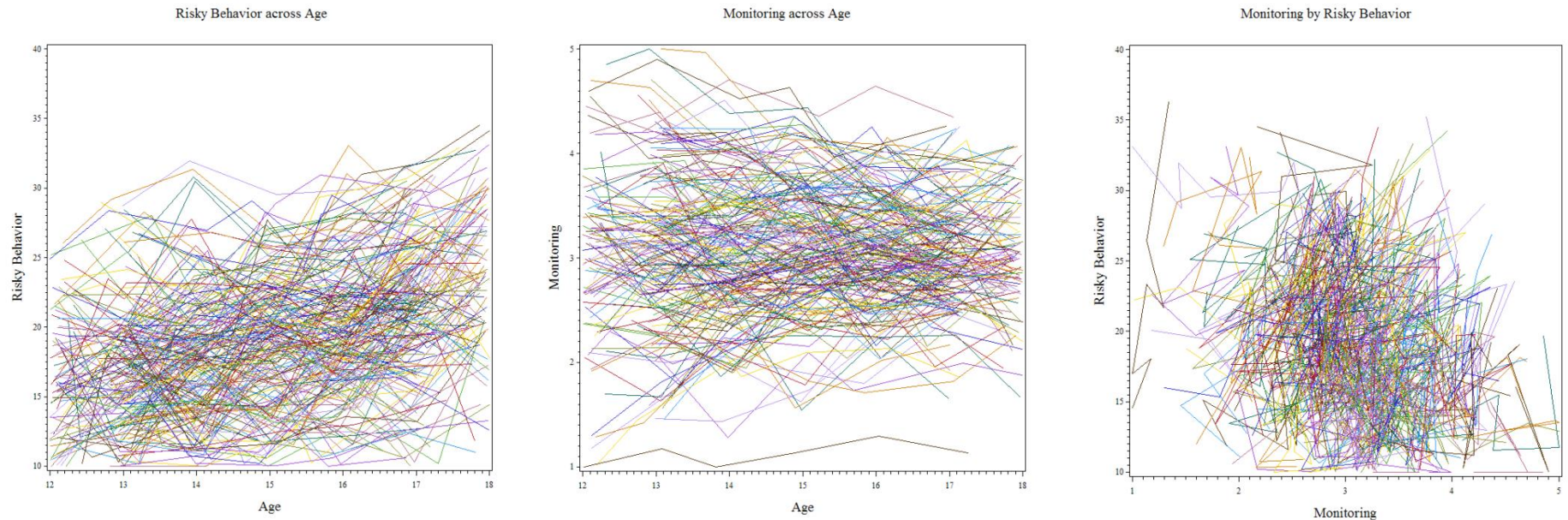


**Example 5b: Multivariate Change and Lagged Effects via Multivariate MLM:
Multilevel SEM (“MLM” below) and Single-Level SEM in Mplus v. 8.1+ (complete syntax and output available electronically)**

These simulated data are from Hoffman (2015) chapter 9 and include 200 girls measured approximately annually from ages 12–18 (time 0 = age 18) on their risky behavior (the outcome, a sum ranging from 10 to 50) and the extent to which their mothers monitored their activities (the time-varying predictor, a mean ranging from 1 to 5, centered at 3). A time-invariant predictor of the conservativeness of mothers’ attitudes about the smoking and drinking (a mean ranging from 1 to 5, centered at 4) was also collected at the age 12 occasion. Here are the individual change trajectories for risky behavior and monitoring:

**Level 1 :****Multivariate Multilevel Model 1**

$$y_{tid} = dvR \left[\beta_{0iR} + \beta_{1iR} (Age_{tiR} - 18) + \beta_{2iR} (Age_{tiR} - 18)^2 + e_{tiR} \right] + dvM \left[\beta_{0iM} + \beta_{1iM} (Age_{tiM} - 18) + e_{tiM} \right]$$

Level 2 :

Risky Intercept: $\beta_{0iR} = \gamma_{00R} + \gamma_{01R} (Attitudes12_i - 4) + U_{0iR}$

Risky Age: $\beta_{1iR} = \gamma_{10R} + \gamma_{11R} (Attitudes12_i - 4) + U_{1iR}$

Risky Age²: $\beta_{2iR} = \gamma_{20R}$

Monitor Intercept: $\beta_{0iM} = \gamma_{00M} + \gamma_{01M} (Attitudes12_i - 4) + U_{0iM}$

Monitor Age: $\beta_{1iM} = \gamma_{10M} + \gamma_{11M} (Attitudes12_i - 4) + U_{1iM}$

The best-fitting unconditional longitudinal models included fixed quadratic and random linear effects of age for risky behavior, but a random linear effect of age for monitoring (although the fixed linear age slope was nonsignificant). In addition, mother’s attitudes significantly predicted the intercept and linear age slope for risky behavior. Although they did not significantly predict monitoring, I have added them here to illustrate how to compute indirect effects.

Chapter 9 began with person-mean-centering and baseline-centering of monitoring as a time-varying predictor of risky behavior. Both were shown to be inadequate because they do not properly distinguish the intercept, linear age slope, and residual variance contained in the monitoring predictor, each of which could potentially relate to those of risky behavior. So the purpose of this example is to demonstrate alternative software methods of estimating models of multivariate change so that you can decide what approach (software and syntax combination) will be most optimal for your own data. See chapter 9 for the results from a directed path model very similar to 2c.

In Mplus, Model 1 as an undirected multivariate MLM via long M-SEM:

```

TITLE: Model 1: Undirected Multivariate Change Model as Long M-SEM
DATA: FILE = Example5b.csv;    ! Syntax in same folder as data
VARIABLE:
! List of variables in data file
  NAMES = PersonID Att12 occasion age risky mon roundage
         time att4 timesq mon3;
! Variables to be analyzed in this model
  USEVARIABLE = time timesq att4 risky mon3;
  MISSING ARE ALL (-999);    ! Missing data identifier
! MLM options
  CLUSTER = PersonID;        ! Level-2 ID
  BETWEEN = att4;            ! Observed ONLY level-2 predictors
  WITHIN = time timesq;      ! Observed ONLY level-1 predictors

ANALYSIS: TYPE = TWOLEVEL RANDOM; ESTIMATOR = ML;

MODEL: ! R = risky behavior, M = monitoring
%WITHIN%
  risky mon3 (Rresvar Mresvar); ! L1 R: WP residual variances (labels)
  Rlin | risky ON time;          ! Placeholder for R linear change
  Rquad | risky ON timesq;       ! Placeholder for R quadratic change
  Mlin | mon3 ON time;           ! Placeholder for M linear change
  risky WITH mon3 (ResCov);      ! L1 R: WP residual covariance

%BETWEEN%
[risky mon3 Rlin Rquad Mlin];   ! Fixed intercepts, fixed change slopes
risky mon3 (Rintvar Mintvar);   ! L2 G: Random intercept variances (labels)
Rlin Mlin (Rlinvar Mlinvar);    ! L2 G: Random linear change variances
Rquad@0;                        ! No quadratic change variance

risky Rlin ON att4;             ! Att-> R int and R linear change
mon3 Mlin ON att4;              ! Att-> M int and M linear change
risky WITH Rlin (RIntLin);      ! L2 G: R Int-change covariance (label)
mon3 WITH Mlin (MIntLin);       ! L2 G: M Int-change covariance (label)

risky WITH mon3 (IntCov);       ! L2 G: Random int-int covariance
Rlin WITH Mlin (LinCov);        ! L2 G: Random change-change covariance
mon3 WITH Rlin (Int2Lin);       ! L2 G: M int, R change covariance
Mlin WITH risky (Lin2Int);      ! L2 G: M change, R int covariance

MODEL CONSTRAINT: ! Linear combinations of any parameter
! First need to name each new combination
NEW(ResCor IntCor LinCor RIScor MIScor I2SCor S2ICor);

! Estimating correlations found in RCORR and GCORR
! Corr = Cov / (SQRT(Yvar)*SQRT(Xvar))
  ResCor = ResCov / (SQRT(Rresvar)*SQRT(Mresvar));
  IntCor = IntCov / (SQRT(Rintvar)*SQRT(Mintvar));
  LinCor = LinCov / (SQRT(Rlinvar)*SQRT(Mlinvar));
  RIScor = RIntLin / (SQRT(Rintvar)*SQRT(Rlinvar));
  MIScor = MIntLin / (SQRT(Mintvar)*SQRT(Mlinvar));
  I2Scor = Int2Lin / (SQRT(Mintvar)*SQRT(Rlinvar));
  S2Icor = Lin2Int / (SQRT(Mlinvar)*SQRT(Rintvar));

