

Example 2a: Predicting Binary Outcomes via SAS GLIMMIX, STATA LOGIT, and R GLM (complete syntax, data, and output available for SAS, STATA, and R electronically))

The (fake) data for this example demonstrating “logistic regression” (i.e., using logit link function and Bernoulli conditional response distribution) came from: <https://stats.idre.ucla.edu/sas/dae/ordinal-logistic-regression/>. In this example we will predict a student’s **binary decision** to apply to grad school (0=no, >0=pry) using undergraduate GPA (centered at 3.0), whether at least one of their parents has a graduate degree (0=no, 1=yes), and whether they attended a private university (0=no, 1=yes). I am using SAS GLIMMIX because it has more helpful options (even though these are not mixed-effects models). Further, because STATA LOGIT does not have denominator degrees of freedom, they were set to “none” in SAS GLIMMIX so that the SAS Wald test results (still labeled as t or F) will match those of STATA (using z or χ^2). I am using the R function GLM in all models for this example (also using z or χ^2).

SAS Syntax for Importing and Preparing Data for Analysis:

```
* Location for original files for these models - change this path;
* \\Client\ precedes path in Virtual Desktop outside H drive;
%LET filesave= C:\Dropbox\22_PSQF6270\PSQF6270_Example2ab;
LIBNAME filesave "&filesave.";

* Import data into work library, transform variables, apply labels;
DATA work.Example2; SET filesave.SAS_Example2;
* Create ID variable;
  PersonID=_N_;
* Rename 3-category outcome;
  apply3=apply;
* Create new binary outcome to demonstrate logistic regression;
  apply2=.; * New empty variable;
  IF apply3=0 THEN apply2=0;
  IF apply3>0 THEN apply2=1;
* Rename and center predictors;
  parD=pared; gpa3=gpa-3;
* Recode ref for public to create positive slope;
  priv=.;
  IF public=1 THEN priv=0;
  IF public=0 THEN priv=1;
* Label variables;
  LABEL apply3= "apply3: 0=Not, 1=Eh, 2=Very"
        apply2= "apply2: 0=No, 1=Pry"
        parD= "parD: Parent Has Graduate Degree (0=N,1=Y)"
        priv= "priv: Student Attends Private University (0=N,1=Y)"
        gpa3= "gpa3: Student GPA (0=3)";
* Filter to only cases complete on all variables to be used below;
  IF NMISS(apply3, parD, priv, gpa3)>0 THEN DELETE;
RUN;

DATA work.FakePeople; * Create new blank dataset;
* INPUT: list variables in order of entry;
  INPUT PersonID gpa3 parD priv;
* Enter data -- each row is a fake person for which to create a predicted outcome;
  DATALINES;
-99 -1 0 0
-99 0 0 0
-99 1 0 0
-99 -1 0 1
-99 0 0 1
-99 1 0 1
-99 -1 1 0
-99 0 1 0
-99 1 1 0
-99 -1 1 1
-99 0 1 1
-99 1 1 1
```

```
; RUN;
* Add fake people to real data;
DATA work.Example2; SET work.FakePeople work.Example2; RUN;
```

STATA Syntax for Importing and Preparing Data for Analysis:

```
// Defining global variable for file location to be replaced in code below
// \\Client\ precedes path in Virtual Desktop outside H drive
global filesave "C:\Dropbox\22_PSQF6270\PSQF6270_Example2ab"

// Open Example 2 STATA dataset and clear away any existing data
use "$filesave\STATA_Example2.dta", clear // Has converted all variables to lower-case

// Create ID variable
gen PersonID = _n
// Rename 3-category outcome
gen apply3=apply

// Create new binary outcome to demonstrate logistic regression
gen apply2=. // New empty variable
replace apply2=0 if apply==0
replace apply2=1 if apply>0

// Rename and center predictors
gen parD=pared
gen gpa3=gpa-3

// Recode ref for public to create positive slope
gen priv=.
replace priv=0 if public==1
replace priv=1 if public==0

// Label variables
label variable apply3 "apply3: 0=Not, 1=Eh, 2=Very"
label variable apply2 "apply2: 0=No, 1=Pry"
label variable parD "parD: Parent Has Graduate Degree (0=N,1=Y)"
label variable priv "priv: Student Attends Private University (0=N,1=Y)"
label variable gpa3 "gpa3: Student GPA (0=3)"

// Filter to only cases complete on all variables to be used below
egen nmiss=rowmiss(apply3 parD priv gpa3)
drop if nmiss>0
```

R Syntax for Importing and Preparing Data for Analysis:

```
# Define variables for working directory and data name
filesave = "C:\\Dropbox\\22_PSQF6270\\PSQF6270_Example2ab/"
filename = "SAS_Example2.sas7bdat"
setwd(dir=filesave)

# Import Example2 SAS data
Example2 = read_sas(data_file=paste0(filesave,filename))
# Convert to data frame without labels to use for analysis
Example2 = as.data.frame(Example2)

# Create ID variable
Example2$PersonID <- seq.int(nrow(Example2))

# Rename 3-category outcome
Example2$apply3=Example2$APPLY
```

```

# Create new binary outcome to demonstrate logistic regression
Example2$apply2 = NA # New empty variable
Example2$apply2[which(Example2$APPLY==0)]=0
Example2$apply2[which(Example2$APPLY>0)]=1

# Rename and center predictors
Example2$parD=Example2$PARED
Example2$gpa3=Example2$GPA-3

# Recode ref for public to create positive slope
Example2$priv = NA # New empty variable
Example2$priv[which(Example2$PUBLIC==1)]=0
Example2$priv[which(Example2$PUBLIC==0)]=1

# Label variables as comments only (not actually added to data)
#apply3= "apply3: 0=Not, 1=Eh, 2=Very"
#apply2= "apply2: 0=No, 1=Pry"
#parD= "parD: Parent Has Graduate Degree (0=N,1=Y)"
#priv= "priv: Student Attends Private University (0=N,1=Y)"
#gpa3= "gpa3: Student GPA (0=3)"

# Filter to only cases complete on all variables to be used below
Example2 = Example2[complete.cases(Example2[,5:9]),]

# Create fake people for use in generating predicted outcomes
FakeGpa3 = c(-1,0,1,-1,0,1,-1,0,1,-1,0,1)
FakeParD = c( 0,0,0, 0,0,0, 1,1,1, 1,1,1)
FakePriv = c( 0,0,0, 1,1,1, 0,0,0, 1,1,1)
# Create dataset using just-created columns and constants for other model variables
FP = data.frame(gpa3=FakeGpa3, parD=FakeParD, priv=FakePriv)

```

Syntax and SAS Output for Descriptive Statistics:

```

TITLE1 "SAS Descriptive Statistics";
PROC MEANS NDEC=2 DATA=work.Example2;
  VAR gpa3 parD priv apply2;
RUN;

display "STATA Descriptive Statistics"
summarize gpa3 parD priv apply2

print("R Descriptive Statistics")
describe(x=Example2[, c("gpa3","parD","priv","apply2")])

```

| Variable | Label | N | Mean | Std Dev | Minimum | Maximum |
|----------|--|-----|-------------|---------|---------|---------|
| gpa3 | gpa3: Student GPA (0=3) | 400 | -0.00 | 0.40 | -1.10 | 1.00 |
| parD | parD: Parent Has Graduate Degree (0=N,1=Y) | 400 | 0.16 | 0.36 | 0.00 | 1.00 |
| priv | priv: Student Attends Private University (0=N,1=Y) | 400 | 0.86 | 0.35 | 0.00 | 1.00 |
| apply2 | apply2: 0=No, 1=Pry | 400 | 0.45 | 0.50 | 0.00 | 1.00 |

So now we know that **55% of the respondents have apply2=0, and 45% have apply2=1**. This information will come in handy in making sure we understand which value our logistic regression models are predicting!

