Two-Level Clustered Data Example: Students within Schools COMPLETED VERSION

These are real data taken from the results of a math test given at the end of 10^{th} grade in a Midwestern Rectangular State. These data include 13,802 students from 94 schools, with 31–515 students in each school (M = 275). We will examine how student gender (0=boy, 1=girl) and student free and reduced lunch status (0=pay for lunch, 1= receive free or reduced lunch) predict math test scores.

SAS Code for Data Manipulation:

```
* Importing data into work library;
%LET example = F:\Example Data\School Data;
LIBNAME example "&example.";
DATA grade10; SET example.grade10;
LABEL studentID=
                    "studentID: Student ID number"
                   "schoolID: School ID number"
      schoolID=
      boyvsgirl= "boyvsgirl: Boy=0, Girl=1"
                   "frlunch: 0=No, 1=Free/Reduced Lunch"
      frlunch=
                    "math: Math Test Score Outcome";
      math=
      KEEP studentID--math;
       * Selecting cases that are complete for analysis variables;
      IF NMISS(studentID, schoolID, boyvsgirl, frlunch, math)>0 THEN DELETE; RUN;
* Getting school means to use as predictors;
PROC SORT DATA=grade10; BY schoolID studentID; RUN;
PROC MEANS NOPRINT N DATA=grade10;
      BY schoolID;
      VAR boyvsgirl frlunch math;
      OUTPUT OUT=SchoolMeans
             MEAN(boyvsgirl frlunch math) = SMboyvsgirl SMfrlunch SMmath; RUN;
* Labeling new school mean variables;
DATA SchoolMeans; SET SchoolMeans;
      SchoolN = _FREQ_; * Saving N per school;
      DROP _TYPE _FREQ ; * Dropping unneeded SAS-created variables;
LABEL SMboyvsgirl= "SMboyvsgirl: School Mean Boy=0, Girl=1"
      SMfrlunch: "SMfrlunch: School Mean 0=No, 1=Free/Reduced Lunch"
                 "SMmath: School Mean Math Outcome"
"SchoolN: # Students Contributing Data"; RUN;
      SMmath=
      SchoolN=
* Merging school means back with individual data;
DATA grade10; MERGE grade10 SchoolMeans; BY schoolID;
       * Selecting only schools with data from at least 30 students;
      IF SchoolN < 31 THEN DELETE; RUN;
* Outputting table of descriptives to rtf document;
ODS RTF FILE="&example.\Descriptive Stats.rtf";
TITLE "Getting means to center predictors with";
PROC MEANS N MEAN STDDEV MIN MAX DATA=grade10;
      VAR math boyvsgirl frlunch SMmath SMboyvsgirl SMfrlunch SchoolN;
RUN; TITLE; ODS RTF CLOSE;
```

Variable	Label	N	Mean	Std Dev	Minimum	Maximum
math	math: Math Test Score Outcome	13082	48.1185599	17.2590473	0	83.0000000
boyvsgirl	boyvsgirl: Boy=0, Girl=1	13082	0.4981654	0.5000157	0	1.0000000
frlunch	frlunch: 0=No, 1=Free/Reduced Lunch	13082	0.3075218	0.4614850	0	1.0000000
SMmath	SMmath: School Mean Math Outcome	13082	48.1185599	6.8181301	29.4509804	61.6136364
SMboyvsgirl	SMboyvsgirl: School Mean Boy=0, Girl=1	13082	0.4981654	0.0422383	0.3333333	0.6842105
SMfrlunch	SMfrlunch: School Mean 0=No, 1=Free/Reduced Lunch	13082	0.3075218	0.2220852	0	0.8032787
SchoolN	SchoolN: # Students Contributing Data	13082	274.9501605	155.3319041	31.0000000	515.0000000

* Centering school mean predictors;

```
DATA grade10; SET grade10;
```

```
SMboyvsgirl50 = SMboyvsgirl - .50; LABEL SMboyvsgirl50= "SMboyvsgirl50: 0=.50";
SMfrlunch30 = SMfrlunch - .30; LABEL SMfrlunch30= "SMfrlunch30: 0=.30"; RUN;
```

* Location for .sas file that holds macro programs; %LET macropath=F:\Example Data; %INCLUDE "¯opath.\MLM_Macros.sas";

Model 1a: Single-Level Empty Means, Residual Variance Only Model

Level 1:	$Math_{ij} = \beta_{0j} + e_{ij}$
Level 2:	$\beta_{0j} = \gamma_{00}$

TITLE1 "Model 1a: Single Level Empty Means, Residual Variance Only";
PROC MIXED DATA=example10 NOCLPRINT COVTEST NAMELEN=100 IC METHOD=ML;
CLASS schoolID studentID;
MODEL math = / SOLUTION DDFM=Satterthwaite; RUN;

DimensionsCovariance Parameters1Columns in X1Columns in Z0Subjects1Max Obs Per Subject13082

	Covariance I	Parameter Est	timates		
		Standard	Z		
Cov Parm	Estimate	Error	Value	Pr > Z	
Residual	297.85	3.6828	80.88	<.0001	

