 + (\beta_1 * \text{Att}_i) + (\beta_2 * \text{Ed}_i) + (\beta_3 * \text{Att}_i * \text{Ed}_i) + e_i$   
 $\text{GPA}_i = 30 + (1 * \text{Att}_i) + (2 * \text{Ed}_i) + (0.5 * \text{Att}_i * \text{Ed}_i) + e_i$
- Interpret  $\beta_0$ :
- Interpret  $\beta_1$ :
- Interpret  $\beta_2$ :
- Interpret  $\beta_3$ : **Attitude** as Moderator:  
**Education** as Moderator:
- **Predicted GPA** for **attitude of 3** and **Ed of 12**?  
 $75 = 30 + 1*(3) + 2*(12) + 0.5*(3)*(12)$

# Model-Implied Simple Main Effects

- **Original:**  $GPA_i = \beta_0 + (\beta_1 * Att_i) + (\beta_2 * Ed_i) + (\beta_3 * Att_i * Ed_i) + e_i$   
 $GPA_i = 30 + (1 * Att_i) + (2 * Ed_i) + (0.5 * Att_i * Ed_i) + e_i$
- Given any values of the predictor variables, the model equation provides predictions for:
  - Value of outcome (model-implied intercept for non-zero predictor values)
  - Any conditional (simple) main effects implied by an interaction term
  - **Simple (Conditional) Main Effect = what it is + what *modifies* it**
- Step 1: **Identify** all terms in model involving the predictor of interest
  - e.g., Effect of Attitudes comes from:  $\beta_1 * Att_i + \beta_3 * Att_i * Ed_i$
- Step 2: **Factor out** common predictor variable
  - Start with  $[\beta_1 * Att_i + \beta_3 * Att_i * Ed_i] \rightarrow [Att_i (\beta_1 + \beta_3 * Ed_i)] \rightarrow Att_i$  (new  $\beta_1$ )
  - Value given by ( ) is then the model-implied coefficient for the predictor
- Step 3: **Calculate** model-implied simple effect and SE
  - Let's try it for a **new reference point of attitude = 3 and education = 12**

# Model-Implied Simple Main Effects

- Old Equation using uncentered predictors:

$$\text{GPA}_i = \beta_0 + (\beta_1 * \text{Att}_i) + (\beta_2 * \text{Ed}_i) + (\beta_3 * \text{Att}_i * \text{Ed}_i) + e_i$$

$$\text{GPA}_i = 30 + (1 * \text{Att}_i) + (2 * \text{Ed}_i) + (0.5 * \text{Att}_i * \text{Ed}_i) + e_i$$

- New equation using centered predictors ( $\text{Att}_i - 3$  and  $\text{Ed}_i - 12$ ):

$$\text{GPA}_i = \_ + \_ * (\text{Att}_i - 3) + \_ * (\text{Ed}_i - 12) + \_ * (\text{Att}_i - 3) * (\text{Ed}_i - 12) + e_i$$

- **Intercept: expected value of GPA when  $\text{Att}_i = 3$  and  $\text{Ed}_i = 12$**

$$\beta_0 = 75$$

- **Simple main effect of Att if  $\text{Ed}_i = 12$**

$$\beta_1 * \text{Att}_i + \beta_3 * \text{Att}_i * \text{Ed}_i \rightarrow \text{Att}_i (\beta_1 + \beta_3 * \text{Ed}_i) \rightarrow \text{Att}_i (1 + 0.5 * 12)$$

- **Simple main effect of Ed if  $\text{Att}_i = 3$**

$$\beta_2 * \text{Ed}_i + \beta_3 * \text{Att}_i * \text{Ed}_i \rightarrow \text{Ed}_i (\beta_2 + \beta_3 * \text{Att}_i) \rightarrow \text{Ed}_i (2 + 0.5 * 3)$$