Empty Model Predicting Binary Apply 2:

$$\text{Logit}(\text{Apply2}_i = 1) = \beta_0 \rightarrow \text{Probability}(\text{Apply2}_i = 1) = \frac{\exp(\beta_0)}{1 + \exp(\beta_0)}$$

SAS Syntax and Output:

```
TITLE1 "SAS Empty Model Predicting Binary Apply2 using MSPL=ML";
PROC GLIMMIX DATA=work.Example2 NOCLPRINT GRADIENT METHOD=MSPL;
* DESCENDING means predict prob=1 rather than prob=0;
MODEL apply2 (DESCENDING) = / SOLUTION DDFM=NONE LINK=LOGIT DIST=BINARY;
* ILINK requests logit estimate to be transformed into probability;
ESTIMATE "Intercept for y=1" int 1 / ILINK;
RUN; TITLE1;
```

The GRADIENT option prints the partial first derivative with respect to each parameter, which is helpful in troubleshooting estimation problems.

The DESCENDING option makes the model predict the **probability of a 1** instead of the probability of a 0.

The GLIMMIX procedure is modeling the probability that apply2='1'.

| Iteration History | | | | | | |
|-------------------|----------|-------------|---------------------|------------|--------------|--|
| Iteration | Restarts | Evaluations | Objective Function | Change | Max Gradient | |
| 0 | 0 | 4 | 275.34490191 | . | 4.200299 | |
| 1 | 0 | 3 | 275.25552601 | 0.08937590 | 0.010221 | |
| 2 | 0 | 3 | 275.25552549 | 0.00000053 | 5.275E-8 | |

Convergence criterion (GCONV=1E-8) satisfied.

Hooray! The overall model converged! But we will still watch for crazy SEs and gradients ≠ 0 for problems with individual model parameters...

Fit Statistics

| | | |
|--------------------------|--------|--|
| -2 Log Likelihood | 550.51 | → -2LL value for model = 2*(-1)*275.256, last value above |
| AIC (smaller is better) | 552.51 | |
| AICC (smaller is better) | 552.52 | |
| BIC (smaller is better) | 556.50 | |
| CAIC (smaller is better) | 557.50 | |
| HQIC (smaller is better) | 554.09 | |
| Pearson Chi-Square | 400.00 | |
| Pearson Chi-Square / DF | 1.00 | → Indicates perfect distribution fit (always happens for binary) |

$$\text{Probability of } (\text{Apply2}_i = 1) = \frac{\exp(-0.2007)}{[1 + \exp(-0.2007)]} = 0.450$$

Parameter Estimates

| Effect | Estimate | Standard Error | DF | t Value | Pr > t | Gradient |
|-----------|----------|----------------|-------|---------|---------|----------|
| Intercept | -0.2007 | 0.1005 | Infty | -2.00 | 0.0465 | 5.275E-8 |

Estimates

| Label | Estimate = predicted logit | | | | | Mean = probability | |
|-------------------|----------------------------|----------------|-------|---------|---------|--------------------|---------------------------------|
| | Estimate | Standard Error | DF | t Value | Pr > t | Mean | Standard Error |
| Intercept for y=1 | -0.2007 | 0.1005 | Infty | -2.00 | 0.0465 | 0.4500 | 0.02487 → 0.450 matches! |

STATA Syntax and Output:

To get STATA to compute $-2LL$ for any model, add this code immediately after the procedure (e.g., logit here) that accesses the saved LL and multiplies by -2 :
// Print $-2LL$ for model
display " $-2LL=" e(11) * -2$ "

```
display "STATA Empty Model Predicting Binary Apply2"
logit apply2
estat ic, n(400) // AIC and BIC to match SAS
margins // Intercept in probability
```

```
Logistic regression
Number of obs = 400
LR chi2(0) = 0.00
Prob > chi2 = .
Pseudo R2 = 0.0000
```

STATA gives LL
(so you need to $*-2$)

Log likelihood = **-275.25553**

| apply2 | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] |
|--------|-------------------|-----------|-------|-------|-----------------------|
| _cons | -0.2006707 | .1005038 | -2.00 | 0.046 | -0.3976545 -0.0036869 |

→ In logits

```
. estat ic, n(400) // AIC and BIC to match SAS
Akaike's information criterion and Bayesian information criterion
```

| Model | Obs | ll(null) | ll(model) | df | AIC | BIC |
|-------|-----|-----------|-----------|----|----------|----------|
| . | 400 | -275.2555 | -275.2555 | 1 | 552.5111 | 556.5025 |

Note: N=400 used in calculating BIC.

```
. margins // Intercept in probability
Expression : Pr(apply2), predict()
```

| | Margin | Delta-method Std. Err. | z | P> z | [95% Conf. Interval] |
|-------|------------|------------------------|-------|-------|----------------------|
| _cons | .45 | .0248747 | 18.09 | 0.000 | .4012465 .4987535 |

→ In probability

R Syntax and Truncated Output:

```
print("R Empty Model Predicting Binary Apply2")
Model2Empty = glm(data=Example2, family=binomial, formula=apply2~1)
summary(Model2Empty); BIC(Model2Empty) # Get BIC too

print("Convert logits to probability to check interpretation")
Model2EmptyProb=1/(1+exp(-1*coefficients(Model2Empty))); Model2EmptyProb
```

```
Deviance Residuals:
Min      1Q  Median      3Q      Max
-1.0935 -1.0935 -1.0935  1.2637  1.2637
```

```
Coefficients:
            Estimate Std. Error z value Pr(>|z|)
(Intercept) -0.20067    0.10050 -1.9967  0.04586 → In logits
```

```
Null deviance: 550.511 on 399 degrees of freedom → Is empty model  $-2LL$ 
Residual deviance: 550.511 on 399 degrees of freedom → Is current model  $-2LL$ 
```

AIC: 552.511

```
> BIC(Model2Empty)
[1] 556.50252
```

```
> Model2EmptyProb
0.45 → In probability
```

Let's add three predictors, starting with main effects only...

$$\text{Logit}(\text{Apply}_{2i} = 1) = \beta_0 + \beta_1(\text{GPA}_i - 3) + \beta_2(\text{ParentGD}_i) + \beta_3(\text{Private}_i)$$

SAS Syntax and Output (condensed and re-arranged for convenience):

```
TITLE1 "SAS Main-Effects-Only Model Predicting Binary Apply2";
PROC GLIMMIX DATA=work.Example2 NOCLPRINT GRADIENT METHOD=MSPL;
MODEL apply2 (DESCENDING) = gpa3 parD priv
  / SOLUTION DDFM=NONE LINK=LOGIT DIST=BINARY ODDSRATIO(AT gpa3=0 LABEL);
CONTRAST "Multiv Wald Test of Model R2" gpa3 1, parD 1, priv 1 / CHISQ;
* Get predicted outcomes (duplicates estimates below with less effort);
OUTPUT OUT=work.Pred2Main PRED(NOILINK)=Ylogit STDERR(NOILINK)=YlogitSE
  PRED(ILINK)=Yprob STDERR(ILINK)=YprobSE;
* Predicted outcomes in logits and ILINK=probability;
ESTIMATE "Yhat: Ndeg Pub GPA=2" int 1 gpa3 -1 parD 0 priv 0 / ILINK;
ESTIMATE "Yhat: Ndeg Pub GPA=3" int 1 gpa3 0 parD 0 priv 0 / ILINK;
ESTIMATE "Yhat: Ndeg Pub GPA=4" int 1 gpa3 1 parD 0 priv 0 / ILINK;
ESTIMATE "Yhat: Ndeg Pri GPA=2" int 1 gpa3 -1 parD 0 priv 1 / ILINK;
ESTIMATE "Yhat: Ndeg Pri GPA=3" int 1 gpa3 0 parD 0 priv 1 / ILINK;
ESTIMATE "Yhat: Ndeg Pri GPA=4" int 1 gpa3 1 parD 0 priv 1 / ILINK;
ESTIMATE "Yhat: Ydeg Pub GPA=2" int 1 gpa3 -1 parD 1 priv 0 / ILINK;
ESTIMATE "Yhat: Ydeg Pub GPA=3" int 1 gpa3 0 parD 1 priv 0 / ILINK;
ESTIMATE "Yhat: Ydeg Pub GPA=4" int 1 gpa3 1 parD 1 priv 0 / ILINK;
ESTIMATE "Yhat: Ydeg Pri GPA=2" int 1 gpa3 -1 parD 1 priv 1 / ILINK;
ESTIMATE "Yhat: Ydeg Pri GPA=3" int 1 gpa3 0 parD 1 priv 1 / ILINK;
ESTIMATE "Yhat: Ydeg Pri GPA=4" int 1 gpa3 1 parD 1 priv 1 / ILINK;
ESTIMATE "GPA Slope" gpa3 1 / ILINK; * Example of non-sense ILINK for a slope;
RUN;
TITLE1 "Predicted Logits and Probabilities for Fake People";
PROC PRINT NOOBS DATA=work.Pred2Main; WHERE PersonID=-99;
  VAR gpa3 parD priv Ylogit--YprobSE;
RUN; TITLE1;
```