```

Number of Free Parameters	22			
Loglikelihood				
H0 Value	-4391.884			
Information Criteria				
Akaike (AIC)	8827.768			
Bayesian (BIC)	8943.141			
Sample-Size Adjusted BIC	8873.255			
(n* = (n + 2) / 24)				
	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Within Level				
RISKY WITH				
MON3	0.287	0.028	10.441	0.000
Residual Variances				
RISKY	8.352	0.374	22.351	0.000
MON3	0.081	0.004	22.354	0.000
Between Level				
RLIN ON				
ATT4	-0.518	0.104	-4.963	0.000
MLIN ON				
ATT4	0.003	0.014	0.240	0.810
RISKY ON				
ATT4	-3.160	0.551	-5.737	0.000
MON3 ON				
ATT4	-0.044	0.057	-0.779	0.436
RISKY WITH				
RLIN	1.878	0.356	5.273	0.000
MLIN	0.040	0.039	1.044	0.296
MON3 WITH				
MLIN	0.000	0.004	-0.105	0.916
RLIN	-0.106	0.031	-3.449	0.001
RLIN WITH				
MLIN	-0.018	0.007	-2.478	0.013
RISKY WITH				
MON3	-0.853	0.168	-5.076	0.000
Means				
RQUAD	0.147	0.021	7.117	0.000
Intercepts				
RISKY	23.322	0.348	67.075	0.000
MON3	0.063	0.034	1.839	0.066
RLIN	1.975	0.138	14.259	0.000
MLIN	-0.003	0.008	-0.380	0.704
Variances				
RQUAD	0.000	0.000	999.000	999.000
Residual Variances				
RISKY	18.049	2.202	8.198	0.000
MON3	0.195	0.023	8.371	0.000
RLIN	0.484	0.080	6.071	0.000
MLIN	0.010	0.001	7.802	0.000
New/Additional Parameters				
RESCOR	0.350	0.028	12.607	0.000
INTCOR	-0.455	0.074	-6.119	0.000
LINCOR	-0.255	0.103	-2.483	0.013
RISCOR	0.635	0.057	11.088	0.000
MISCOR	-0.009	0.089	-0.105	0.917
I2SCOR	-0.346	0.095	-3.646	0.000
S2ICOR	0.093	0.087	1.066	0.286

In Mplus, the same Model 1 as an undirected single-level wide SEM:

```

TITLE: Model 1: Undirected Multivariate Change Model as Wide Single-Level SEM
DATA: FILE = Example5b.csv; ! Syntax in same folder as data
! Unstacking to wide format
DATA LONGTOWIDE:
! Names of old stacked former variables (without numbers)
LONG = risky|mon3|time;
! Names of new wide variables (that use numbers)
WIDE = risky12-risky18|mon12-mon18|age12-age18;
! Variable with level-2 ID info
IDVARIABLE = PersonID;
! Old level-1 identifier
REPETITION = roundage (12 13 14 15 16 17 18);
VARIABLE:
! List of variables in original data file
NAMES = PersonID Att12 occasion age risky mon roundage
time att4 timesq mon3;
! Variables to be analyzed in this model
USEVARIABLE = att4 age12-age18 mon12-mon18 risky12-risky18;
MISSING ARE ALL (-999); ! Missing data identifier
TSCORES = age12-age18; ! Exact time indicator
ANALYSIS: TYPE = RANDOM; ESTIMATOR = ML; MODEL = NOCOVARIANCES;
MODEL: ! R = risky behavior, M = monitoring
[risky12-risky18@0 mon12-mon18@0]; ! All variable intercepts fixed to 0
risky12-risky18 (Rresvar); ! L1 R: R WP residual variances held equal
mon12-mon18 (Mresvar); ! L1 M: M WP residual variances held equal

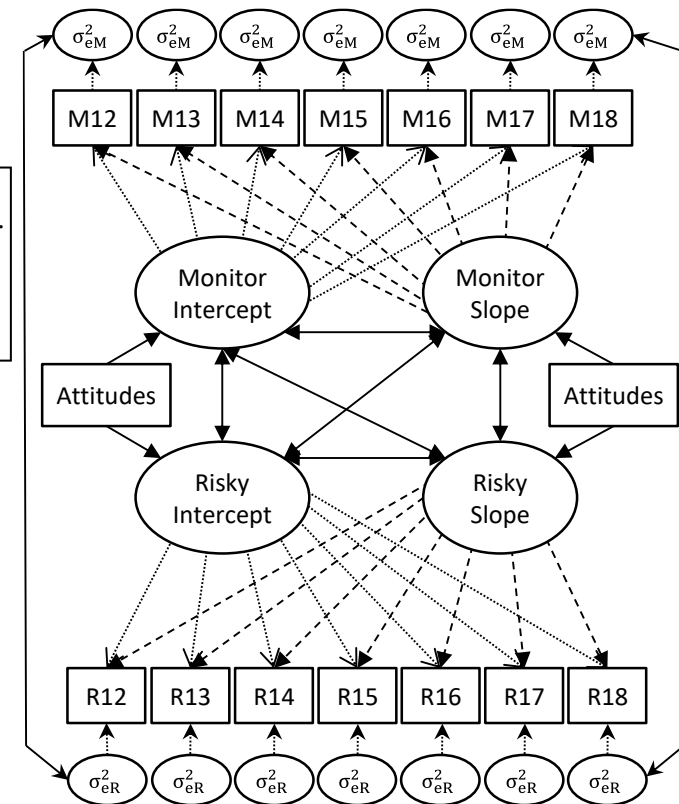
! Risky behavior quadratic change model using exact age as loadings
Rint Rlin Rquad | risky12-risky18 AT age12-age18;
! Monitoring linear change model using exact age as loadings
Mint Mlin | mon12-mon18 AT age12-age18;
! Fixed intercept and change effects for R and M
[Rint Rlin Rquad Mint Mlin];
! L2 G: Random int and linear change variances, no quad change variance
Rint Rlin Mint Mlin (Rintvar Rlinvar Mintvar Mlinvar); Rquad@0;
! L2 G: Within-variable random int-change covariances for R, M
Rint WITH Rlin (Rintlin); Mint WITH Mlin (Mintlin);
! Attitudes --> R int, R linear change, M int, M linear change
Rint Rlin Mint Mlin ON att4;
! L2 G: Covariances between outcomes
Rint WITH Mint (IntCov); ! L2 G: Random int-int covariance
Rlin WITH Mlin (LinCov); ! L2 G: Random change-change covariance
Mint WITH Rlin (Int2Lin); ! L2 G: M int, R change covariance
Mlin WITH Rint (Lin2Int); ! L2 G: M change, R int covariance
! L1 R: WP residual covariance between same ages, held equal across age
risky12-risky18 PWITH mon12-mon18 (ResCov);

MODEL CONSTRAINT: ! Linear combinations of any parameter
NEW(ResCor IntCor LinCor); ! First need to name each new created effect
! Estimating correlations found in RCORR and GCORR
! Cor = Cov / (SQRT(Yvar)*SQRT(Xvar))
ResCor = ResCov / (SQRT(Rresvar)*SQRT(Mresvar));
IntCor = IntCov / (SQRT(Rintvar)*SQRT(Mintvar));
LinCor = LinCov / (SQRT(Rlinvar)*SQRT(Mlinvar));

```

MODEL = NOCOVARIANCES
removes all default covariances.
This will be especially useful
for wide-format SEMs with
“structured residuals”
(later in this example)

Quadratic fixed effect for risky outcome not shown in diagram...
and “slope” factors refer to age change (since there are many slopes!)