- **Two-way interaction of Att and Ed:**

$$(0.5 * \text{Att}_i * \text{Ed}_i)$$

# Significance of Model-Implied Fixed Effects

- We now know how to calculate simple (conditional) main effects:  
**Effect of interest = what it is + what *modifies* it**  
e.g., **Effect of Attitudes =  $\beta_1 + \beta_3 * Ed$**
- But if we want to test whether that new effect is  $\neq 0$ , we also need its **standard error (SE)** to get Wald test *t*- or *z*-value  $\rightarrow$  *p*-value)
- Even if the simple (conditional) main effect is not *directly* given by the model, its estimate and SE are still *implied* by the model
- **3 options** to get the new simple (conditional) main effect estimate and SE (in order of least to most annoying):
  1. **Ask the software to give it to you** using your original model (e.g., ESTIMATE in SAS, TEST in SPSS, LINCOM in STATA, NEW in Mplus... most programs will do this if you know how to ask)

# Significance of Model-Implied Fixed Effects

2. **Re-center your predictors** to the interacting value of interest (e.g., make attitudes=3 the new 0 for attitudes) and **re-estimate** your model; repeat as needed for each value of interest
3. **Hand calculations** (what the program does for you in option #1)

For example: **Effect of Attitudes** =  $\beta_1 + \beta_3 * Ed$

- $SE^2$  = sampling variance of estimate  $\rightarrow$  e.g.,  $Var(\beta_1) = SE_{\beta_1}^2$
- $SE_{\beta_1}^2 = Var(\beta_1) + Var(\beta_3) * Ed + 2Cov(\beta_1, \beta_3) * Ed$ 
  - Values come from "asymptotic (sampling) covariance matrix" (COVB)
  - Variance of a sum of terms always includes 2\*covariance among them
  - Here, this is because what each main effect estimate could be is related to what the other main effect estimates could be
  - Note that if a main effect is unconditional, its  $SE^2 = Var(\beta)$  only

# 1. Model-Implied Simple Main Effects

- Old Equation using uncentered predictors:  
$$\text{GPA}_i = \beta_0 + (\beta_1 * \text{Att}_i) + (\beta_2 * \text{Ed}_i) + (\beta_3 * \text{Att}_i * \text{Ed}_i) + e_i$$
$$\text{GPA}_i = 30 + (1 * \text{Att}_i) + (2 * \text{Ed}_i) + (0.5 * \text{Att}_i * \text{Ed}_i) + e_i$$
- Intercept: predicted GPA if  $\text{Att}_i=3$  and  $\text{Ed}_i=12$  ?
- Simple main effect of Att if  $\text{Ed}_i=12$  ?  $\text{Att}_i(\beta_1 + \beta_3 * \text{Ed}_i)$
- Simple main effect of Ed if  $\text{Att}_i=3$  ?  $\text{Ed}_i(\beta_2 + \beta_3 * \text{Att}_i)$

```
ECHO 'Requesting Model-Implied Fixed Effects From Previous Slide'.
MIXED y WITH att ed /METHOD = REML /PRINT = SOLUTION TESTCOV
/FIXED = att ed att*ed
/TEST = "Pred GPA if Att=3 Ed=12" intercept 1 att 3 ed 12 att*ed 36
/TEST = "Effect of Att if Ed=12" att 1 att*ed 12
/TEST = "Effect of Ed if Att=3" ed 1 att*ed 3.
```

In TEST statements, the variables refer to their betas; the numbers refer to the operations of their betas.

These estimates would be given directly by the fixed effects instead if you re-centered the predictors as: Att-3, Ed-12.

# Requesting Model-Implied Fixed Effects

- **To request predicted outcomes (= intercepts):**
  - Need to start with "intercept 1" (for  $\beta_0$ )
  - ALL model effects must be included or else are held = 0
  - Note: predictors on CLASS/BY statements must be given a value (more on this next week)

- **For example: regression after centering both predictors**

$$\text{GPA}_i = \beta_0 + (\beta_1 * \text{Att}_i - 3) + (\beta_2 * \text{Ed}_i - 12) + (\beta_3 * \text{Att}_i - 3 * \text{Ed}_i - 12) + e_i$$

```
"GPA if Att=5 Ed=16" intercept 1 att ___ ed ___ att*ed ___
```

```
"GPA if Att=1 Ed=12" intercept 1 att ___ ed ___ att*ed ___
```

```
"GPA if Att=3 Ed=20" intercept 1 att ___ ed ___ att*ed ___
```