Fit Statistics

| | |
|-------------------------|---------------|
| -2 Log Likelihood | 529.92 |
| AIC (smaller is better) | 537.92 |
| BIC (smaller is better) | 553.89 |

Contrasts

| Label | Num DF | Den DF | Chi-Square | F Value | Pr > ChiSq | Pr > F |
|------------------------------|-----------|-----------|--------------|---------|------------|--------|
| Multiv Wald Test of Model R2 | 3 | Infy | 18.95 | 6.32 | 0.0003 | 0.0003 |

Parameter Estimates

| Effect | Estimate | Standard Error | DF | t Value | Pr > t | Gradient | |
|-----------|----------|-------------------|------|---------|---------|----------|--------------|
| Intercept | -0.5388 | 0.2874 | Infy | -1.87 | 0.0608 | -7.76E-6 | Beta0 |
| gpa3 | 0.5482 | 0.2724 | Infy | 2.01 | 0.0442 | -0.00002 | Beta1 |
| parD | 1.0596 | 0.2974 | Infy | 3.56 | 0.0004 | -0.00003 | Beta2 |
| priv | 0.2006 | 0.3053 | Infy | 0.66 | 0.5113 | -6.75E-6 | Beta3 |

Interpret each fixed effect...

Intercept:

GPA3:

parentGD:

private:

Odds Ratio Estimates → from exp(logit)

| Comparison | Estimate | DF | 95% Confidence | | |
|---------------------------------|----------|-------|----------------|-------|-------------------|
| | | | Limits | | |
| unit change of gpa3 from gpa3=0 | 1.730 | Infty | 1.014 | 2.951 | exp(Beta1) |
| unit change of parD from gpa3=0 | 2.885 | Infty | 1.611 | 5.168 | exp(Beta2) |
| unit change of priv from gpa3=0 | 1.222 | Infty | 0.672 | 2.223 | exp(Beta3) |

Effects of continuous variables are assessed as units offsets from the reference value.

Estimates

| Label | Estimate | Standard Error | DF | t Value | Pr > t | Mean | Standard |
|----------------------|----------|----------------|-------|---------|---------|---------------|----------------------|
| | | | | | | | Error |
| Yhat: Ndeg Pub GPA=2 | -1.0870 | 0.4312 | Infty | -2.52 | 0.0117 | 0.2522 | 0.08132 |
| Yhat: Ndeg Pub GPA=3 | -0.5388 | 0.2874 | Infty | -1.87 | 0.0608 | 0.3685 | 0.06688 |
| Yhat: Ndeg Pub GPA=4 | 0.009454 | 0.3574 | Infty | 0.03 | 0.9789 | 0.5024 | 0.08934 |
| Yhat: Ndeg Pri GPA=2 | -0.8865 | 0.2843 | Infty | -3.12 | 0.0018 | 0.2918 | 0.05877 |
| Yhat: Ndeg Pri GPA=3 | -0.3382 | 0.1187 | Infty | -2.85 | 0.0044 | 0.4162 | 0.02885 |
| Yhat: Ndeg Pri GPA=4 | 0.2100 | 0.3095 | Infty | 0.68 | 0.4974 | 0.5523 | 0.07652 |
| Yhat: Ydeg Pub GPA=2 | -0.02742 | 0.5123 | Infty | -0.05 | 0.9573 | 0.4931 | 0.1280 |
| Yhat: Ydeg Pub GPA=3 | 0.5208 | 0.3714 | Infty | 1.40 | 0.1608 | 0.6273 | 0.08683 |
| Yhat: Ydeg Pub GPA=4 | 1.0691 | 0.4024 | Infty | 2.66 | 0.0079 | 0.7444 | 0.07656 |
| Yhat: Ydeg Pri GPA=2 | 0.1731 | 0.4078 | Infty | 0.42 | 0.6712 | 0.5432 | 0.1012 |
| Yhat: Ydeg Pri GPA=3 | 0.7214 | 0.2801 | Infty | 2.58 | 0.0100 | 0.6729 | 0.06164 |
| Yhat: Ydeg Pri GPA=4 | 1.2696 | 0.3728 | Infty | 3.41 | 0.0007 | 0.7807 | 0.06383 |
| GPA Slope | 0.5482 | 0.2724 | Infty | 2.01 | 0.0442 | 0.6337 | 0.06324 Nope! |

The last line illustrates why you cannot “un-logit” a slope all the way back into probability... the difference between the intercepts per unit GPA in logits is a constant 0.5482, but the corresponding difference in probability is not constant between GPA units. Similarly, the difference between the groups is constant in logits, but it is NOT constant in probability—it depends where you are on the probability scale.

Predicted Logits and Probabilities for Fake People

| gpa3 | parD | priv | Ylogit | Yprob | YlogitSE | YprobSE |
|------|------|------|----------|---------|----------|---------|
| -1 | 0 | 0 | -1.08703 | 0.25218 | 0.43120 | 0.08132 |
| 0 | 0 | 0 | -0.53879 | 0.36847 | 0.28742 | 0.06688 |
| 1 | 0 | 0 | 0.00945 | 0.50236 | 0.35738 | 0.08934 |
| -1 | 0 | 1 | -0.88648 | 0.29184 | 0.28435 | 0.05877 |
| 0 | 0 | 1 | -0.33823 | 0.41624 | 0.11871 | 0.02885 |
| 1 | 0 | 1 | 0.21001 | 0.55231 | 0.30947 | 0.07652 |
| -1 | 1 | 0 | -0.02742 | 0.49314 | 0.51225 | 0.12804 |
| 0 | 1 | 0 | 0.52082 | 0.62734 | 0.37140 | 0.08683 |
| 1 | 1 | 0 | 1.06906 | 0.74442 | 0.40238 | 0.07656 |
| -1 | 1 | 1 | 0.17313 | 0.54318 | 0.40785 | 0.10120 |
| 0 | 1 | 1 | 0.72138 | 0.67291 | 0.28007 | 0.06164 |
| 1 | 1 | 1 | 1.26962 | 0.78068 | 0.37281 | 0.06383 |

STATA Syntax and Output (condensed and re-arranged for convenience):

```
display "STATA Main-Effects-Only Model Predicting Binary Apply2"
logit apply2 c.gpa3 c.parD c.priv
estat ic, n(400) // AIC and BIC to match SAS
test (c.gpa3=0) (c.parD=0) (c.priv=0) // Multiv Wald Test of Model R2
// For at, (from(by)to) for range of predictors
margins, at(c.gpa3=(-1(1)1) c.parD=(0(1)1) c.priv=(0(1)1)) predict(xb) // Yhat in logits
margins, at(c.gpa3=(-1(1)1) c.parD=(0(1)1) c.priv=(0(1)1)) // Yhat in probability

// Must re-estimate with 'or' added to first line to get odds ratios
display "STATA Main-Effects-Only Model Predicting Binary Apply2"
display "Get Odds Ratios Instead of Logit Fixed Effects"
logit apply2 c.gpa3 c.parD c.priv, or
```

Logistic regression

Number of obs = 400
 LR chi2(3) = 20.59 → From LRT
 Prob > chi2 = 0.0001
 Pseudo R2 = 0.0374

Log likelihood = -264.9624 * -2 = -2LL

| apply2 | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|--------|-----------|-----------|-------|-------|----------------------|--------------|
| gpa3 | .5482457 | .2724341 | 2.01 | 0.044 | .0142846 1.082207 | Beta1 |
| parD | 1.059612 | .2973854 | 3.56 | 0.000 | .4767471 1.642476 | Beta2 |
| priv | .2005571 | .3053354 | 0.66 | 0.511 | -.3978894 .7990035 | Beta3 |
| _cons | -.5387909 | .287416 | -1.87 | 0.061 | -1.102116 .0245341 | Beta0 |

. estat ic, n(400) // AIC and BIC to match SAS
 Akaike's information criterion and Bayesian information criterion

| Model | Obs | ll(null) | ll(model) | df | AIC | BIC |
|-------|-----|-----------|-----------|----|----------|----------|
| . | 400 | -275.2555 | -264.9624 | 4 | 537.9248 | 553.8907 |

Note: N=400 used in calculating BIC.