Information Criteria									
Parms	AIC	AICC	HQIC	BIC	CAIC				
2	111652	111652	111657	111667	111669				
Solution for Fixed Effects									
	2	Parms AIC 2 111652	Parms AIC AICC 2 111652 111652 Solution for Fixed Effects	Parms AIC AICC HQIC 2 111652 111652 111657 Solution for Fixed Effects	Parms AIC AICC HQIC BIC 2 111652 111652 111657 111667 Solution for Fixed Effects				

		Standard			
Effect	Estimate	Error	DF	t Value	Pr > t
Intercept	48.1186	0.1509	13E3	318.90	<.0001

Model 1b: Empty Means, Random Intercept Model

```
TITLE1 "Model 1b: 2-Level Empty Means, Random Intercept";
PROC MIXED DATA=example10 NOCLPRINT COVTEST NAMELEN=100 IC METHOD=ML;
        CLASS schoolID studentID;
        MODEL math = / SOLUTION DDFM=Satterthwaite;
        RANDOM INTERCEPT / TYPE=UN SUBJECT=schoolID;
        ODS OUTPUT CovParms=CovEmpty InfoCrit=FitEmpty;
RUN;
```

Level 1: Math_{ij} = β_{0j} + e_{ij} Level 2: β_{0j} = γ_{00} + U_{0j}

Dimensions Calculate the ICC (correlation of students in same Covariance Parameters 2 school in math): ICC = 45 / (45 + 253) = .15Columns in X 1 Columns in Z Per Subject 1 **Calculate the design effect:** = 1 + [(n-1) * ICC]Subjects 94 Max Obs Per Subject 515 DE = 1 + [(275-1)*.15] = 42.1Null Model Likelihood Ratio Test **Calculate the effective N:** $N_{effective} = (\#Total Obs) / DE$ DF Chi-Square Pr > ChiSq 13,082 / 42.1 = **311!!!** 1 1857.08 <.0001 Information Criteria AICC HQIC Neg2LogLike Parms ATC BIC CATC 109791 З 109797 109797 109800 109805 109808 Coloulate the 05% wandow offects . . .

	Solut	ion for Fix	Calculate the 95% random effects			
		CI for the intercept across schools:				
Effect	Estimate	Error	DF	t Value	Pr > t	Fixed effect ± 1.96 *SQRT(variance)
Intercept	47.7561	0.7192	94.9	66.40	<.0001	$48 \pm 1.96 \text{*}\text{SQRT}(45) = 35 \text{ to } 61$

Model 2a: Adding a Fixed Effect of Student Gender

Level 1: $Math_{ij} = \beta_{0j} + \beta_{1j}(BoyvsGirl_{ij}) + e_{ij}$ Level 2: Intercept: $\beta_{0j} = \gamma_{00} + U_{0j}$ Student Gender: $\beta_{1j} = \gamma_{10}$

TITLE1 "Model 2a: Adding a Fixed Effect of Student Gender";
PROC MIXED DATA=work.grade10 NOCLPRINT COVTEST NAMELEN=100 IC METHOD=ML;
CLASS schoolID studentID;
MODEL math = boyvsgirl / SOLUTION DDFM=Satterthwaite;
RANDOM INTERCEPT / TYPE=UN SUBJECT=schoolID;
ODS OUTPUT CovParms=CovGen1 InfoCrit=FitGen1;

```
RUN;
```

Covariance Parameter Estimates									
Standard Z									
Cov Parm	Subject	Estimate	Error	• Value	e Pr > Z				
UN(1,1)	schoolID	44.8203	7.0210	6.38	<.0001				
Residual		253.00	3.1394	80.59	<.0001				
	Information Criteria								
Neg2LogLike	Parms	AIC	AICC	HQIC	BIC	CAIC			
109782	4	109790	109790	109794	109800	109804			
	Solu	tion for Fix	ed Effects	6					
		Standard							
Effect	Estimate	Error	DF	t Value	Pr > t				
Intercept	47.3300	0.7319	102	64.66	<.0001				
boyvsgirl	0.8462	0.2791	13E3	3.03	0.0024				

What does the main effect of student gender represent in this model?

Girls score an average of 0.8462 higher than boys.

What are we assuming about the main effect of student gender?

We are assuming no contextual effect (that the between- and within-school effects of gender are equal).

Proportion reduction in each variance relative to the 2-level empty model:

* Calculate PseudoR2 relative to empty means model; %PseudoR2(NCov=2, CovFewer=CovEmpty, NameFewer=Empty, CovMore=CovGen1, NameMore=GenderL1);							
PsuedoR2 (%	<pre>& Reduction)</pre>	for Empty \	/s. GenderL1				
Name	CovParm	Subject	Estimate	StdErr	ZValue	ProbZ	PseudoR2
Empty	UN(1,1)	schoolID	44.9335	7.0391	6.38	<.0001	
Empty	Residual		253.18	3.1415	80.59	<.0001	
GenderL1	UN(1,1)	schoolID	44.8203	7.0210	6.38	<.0001	.002520558
GenderL1	Residual		253.00	3.1394	80.59	<.0001	.000690265

Why were both variances reduced when gender is a level-1 predictor?