- > Indicates paths fixed = 1
- > Indicates paths fixed = time values
- > Indicates paths freely estimated
- > Indicates paths freely estimated between residuals at the same occasion but held equal over time

For balanced time, a linear change model would look like this instead (add Mquad as third variable before |):

```

Mint Mlin | mon12@-6 mon13@-5 mon14@-4 mon15@-3
mon16@-2 mon17@-1 mon18@0;

```

MODEL FIT INFORMATION						Means				
Number of Free Parameters						RQUAD	0.147	0.021	7.117	0.000
Loglikelihood						Intercepts				
H0 Value						MON12	0.000	0.000	999.000	999.000
Information Criteria						MON13	0.000	0.000	999.000	999.000
Akaike (AIC)						MON14	0.000	0.000	999.000	999.000
Bayesian (BIC)						MON15	0.000	0.000	999.000	999.000
Sample-Size Adjusted BIC						MON16	0.000	0.000	999.000	999.000
(n* = (n + 2) / 24)						MON17	0.000	0.000	999.000	999.000
MON18						MON18	0.000	0.000	999.000	999.000
MODEL RESULTS						RISKY12	0.000	0.000	999.000	999.000
		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value	RISKY13	0.000	0.000	999.000	999.000
RINT	ON					RISKY14	0.000	0.000	999.000	999.000
ATT4		-3.160	0.551	-5.737	0.000	RISKY15	0.000	0.000	999.000	999.000
RLIN	ON					RISKY16	0.000	0.000	999.000	999.000
ATT4		-0.518	0.104	-4.963	0.000	RISKY17	0.000	0.000	999.000	999.000
MINT	ON					RISKY18	0.000	0.000	999.000	999.000
ATT4		-0.044	0.057	-0.779	0.436	RINT	23.322	0.348	67.074	0.000
MLIN	ON					RLIN	1.975	0.138	14.259	0.000
ATT4		0.003	0.014	0.240	0.810	MINT	0.063	0.034	1.839	0.066
RINT	WITH					MLIN	-0.003	0.008	-0.380	0.704
RLIN		1.878	0.356	5.273	0.000	Variances				
MINT		-0.853	0.168	-5.076	0.000	RQUAD	0.000	0.000	999.000	999.000
MLIN		0.040	0.039	1.044	0.296	Residual Variances				
MINT	WITH					MON12	0.081	0.004	22.354	0.000
MLIN		0.000	0.004	-0.105	0.916	MON13	0.081	0.004	22.354	0.000
RLIN		-0.106	0.031	-3.449	0.001	MON14	0.081	0.004	22.354	0.000
RLIN	WITH					MON15	0.081	0.004	22.354	0.000
MLIN		-0.018	0.007	-2.478	0.013	MON16	0.081	0.004	22.354	0.000
RISKY12	WITH					MON17	0.081	0.004	22.354	0.000
MON12		0.287	0.028	10.441	0.000	MON18	0.081	0.004	22.354	0.000
RISKY13	WITH					RISKY12	8.352	0.374	22.351	0.000
MON13		0.287	0.028	10.441	0.000	RISKY13	8.352	0.374	22.351	0.000
RISKY14	WITH					RISKY14	8.352	0.374	22.351	0.000
MON14		0.287	0.028	10.441	0.000	RISKY15	8.352	0.374	22.351	0.000
RISKY15	WITH					RISKY16	8.352	0.374	22.351	0.000
MON15		0.287	0.028	10.441	0.000	RISKY17	8.352	0.374	22.351	0.000
RISKY16	WITH					RISKY18	8.352	0.374	22.351	0.000
MON16		0.287	0.028	10.441	0.000	RINT	18.049	2.202	8.198	0.000
RISKY17	WITH					RLIN	0.484	0.080	6.071	0.000
MON17		0.287	0.028	10.441	0.000	MINT	0.195	0.023	8.371	0.000
RISKY18	WITH					MLIN	0.010	0.001	7.802	0.000
MON18		0.287	0.028	10.441	0.000	New/Additional Parameters				
						RESCOR	0.350	0.028	12.607	0.000
						INTCOR	-0.455	0.074	-6.119	0.000
						LINCOR	-0.255	0.103	-2.483	0.013

Model 2a: Partially Directed Path Multivariate MLM in Mplus as Long M-SEM:
Monitor → Risky for BP intercepts and slopes, but WP residuals covary

TITLE: Model 2a: Partially Directed Multivariate Change Model as Long M-SEM	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
L1 WP relation still as residual covariance				
(DATA, VARIABLE, and ANALYSIS are the same as for Model 1 MLM)				
MODEL: ! R = risky behavior, M = monitoring				
%WITHIN%				
risky mon3 (Rresvar Mresvar); ! L1 R: WP residual variances (labels)				
Rlin risky ON time; ! Placeholder for R linear change				
Rquad risky ON timesq; ! Placeholder for R quadratic change				
Mlin mon3 ON time; ! Placeholder for M linear change				
risky WITH mon3 (ResCov); ! L1 R: Still WP residual covariance				
%BETWEEN%				
[risky mon3 Rlin Rquad Mlin]; ! Fixed intercepts, fixed change slopes				
risky mon3 (Rintvar Mintvar); ! L2 G: Random intercept variances (labels)				
Rlin Mlin (Rlinvar Mlinvar); ! L2 G: Random linear change variances				
Rquad@0; ! No quadratic change variance				
risky Rlin ON att4; ! Att-> R int, R linear change				
mon3 Mlin ON att4; ! Att-> M int, M linear change				
risky WITH Rlin (RIntLin); ! L2 G: R int-change covariance (label)				
mon3 WITH Mlin (MIntLin); ! L2 G: M int-change covariance (label)				
! Although we have changed the int-int and change-change relations to direct				
! paths from M -> R instead of covariances, they still represent total L2				
! BP relationships because the L1 relationship is still a covariance				
risky ON mon3 (BPIntEff); ! L2 BP int-to-int effect				
Rlin ON Mlin (BPLinEff); ! L2 BP change-to-change effect				
mon3 WITH Rlin (Int2Lin); ! L2 G: M int, R change covariance				
Mlin WITH risky (Lin2Int); ! L2 G: M change, R int covariance				
MODEL CONSTRAINT:				
! All values below are variances from undirected model 1				
NEW(ResCor IntStd LinStd);				
! Corr = Cov / (SQRT(Yvar)*SQRT(Xvar))				
ResCor = ResCov / (SQRT(8.3538)*SQRT(0.08077)); ! L1 WP res corr				
! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)				
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP int-int effect				
LinStd = BPLinEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP change-change				
This is an equivalent model, just with a different way of specifying the				
level-2 BP intercept-to-intercept and change-to-change relationships.				

Within Level				
RISKY WITH				
MON3	0.287	0.028	10.441	0.000
Residual Variances				
RISKY	8.352	0.374	22.351	0.000
MON3	0.081	0.004	22.354	0.000
Between Level - changed parameters are in BOLD				
RLIN ON				
MLIN	-1.736	0.713	-2.434	0.015
RLIN ON				
ATT4	-0.512	0.105	-4.869	0.000
MLIN ON				
ATT4	0.003	0.014	0.240	0.810
RISKY ON				
ATT4	-3.354	0.528	-6.348	0.000
MON3	-4.380	0.797	-5.497	0.000
MON3 ON				
ATT4	-0.044	0.057	-0.779	0.436
RISKY WITH				
RLIN	1.480	0.345	4.285	0.000
MLIN	0.039	0.038	1.023	0.306
MON3 WITH				
MLIN	0.000	0.004	-0.105	0.916
RLIN	-0.107	0.031	-3.454	0.001
Means				
RQUAD	0.147	0.021	7.117	0.000
Intercepts				
RISKY	23.598	0.338	69.837	0.000
MON3	0.063	0.034	1.839	0.066
RLIN	1.969	0.139	14.195	0.000
MLIN	-0.003	0.008	-0.380	0.704
Variances				
RQUAD	0.000	0.000	999.000	999.000
Residual Variances				
RISKY	14.315	2.030	7.053	0.000
MON3	0.195	0.023	8.371	0.000
RLIN	0.453	0.081	5.564	0.000
MLIN	0.010	0.001	7.802	0.000
New/Additional Parameters				
RESCOR	0.350	0.034	10.441	0.000
INTSTD	-0.455	0.083	-5.497	0.000
LINSTD	-0.254	0.104	-2.434	0.015

**In Mplus, the same Model 2a as a partially directed wide single-level SEM:
Monitor → Risky for intercepts and slopes, but residuals covary**

TITLE: Model 2a: Partially Directed Multivariate Change Model as Wide
Single-Level SEM, L1 WP relation still as residual covariance

(DATA, VARIABLE, and ANALYSIS are the same as for Model 1 SEM)

MODEL: ! R = risky behavior, M = monitoring
[risky12-risky18@0 mon12-mon18@0]; ! All variable intercepts fixed to 0
risky12-risky18 (Rresvar); ! L1 R: WP R residual variances held equal
mon12-mon18 (Mresvar); ! L1 R: WP M residual variances held equal