. test (c.gpa3=0) (c.parD=0) (c.priv=0) // **Multiv Wald Test of Model R2**
 (1) [apply2]gpa3 = 0
 (2) [apply2]parD = 0
 (3) [apply2]priv = 0

chi2(3) = **18.95**
 Prob > chi2 = 0.0003

Long table that labels the predictors for each predicted value was omitted

| Predicted Logit | | Delta-method | | | | |
|--------------------|-----------|--------------|-----------|-------|---------------------|----------------------|
| | | Margin | Std. Err. | z | P> z | [95% Conf. Interval] |
| 1 | -1.087037 | .4312041 | -2.52 | 0.012 | -1.932181 -.241892 | |
| 2 | -.8864795 | .2843452 | -3.12 | 0.002 | -1.443786 -.329173 | |
| 3 | -.0274248 | .5122537 | -0.05 | 0.957 | -1.031424 .9765739 | |
| 4 | .1731323 | .4078464 | 0.42 | 0.671 | -.6262321 .9724966 | |
| 5 | -.5387909 | .287416 | -1.87 | 0.061 | -1.102116 .0245341 | |
| 6 | -.3382338 | .1187131 | -2.85 | 0.004 | -.5709072 -.1055604 | |
| 7 | .5208208 | .3713982 | 1.40 | 0.161 | -.2071063 1.248748 | |
| 8 | .7213779 | .2800735 | 2.58 | 0.010 | .1724439 1.270312 | |
| 9 | .0094548 | .3573788 | 0.03 | 0.979 | -.6909948 .7099044 | |
| 10 | .2100119 | .3094739 | 0.68 | 0.497 | -.3965457 .8165695 | |
| 11 | 1.069067 | .4023806 | 2.66 | 0.008 | .280415 1.857718 | |
| 12 | 1.269624 | .372806 | 3.41 | 0.001 | .5389373 2.00031 | |

| Predicted Probability | | Delta-method | | | | |
|--------------------------|----------|--------------|-----------|-------|-------------------|----------------------|
| | | Margin | Std. Err. | z | P> z | [95% Conf. Interval] |
| 1 | .2521767 | .081318 | 3.10 | 0.002 | .0927963 .4115571 | |
| 2 | .2918369 | .0587651 | 4.97 | 0.000 | .1766594 .4070143 | |
| 3 | .4931442 | .1280393 | 3.85 | 0.000 | .2421917 .7440967 | |
| 4 | .5431753 | .1012013 | 5.37 | 0.000 | .3448243 .7415263 | |
| 5 | .3684689 | .0668816 | 5.51 | 0.000 | .2373834 .4995544 | |
| 6 | .4162386 | .0288454 | 14.43 | 0.000 | .3597027 .4727745 | |
| 7 | .6273397 | .0868272 | 7.23 | 0.000 | .4571615 .7975178 | |
| 8 | .6729104 | .0616447 | 10.92 | 0.000 | .5520889 .7937318 | |
| 9 | .5023637 | .0893427 | 5.62 | 0.000 | .3272552 .6774722 | |
| 10 | .5523108 | .0765216 | 7.22 | 0.000 | .4023312 .7022905 | |
| 11 | .7444194 | .0765566 | 9.72 | 0.000 | .5943712 .8944675 | |
| 12 | .7806783 | .0638317 | 12.23 | 0.000 | .6555704 .9057862 | |


```

Logistic regression (ODDS RATIOS VERSION)      Number of obs   =      400
                                                LR chi2(3)      =      20.59
                                                Prob > chi2     =      0.0001
Log likelihood = -264.9624                    Pseudo R2       =      0.0374
    
```

| apply2 | Odds Ratio | Std. Err. | z | P> z | [95% Conf. Interval] | | |
|--------|------------|-----------|-------|-------|----------------------|----------|-------------------|
| gpa3 | 1.730215 | .4713696 | 2.01 | 0.044 | 1.014387 | 2.951185 | exp(Beta1) |
| parD | 2.885251 | .8580314 | 3.56 | 0.000 | 1.610826 | 5.167952 | exp(Beta2) |
| priv | 1.222083 | .3731454 | 0.66 | 0.511 | .6717363 | 2.223324 | exp(Beta3) |
| _cons | .5834533 | .1676938 | -1.87 | 0.061 | .3321675 | 1.024838 | exp(Beta0) |

R Syntax and Output (condensed for convenience):

```

print("R Main-Effects-Only Model Predicting Binary Apply2")
Model2Main = glm(data=Example2, family=binomial,
                formula=apply2~1+gpa3+parD+priv
summary(Model2Main); BIC(Model2Main) # Get BIC too

print("Odds ratios -- 95% CIs will not match SAS,STATA exactly")
exp(cbind(OR = coef(Model2Main), confint(Model2Main)))

print("Multiv Wald Test of Model R2")
Main2R2 = glht(model=Model2Main, linfct=c("gpa3=0","parD=0","priv=0"))
summary(Main2R2, test=Ftest()) # Joint F-test instead of separate
summary(Main2R2, test=Chisqtest()) # Joint chi-square test

print("Yhat in logits for specific values of predictors")
Main2Logits = prediction(model=Model2Main, type="link",
                        at=list(gpa3=-1:1,parD=0:1,priv=0:1))
summary(Main2Logits)
print("Yhat in probability for specific values of predictors")
Main2Probs = prediction(model=Model2Main, type="response",
                       at=list(gpa3=-1:1,parD=0:1,priv=0:1))
summary(Main2Probs)
    
```

```

Deviance Residuals:
  Min       1Q   Median       3Q      Max
-1.67299 -1.03536 -0.89593  1.27410  1.55108
    
```

```

Coefficients:
              Estimate Std. Error z value Pr(>|z|)
(Intercept) -0.53879    0.28742 -1.8746 0.0608473 Beta0
gpa3         0.54825    0.27243  2.0124 0.0441780 Beta1
parD         1.05961    0.29739  3.5631 0.0003665 Beta2
priv         0.20056    0.30534  0.6568 0.5112826 Beta3
    
```

```

Null deviance: 550.511 on 399 degrees of freedom → Is empty model -2LL
Residual deviance: 529.925 on 396 degrees of freedom → Is current model -2LL
    
```

AIC: 537.925

```

> BIC(Model2Main)
[1] 553.89066
    
```

```

[1] "Odds ratios -- 95% CIs will not match SAS,STATA exactly"
              OR      2.5 %      97.5 %
(Intercept) 0.58345329 0.32846931 1.0192795 exp(Beta0)
gpa3        1.73021501 1.01774088 2.9675894 exp(Beta1)
parD        2.88525056 1.62940440 5.2560216 exp(Beta2)
priv        1.22208338 0.67428604 2.2423595 exp(Beta3)
    
```

[1] "Multiv Wald Test of Model R2"

General Linear Hypotheses
 Linear Hypotheses:
 Estimate
 gpa3 == 0 0.54825
 parD == 0 1.05961
 priv == 0 0.20056