The smushed effect of gender contains part of the level-2 gender effect, too.

Model 2b: Adding a Fixed Effect of School Proportion of Girls

Level 1: Math_{ij} = $\beta_{0j} + \beta_{1j}$ (BoyVsGirl_{ij}) + e_{ij} Level 2: Intercept: $\beta_{0j} = \gamma_{00} + \gamma_{01}$ (SchoolGender_j) + U_{0j} Student Gender: $\beta_{1j} = \gamma_{10}$

TITLE1 "Model 2b: Adding Fixed Effect of School Proportion Girls";
PROC MIXED DATA=work.grade10 NOCLPRINT COVTEST NAMELEN=100 IC METHOD=ML;
 CLASS schoolID studentID;
 MODEL math = boyvsgirl SMboyvsgirl50 / SOLUTION DDFM=Satterthwaite OUTPM=PredGen2;
 RANDOM INTERCEPT / TYPE=UN SUBJECT=schoolID;
 ODS OUTPUT CovParms=CovGen2 InfoCrit=FitGen2;
 ESTIMATE "Gender Between-School Effect" boyvsgirl 1 SMboyvsgirl50 1;

RUN;

Covariance Parameter Estimates									
			Standard	Z					
Cov Parm	Subject	Estimate	Error	Value	Pr > Z				
UN(1,1)	schoolID	43.4390	6.7859	6.40	<.0001				
Residual		253.00	3.1392	80.59	<.0001				
		Informa	tion Criter	ria					
Neg2LogLike	Parms	AIC	AICC	HQIC	BIC	CAIC			
109779	5	109789	109789	109794	109801	109806			
Solution for Fixed Effects									
		Standa	ırd						
Effect	Estimate	Err	or DF	t Value	Pr > t				
Intercept	47.2605	0.72	29 103	65.37	<.0001				
boyvsgirl	0.8352	0.27	'91 13E3	2.99	0.0028				
SMboyvsgirl5	0 20.8313	11.96	611 103	1.74	0.0846				
Estimates									
			St	andard					
Label		Est	imate	Error	DF tVa	lue Pr	> t		
Gender Betwe	Gender Between-School Effect 21.6665 11.9578 103 1.81 0.0729								

What does the main effect of school mean gender represent in this model?

After controlling for kid gender, there is no contextual (incremental between-school) effect of gender; the additional increment to school mean math scores for a one-unit difference in proportion of students who are girls of 20.83 (the difference between 0 and 100% girls) is nonsignificant. If we don't control for kid gender, the between-school gender effect of 21.67 is still nonsignificant.

Proportion reduction in each variance due to effect of school mean gender:

```
* Calculate PseudoR2 relative to level-1 gender only model;
%PseudoR2(NCov=2, CovFewer=CovGen1, NameFewer=GenderL1, CovMore=CovGen2, NameMore=GenderL2);
PsuedoR2 (% Reduction) for GenderL2 vs. GenderL2
 Name
           CovParm
                       Subject
                                   Estimate
                                                StdErr
                                                          ZValue
                                                                     ProbZ
                                                                              PseudoR2
GenderL1
           UN(1,1)
                       schoolID
                                    44.8203
                                                7.0210
                                                           6.38
                                                                    <.0001
                                                                               .
GenderL1
                                    253.00
                                                3.1394
                                                           80.59
                                                                    <.0001
           Residual
                                                                    <.0001
                                                                              0.030817
GenderL2
                       schoolID
                                    43.4390
                                                6.7859
                                                           6.40
           UN(1,1)
GenderL2
           Residual
                                     253.00
                                                3.1392
                                                           80.59
                                                                    <.0001
                                                                              0.000021
* Calculate Total R2; PROC CORR NOSIMPLE DATA=PredGen2; VAR math pred; RUN;
                                      math
                                                   Pred
math
                                   1.00000
                                                0.03016
                                                             R = .03016, so total R^2 = .009
math: Math Test Score Outcome
                                                 0.0006
```