# Requesting Model-Implied Fixed Effects

- To request predicted slopes (= simple main effects):
  - **DO NOT** start with "intercept 1" ( $\beta_0$  does **not** contribute to slopes)
  - **NOT ALL** model effects must be included (**only** what modifies the slope)
  - Note: predictors on CLASS/BY statements must be given a value if they modify the slope in an interaction (more on this next week)
- For example: regression after centering both predictors

$$\text{GPA}_i = \beta_0 + (\beta_1 * \text{Att}_i - 3) + (\beta_2 * \text{Ed}_i - 12) + (\beta_3 * \text{Att}_i - 3 * \text{Ed}_i - 12) + e_i$$

"Att Slope if Ed=10" intercept 0 att \_\_\_ ed \_\_\_ att\*ed \_\_\_

"Att Slope if Ed=18" intercept 0 att \_\_\_ ed \_\_\_ att\*ed \_\_\_

"Ed Slope if Att=2" intercept 0 att \_\_\_ ed \_\_\_ att\*ed \_\_\_

"Ed Slope if Att=5" intercept 0 att \_\_\_ ed \_\_\_ att\*ed \_\_\_

# Regions of Significance for Main Effects

- For continuous predictors, there may not be specific values of the moderator at which you want to know significance...

- For example, age\*woman (in which 0=man, 1=woman):

$$y_i = \beta_0 + (\beta_1 * \text{Age}_i - 85) + (\beta_2 * \text{Woman}_i) + (\beta_3 * \text{Age}_i - 85 * \text{Woman}_i) + e_i$$

- Age slopes are only possible for two specific values of woman:

"Age Slope for Men"      age85 \_\_\_ woman \_\_\_ age85\*woman \_\_\_

"Age Slope for Women"      age85 \_\_\_ woman \_\_\_ age85\*woman \_\_\_

- But there are many ages to request gender differences for...

"Gender Diff at Age=80"      age85 \_\_\_ woman \_\_\_ age85\*woman \_\_\_

"Gender Diff at Age=90"      age85 \_\_\_ woman \_\_\_ age85\*woman \_\_\_

# Regions of Significance for Main Effects

- An alternative approach for continuous moderators is known as **regions of significance** (see Hoffman 2014 ch 2. for refs)
- Rather than asking if the simple main effect of gender is still significant at a particular age, we can find the **boundary ages** at which the gender effect becomes non-significant
- We know that:  $EST / SE = t\text{-value} \rightarrow$  if  $|t| > |1.96|$ , then  $p < .05$
- So we work backwards to find the EST and SE such that:

$$\pm t = \pm 1.96 = \frac{\text{Slope Estimate}}{\sqrt{\text{Variance of Slope Estimate}}}, \text{ where:}$$

$$\text{Gender Slope (Gender Difference) Estimate} = \beta_2 + \beta_3 (\text{Age} - 85)$$

$$\text{Variance of Slope Estimate} = \text{Var}(\beta_2) + 2\text{Cov}(\beta_2\beta_3)(\text{Age} - 85) + \text{Var}(\beta_3)(\text{Age} - 85)^2$$


- Need to request "asymptotic covariance matrix" (COVB)
  - Covariance matrix of fixed effect estimates (SE<sup>2</sup> on diagonal)

# Regions of Significance for Main Effects

$$\pm t = \pm 1.96 = \frac{\text{Slope Estimate}}{\sqrt{\text{Variance of Slope Estimate}}}, \text{ where:}$$

$$\text{Gender Slope (Gender Difference) Estimate} = \beta_2 + \beta_3 (\text{Age} - 85)$$

$$\text{Variance of Slope Estimate} = \text{Var}(\beta_2) + 2\text{Cov}(\beta_2\beta_3)(\text{Age} - 85) + \text{Var}(\beta_3)(\text{Age} - 85)^2$$

- For example, age\*woman (0=man, 1=woman), age = moderator:  
 $y_i = \beta_0 + (\beta_1 * \text{Age}_i - 85) + (\beta_2 * \text{Woman}_i) + (\beta_3 * \text{Age}_i - 85 * \text{Woman}_i) + e_i$
- $\beta_2 = -0.5306^*$  at age=85,  $\text{Var}(\beta_2) \rightarrow \text{SE}^2$  for  $\beta_2 = 0.06008$
- $\beta_3 = -0.1104^*$  unconditional,  $\text{Var}(\beta_3) \rightarrow \text{SE}^2$  for  $\beta_3 = 0.00178$
- Covariance of  $\beta_2$  SE and  $\beta_3$  SE = 0.00111
- **Regions of Significance for Moderator of Age = 60.16 to 79.52**
  - The gender effect  $\beta_2$  is predicted to be significantly negative above age 79.52, non-significant from ages 79.52 to 60.16, and significantly positive below age 60.16 (because non-parallel lines will cross eventually).