Global Test:
 F DF1 DF2 Pr(>F)
 1 6.3157 3 396 0.00034174

Global Test:
 Chisq DF Pr(>Chisq)
 1 **18.947** 3 0.00028037

[1] "Yhat in logits for specific values of predictors"

| at(gpa3) | at(parD) | at(priv) | Prediction | SE | z | p | lower | upper |
|----------|----------|----------|------------|--------|----------|-----------|---------|----------|
| -1 | 0 | 0 | -1.087037 | 0.4312 | -2.52093 | 0.0117044 | -1.9322 | -0.24189 |
| 0 | 0 | 0 | -0.538791 | 0.2874 | -1.87460 | 0.0608473 | -1.1021 | 0.02453 |
| 1 | 0 | 0 | 0.009455 | 0.3574 | 0.02646 | 0.9788937 | -0.6910 | 0.70990 |
| -1 | 1 | 0 | -0.027425 | 0.5123 | -0.05354 | 0.9573036 | -1.0314 | 0.97657 |
| 0 | 1 | 0 | 0.520821 | 0.3714 | 1.40232 | 0.1608183 | -0.2071 | 1.24875 |
| 1 | 1 | 0 | 1.069067 | 0.4024 | 2.65685 | 0.0078873 | 0.2804 | 1.85772 |
| -1 | 0 | 1 | -0.886479 | 0.2843 | -3.11762 | 0.0018232 | -1.4438 | -0.32917 |
| 0 | 0 | 1 | -0.338234 | 0.1187 | -2.84917 | 0.0043833 | -0.5709 | -0.10556 |
| 1 | 0 | 1 | 0.210012 | 0.3095 | 0.67861 | 0.4973853 | -0.3965 | 0.81657 |
| -1 | 1 | 1 | 0.173132 | 0.4078 | 0.42450 | 0.6711986 | -0.6262 | 0.97250 |
| 0 | 1 | 1 | 0.721378 | 0.2801 | 2.57567 | 0.0100045 | 0.1724 | 1.27031 |
| 1 | 1 | 1 | 1.269624 | 0.3728 | 3.40559 | 0.0006602 | 0.5389 | 2.00031 |

[1] "Yhat in probability for specific values of predictors"

| at(gpa3) | at(parD) | at(priv) | Prediction | SE | z | p | lower | upper |
|----------|----------|----------|------------|---------|--------|-----------|--------|--------|
| -1 | 0 | 0 | 0.2522 | 0.08132 | 3.101 | 1.928e-03 | 0.0928 | 0.4116 |
| 0 | 0 | 0 | 0.3685 | 0.06688 | 5.509 | 3.603e-08 | 0.2374 | 0.4996 |
| 1 | 0 | 0 | 0.5024 | 0.08934 | 5.623 | 1.878e-08 | 0.3273 | 0.6775 |
| -1 | 1 | 0 | 0.4931 | 0.12804 | 3.852 | 1.174e-04 | 0.2422 | 0.7441 |
| 0 | 1 | 0 | 0.6273 | 0.08683 | 7.225 | 5.005e-13 | 0.4572 | 0.7975 |
| 1 | 1 | 0 | 0.7444 | 0.07656 | 9.724 | 2.388e-22 | 0.5944 | 0.8945 |
| -1 | 0 | 1 | 0.2918 | 0.05877 | 4.966 | 6.829e-07 | 0.1767 | 0.4070 |
| 0 | 0 | 1 | 0.4162 | 0.02885 | 14.430 | 3.351e-47 | 0.3597 | 0.4728 |
| 1 | 0 | 1 | 0.5523 | 0.07652 | 7.218 | 5.287e-13 | 0.4023 | 0.7023 |
| -1 | 1 | 1 | 0.5432 | 0.10120 | 5.367 | 7.994e-08 | 0.3448 | 0.7415 |
| 0 | 1 | 1 | 0.6729 | 0.06164 | 10.916 | 9.672e-28 | 0.5521 | 0.7937 |
| 1 | 1 | 1 | 0.7807 | 0.06383 | 12.230 | 2.143e-34 | 0.6556 | 0.9058 |

Let's examine a model with two new interactions...

$$\text{Logit}(\text{Apply}_i = 1) = \beta_0 + \beta_1(\text{GPA}_i - 3) + \beta_2(\text{ParentGD}_i) + \beta_3(\text{Private}_i) + \beta_4(\text{GPA}_i - 3)(\text{ParentGD}_i) + \beta_5(\text{GPA}_i - 3)(\text{Private}_i)$$

Model-implied GPA Slope: $\beta_1 + \beta_4(\text{ParentGD}_i) + \beta_5(\text{Private}_i)$

SAS Syntax and Output (condensed and re-arranged for convenience):

```
TITLE1 "SAS Interaction Model Predicting Binary Apply2";
PROC GLIMMIX DATA=work.Example2 NOCLPRINT GRADIENT METHOD=MSPL;
MODEL apply2 (DESCENDING) = gpa3 parD priv gpa3*parD gpa3*priv
  / SOLUTION DDFM=NONE LINK=LOGIT DIST=BINARY ODDSRATIO(AT gpa3=0 DIFF=ALL LABEL);
CONTRAST "Multiv Wald Test of R2 Change" gpa3*parD 1, gpa3*priv 1 / CHISQ;
* Get predicted outcomes (duplicates estimates below with less effort;
OUTPUT OUT=work.Pred2Int PRED(NOILINK)=Ylogit STDERR(NOILINK)=YlogitSE
  PRED(ILINK)=Yprob STDERR(ILINK)=YprobSE;
* Predicted outcomes in logits and ILINK=probability;
ESTIMATE "Yhat: Ndeg Pub GPA=2" int 1 gpa3 -1 parD 0 priv 0 gpa3*parD 0 gpa3*priv 0 / ILINK;
ESTIMATE "Yhat: Ndeg Pub GPA=3" int 1 gpa3 0 parD 0 priv 0 gpa3*parD 0 gpa3*priv 0 / ILINK;
ESTIMATE "Yhat: Ndeg Pub GPA=4" int 1 gpa3 1 parD 0 priv 0 gpa3*parD 0 gpa3*priv 0 / ILINK;
ESTIMATE "Yhat: Ndeg Pri GPA=2" int 1 gpa3 -1 parD 0 priv 1 gpa3*parD 0 gpa3*priv -1 / ILINK;
ESTIMATE "Yhat: Ndeg Pri GPA=3" int 1 gpa3 0 parD 0 priv 1 gpa3*parD 0 gpa3*priv 0 / ILINK;
ESTIMATE "Yhat: Ndeg Pri GPA=4" int 1 gpa3 1 parD 0 priv 1 gpa3*parD 0 gpa3*priv 1 / ILINK;
ESTIMATE "Yhat: Ydeg Pub GPA=2" int 1 gpa3 -1 parD 1 priv 0 gpa3*parD -1 gpa3*priv 0 / ILINK;
ESTIMATE "Yhat: Ydeg Pub GPA=3" int 1 gpa3 0 parD 1 priv 0 gpa3*parD 0 gpa3*priv 0 / ILINK;
ESTIMATE "Yhat: Ydeg Pub GPA=4" int 1 gpa3 1 parD 1 priv 0 gpa3*parD 1 gpa3*priv 0 / ILINK;
ESTIMATE "Yhat: Ydeg Pri GPA=2" int 1 gpa3 -1 parD 1 priv 1 gpa3*parD -1 gpa3*priv -1 / ILINK;
ESTIMATE "Yhat: Ydeg Pri GPA=3" int 1 gpa3 0 parD 1 priv 1 gpa3*parD 0 gpa3*priv 0 / ILINK;
ESTIMATE "Yhat: Ydeg Pri GPA=4" int 1 gpa3 1 parD 1 priv 1 gpa3*parD 1 gpa3*priv 1 / ILINK;
* Simple slopes for GPA by moderators (EXP to get odds ratios, add CL for CIs);
ESTIMATE "GPA Slope: Ndeg Pub" gpa3 1 gpa3*parD 0 gpa3*priv 0 / EXP;
ESTIMATE "GPA Slope: Ndeg Pri" gpa3 1 gpa3*parD 0 gpa3*priv 1 / EXP;
ESTIMATE "GPA Slope: Ydeg Pub" gpa3 1 gpa3*parD 1 gpa3*priv 0 / EXP;
ESTIMATE "GPA Slope: Ydeg Pri" gpa3 1 gpa3*parD 1 gpa3*priv 1 / EXP;
RUN;
TITLE1 "Predicted Logits and Probabilities for Fake People";
PROC PRINT NOOBS DATA=work.Pred2Int; WHERE PersonID=-99;
  VAR gpa3 parD priv Ylogit--YprobSE;
RUN; TITLE1;
```

| | | | | | | | |
|-------------------------------|----------|-----------|-------|-------------|---------|------------|---------------|
| Fit Statistics | | | | | | | |
| -2 Log Likelihood | | | | | | | 528.17 |
| AIC (smaller is better) | | | | | | | 540.17 |
| BIC (smaller is better) | | | | | | | 564.12 |
| | | Contrasts | | | | | |
| | | Num | Den | | | | |
| Label | | DF | DF | Chi-Square | F Value | Pr > ChiSq | Pr > F |
| Multiv Wald Test of R2 Change | | 2 | Infty | 1.70 | 0.85 | 0.4272 | 0.4272 |
| Parameter Estimates | | | | | | | |
| | | Standard | | | | | |
| Effect | Estimate | Error | DF | t Value | Pr > t | Gradient | |
| Intercept | -0.6594 | 0.3374 | Infty | -1.95 | 0.0506 | 3.151E-6 | Beta0 |
| gpa3 | 1.2564 | 0.7737 | Infty | 1.62 | 0.1044 | -0.00001 | Beta1 |
| parD | 1.1623 | 0.3197 | Infty | 3.64 | 0.0003 | -0.00002 | Beta2 |
| priv | 0.3198 | 0.3518 | Infty | 0.91 | 0.3634 | -3.2E-6 | Beta3 |
| gpa3*parD | -0.8359 | 0.7696 | Infty | -1.09 | 0.2774 | -1.44E-6 | Beta4 |
| gpa3*priv | -0.6821 | 0.8077 | Infty | -0.84 | 0.3984 | -9E-6 | Beta5 |

Interpret each simple effect and interaction...