Model 2c: Adding a Random Effect of Student Gender

Level 1: M	$ath_{ii} = \beta_{0i} + \beta_{0i}$	$-\beta_{1j}(BoyVsO)$	Girl_{ii}) + e_{ii}			
		$=\gamma_{00}+\gamma_{01}(\overline{S})$		\overline{der}_i) + U ₀ ;		
	Gender: β_{1i}			J7 - 0J		
		710 · C1j				
CLAS MODE RAND	DATA=work S schoolII L math = b OM INTERCE	grade10 NO studentID;	CLPRINT CC : fboyvsgirl :l / G TYP	OVTEST NAME 50 / SOLUT E=UN SUBJE	LEN=100 IC ION DDFM=S CT=schoolI	<pre>METHOD=ML; atterthwaite; D;</pre>
	Estima	ted G Matrix				
	Sc	hool				
Davis Effe	ID		0-14	0-10		
Row Effe 1 Inte			Col1 0992 -0	Col2 .7533		
	•			0.6203		
,	U					
	Covar	iance Parame	ter Estimat	es		
			Standard	Z		
Cov Parm	Subject	Estimate	Error	Value	Pr Z	
UN(1,1)	schoolID	44.0992	7.1632	6.16	<.0001	
UN(2,1) UN(2,2)	schoolID schoolID	-0.7533 0.6203	2.1120 0.8957	-0.36 0.69	0.7213 0.2443	
Residual	SCHOOLID	252.84	3.1432	80.44	<.0001	
neoradar		202104	011402	00144	10001	
		Informa	tion Criter	ia		
Neg2LogLike	Parms	AIC	AICC	HQIC	BIC	CAIC
109778	7	109792	109792	109799	109810	109817
	So	lution for F Standa		:s		
Effect	Estima	te Err	or DF	t Value	Pr > t	
Intercept	47.25	21 0.72	89 94.9	64.82	<.0001	
boyvsgirl	0.85	53 0.30	05 55.1	2.85	0.0062	
SMboyvsgirl5	0 20.54	25 11.96	66 103	1.72	0.0890	
* Calculate	a differend	ce in model	fit relat	ive to bot	h fixed ge	el fit is not better. ender effects model; GenRand, NameMore=Ra

%FitTest(FitFewer=FitGen2, NameFewer=FixedGender, FitMore=FitGenRand, NameMore=RandomGender);

Likelihood Ratio Test for FixedGender vs. RandomGender							
	Neg2Log						
Name	Like	Parms	AIC	BIC	DevDiff	DFdiff	Pvalue
FixedGender	109779	5	109789	109801			
RandomGender	109778	7	109792	109810	0.63311	2	0.72865

So what does this mean about the effect of student gender?

The 0.86 advantage for girls is approximately the same across schools.

Given the non-significant improvement in model fit, we will remove the random slope for gender. However, we will continue to control for gender at both levels of the model.

Model 3a: Adding a Fixed Effect of Student Free/Reduced Lunch

Level 1: $Math_{ij} = \beta_{0j} + \beta_{1j}(BoyVsGirl_{ij}) + \beta_{2j}(FRlunch_{ij}) + e_{ij}$ Level 2: Intercept: $\beta_{0j} = \gamma_{00} + \gamma_{01}(SchoolGender_j) + U_{0j}$ Student Gender: $\beta_{1j} = \gamma_{10}$ Free/Reduced Lunch: $\beta_{2j} = \gamma_{20}$

<pre>TITLE1 "Model 3a: Adding Fixed Effect of Student Free/Reduced Lunch"; PROC MIXED DATA=work.gradel0 NOCLPRINT COVTEST NAMELEN=100 IC METHOD=ML; CLASS schoolID studentID; MODEL math = boyvsgirl SMboyvsgirl50 frlunch / SOLUTION DDFM=Satterthwaite; RANDOM INTERCEPT / TYPE=UN SUBJECT=schoolID; ODS OUTPUT CovParms=CovFR1 InfoCrit=FitFR1; RUN;</pre>									
Covariance Parameter Estimates									
			Standard	Z					
Cov Parm	Subject	Estimate	Error	Value	Pr > Z				
UN(1,1)	schoolID	26.1636	4.3105	6.07	<.0001				
Residual		239.14	2.9675	80.59	<.0001				
		Informa	tion Criter	ia					
Neg2LogLike	Parms	AIC	AICC	HQIC	BIC	CAIC			
109003	6	109015	109015	109022	109031	109037			
	Solution for Fixed Effects Standard								
Effect	Estimat	e Err	or DF	t Value	Pr > t				
Intercept	50.132	5 0.58	68 110	85.44	<.0001				
boyvsgirl	0.855	5 0.27	14 13E3	3.15	0.0016				
SMboyvsgirl5	50 15.079	6 9.56	69 104	1.58	0.1180				
frlunch	-9.429	B 0.33	16 13E3	-28.43	<.0001				

What does the main effect of student free/reduced lunch represent in this model?

Kids who get free/reduced lunch score 9.4 points lower than kids who don't.

What are we assuming about the main effect of student free/reduced lunch?

We are assuming no contextual effect (that the between- and within-school effects of FRlunch are equal).

Proportion reduction in each variance relative to Model 2b with both fixed gender effects:

* Calculate PseudoR2 relative to both fixed gender effects model; %PseudoR2(NCov=2, CovFewer=CovGen2, NameFewer=Gender, CovMore=CovFR1, NameMore=FRLunch1);

PsuedoR2 (% Reduction)	for Gender	vs. FRLunch1				
Name	CovParm	Subject	Estimate	StdErr	ZValue	ProbZ	PseudoR2
Gender	UN(1,1)	schoolID	43.4390	6.7859	6.40	<.0001	
Gender	Residual		253.00	3.1392	80.59	<.0001	
FRLunch1	UN(1,1)	schoolID	26.1636	4.3105	6.07	<.0001	0.39769
FRLunch1	Residual		239.14	2.9675	80.59	<.0001	0.05478

Why were both variances reduced when FRlunch is a level-1 predictor?

The smushed effect of free/reduced lunch contains part of the free/reduced lunch level-2 effect, too.