! Risky behavior quadratic change model using exact age as loadings
Rint Rlin Rquad | risky12-risky18 AT age12-age18;
! Monitoring linear change model using exact age as loadings
Mint Mlin | mon12-mon18 AT age12-age18;
! Fixed intercept and change effects for R and M
[Rint Rlin Rquad Mint Mlin];
! L2 G: Random int and linear change variances, no quad change variance
Rint Rlin Mint Mlin (Rintvar Rlinvar Mintvar Mlinvar); Rquad@0;
! L2 G: Within-variable random int-change covariances for R, M
Rint WITH Rlin (Rintlin); Mint WITH Mlin (Mintlin);
! Attitudes --> R int, R linear change, M int, M change slope
Rint Rlin Mint Mlin ON att4;

! Although we have changed the int-int and change-change relations to direct
! paths from M -> R instead of covariances, they still represent total L2
! BP relationships because the L1 relationship is still a covariance

Rint ON Mint (BPIntEff); ! L2 BP int-to-int effect
Rlin ON Mlin (BPLinEff); ! L2 BP change-to-change effect

Mint WITH Rlin (Int2Lin); ! L2 G: M int, R change covariance
Mlin WITH Rint (Lin2Int); ! L2 G: M change, R int covariance

! L1 R: WP residual covariance between same ages, held equal across age
risky12-risky18 PWITH mon12-mon18 (ResCov);

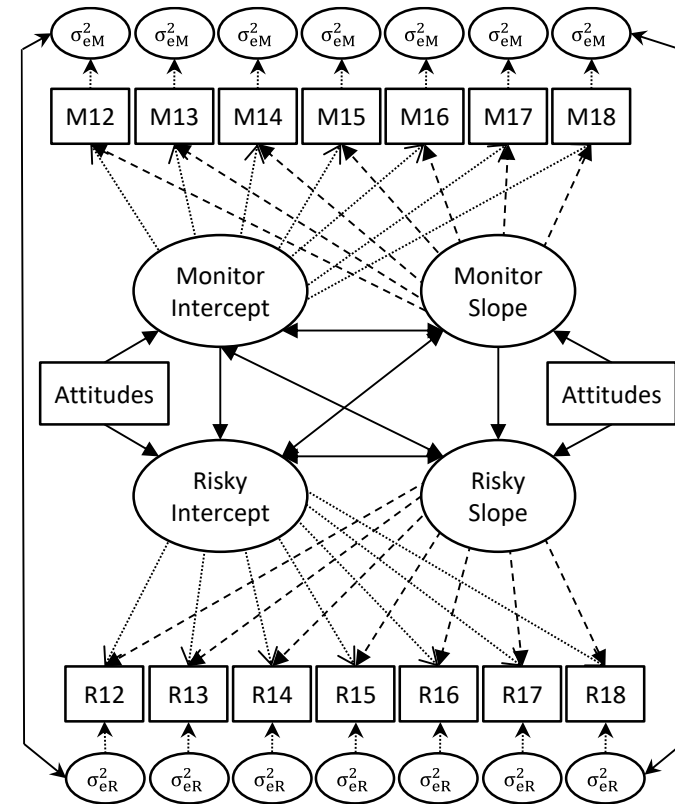
MODEL CONSTRAINT: ! All values below are variances from undirected model 1
NEW(ResCor IntStd LinStd);

! Corr = Cov / (SQRT(Yvar)*SQRT(Xvar))
ResCor = ResCov / (SQRT(8.3538)*SQRT(0.08077)); ! L1 WP res corr

! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP int-int effect
LinStd = BPLinEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP change-change

This is an equivalent model, just with a different way of specifying the level-2 BP
intercept-to-intercept and change-to-change relationships.

Quadratic fixed effect for risky outcome not shown in diagram...
and “slope” factors refer to age change (since there are many slopes!)



-→ Indicates paths fixed = 1
- Indicates paths fixed = time values
- ◄-----► Indicates paths freely estimated
- ◄-----► Indicates paths freely estimated between residuals at the same occasion but held equal over time

For balanced time, a linear change model would look like this instead (add Mquad as third variable before |):

Mint Mlin | mon12@-6 mon13@-5 mon14@-4 mon15@-3
mon16@-2 mon17@-1 mon18@0;

MODEL RESULTS - changed parameters are in BOLD						Means			
		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value	RQUAD	0.147	0.021	7.117
RINT	ON								
MINT		-4.380	0.797	-5.496	0.000				
RLIN	ON								
MLIN		-1.735	0.713	-2.434	0.015				
RINT	ON								
ATT4		-3.354	0.528	-6.348	0.000				
RLIN	ON								
ATT4		-0.512	0.105	-4.869	0.000				
MINT	ON								
ATT4		-0.044	0.057	-0.779	0.436				
MLIN	ON								
ATT4		0.003	0.014	0.240	0.810				
RINT	WITH								
RLIN		1.480	0.345	4.285	0.000				
MLIN		0.039	0.038	1.023	0.306				
MINT	WITH								
MLIN		0.000	0.004	-0.105	0.916				
RLIN		-0.107	0.031	-3.454	0.001				
RISKY12	WITH								
MON12		0.287	0.028	10.441	0.000				
RISKY13	WITH								
MON13		0.287	0.028	10.441	0.000				
RISKY14	WITH								
MON14		0.287	0.028	10.441	0.000				
RISKY15	WITH								
MON15		0.287	0.028	10.441	0.000				
RISKY16	WITH								
MON16		0.287	0.028	10.441	0.000				
RISKY17	WITH								
MON17		0.287	0.028	10.441	0.000				
RISKY18	WITH								
MON18		0.287	0.028	10.441	0.000				
						Variances			
						RQUAD	0.000	0.000	999.000
						Residual Variances			
						RISKY12	8.352	0.374	22.351
						RISKY13	8.352	0.374	22.351
						RISKY14	8.352	0.374	22.351
						RISKY15	8.352	0.374	22.351
						RISKY16	8.352	0.374	22.351
						RISKY17	8.352	0.374	22.351
						RISKY18	8.352	0.374	22.351
						MON12	0.081	0.004	22.354
						MON13	0.081	0.004	22.354
						MON14	0.081	0.004	22.354
						MON15	0.081	0.004	22.354
						MON16	0.081	0.004	22.354
						MON17	0.081	0.004	22.354
						MON18	0.081	0.004	22.354
						RINT	14.315	2.030	7.053
						RLIN	0.453	0.081	5.564
						MINT	0.195	0.023	8.371
						MLIN	0.010	0.001	7.802
						New/Additional Parameters			
						RESCOR	0.350	0.034	10.441
						INTSTD	-0.455	0.083	-5.496
						LINSTD	-0.254	0.104	-2.434

Model 2b: Partially Directed Path Multivariate MLM in Mplus as Long M-SEM: Monitor → Risky for WP residuals within L1 model**Also demonstrating how to request BP indirect effects**

TITLE: Model 2b: Partially Directed Multivariate Change Model as Long M-SEM L1 WP relation NOW as direct path specified in L1 WITHIN	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
(DATA, VARIABLE, and ANALYSIS are the same as for Model 1 MLM)				
MODEL: ! R = risky behavior, M = monitoring				
%WITHIN%				
risky mon3 (Rresvar Mresvar); ! L1 R: Residual variances (labels)				
Rlin risky ON time; ! Placeholder for R linear change				
Rquad risky ON timesq; ! Placeholder for R quadratic change				
Mlin mon3 ON time; ! Placeholder for M linear change				
risky ON mon3 (ResEff); ! L1 WP fixed effect M->R here (label)				
%BETWEEN%				
[risky mon3 Rlin Rquad Mlin]; ! Fixed intercepts, fixed change slopes				
risky mon3 (Rintvar Mintvar); ! L2 G: Random intercept variances (labels)				
Rlin Mlin (Rlinvar Mlinvar); ! L2 G: Random linear change variances				
Rquad@0; ! No quadratic change variance				
risky Rlin ON att4 (XtoYint XtoYlin); ! Att-> R int, R linear change				
mon3 Mlin ON att4 (XtoMint XtoMlin); ! Att-> M int, M linear change				
risky WITH Rlin (RintLin); ! L2 G: R int-change covariance (label)				
mon3 WITH Mlin (MintLin); ! L2 G: M int-change covariance (label)				
! Although the intercept -> intercept path remains the total L2 BP effect, ! now the change -> change path becomes the L2 contextual effect instead				
risky ON mon3 (BPIntEff); ! STILL L2 BP int-to-int effect				
Rlin ON Mlin (LinCont); ! NOW L2 contextual change-to-change effect				
mon3 WITH Rlin (Int2Lin); ! L2 G: M int, R change covariance				
Mlin WITH risky (Lin2Int); ! L2 G: M change, R int covariance				
MODEL CONSTRAINT:				
NEW(ResStd IntStd BPLinEff LinStd indBPint indBplin);				
! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)				
ResStd = ResEff * SQRT(0.08077) / SQRT(8.3538); ! STD L1 WP res-res effect				
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP int-int effect				
BPLinEff = ResEff + LinCont; ! WP + Context = BP change-change				
LinStd = BPLinEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP change effect				
indBPint = XtoMint * BPIntEff; ! BP intercept indirect effect				
indBplin = XtoMlin * BPLinEff; ! BP change indirect effect				
This is still an equivalent model, just with a different way of specifying the level-1 WP residual-to-residual relationship. This syntax method will only work for level-1 effects that are only fixed, though—adding random effects and/or cross-level interactions requires the level-1 placeholder syntax.				