GPA3:

parentGD:

private:

GPA3*parentGD:

GPA3*private:

| Odds Ratio Estimates | | | | 95% Confidence Limits | | This GPA odds ratio appears to be marginalized across interacting predictors |
|---------------------------------|----------|-------|----------|-----------------------|-------------------|--|
| Comparison | Estimate | DF | Estimate | Lower | Upper | |
| unit change of gpa3 from gpa3=0 | 1.714 | Infty | 0.998 | 2.944 | NOT = exp(Beta1)! | |
| unit change of parD from gpa3=0 | 3.197 | Infty | 1.709 | 5.983 | is = exp(Beta2) | |
| unit change of priv from gpa3=0 | 1.377 | Infty | 0.691 | 2.744 | is = exp(Beta3) | |

Effects of continuous variables are assessed as units offsets from the reference value.

| Label | Estimates | | DF | t Value | Pr > t | Standard Error | | Exponentiated Estimate |
|----------------------|-----------|--------|-------|---------|---------|----------------|---------|------------------------|
| | Estimate | Error | | | | Mean | Mean | |
| Yhat: Ndeg Pub GPA=2 | -1.9159 | 0.9908 | Infty | -1.93 | 0.0532 | 0.1283 | 0.1108 | . |
| Yhat: Ndeg Pub GPA=3 | -0.6594 | 0.3374 | Infty | -1.95 | 0.0506 | 0.3409 | 0.07580 | . |
| Yhat: Ndeg Pub GPA=4 | 0.5970 | 0.6656 | Infty | 0.90 | 0.3698 | 0.6450 | 0.1524 | . |
| Yhat: Ndeg Pri GPA=2 | -0.9140 | 0.3194 | Infty | -2.86 | 0.0042 | 0.2862 | 0.06525 | . |
| Yhat: Ndeg Pri GPA=3 | -0.3397 | 0.1187 | Infty | -2.86 | 0.0042 | 0.4159 | 0.02884 | . |
| Yhat: Ndeg Pri GPA=4 | 0.2347 | 0.3422 | Infty | 0.69 | 0.4929 | 0.5584 | 0.08439 | . |
| Yhat: Ydeg Pub GPA=2 | 0.08233 | 1.2300 | Infty | 0.07 | 0.9466 | 0.5206 | 0.3070 | . |
| Yhat: Ydeg Pub GPA=3 | 0.5029 | 0.4289 | Infty | 1.17 | 0.2410 | 0.6231 | 0.1007 | . |
| Yhat: Ydeg Pub GPA=4 | 0.9234 | 0.8068 | Infty | 1.14 | 0.2524 | 0.7157 | 0.1641 | . |
| Yhat: Ydeg Pri GPA=2 | 1.0842 | 0.8967 | Infty | 1.21 | 0.2266 | 0.7473 | 0.1693 | . |
| Yhat: Ydeg Pri GPA=3 | 0.8227 | 0.3034 | Infty | 2.71 | 0.0067 | 0.6948 | 0.06433 | . |
| Yhat: Ydeg Pri GPA=4 | 0.5611 | 0.6796 | Infty | 0.83 | 0.4090 | 0.6367 | 0.1572 | . |
| GPA Slope: Ndeg Pub | 1.2564 | 0.7737 | Infty | 1.62 | 0.1044 | Non-est | . | 3.5129 |
| GPA Slope: Ndeg Pri | 0.5743 | 0.3090 | Infty | 1.86 | 0.0631 | Non-est | . | 1.7759 |
| GPA Slope: Ydeg Pub | 0.4206 | 0.9476 | Infty | 0.44 | 0.6572 | Non-est | . | 1.5228 |
| GPA Slope: Ydeg Pri | -0.2616 | 0.7355 | Infty | -0.36 | 0.7221 | Non-est | . | 0.7699 |

Predicted Logits and Probabilities for Fake People

| gpa3 | parD | priv | Ylogit | Yprob | YlogitSE | YprobSE |
|------|------|------|----------|---------|----------|---------|
| -1 | 0 | 0 | -1.91589 | 0.12832 | 0.99084 | 0.11083 |
| 0 | 0 | 0 | -0.65944 | 0.34086 | 0.33740 | 0.07580 |
| 1 | 0 | 0 | 0.59700 | 0.64497 | 0.66562 | 0.15242 |
| -1 | 0 | 1 | -0.91401 | 0.28618 | 0.31943 | 0.06525 |
| 0 | 0 | 1 | -0.33967 | 0.41589 | 0.11874 | 0.02884 |
| 1 | 0 | 1 | 0.23466 | 0.55840 | 0.34224 | 0.08439 |
| -1 | 1 | 0 | 0.08233 | 0.52057 | 1.23003 | 0.30699 |
| 0 | 1 | 0 | 0.50289 | 0.62314 | 0.42891 | 0.10072 |
| 1 | 1 | 0 | 0.92345 | 0.71574 | 0.80679 | 0.16415 |
| -1 | 1 | 1 | 1.08421 | 0.74729 | 0.89666 | 0.16933 |
| 0 | 1 | 1 | 0.82266 | 0.69480 | 0.30336 | 0.06433 |
| 1 | 1 | 1 | 0.56111 | 0.63671 | 0.67958 | 0.15719 |

Here are the "simple" GPA odds ratios (for each interacting predictor value)


```
. estat ic, n(400) // AIC and BIC to match SAS
Akaike's information criterion and Bayesian information criterion
```

| Model | Obs | ll(null) | ll(model) | df | AIC | BIC |
|-------|-----|-----------|-----------|----|----------|----------|
| . | 400 | -275.2555 | -264.0868 | 6 | 540.1737 | 564.1225 |

Note: N=400 used in calculating BIC.

```
. test (c.gpa3#c.parD=0) (c.gpa3#c.priv=0) // Multiv Wald Test of R2 Change
```

- (1) [apply2]c.gpa3#c.parD = 0
- (2) [apply2]c.gpa3#c.priv = 0

```
chi2( 2) = 1.70
Prob > chi2 = 0.4272
```