Model 3b: Adding a Fixed Effect of School Proportion Free/Reduced Lunch

Level 1: $\operatorname{Math}_{ij} = \beta_{0j} + \beta_{1j}(\operatorname{BoyVsGirl}_{ij}) + \beta_{2j}(\operatorname{FRlunch}_{ij}) + e_{ij}$							
Level 2: Intercept: $\beta_{0j} = \gamma_{00} + \gamma_{01}(\overline{\text{SchoolGender}}_j) + \gamma_{02}(\overline{\text{SchoolFRLunch}}_j) + U_{0j}$							
Student Gender: $\beta_{1j} = \gamma_{10}$							
Free/Reduced Lunch: $\beta_{2j} = \gamma_{20}$							

<pre>TITLE1 "Model 3b: Adding Fixed Effect of School Proportion Free/Reduced Lunch"; PROC MIXED DATA=work.grade10 NOCLPRINT COVTEST NAMELEN=100 IC METHOD=ML; CLASS schoolID studentID; MODEL math = boyvsgirl SMboyvsgirl50 frlunch SMfrlunch30</pre>										
	Covar	riance Parame	eter Estimat	es						
	oorar		Standard	Z						
Cov Parm	Subject	Estimate	Error	Value	Pr > Z					
UN(1,1)	schoolID	13.3767	2.5214	5.31	<.0001					
Residual		239.21	2.9693	80.56	<.0001					
Information Criteria										
Neg2LogLike	Parms	AIC	AICC	HQIC	BIC	CAIC				
108954	7	108968	108968	108975	108986	108993				
	Sc	olution for H	- ixed Effect	s						
		Standa	ard							
Effect	Estima	ate Eri	ror DF	t Value	Pr > t					
Intercept	50.15	574 0.4	542 113	110.44	<.0001					
boyvsgirl	0.85	649 0.2	714 13E3	3.15	0.0016					
SMboyvsgirl				0.77	0.4448					
frlunch	-9.17	757 0.33	343 13E3	-27.45	<.0001					
SMfrlunch30	-16.59	983 2.0	143 83.8	-8.24	<.0001					
			Estimate	s						
				Standard						
Label	Label Estimate Error DF t Value Pr > t									
FR Lunch Between-School Effect -25.7740 1.9863 79.2 -12.98 <.0001										

What does the main effect of school proportion free/reduced lunch represent in this model?

This is the contextual (incremental between-school) effect for FRlunch: after controlling for kid free/reduced lunch status, for every 10% more kids in their school who receive free/reduced lunch, school mean math is lower by 1.66. If we don't control for kid FRlunch, the between-school effect of FRlunch of -2.5.77 per 10% is still significant.

What does the effect of student free/reduced lunch NOW represent in this model?

This is the pure within-school effect: holding school free/reduced lunch status constant (i.e., within the same school), kids who receive free/reduced lunch score 9.2 points lower than kids who don't.

Proportion reduction in each variance due to effect of school mean FRlunch:

* Calculate PseudoR2 relative to level-1 FRlunch only model; %PseudoR2(NCov=2, CovFewer=CovFR1, NameFewer=FRLunch1, CovMore=CovFR2, NameMore=FRLunch2);

PsuedoR2 (% Reduction) for FRLunch1 vs. FRLunch2

Name	CovParm	Subject	Estimate	StdErr	ZValue	ProbZ	PseudoR2
FRLunch1	UN(1,1)	schoolID	26.1636	4.3105	6.07	<.0001	
FRLunch1	Residual		239.14	2.9675	80.59	<.0001	
FRLunch2	UN(1,1)	schoolID	13.3767	2.5214	5.31	<.0001	0.48873
FRLunch2	Residual		239.21	2.9693	80.56	<.0001	-0.00031

Proportion reduction in each variance due to **both** effects of FRlunch:

* Calculate PseudoR2 for both FRlunch effects relative to both gender effects; %PseudoR2(NCov=2, CovFewer=CovGen2, NameFewer=Gender, CovMore=CovFR2, NameMore=FRLunch2);

PsuedoR2 (% Reduction) for Gender vs. FRLunch2

Name	CovParm	Subject	Estimate	StdErr	ZValue	ProbZ	PseudoR2
Gender	UN(1,1)	schoolID	43.4390	6.7859	6.40	<.0001	
Gender	Residual		253.00	3.1392	80.59	<.0001	
FRLunch2	UN(1,1)	schoolID	13.3767	2.5214	5.31	<.0001	0.69206
FRLunch2	Residual		239.21	2.9693	80.56	<.0001	0.05448

Total reduction in math variance due to both effects of FRlunch:

```
* Calculate Total R2 change relative to both gender effects only model;
%TotalR2(DV=math, PredFewer=PredGen2, NameFewer=Gender, PredMore=PredLunch2,
NameMore=FRlunch2);
```

Total R2 (% Reduction) for Gender vs. FRlunch2 Pred Total Name Corr TotalR2 R2Diff Gender 0.03016 0.00091 . FRlunch2 0.40455 0.16366 0.1627