# More Generally...

- Can decompose a **2-way interaction** by testing the simple effect of X at different levels of Z (and vice-versa)
  - Use TESTs to request simple effects at any point of the interacting predictor
  - Regions of significance are useful for continuous interacting predictors
- More general rules of interpretation, given a **3-way interaction**:
  - *Simple (main) effects move the intercept*
    - 1 possible interpretation for each simple main effect
    - Each simple main effect is conditional on other two variables = 0
  - *The 2-way interactions (3 of them in a 3-way model) move the simple effects*
    - 2 possible interpretations for each 2-way interaction
    - Each simple 2-way interaction is conditional on third variable = 0
  - *The 3-way interaction moves each of the 2-way interactions*
    - 3 possible interpretations of the 3-way interaction
    - Is highest-order term in model, so is unconditional (applies always)

# Practice with 3-Way Interactions

- Intercept = 5, Effect of X = 1.0, Effect of Z = 0.50, Effect of W = 0.20
- **X\*Z = .10 (applies specifically when W is 0)**
  - #1: for every 1-unit  $\Delta X$ ,
  - #2: for every 1-unit  $\Delta Z$ ,
- **X\*W = .01 (applies specifically when Z is 0)**
  - #1: for every 1-unit  $\Delta X$ ,
  - #2: for every 1-unit  $\Delta W$ ,
- **Z\*W = .05 (applies specifically when X is 0)**
  - #1: for every 1-unit  $\Delta Z$ ,
  - #2: for every 1-unit  $\Delta W$ ,
- **X\*Z\*W = .001 (unconditional because is highest order)**
  - #1: for every 1-unit  $\Delta X$ ,
  - #2: for every 1-unit  $\Delta Z$ ,
  - #3: for every 1-unit  $\Delta W$ ,

# Practice with 3-Way Interactions

- Model:  $y_i = \beta_0 + \beta_1 X_i + \beta_2 Z_i + \beta_3 W_i + \beta_4 X_i W_i + \beta_5 X_i Z_i + \beta_6 Z_i W_i + \beta_7 X_i Z_i W_i + e_i$
- Formula to get simple main effects:
  - Simple effect of X =
  - Simple effect of Z =
  - Simple effect of W =
- Formula to get simple 2-way interactions:
  - Simple X\*Z =
  - Simple X\*W =
  - Simple Z\*W =

# Interpreting Interactions: Summary

- Interactions represent “moderation” – the idea that the effect of one variable depends upon the level of other(s)
- The main effects WILL CHANGE in once an interaction with them is added, because they now mean different things:
  - Main effect → Simple effect specifically when interacting predictor = 0
  - Need to have 0 as a meaningful value for each predictor for that reason
- Rules for interpreting conditional (simple) fixed effects:
  - Intercepts are conditional on (i.e., get adjusted by) main effects
  - Main effects are conditional on two-ways (become ‘simple main effects’)
  - Two-ways are conditional on three-ways... And so forth
  - Highest-order term is unconditional – same regardless of centering

# Creating Predicted Outcomes

- Figures of predicted outcomes will be essential in describing interaction terms (especially in talks and posters)
- Three ways to get them (in order of most to least painful):
  1. In excel: input fixed effects, input variable values, write equation to create predicted outcomes for each row
    - Good for pedagogy, but gets old quickly (and error-prone)
  2. Via programming statements:
    - Per prediction: Use SAS ESTIMATE or SPSS TEST
    - For a range of predictor values: Use STATA MARGINS
  3. Via “fake people” (most useful in SPSS and SAS)
    - Add cases to your data with desired predictor values (no outcomes)
    - Ask program to save predicted outcomes for all cases (FIXPRED in SPSS)
    - Fake cases won't contribute to model, but will get predicted outcomes