Long table that labels the predictors for each predicted value was omitted

| Predicted Logit | Delta-method | | | | | |
|-----------------|--------------|-----------|-------|-------|----------------------|-----------|
| | Margin | Std. Err. | z | P> z | [95% Conf. Interval] | |
| 1 | -1.915897 | .9908357 | -1.93 | 0.053 | -3.857899 | .0261053 |
| 2 | -.9140066 | .3194316 | -2.86 | 0.004 | -1.540081 | -.2879321 |
| 3 | .0823253 | 1.230028 | 0.07 | 0.947 | -2.328486 | 2.493137 |
| 4 | 1.084216 | .8966626 | 1.21 | 0.227 | -.6732107 | 2.841642 |
| 5 | -.6594466 | .3373972 | -1.95 | 0.051 | -1.320733 | .0018397 |
| 6 | -.3396724 | .1187386 | -2.86 | 0.004 | -.5723958 | -.106949 |
| 7 | .502887 | .4289053 | 1.17 | 0.241 | -.3377518 | 1.343526 |
| 8 | .8226612 | .3033583 | 2.71 | 0.007 | .2280899 | 1.417233 |
| 9 | .5970039 | .6656165 | 0.90 | 0.370 | -.7075806 | 1.901588 |
| 10 | .2346618 | .3422373 | 0.69 | 0.493 | -.4361109 | .9054345 |
| 11 | .9234487 | .8067913 | 1.14 | 0.252 | -.6578331 | 2.504731 |
| 12 | .5611067 | .6795829 | 0.83 | 0.409 | -.7708513 | 1.893065 |

| Predicted Probability | Delta-method | | | | | |
|-----------------------|--------------|-----------|-------|-------|----------------------|----------|
| | Margin | Std. Err. | z | P> z | [95% Conf. Interval] | |
| 1 | .1283198 | .1108288 | 1.16 | 0.247 | -.0889006 | .3455402 |
| 2 | .2861807 | .0652539 | 4.39 | 0.000 | .1582854 | .414076 |
| 3 | .5205697 | .3069867 | 1.70 | 0.090 | -.0811131 | 1.122253 |
| 4 | .747291 | .1693322 | 4.41 | 0.000 | .4154059 | 1.079176 |
| 5 | .3408639 | .0758049 | 4.50 | 0.000 | .192289 | .4894389 |
| 6 | .4158891 | .0288446 | 14.42 | 0.000 | .3593547 | .4724235 |
| 7 | .6231376 | .1007229 | 6.19 | 0.000 | .4257243 | .8205508 |
| 8 | .694801 | .0643279 | 10.80 | 0.000 | .5687206 | .8208813 |
| 9 | .6449705 | .1524152 | 4.23 | 0.000 | .3462422 | .9436989 |
| 10 | .5583977 | .0843922 | 6.62 | 0.000 | .3929921 | .7238034 |
| 11 | .7157443 | .1641452 | 4.36 | 0.000 | .3940255 | 1.037463 |
| 12 | .6367086 | .1571948 | 4.05 | 0.000 | .3286123 | .9448048 |

```
. lincom c.gpa3*1 + c.gpa3#c.parD*0 + c.gpa3#c.priv*0 // GPA slope: Ndeg Pub
```

| apply2 | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|--------|---------|-----------|------|-------|----------------------|----------|
| (1) | 1.25645 | .7736689 | 1.62 | 0.104 | -.2599127 | 2.772814 |

```
. lincom c.gpa3*1 + c.gpa3#c.parD*0 + c.gpa3#c.priv*1 // GPA slope: Ndeg Pri
```

| apply2 | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|--------|----------|-----------|------|-------|----------------------|----------|
| (1) | .5743342 | .3090026 | 1.86 | 0.063 | -.0312998 | 1.179968 |

```
. lincom c.gpa3*1 + c.gpa3#c.parD*1 + c.gpa3#c.priv*0 // GPA slope: Ydeg Pub
-----+-----
apply2 |      Coef.   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
(1) |      .4205617   .9476188    0.44   0.657   -1.436737    2.27786
-----+-----
```

```
. lincom c.gpa3*1 + c.gpa3#c.parD*1 + c.gpa3#c.priv*1 // GPA slope: Ydeg Pri
-----+-----
apply2 |      Coef.   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
(1) |     -.2615545   .7354537   -0.36   0.722   -1.703017    1.179908
-----+-----
```

```
Logistic regression (ODDS RATIOS VERSION)      Number of obs   =      400
                                                LR chi2(5)      =      22.34
                                                Prob > chi2     =      0.0005
Log likelihood = -264.08684                    Pseudo R2       =      0.0406
-----+-----
```

```
apply2 | Odds Ratio   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
gpa3 |      3.51293   2.717845     1.62   0.104    .7711189    16.0036 exp(Beta1)
parD |      3.197386  1.022105     3.64   0.000    1.708805    5.982706 exp(Beta2)
priv |      1.376817  .4843848     0.91   0.363    .6908928    2.743732 exp(Beta3)
c.gpa3#c.parD | .433489   .3335948    -1.09   0.277    .0959245    1.958966 exp(Beta4)
c.gpa3#c.priv | .505546   .4083531    -0.84   0.398    .1038019    2.462158 exp(Beta5)
_cons | .5171374   .1744807    -1.95   0.051    .2669396    1.001841 exp(Beta0)
-----+-----
```

```
. lincom c.gpa3*1 + c.gpa3#c.parD*0 + c.gpa3#c.priv*0, or // GPA slope: Ndeg Pub
-----+-----
apply2 | Odds Ratio   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
(1) |      3.51293   2.717845     1.62   0.104    .7711189    16.0036
-----+-----
```

```
. lincom c.gpa3*1 + c.gpa3#c.parD*0 + c.gpa3#c.priv*1, or // GPA slope: Ndeg Pri
-----+-----
apply2 | Odds Ratio   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
(1) |      1.775948  .5487725     1.86   0.063    .969185     3.254271
-----+-----
```

```
. lincom c.gpa3*1 + c.gpa3#c.parD*1 + c.gpa3#c.priv*0, or // GPA slope: Ydeg Pub
-----+-----
apply2 | Odds Ratio   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
(1) |      1.522817  1.44305     0.44   0.657    .2377021    9.755784
-----+-----
```

```
. lincom c.gpa3*1 + c.gpa3#c.parD*1 + c.gpa3#c.priv*1, or // GPA slope: Ydeg Pri
-----+-----
apply2 | Odds Ratio   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
(1) |      .7698539  .5661919    -0.36   0.722    .1821331    3.254075
-----+-----
```

R Syntax and Output (condensed and re-arranged for convenience):

```

print("R Interaction Model Predicting Binary Apply2")
Model2Int = glm(data=Example2, family=binomial,
               formula=apply2~1+gpa3+parD+priv+gpa3:parD+gpa3:priv)
summary(Model2Int); BIC(Model2Int) # Get BIC too

print("Get odds ratios -- 95% CIs will not match SAS,STATA exactly")
exp(cbind(OR = coef(Model2Int), confint(Model2Int)))

print("Multiv Wald Test of R2 Change")
Int2CR2 = glht(model=Model2Int, linfct=c("gpa3:parD=0", "gpa3:priv=0"))
summary(Int2CR2, test=Chisqtest()) # Joint chi-square test

print("Yhat in logits for specific values of predictors")
Int2Logits = prediction(model=Model2Int, type="link",
                       at=list(gpa3=-1:1, parD=0:1, priv=0:1))
summary(Int2Logits)
print("Yhat in probability for specific values of predictors")
Int2Probs = prediction(model=Model2Int, type="response",
                       at=list(gpa3=-1:1, parD=0:1, priv=0:1))
summary(Int2Probs)

print("Simple slopes for GPA by moderators")
Int2Slopes = (summary(glht(model=Model2Int, linfct=rbind(
  "GPA Slope: Ndeg Pub" = c(0,1,0,0,0,0), # in order of fixed effects
  "GPA Slope: Ndeg Pri" = c(0,1,0,0,0,1),
  "GPA Slope: Ydeg Pub" = c(0,1,0,0,1,0),
  "GPA Slope: Ydeg Pri" = c(0,1,0,0,1,1))), test=adjusted("none")))
Int2Slopes; print("Odds ratios for simple slopes")
data.frame(OR=exp(Int2Slopes$test$coefficients))

```

Deviance Residuals:

| | Min | 1Q | Median | 3Q | Max |
|--|----------|----------|----------|---------|---------|
| | -1.62915 | -1.04275 | -0.87951 | 1.26659 | 1.78376 |

Coefficients:

| | Estimate | Std. Error | z value | Pr(> z) | |
|-------------|----------|------------|---------|-----------|--------------|
| (Intercept) | -0.65945 | 0.33740 | -1.9545 | 0.0506406 | Beta0 |
| gpa3 | 1.25645 | 0.77367 | 1.6240 | 0.1043722 | Beta1 |
| parD | 1.16233 | 0.31967 | 3.6361 | 0.0002768 | Beta2 |
| priv | 0.31977 | 0.35181 | 0.9089 | 0.3633882 | Beta3 |
| gpa3:parD | -0.83589 | 0.76956 | -1.0862 | 0.2773930 | Beta4 |
| gpa3:priv | -0.68212 | 0.80775 | -0.8445 | 0.3984075 | Beta5 |