Model 3c: Adding a Random Effect of Student Free/Reduced Lunch

Level 1: $\operatorname{Math}_{ij} = \beta_{0j} + \beta_{1j}(\operatorname{BoyVsGirl}_{ij}) + \beta_{2j}(\operatorname{FRlunch}_{ij}) + e_{ij}$ Level 2: Intercept: $\beta_{0j} = \gamma_{00} + \gamma_{01}(\operatorname{SchoolGender}_{j}) + \gamma_{02}(\operatorname{SchoolFRLunch}_{j}) + U_{0j}$ Student Gender: $\beta_{1j} = \gamma_{10}$ Free/Reduced Lunch: $\beta_{2j} = \gamma_{20}$ $+ U_{2j}$

```
TITLE1 "Model 3c: Adding Random Effect of Student Free/Reduced Lunch";
PROC MIXED DATA=work.grade10 NOCLPRINT COVTEST NAMELEN=100 IC METHOD=ML;
        CLASS schoolID studentID;
        MODEL math = boyvsgirl SMboyvsgirl50 frlunch SMfrlunch30 / SOLUTION DDFM=Satterthwaite;
        RANDOM INTERCEPT frlunch / G TYPE=UN SUBJECT=schoolID;
        ODS OUTPUT CovParms=CovFRRand InfoCrit=FitFRRand;
RUN;
```

	Est	imated G	Matrix	
		School		
		ID		
Row	Effect	number	Col1	Col2
1	Intercept	125	19.7443	-11.7978
2	frlunch	125	-11.7978	12.7448

Covariance Parameter Estimates									
			Standard	Z					
Cov Parm	Subject	Estimate	Error	Value	Pr Z				
UN(1,1)	schoolID	19.7443	3.7094	5.32	<.0001				
UN(2,1)	schoolID	-11.7978	3.1478	-3.75	0.0002				
UN(2,2)	schoolID	12.7448	3.3242	3.83	<.0001				
Residual		236.63	2.9443	80.37	<.0001				
Information Criteria									
Neg2LogLike	Parms	AIC	AICC	HQIC	BIC	CAIC			
108866	9	108884	108884	108893	108907	108916			
	So	lution for F	ixed Effect	S					
		Standa	ard						
Effect	Estima	te Err	ror DF	t Value	Pr > t				
Intercept	49.79	10 0.53	97.4	93.79	<.0001				
boyvsgirl	0.89	61 0.27	703 13E3	3.31	0.0009				
SMboyvsgirl5	0 4.32	23 7.23	337 104	0.60	0.5515				
frlunch	-8.45	52 0.56	98.2	-15.04	<.0001				
SMfrlunch30	-16.84	99 1.94	41 77.7	-8.67	<.0001				

Does the random slope for student gender help the model? Yes, model fit is better.

* Calculate difference in model fit relative to both fixed FRlunch effects model; %FitTest(FitFewer=FitFR2, NameFewer=FixedFRlunch, FitMore=FitFRRand, NameMore=RandomFRLunch);

Likelihood Ratio Test for FixedFRlunch vs. RandomFRLunch

	Neg2Log						
Name	Like	Parms	AIC	BIC	DevDiff	DFdiff	Pvalue
FixedFRlunch	108954	7	108968	108986			
RandomFRLunch	108866	9	108884	108907	88.2542	2	0

So what does this mean about the effect of student free/reduced lunch?

The difference in math between kids who get free/reduced lunch and kids who don't varies significantly over schools.

Calculate and interpret a 95% random effects confidence interval for the random slope:

 -8.4552 ± 1.96 *SQRT(12.7488) = -15.45 to -1.46On average, the gap in math performance related to kid free/reduced lunch status is 8.4 points, but across

95% of the schools, that gap is predicted to be anywhere from 1.46 to 15.45 points.

Model 3d: Adding a Cross-Level Interaction of Student by School Free/Reduced Lunch

Level 1: $\operatorname{Math}_{ij} = \beta_{0j} + \beta_{1j}(\operatorname{BoyVsGirl}_{ij}) + \beta_{2j}(\operatorname{FRlunch}_{ij}) + e_{ij}$ Level 2: Intercept: $\beta_{0j} = \gamma_{00} + \gamma_{01}(\operatorname{SchoolGender}_{j}) + \gamma_{02}(\operatorname{SchoolFRLunch}_{j}) + U_{0j}$ Student Gender: $\beta_{1j} = \gamma_{10}$ Free/Reduced Lunch: $\beta_{2j} = \gamma_{20}$ $+ \gamma_{22}(\operatorname{SchoolFRLunch}_{j}) + U_{2j}$