Within Level - changed parameters are in BOLD	
RISKY ON	
MON3	3.559 0.301 11.809 0.000
Residual Variances	
RISKY	7.329 0.328 22.353 0.000
MON3	0.081 0.004 22.354 0.000
Between Level - changed parameters are in BOLD	
RLIN ON	
MLIN	-5.294 0.806 -6.569 0.000
RLIN ON	
ATT4	-0.512 0.105 -4.869 0.000
MLIN ON	
ATT4	0.003 0.014 0.240 0.810
RISKY ON	
ATT4	-3.354 0.528 -6.348 0.000
MON3	-4.380 0.797 -5.497 0.000
MON3 ON	
ATT4	-0.044 0.057 -0.779 0.436
RISKY WITH	
RLIN	1.480 0.345 4.285 0.000
MLIN	0.039 0.038 1.023 0.306
MON3 WITH	
MLIN	0.000 0.004 -0.105 0.916
RLIN	-0.107 0.031 -3.454 0.001
Means	
RQUAD	0.147 0.021 7.117 0.000
Intercepts	
RISKY	23.598 0.338 69.837 0.000
MON3	0.063 0.034 1.839 0.066
RLIN	1.969 0.139 14.195 0.000
MLIN	-0.003 0.008 -0.380 0.704
Variances	
RQUAD	0.000 0.000 999.000 999.000
Residual Variances	
RISKY	14.315 2.030 7.053 0.000
MON3	0.195 0.023 8.371 0.000
RLIN	0.453 0.081 5.564 0.000
MLIN	0.010 0.001 7.802 0.000
New/Additional Parameters	
RESSTD	0.350 0.034 10.441 0.000
INTSTD	-0.455 0.083 -5.497 0.000
BPLINEFF	-1.736 0.713 -2.434 0.015
LINSTD	-0.254 0.104 -2.434 0.015
INDBPINT	0.194 0.251 0.772 0.440
INDBPLIN	-0.006 0.024 -0.239 0.811

This is still an equivalent model, just with a different way of specifying the level-1 WP residual-to-residual relationship. This syntax method will only work for level-1 effects that are only fixed, though—adding random effects and/or cross-level interactions requires the level-1 placeholder syntax.

**In Mplus, the same Model 2b as a partially directed wide single-level SEM:
Monitor → Risky for WP residuals using structured residuals**

TITLE: Model 2b: Partially Directed Multivariate Change Model as Wide
Single-Level SEM, L1 WP effect as direct path via structured residuals
(DATA, VARIABLE, and ANALYSIS are the same as for Model 1 SEM)

MODEL: ! R = risky behavior, M = monitoring
[riskyl2-riskyl18@0 mon12-mon18@0]; ! All variable intercepts fixed to 0

```
! Risky behavior quadratic change model using exact age as loadings
Rint Rlin Rquad | riskyl2-riskyl18 AT age12-age18;
! Monitoring linear change model using exact age as loadings
Mint Mlin | mon12-mon18 AT age12-age18;
! Fixed intercept and change effects for R and M
[Rint Rlin Rquad Mint Mlin];
! L2 G: Random int and linear change variances, no quad change variance
Rint Rlin Mint Mlin (Rintvar Rlinvar Mintvar Mlinvar); Rquad@0;
! L2 G: Within-variable random int-change covariances for R, M
Rint WITH Rlin (Rintlin); Mint WITH Mlin (Mintlin);
! Attitudes --> R int, R change, M int, M change
Rint Rlin Mint Mlin ON att4 (XtoYint XtoYlin XtoMint XtoMlin);
Rint ON Mint (BPIntEff); ! BP int-to-int effect
Rlin ON Mlin (BPLinEff); ! BP change-to-change effect
Mint WITH Rlin (Int2Lin); ! L2 G: M int, R change covariance
Mlin WITH Rint (Lin2Int); ! L2 G: M change, R int covariance
```

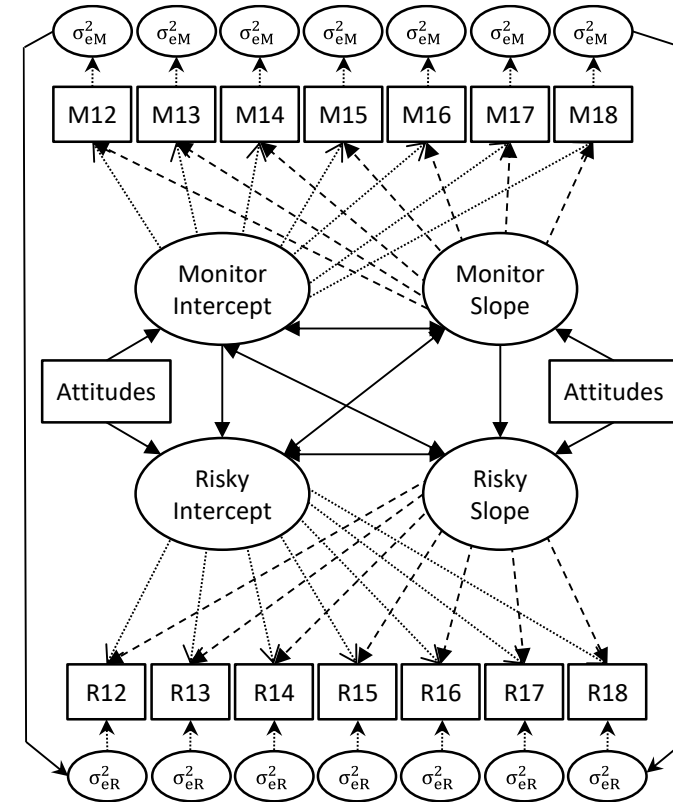
```
! Define new latent factors for residuals at each occasion
Frisky12 BY riskyl2@1; Frisky13 BY riskyl3@1; Frisky14 BY riskyl4@1;
Frisky15 BY riskyl5@1; Frisky16 BY riskyl6@1; Frisky17 BY riskyl7@1;
Frisky18 BY riskyl8@1; Fmon12 BY mon12@1; Fmon13 BY mon13@1;
Fmon14 BY mon14@1; Fmon15 BY mon15@1; Fmon16 BY mon16@1;
Fmon17 BY mon17@1; Fmon18 BY mon18@1;
! All factor means fixed to 0
[Frisky12-Frisky18@0 Fmon12-Fmon18@0];
! Shut off old residual variances
riskyl2-riskyl18@0 mon12-mon18@0;
! Hold new residual variances equal over time
Frisky12-Frisky18 (Rresvar); ! L1 R: R WP residual variances held equal
Fmon12-Fmon18 (Mresvar); ! L1 R: M WP residual variances held equal
! Factor residual WP effect between same ages, held equal across age
Frisky12-Frisky18 PON Fmon12-Fmon18 (ResEff);
```

MODEL CONSTRAINT:

```
NEW(ResStd IntStd LinStd indBPint indBplin);
! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
ResStd = ResEff * SQRT(0.08077) / SQRT(8.3538); ! STD WP res-res effect
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP int-int effect
LinStd = BPLinEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP change-change
indBPint = XtoMint * BPIntEff; ! BP intercept indirect effect
indBplin = XtoMlin * BPLinEff; ! BP change indirect effect
```

This is still an equivalent model, just with a different way of specifying the level-1 WP residual-to-residual relationship. This syntax method will only work for level-1 effects that are only fixed, though...

Quadratic fixed effect for risky outcome not shown in diagram...
and “slope” factors refer to age change (since there are many slopes!)