Null deviance: 550.511 on 399 degrees of freedom → **Is empty model -2LL**
 Residual deviance: 528.174 on 394 degrees of freedom → **Is current model -2LL**

AIC: 540.174

> BIC(Model2Int)

[1] 564.12247

[1] "Get odds ratios -- 95% CIs will not match SAS,STATA exactly"

| | OR | 2.5 % | 97.5 % | |
|-------------|------------|-------------|-------------|-------------------|
| (Intercept) | 0.51713738 | 0.257469434 | 0.97967885 | exp(Beta0) |
| gpa3 | 3.51293103 | 0.813326564 | 17.54149233 | exp(Beta1) |
| parD | 3.19738627 | 1.740136424 | 6.15935755 | exp(Beta2) |
| priv | 1.37681699 | 0.704453915 | 2.83251229 | exp(Beta3) |
| gpa3:parD | 0.43348900 | 0.092168868 | 1.93822999 | exp(Beta4) |
| gpa3:priv | 0.50554586 | 0.095682692 | 2.34652419 | exp(Beta5) |

[1] "Multiv Wald Test of R2 Change"

General Linear Hypotheses
 Linear Hypotheses:
 Estimate
 gpa3:parD == 0 -0.83589
 gpa3:priv == 0 -0.68212

Global Test:
 Chisq DF Pr(>Chisq)
 1 **1.701** 2 0.42721

[1] "Yhat in logits for specific values of predictors"

| at(gpa3) | at(parD) | at(priv) | Prediction | SE | z | p | lower | upper |
|----------|----------|----------|------------|--------|----------|----------|---------|-----------|
| -1 | 0 | 0 | -1.91590 | 0.9908 | -1.93362 | 0.053160 | -3.8579 | 0.026104 |
| 0 | 0 | 0 | -0.65945 | 0.3374 | -1.95451 | 0.050641 | -1.3207 | 0.001839 |
| 1 | 0 | 0 | 0.59700 | 0.6656 | 0.89692 | 0.369762 | -0.7076 | 1.901588 |
| -1 | 1 | 0 | 0.08233 | 1.2300 | 0.06693 | 0.946638 | -2.3285 | 2.493135 |
| 0 | 1 | 0 | 0.50289 | 0.4289 | 1.17249 | 0.241000 | -0.3378 | 1.343526 |
| 1 | 1 | 0 | 0.92345 | 0.8068 | 1.14459 | 0.252377 | -0.6578 | 2.504730 |
| -1 | 0 | 1 | -0.91401 | 0.3194 | -2.86135 | 0.004218 | -1.5401 | -0.287932 |
| 0 | 0 | 1 | -0.33967 | 0.1187 | -2.86067 | 0.004227 | -0.5724 | -0.106949 |
| 1 | 0 | 1 | 0.23466 | 0.3422 | 0.68567 | 0.492921 | -0.4361 | 0.905434 |
| -1 | 1 | 1 | 1.08422 | 0.8967 | 1.20917 | 0.226598 | -0.6732 | 2.841641 |
| 0 | 1 | 1 | 0.82266 | 0.3034 | 2.71185 | 0.006691 | 0.2281 | 1.417232 |
| 1 | 1 | 1 | 0.56111 | 0.6796 | 0.82566 | 0.408995 | -0.7709 | 1.893064 |

[1] "Yhat in probability for specific values of predictors"

| at(gpa3) | at(parD) | at(priv) | Prediction | SE | z | p | lower | upper |
|----------|----------|----------|------------|---------|--------|-----------|----------|--------|
| -1 | 0 | 0 | 0.1283 | 0.11083 | 1.158 | 2.469e-01 | -0.08890 | 0.3455 |
| 0 | 0 | 0 | 0.3409 | 0.07580 | 4.497 | 6.905e-06 | 0.19229 | 0.4894 |
| 1 | 0 | 0 | 0.6450 | 0.15242 | 4.232 | 2.320e-05 | 0.34624 | 0.9437 |
| -1 | 1 | 0 | 0.5206 | 0.30699 | 1.696 | 8.993e-02 | -0.08111 | 1.1223 |
| 0 | 1 | 0 | 0.6231 | 0.10072 | 6.187 | 6.145e-10 | 0.42572 | 0.8206 |
| 1 | 1 | 0 | 0.7157 | 0.16415 | 4.360 | 1.298e-05 | 0.39403 | 1.0375 |
| -1 | 0 | 1 | 0.2862 | 0.06525 | 4.386 | 1.156e-05 | 0.15829 | 0.4141 |
| 0 | 0 | 1 | 0.4159 | 0.02884 | 14.418 | 3.972e-47 | 0.35935 | 0.4724 |
| 1 | 0 | 1 | 0.5584 | 0.08439 | 6.617 | 3.673e-11 | 0.39299 | 0.7238 |
| -1 | 1 | 1 | 0.7473 | 0.16933 | 4.413 | 1.019e-05 | 0.41541 | 1.0792 |
| 0 | 1 | 1 | 0.6948 | 0.06433 | 10.801 | 3.407e-27 | 0.56872 | 0.8209 |
| 1 | 1 | 1 | 0.6367 | 0.15719 | 4.050 | 5.112e-05 | 0.32861 | 0.9448 |

[1] "Simple slopes for GPA by moderators"

Simultaneous Tests for General Linear Hypotheses

| | Estimate | Std. Error | z value | Pr(> z) |
|--------------------------|----------|------------|---------|----------|
| GPA Slope: Ndeg Pub == 0 | 1.25645 | 0.77367 | 1.6240 | 0.10437 |
| GPA Slope: Ndeg Pri == 0 | 0.57433 | 0.30900 | 1.8587 | 0.06307 |
| GPA Slope: Ydeg Pub == 0 | 0.42056 | 0.94762 | 0.4438 | 0.65718 |
| GPA Slope: Ydeg Pri == 0 | -0.26155 | 0.73545 | -0.3556 | 0.72211 |

(Adjusted p values reported -- none method)

[1] "Odds ratios for simple slopes"

OR
 GPA Slope: Ndeg Pub 3.5129310
 GPA Slope: Ndeg Pri 1.7759477
 GPA Slope: Ydeg Pub 1.5228169
 GPA Slope: Ydeg Pri 0.7698538

Sample results section using SAS output (can replace χ^2 with F when using denominator DF):

We examined the extent to which a binary decision to apply to graduate school (55.00% 0=No, 45.00% 1=Yes) could be predicted by a student's undergraduate GPA ($M = 3.0$, $SD = 0.40$, range = 1.90 to 4.00), whether at least one of their parents has a graduate degree (15.75% 0=No, 84.25% 1=Yes), and whether they attended a private university (14.25% 0=No, 85.75% 1=Yes). Specifically, we estimated generalized linear models using maximum likelihood in SAS GLIMMIX, in which the conditional probability of applying to graduate school was predicted using a logit link function and a conditional Bernoulli distribution (i.e., logistic regressions). The GPA predictor was centered such that 0 indicated a GPA = 3. Effect sizes are provided using odds ratios (OR), in which OR values between 0 and 1 indicate negative effects, 1 indicates no effect, and values above 1 indicate positive effects. SAS ESTIMATE and LSMEANS statements were used to request simple effects and model-implied predicted outcomes.

The first model examined only the main effects of the three predictors, which together resulted in a significant model, $\chi^2(3) = 18.95$, $p = .0003$. GPA had a significantly positive effect, such that for every unit greater GPA, the logit of applying to graduate school was greater by 0.548 ($SE = 0.272$; $OR = 1.730$). Likewise, the logit of applying to graduate school was significantly greater for students for whom at least one parent had a graduate degree by 1.060 ($SE = 0.297$, $OR = 2.882$). However, the logit of applying to graduate school was nonsignificantly greater for students who attended a private university by 0.200 ($SE = 0.305$, $OR = 1.222$).

The second model then included two-way interactions of GPA with parent graduate degree and GPA with university type. This augmented model was not a significant improvement over the main effects model, $\chi^2(2) = 1.70$, $p = .427$. Neither individual interaction term was significant, nor was the simple slope of GPA significant in any of the four subgroups (i.e., formed by parent graduate degree by university type).