Covariance Parameter Estimates									
		Standard		Z					
Subject	Estimate	Error	Value P		Z				
schoolID	19.5909	3.6558	5.3	6 <.00	01				
schoolID	-11.1933	3.0624	-3.6	6 0.00	03				
schoolID	11.8021	3.1679	3.7	3 <.00	01				
	236.62	2.9439	80.3	<.00	01				
	Informa	ation Criter	ia						
Parms	AIC	AICC	HQIC	BIC	CAIC				
10	108883	108883	108894	108909	108919				
	Solution fo	or Fixed Eff	ects						
		Standard							
	Estimate	Error	DF	t Value	Pr > t				
	49.7497	0.5300	99.7	93.87	<.0001				
	0.8996	0.2703	13E3	3.33	0.0009				
SMboyvsgirl50 4.8849 7.2788			103	0.67	0.5037				
	-8.7042	0.5679	112	-15.33	<.0001				
SMfrlunch30 -19.2722 2.4760 89.					<.0001				
lunch30	4.3141	2.6386	103	1.64	0.1051				
	Subject schoolID schoolID schoolID Parms 10	Subject Estimate schoolID 19.5909 schoolID -11.1933 schoolID 11.8021 236.62 Informa Parms AIC 10 108883 Solution fo Estimate 49.7497 0.8996 0 4.8849 -8.7042 -19.2722	Standard Subject Estimate Error schoolID 19.5909 3.6558 schoolID -11.1933 3.0624 schoolID 11.8021 3.1679 236.62 2.9439 Information Criter Parms AIC 10 108883 Solution for Fixed Eff Standard Estimate Error 49.7497 0.5300 0 4.8849 7.2788 -8.7042 0.5679 -19.2722 2.4760	Standard Subject Estimate Error Valu schoolID 19.5909 3.6558 5.3 schoolID -11.1933 3.0624 -3.6 schoolID 11.8021 3.1679 3.7 236.62 2.9439 80.3 Information Criteria Parms AIC AIC HQIC 10 108883 108883 108894 Solution for Fixed Effects Standard Estimate Error DF 49.7497 0.5300 99.7 0.8996 0.2703 13E3 0 4.8849 7.2788 103 -8.7042 0.5679 112 -19.2722 2.4760 89.2 -19.2722 -14760 89.2	Standard Z Subject Estimate Error Value Pr schoolID 19.5909 3.6558 5.36 <.00				

What does the effect of student free/reduced lunch NOW represent in this model?

This is the difference between kids who get free/reduced lunch and those who don't in schools where 30% of the kids get free/reduced lunch: those kids who get free/reduced lunch are lower by 8.70.

What does the effect of school proportion free/reduced lunch NOW represent in this model?

This is the contextual (incremental between-school) effect for a kid who does not receive free/reduced lunch: for those kids, for every 10% more kids in their school that receive free/reduced lunch, their school mean math is lower by 1.9.

What does the cross-level interaction of student by school free/reduced lunch represent?

The effect of being a kid who receives free/reduced lunch is reduced nonsignificantly by 0.4 for every 10% more children in their school who get free/reduced lunch. But this effect is currently smushed—it assumes without testing that school FRlunch moderates the within-school and between-school effects of FRlunch to the same extent.

Proportion reduction in each variance relative to Model 3c with random FRLunch:

```
* Calculate PseudoR2 for both FRlunch effects relative to random FRlunch;
%PseudoR2(NCov=4, CovFewer=CovFRRand, NameFewer=LunchMain, CovMore=CovInt1,
NameMore=LunchInt1);
```

Name	CovParm	Subject	Estimate	StdErr	ZValue	ProbZ	PseudoR2
LunchMain	UN(1,1)	schoolID	19.7443	3.7094	5.32	<.0001	
LunchMain	UN(2,2)	schoolID	12.7448	3.3242	3.83	<.0001	
LunchMain	Residual		236.63	2.9443	80.37	<.0001	
LunchInt1	UN(1,1)	schoolID	19.5909	3.6558	5.36	<.0001	0.007768
LunchInt1	UN(2,2)	schoolID	11.8021	3.1679	3.73	<.0001	0.073968
LunchInt1	Residual		236.62	2.9439	80.37	<.0001	0.000060

Which variance component should be reduced (non-significantly)? Random slope for FRlunch