For simpler residuals code using Mplus
v. 8.8+, see [Hoffman & Hall \(2024\)](#)
[Example 1 \(supplemental materials .zip\)](#)

-> Indicates paths fixed = 1
- > Indicates paths fixed = time values
- ←-----> Indicates paths freely estimated
- > Indicates paths freely estimated
- > Indicates paths freely estimated between residuals at the same occasion but held equal over time

For balanced time, a linear change model would look like this instead (add Mquad as third variable before !):

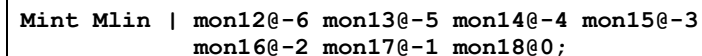
```
Mint Mlin | mon12@-6 mon13@-5 mon14@-4 mon15@-3
mon16@-2 mon17@-1 mon18@0;
```

MODEL RESULTS - changed parameters are in BOLD						Means				
		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value	RQUAD	0.147	0.021	7.117	0.000
Factor loadings set to 1 omitted						Intercepts				
						(Intercepts fixed to 0 omitted)				
RINT	ON					RINT	23.598	0.338	69.836	0.000
MINT		-4.380	0.797	-5.496	0.000	RLIN	1.969	0.139	14.195	0.000
RLIN	ON					MINT	0.063	0.034	1.839	0.066
MLIN		-1.736	0.713	-2.434	0.015	MLIN	-0.003	0.008	-0.380	0.704
FRISKY12	ON					Variances				
FMON12		3.563	0.302	11.810	0.000	FMON12	0.081	0.004	22.327	0.000
FRISKY13	ON					FMON13	0.081	0.004	22.327	0.000
FMON13		3.563	0.302	11.810	0.000	FMON14	0.081	0.004	22.327	0.000
FRISKY14	ON					FMON15	0.081	0.004	22.327	0.000
FMON14		3.563	0.302	11.810	0.000	FMON16	0.081	0.004	22.327	0.000
FRISKY15	ON					FMON17	0.081	0.004	22.327	0.000
FMON15		3.563	0.302	11.810	0.000	FMON18	0.081	0.004	22.327	0.000
FRISKY16	ON					RQUAD	0.000	0.000	999.000	999.000
FMON16		3.563	0.302	11.810	0.000	Residual Variances				
FRISKY17	ON					(Residual variances fixed to 0 omitted)				
FMON17		3.563	0.302	11.810	0.000	FRISKY12	7.328	0.328	22.349	0.000
FRISKY18	ON					FRISKY13	7.328	0.328	22.349	0.000
FMON18		3.563	0.302	11.810	0.000	FRISKY14	7.328	0.328	22.349	0.000
RINT	ON					FRISKY15	7.328	0.328	22.349	0.000
ATT4		-3.354	0.528	-6.348	0.000	FRISKY16	7.328	0.328	22.349	0.000
RLIN	ON					FRISKY17	7.328	0.328	22.349	0.000
ATT4		-0.512	0.105	-4.869	0.000	FRISKY18	7.328	0.328	22.349	0.000
MINT	ON					RINT	14.315	2.030	7.053	0.000
ATT4		-0.044	0.057	-0.779	0.436	RLIN	0.453	0.081	5.564	0.000
MLIN	ON					MINT	0.195	0.023	8.371	0.000
ATT4		0.003	0.014	0.240	0.810	MLIN	0.010	0.001	7.802	0.000
RINT	WITH					New/Additional Parameters				
RLIN		1.480	0.345	4.285	0.000	RESSTD	0.350	0.034	10.441	0.000
MLIN		0.039	0.038	1.023	0.306	INTSTD	-0.455	0.083	-5.496	0.000
MINT	WITH					LINSTD	-0.254	0.104	-2.434	0.015
MLIN		0.000	0.004	-0.105	0.916	INDBPINT	0.194	0.251	0.772	0.440
RLIN		-0.107	0.031	-3.454	0.001	INDBPLIN	-0.006	0.024	-0.239	0.811

Model 2c: Partially Directed Path Multivariate MLM in Mplus as Long M-SEM: Monitor → Risky for WP residuals within L2 model
via placeholder syntax; also demonstrating how to request BP indirect effects

TITLE: Model 2c: Partially Directed Multivariate Change Model as Long M-SEM		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
L1 WP relation using placeholder specified in L2 BETWEEN					
(DATA, VARIABLE, and ANALYSIS are the same as for Model 1 MLM)					
MODEL: ! R = risky behavior, M = monitoring					
%WITHIN%					
risky mon3 (Rresvar Mresvar); ! L1 R: WP residual variances (labels)					
Rlin risky ON time; ! Placeholder for R linear change					
Rquad risky ON timesq; ! Placeholder for R quadratic change					
Mlin mon3 ON time; ! Placeholder for M linear change					
WPRES risky ON mon3; ! NEW placeholder for L1 WP effect M->R					
%BETWEEN%					
[risky mon3 Rlin Rquad Mlin]; ! Fixed intercepts, fixed change slopes					
risky mon3 (Rintvar Mintvar); ! L2 G: Random intercept variances (labels)					
Rlin Mlin (Rlinvar Mlinvar); ! L2 G: Random linear change variances					
Rquad@0; ! No quadratic change variance					
risky Rlin ON att4 (XtoYint XtoYlin); ! Att-> R int, R linear change					
mon3 Mlin ON att4 (XtoMint XtoMlin); ! Att-> M int, M linear change					
risky WITH Rlin (Int2Lin); ! L2 G: R Int-change covariance (label)					
mon3 WITH Mlin (MIntLin); ! L2 G: M Int-change covariance (label)					
! And now both the intercept -> intercept path and the change -> change path					
! are L2 contextual effects given the L1 placeholder for WP M -> R effect					
risky ON mon3 (IntCont); ! NOW L2 contextual int-to-int effect					
Rlin ON Mlin (LinCont); ! STILL L2 contextual change-to-change effect					
mon3 WITH Rlin (Int2Lin); ! L2 G: M int, R change covariance					
Mlin WITH risky (Lin2Int); ! L2 G: M change, R int covariance					
[WPRES] (ResEff); ! Fixed effect for L1 WP M->R (as defined earlier)					
WPRES@0; ! No random L1 WP M->R slope variance					
MODEL CONSTRAINT:					
NEW(ResStd BPIntEff IntStd BPLinEff LinStd indBPint indBPLin);					
! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)					
ResStd = ResEff * SQRT(0.08077) / SQRT(8.3538); ! STD WP effect					
BPIntEff = ResEff + IntCont; ! WP + Context = BP int-int					
IntStd = BPIntEff * SQRT(0.19530) / SQRT(18.0644); ! STD BP int-int effect					
BPLinEff = ResEff + LinCont; ! WP + Context = BP change-change					
LinStd = BPLinEff * SQRT(0.01049) / SQRT(0.48830); ! STD BP change-change					
indBPint = XtoMint * BPIntEff; ! BP intercept indirect effect					
indBPLin = XtoMlin * BPLinEff; ! BP change indirect effect					
This is still an equivalent model, just with a different syntax for specifying the same					
level-1 WP residual-to-residual directed relationship. This is the version that is					
necessary in order to have the level-1 effect become random or systematically					
varying (i.e., add cross-level interactions). But if you switch to Bayes estimation,					
then all the L2 effects are BP instead (see Hoffman 2019)!!!					
Within Level					
Residual Variances					
RISKY		7.329	0.328	22.353	0.000
MON3		0.081	0.004	22.354	0.000
Between Level - changed parameters are in BOLD					
RLIN ON					
MLIN		-5.294	0.806	-6.569	0.000
RLIN ON					
ATT4		-0.512	0.105	-4.869	0.000
MLIN ON					
ATT4		0.003	0.014	0.240	0.810
RISKY ON					
ATT4		-3.354	0.528	-6.348	0.000
MON3		-7.938	0.872	-9.099	0.000
MON3 ON					
ATT4		-0.044	0.057	-0.779	0.436
RISKY WITH					
RLIN		1.480	0.345	4.285	0.000
MLIN		0.039	0.038	1.023	0.306
MON3 WITH					
MLIN		0.000	0.004	-0.105	0.916
RLIN		-0.107	0.031	-3.454	0.001
Means					
RQUAD		0.147	0.021	7.117	0.000
WPRES (here now)		3.559	0.301	11.810	0.000
Intercepts					
RISKY		23.598	0.338	69.837	0.000
MON3		0.063	0.034	1.839	0.066
RLIN		1.969	0.139	14.195	0.000
MLIN		-0.003	0.008	-0.380	0.704
Variances					
RQUAD		0.000	0.000	999.000	999.000
WPRES		0.000	0.000	999.000	999.000
Residual Variances					
RISKY		14.315	2.030	7.053	0.000
MON3		0.195	0.023	8.371	0.000
RLIN		0.453	0.081	5.564	0.000
MLIN		0.010	0.001	7.802	0.000
New/Additional Parameters					
RESSTD		0.350	0.030	11.810	0.000
BPINTEFF		-4.380	0.797	-5.496	0.000
INTSTD		-0.455	0.083	-5.496	0.000
BPLINEFF		-1.735	0.713	-2.434	0.015
LINSTD		-0.254	0.104	-2.434	0.015
INDBPINT		0.194	0.251	0.772	0.440
INDBPLIN		-0.006	0.024	-0.239	0.811

This is still an equivalent model, just with a different syntax for specifying the same level-1 WP directed relationship. The consequence is that the intercept-to-intercept and change-to-change relationships become L2 contextual effects (as in the MLM version). Oddly, if we were to switch to ON for the intercept–change cross-variable relationships, those stay L2 BP (see chapter 9 for an example using this version of the model).



MODEL RESULTS - changed parameters are in BOLD						Means				
		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value	RQUAD	0.147	0.021	7.117	0.000
Factor loadings set to 1 omitted						Intercepts				
						Intercepts fixed to 0 omitted				
RINT	ON					RINT	23.598	0.338	69.836	0.000
MINT		-7.939	0.872	-9.099	0.000	RLIN	1.969	0.139	14.195	0.000
RLIN	ON					MINT	0.063	0.034	1.839	0.066
MLIN		-5.294	0.806	-6.569	0.000	MLIN	-0.003	0.008	-0.380	0.704
RINT	ON					Variances				
ATT4		-3.354	0.528	-6.348	0.000	RQUAD	0.000	0.000	999.000	999.000
RLIN	ON					Residual Variances				
ATT4		-0.512	0.105	-4.869	0.000	RISKY12	7.329	0.328	22.353	0.000
MINT	ON					RISKY13	7.329	0.328	22.353	0.000
ATT4		-0.044	0.057	-0.779	0.436	RISKY14	7.329	0.328	22.353	0.000
MLIN	ON					RISKY15	7.329	0.328	22.353	0.000
ATT4		0.003	0.014	0.240	0.810	RISKY16	7.329	0.328	22.353	0.000
RISKY12	ON					RISKY17	7.329	0.328	22.353	0.000
MON12		3.559	0.301	11.810	0.000	RISKY18	7.329	0.328	22.353	0.000
RISKY13	ON					MON12	0.081	0.004	22.354	0.000
MON13		3.559	0.301	11.810	0.000	MON13	0.081	0.004	22.354	0.000
RISKY14	ON					MON14	0.081	0.004	22.354	0.000
MON14		3.559	0.301	11.810	0.000	MON15	0.081	0.004	22.354	0.000
RISKY15	ON					MON16	0.081	0.004	22.354	0.000
MON15		3.559	0.301	11.810	0.000	MON17	0.081	0.004	22.354	0.000
RISKY16	ON					MON18	0.081	0.004	22.354	0.000
MON16		3.559	0.301	11.810	0.000	RINT	14.315	2.030	7.053	0.000
RISKY17	ON					RLIN	0.453	0.081	5.564	0.000
MON17		3.559	0.301	11.810	0.000	MINT	0.195	0.023	8.371	0.000
RISKY18	ON					MLIN	0.010	0.001	7.802	0.000
MON18		3.559	0.301	11.810	0.000	New/Additional Parameters				
RINT	WITH					RESSTD	0.350	0.030	11.810	0.000
RLIN		1.480	0.345	4.285	0.000	BPINTEFF	-4.380	0.797	-5.496	0.000
MLIN		0.039	0.038	1.023	0.306	INTSTD	-0.455	0.083	-5.496	0.000
MINT	WITH					BPLINEFF	-1.735	0.713	-2.434	0.015
MLIN		0.000	0.004	-0.105	0.916	LINSTD	-0.254	0.104	-2.434	0.015
RLIN		-0.107	0.031	-3.454	0.001	INDBPINT	0.194	0.251	0.772	0.440
						INDBPLIN	-0.006	0.024	-0.239	0.811

Here is an example of how to use “structured residuals” to fit **two cross-lagged effects at level 1** using wide single-level SEM:
Model 3a, which switches to covariances at level 2 when fitting these models (per convention, to be agnostic as to which comes first)

TITLE: Model 3a: Wide SEM Structured Residuals to Fit 2 Cross-Lagged Paths (DATA, VARIABLE, and ANALYSIS are the same as for Model 1 SEM)		MODEL RESULTS - Parameters fixed to 0 or 1 are omitted for brevity changed parameters in bold				
MODEL: ! R = risky behavior, M = monitoring [risky12-risky18@0 mon12-mon18@0]; ! All variable intercepts fixed to 0			Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
! Risky behavior quadratic change model using exact age as loadings Rint Rlin Rquad risky12-risky18 AT age12-age18;		FRISKY13 ON				
		FMON12	-0.171	0.368	-0.465	0.642
! Monitoring linear change model using exact age as loadings Mint Mlin mon12-mon18 AT age12-age18;		FRISKY14 ON				
		FMON13	-0.171	0.368	-0.465	0.642
! Fixed intercept and change effects for R and M [Rint Rlin Rquad Mint Mlin];		FRISKY15 ON				
		FMON14	-0.171	0.368	-0.465	0.642
! L2 G: Random int and linear change variances, no quad change variance Rint Rlin Mint Mlin (Rintvar Rlinvar Mintvar Mlinvar); Rquad@0;		FRISKY16 ON				
		FMON15	-0.171	0.368	-0.465	0.642
! Attitudes --> R int, R change, M int, M change Rint Rlin Mint Mlin ON att4 (XtoYint XtoYlin XtoMint XtoMlin);		FRISKY17 ON				
		FMON16	-0.171	0.368	-0.465	0.642
! L2 G: covariances for random intercepts and slopes across outcomes Rint Rlin Mint Mlin WITH Rint Rlin Mint Mlin;		FRISKY18 ON				
		FMON17	-0.171	0.368	-0.465	0.642
! Define new latent factors for residuals at each occasion Frisky12 BY risky12@1; Frisky13 BY risky13@1; Frisky14 BY risky14@1; Frisky15 BY risky15@1; Frisky16 BY risky16@1; Frisky17 BY risky17@1; Frisky18 BY risky18@1; Fmon12 BY mon12@1; Fmon13 BY mon13@1; Fmon14 BY mon14@1; Fmon15 BY mon15@1; Fmon16 BY mon16@1; Fmon17 BY mon17@1; Fmon18 BY mon18@1;		FMON13 ON				
		FRISKY12	0.008	0.004	2.131	0.033
! All factor means fixed to 0 [Frisky12-Frisky18@0 Fmon12-Fmon18@0];		FMON14 ON				
! Shut off old residual variances risky12-risky18@0 mon12-mon18@0;		FRISKY13	0.008	0.004	2.131	0.033
! Hold new residual variances equal over time if predicted Frisky13-Frisky18 (Rresvar); ! L1 R: R WP residual variances held equal Fmon13-Fmon18 (Mresvar); ! L1 R: M WP residual variances held equal		FMON15 ON				
		FRISKY14	0.008	0.004	2.131	0.033
! WP residual covariance for unpredicted occasion between same age Frisky12 WITH Fmon12;		FMON16 ON				
		FRISKY15	0.008	0.004	2.131	0.033
! WP residual covariance for predicted occasions held equal across age Frisky13-Frisky18 PWITH Fmon13-Fmon18 (ResCov);		FMON17 ON				
		FRISKY16	0.008	0.004	2.131	0.033
		FMON18 ON				
		FRISKY17	0.008	0.004	2.131	0.033
! Cross-lagged WP effects predicting next occasion, held equal across age Frisky13-Frisky18 PON Fmon12-Fmon17 (MR2RR); Fmon13-Fmon18 PON Frisky12-Frisky17 (RR2MR);		RINT ON				
		ATT4	-3.153	0.552	-5.713	0.000
		RLIN ON				
		ATT4	-0.514	0.105	-4.877	0.000
		MINT ON				
		ATT4	-0.043	0.057	-0.765	0.444
		MLIN ON				
		ATT4	0.004	0.014	0.309	0.757
		RINT WITH				
		RLIN	1.945	0.363	5.365	0.000
		MINT	-0.884	0.170	-5.213	0.000
		MLIN	0.038	0.039	0.976	0.329
		RLIN WITH				
		MINT	-0.110	0.031	-3.511	0.000
		MLIN	-0.018	0.008	-2.375	0.018
		MINT WITH				
		MLIN	-0.001	0.004	-0.310	0.756
MODEL CONSTRAINT: NEW(MR2RRsd RR2MRsd); ! STD = Unstd * SQRT(Xvar) / SQRT(Yvar) MR2RRsd = MR2RR * SQRT(0.08077) / SQRT(8.3538); ! STD M->R lagged effect RR2MRsd = RR2MR * SQRT(8.3538) / SQRT(0.08077); ! STD R->M lagged effect						
This is NOT an equivalent model given the two new cross-lagged effects, along with the two separate residual variances and covariance for age 12.						

FRISKY12 WITH FMON12	0.053	0.087	0.608	0.543
FRISKY13 WITH FMON13	0.329	0.033	9.970	0.000
FRISKY14 WITH FMON14	0.329	0.033	9.970	0.000
FRISKY15 WITH FMON15	0.329	0.033	9.970	0.000
FRISKY16 WITH FMON16	0.329	0.033	9.970	0.000
FRISKY17 WITH FMON17	0.329	0.033	9.970	0.000
FRISKY18 WITH FMON18	0.329	0.033	9.970	0.000
Means				
RQUAD	0.145	0.021	7.041	0.000
Intercepts				
RINT	23.322	0.347	67.162	0.000
RLIN	1.971	0.137	14.378	0.000
MINT	0.064	0.034	1.873	0.061
MLIN	-0.003	0.008	-0.326	0.745
Variances				
FRISKY12	9.010	1.277	7.055	0.000
FMON12	0.061	0.011	5.372	0.000
RQUAD	0.000	0.000	999.000	999.000
Residual Variances				
FRISKY13	8.188	0.395	20.711	0.000
FRISKY14	8.188	0.395	20.711	0.000
FRISKY15	8.188	0.395	20.711	0.000
FRISKY16	8.188	0.395	20.711	0.000
FRISKY17	8.188	0.395	20.711	0.000
FRISKY18	8.188	0.395	20.711	0.000
FMON13	0.084	0.004	20.580	0.000
FMON14	0.084	0.004	20.580	0.000
FMON15	0.084	0.004	20.580	0.000
FMON16	0.084	0.004	20.580	0.000
FMON17	0.084	0.004	20.580	0.000
FMON18	0.084	0.004	20.580	0.000
RINT	18.257	2.213	8.249	0.000
RLIN	0.513	0.083	6.180	0.000
MINT	0.191	0.023	8.259	0.000
MLIN	0.010	0.001	7.859	0.000
New/Additional Parameters				
MR2RRSD	-0.017	0.036	-0.465	0.642
RR2MRSD	0.079	0.037	2.131	0.033

What if we controlled for the concurrent effect of Monitor → Risky before examining the lagged effect of Monitor → Risky (my own preference)?

TITLE: Model 3b: Example of Structured Residuals to Fit M→R Cross-Lagged Path that controls for concurrent effect before fitting the lagged effect

All else is the same until here...

! Residual WP effect between same ages, held equal across age
Frisky12-Frisky18 PON Fmon12-Fmon18 (ResEff);

! Cross-lagged WP effects predicting next occasion, held equal across age
Frisky13-Frisky18 PON Fmon12-Fmon17 (MR2RR);

MODEL CONSTRAINT:

NEW(ResStd MR2RRsd);

! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)

ResStd = ResEff * SQRT(Mresvar) / SQRT(Rresvar); ! STD M→R concurrent

MR2RRsd = MR2RR * SQRT(Mresvar) / SQRT(Rresvar); ! STD M→R lagged effect

	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
New/Additional Parameters				
RESSTD	0.339	0.031	10.952	0.000
MR2RRSD	-0.054	0.033	-1.610	0.107

TITLE: Model 3c: Example of Structured Residuals to Fit R→M Cross-Lagged Path that controls for concurrent effect before fitting the lagged effect

All else is the same until here...

! Residual WP effect between same ages, held equal across age
Fmon12-Fmon18 PON Frisky12-Frisky18 (ResEff);

! Cross-lagged WP effects predicting next occasion, held equal across age
Fmon13-Fmon18 PON Frisky12-Frisky17 (RR2MR);

MODEL CONSTRAINT:

NEW(ResStd RR2MRsd);

! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)

ResStd = ResEff * SQRT(8.3538) / SQRT(0.08077); ! STD R→M concurrent

RR2MRsd = RR2MR * SQRT(8.3538) / SQRT(0.08077); ! STD R→M lagged effect

	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
New/Additional Parameters				
RESSTD	0.371	0.031	11.846	0.000
RR2MRSD	0.087	0.034	2.547	0.011

It looks like evidence for a lagged Risky → Monitor lagged effect is a little stronger after controlling for the concurrent effect (and vice-versa).

Here is a comparison of the SEM cross-lagged effects to those using long M-SEM with the **LAGGED residual option** (available in Mplus 8.1+ with BAYES estimation).

```

TITLE:  Model 3a: Undirected Directed Multivariate Change Model as Long M-SEM
All fixed L1 effects specified in WITHIN
Adding lagged effects of both M -> R and R --> M
using new residual LAGGED option and BAYES estimation

DATA:  FILE = Example5b.csv;    ! Syntax in same folder as data
VARIABLE: ! List of variables in data file
NAMES = PersonID Att12 occasion age risky mon roundage
        time att4 timesq mon3;
! Variables to be analyzed in this model
USEVARIABLE = time timesq att4 risky mon3;
MISSING ARE ALL (-999);    ! Missing data identifier
! MLM options
CLUSTER = PersonID;        ! Level-2 ID
BETWEEN = att4;            ! Observed ONLY level-2 predictors
WITHIN = time timesq;      ! Observed ONLY level-1 predictors
LAGGED = risky(1) mon3(1);    ! Create Mplus lag-1 variables

ANALYSIS:  TYPE = TWOLEVEL RANDOM; ESTIMATOR = BAYES; BITERATIONS = 50000;
ANALYSIS:  TECH8;    ! Used to examine convergence
MODEL:    ! R = risky behavior, M = monitoring
%WITHIN%
risky mon3 (Rresvar Mresvar); ! L1 R: Residual variances (labels)
Rlin | risky ON time;          ! Placeholder for R linear change
Rquad | risky ON timesq;       ! Placeholder for R quadratic change
Mlin | mon3 ON time;           ! Placeholder for M linear change
risky WITH mon3 (ResCov);      ! L1 WP covariance for concurrent M->R
risky^ ON mon3^1 (MRLagEff); ! L1 WP fixed effect of ^residual lagged M->R
mon3^ ON risky^1 (RMLagEff); ! L1 WP fixed effect of ^residual lagged R->M

%BETWEEN%
[risky mon3 Rlin Rquad Mlin]; ! Fixed intercepts, fixed change slopes
risky mon3 (Rintvar Mintvar); ! L2 G: Random intercept variances (labels)
Rlin Mlin (Rlinvar Mlinvar); ! L2 G: Random linear change variances
risky Rlin ON att4 (XtoYint XtoYlin); ! Att-> R int, R linear change
mon3 Mlin ON att4 (XtoMint XtoMlin); ! Att-> M int, M linear change
! L2 G: covariances for random intercepts and slopes across outcomes
risky Rlin mon3 Mlin WITH risky Rlin mon3 Mlin;
Rquad@0; ! No quadratic change variance

MODEL CONSTRAINT:
NEW(MR2RRsd RR2MRsd);
! STD = Unstd * SQRT(Xvar) / SQRT(Yvar)
MR2RRsd = MRLagEff * SQRT(0.08077) / SQRT(8.3538); ! STD M->R lagged effect
RR2MRsd = RMLagEff * SQRT(8.3538) / SQRT(0.08077); ! STD R->M lagged effect

```

Note that this Model 3a is not fully equivalent to the wide single-level SEM version, which constrained the level-1 residual variances and same-occasion covariance to be equal over time (even though the age 12 versions are unpredicted).

3a SEM via ML and its lagged effects

	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
New/Additional Parameters				
MR2RRSD	-0.017	0.036	-0.465	0.642
RR2MRSD	0.079	0.037	2.131	0.033 *

3a MLM via Bayes and their residual lagged effects

	Estimate	Posterior S.D.	One-Tailed P-Value	Significance
New/Additional Parameters				
MR2RRSD	-0.017	0.037	0.303	
RR2MRSD	0.081	0.036	0.013	*

For example results sections using this data story and these types of models, please see [Hoffman \(2015 chapter 9\)](#) or [Hoffman & Hall \(2024\)](#)