Model 3e: Adding a Level-2 Interaction of Quadratic School Free/Reduced Lunch

Level 1: M	$\operatorname{Iath}_{ij} = \beta_{0j} + \beta_{0j}$	β_{1j} (BoyVsC	$\operatorname{Birl}_{ij} + \beta_{2j}$	FRlunch _{ij})	+e _{ij}				
Level 2:		- 5)1 (SchoolGe	$\overline{\text{ender}}_{j}$) + γ_{0}	$_{02}(\overline{\text{SchoolF}})$	RLunch _j)+	γ_{03} (SchoolFRI	$\operatorname{Lunch}_{j})^{2} + \mathrm{U}_{0j}$	
Studer	nt Gender: β	$\gamma_{1j} = \gamma_{10}$							
Free/Reduc	ed Lunch: β	$\beta_{2j} = \gamma_{20}$		$+\gamma$	22(SchoolF	RLunch _j)		$+U_{2j}$	
<pre>TITLE1 "Model 3e: Adding Level-2 Interaction of Quadratic School Free/Reduced Lunch"; PROC MIXED DATA=work.grade10 NOCLPRINT COVTEST NAMELEN=100 IC METHOD=ML; CLASS schoolID studentID; MODEL math = boyvsgirl SMboyvsgirl50 frlunch SMfrlunch30 frlunch*SMfrlunch30 SMfrlunch30*SMfrlunch30 / SOLUTION DDFM=Satterthwaite OUTPM=PredInt2; RANDOM INTERCEPT frlunch / G TYPE=UN SUBJECT=schoolID; ODS OUTPUT CovParms=CovInt2 InfoCrit=FitInt2; ESTIMATE "FR Lunch Between-School Main Effect" frlunch 1 SMfrlunch30 1; ESTIMATE "FR Lunch Between-School Interact" frlunch*SMfrlunch30 1 SMfrlunch30*SMfrlunch30 1; RUN; Covariance Parameter Estimates</pre>									
	Covari	iance Paramet	ter Estimate	s					
			Standard	Z					
Cov Parm	Subject	Estimate	Error	Value	Pr Z				
UN(1,1)	schoolID	18.8449	3.5480	5.31	<.0001				
UN(2,1)	schoolID	-10.8805	3.0207	-3.60	0.0003				
UN(2,2)	schoolID	11.8613	3.1882	3.72	<.0001				
Residual		236.61	2.9439	80.37	<.0001				
		Informat	tion Criteri	.a					
Neg2LogLike	Parms	AIC	AICC	HQIC	BIC	CAIC			
108861	11	108883	108883	108894	108911	108922			
		Solution 1	for Fixed Ef Standard						
Effect		Estimate	Error		t Value	Pr > t			
Intercept		50.3622	0.6594	114	76.37	<.0001			
boyvsgirl		0.9016	0.2703	13E3	3.34	0.0009			
SMboyvsgirl5	50	3.0898	7.2988	100	0.42	0.6730			
frlunch		-8.8417	0.5775	5 114	-15.31	<.0001			
SMfrlunch30		-17.9370	2.5918	88.1	-6.92	<.0001			
frlunch*SMfr	lunch30	5.5170	2.7679	108	1.99	0.0488			
SMfrlunch30*	SMfrlunch30	-13.5458	8.9313	86.4	-1.52	0.1330			
			Estimat	es					
				Standa	rd				
Label			Estimate	e Err	or DF	t Value	Pr > t		
FR Lunch Bet	tween-School	Main Effect	-26.7787	2.60	01 90.5	-10.30	<.0001		
FR Lunch Bet	tween-School	Interaction	-8.0288	8.53	65 74.4	-0.94	0.3500		

What does the cross-level interaction of student by school free/reduced lunch NOW represent?

The effect of being a kid who receives free/reduced lunch (now after allowing for differential moderation across levels of the effects of free/reduced lunch at both levels by school mean free/reduced lunch) is reduced significantly by 0.55 for every 10% more children in their school who get free/reduced lunch.

What does the level-2 interaction of quadratic school free/reduced lunch represent?

After controlling for kid free/reduced lunch status, the contextual (incremental between-school) effect of school mean free/reduced lunch as evaluated at 30% FRlunch becomes nonsignificantly more negative by 2*1.3 for every 10% more kids in their school with free/reduced lunch.

If we don't control for kid free/reduced lunch, the between-school effect of -2.68 per 10% of school mean free/reduced lunch as evaluated at 30% FRlunch becomes nonsignificantly more negative by 2*0.80 for every 10% more kids in their school with free/reduced lunch.

So school mean free/reduced lunch moderates the within-school FRlunch effect, but not the contextual (incremental between-school) or between-school effects.

Proportion reduction in each variance due to the quadratic school free/reduced lunch effect:

* Calculate PseudoR2 relative to level-1 lunch interaction only model; %PseudoR2(NCov=4, CovFewer=CovInt1, NameFewer=LunchInt1, CovMore=CovInt2, NameMore=LunchInt2);									
PsuedoR2 (%	Reduction)	for LunchInt1	vs. LunchInt2	2					
Name	CovParm	Subject	Estimate	StdErr	ZValue	ProbZ	PseudoR2		
LunchInt1	UN(1,1)	schoolID	19.5909	3.6558	5.36	<.0001			
LunchInt1	UN(2,2)	schoolID	11.8021	3.1679	3.73	<.0001			
LunchInt1	Residual		236.62	2.9439	80.37	<.0001			
LunchInt2	UN(1,1)	schoolID	18.8449	3.5480	5.31	<.0001	0.038080		
LunchInt2	UN(2,2)	schoolID	11.8613	3.1882	3.72	<.0001	-0.005017		
LunchInt2	Residual		236.61	2.9439	80.37	<.0001	0.000015		

Total reduction in math variance due to both intra-variable interactions of FRlunch:

```
* Calculate Total R2 change relative to both gender effects only model;
%TotalR2(DV=math, PredFewer=PredLunch2, NameFewer=LunchMain, PredMore=PredInt2,
NameMore=LunchInt2);
```

Total R2	(% Reduction)	for LunchMain	vs. LunchInt2
	Pred		Total
Name	Corr	TotalR2	R2Diff
LunchMain	0.40455	0.16366 .	
LunchInt2	0.40589	0.16475 .	001088340

Plot of model-predicted math by free/reduced lunch status to illustrate interactions:

