

Longitudinal Invariance CFA (using MLR) Example in Mplus v. 8.4 (N = 151; 6 items over 3 occasions)

These data measuring a latent trait of social functioning were collected at a Psychiatric Rehabilitation center, in which time 1 was admittance, and times 2 and 3 were collected at six-month intervals. There were six subscales that were completed by the hospital staff for each patient, including positively-oriented measures of Social Competence, Social Interest, and Personal Neatness, and negatively-oriented measures of Psychoticism, Motor Retardation, and Irritability. The negatively-oriented subscales were reflected ($*-1$) prior to analysis. Initial models examined the fit of one-factor versus two-factor models given the two valences of the subscales, but the fit of the two-factor model was not a significant improvement, and thus a one-factor model with all six items was used here.

Mplus Code to Read in Data:

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TITLE:      Longitudinal Invariance
DATA:      FILE = Example7b.csv;           ! Don't need path if in same folder
              FORMAT = free; TYPE = INDIVIDUAL; ! Defaults

VARIABLE:  NAMES = ID v1T1 v1T2 v1T3 v2T1 v2T2 v2T3           ! Every variable in data set
              v3T1 v3T2 v3T3 v4T1 v4T2 v4T3
              v5T1 v5T2 v5T3 v6T1 v6T2 v6T3;

              USEVARIABLES = v1T1 v1T2 v1T3 v2T1 v2T2 v2T3       ! Every variable in MODEL
              v3T1 v3T2 v3T3 v4T1 v4T2 v4T3
              v5T1 v5T2 v5T3 v6T1 v6T2 v6T3;

              MISSING = ALL (9999);          ! Specify all missing values
              IDVARIABLE = ID;              ! Specify person ID variable

! Reverse-coding items so that higher = better
DEFINE:    v4T1 = v4T1*(-1);
              v4T2 = v4T2*(-1);
              v4T3 = v4T3*(-1);
              v5T1 = v5T1*(-1);
              v5T2 = v5T2*(-1);
              v5T3 = v5T3*(-1);
              v6T1 = v6T1*(-1);
              v6T2 = v6T2*(-1);
              v6T3 = v6T3*(-1);

ANALYSIS:  ESTIMATOR = MLR;              ! For continuous items whose residuals may not be normal

OUTPUT:   MODINDICES(3.84);             ! For modification indices of p<.05 for DF=1
              STDYX RESIDUAL;             ! Fully standardized solution, local model fit

MODEL:    ! Model syntax goes here, to be changed for each model

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Model 1. Configural Longitudinal Invariance Model (everything separate across time)

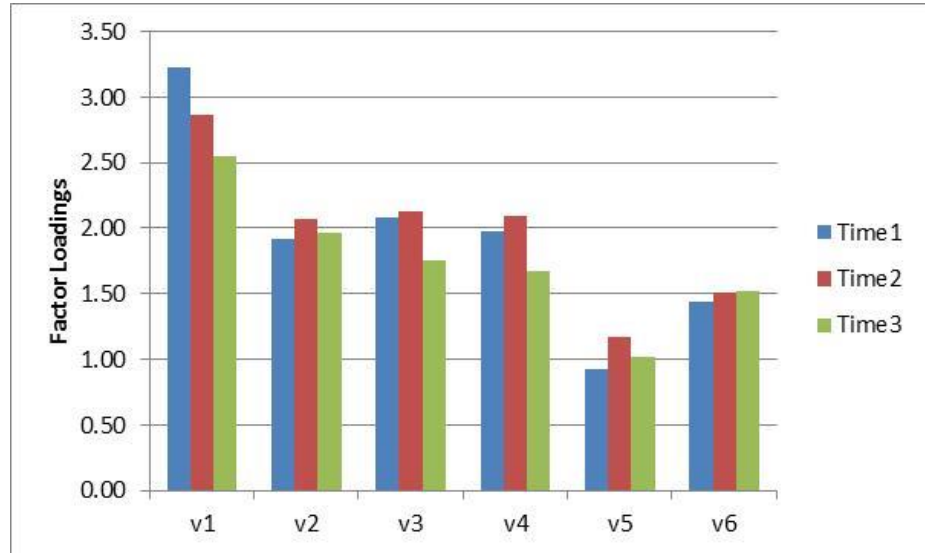
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MODEL:
!!!! Model 1: Configural Longitudinal Invariance

! Factor loadings all freely estimated, not labeled
Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1*;
Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2*;
Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3*;
! Item intercepts all freely estimated, not labeled
[v1T1* v1T2* v1T3*]; [v2T1* v2T2* v2T3*];
[v3T1* v3T2* v3T3*]; [v4T1* v4T2* v4T3*];
[v5T1* v5T2* v5T3*]; [v6T1* v6T2* v6T3*];
! Residual variances all freely estimated, not labeled
v1T1* v1T2* v1T3*; v2T1* v2T2* v2T3*;
v3T1* v3T2* v3T3*; v4T1* v4T2* v4T3*;
v5T1* v5T2* v5T3*; v6T1* v6T2* v6T3*;
! Factor variances all fixed=1 for identification
Time1@1 Time2@1 Time3@1;
! Factor means all fixed=0 for identification
[Time1@0 Time2@0 Time3@0];
! Factor covariances all freely estimated
Time1 Time2 Time3 WITH Time1* Time2* Time3*;
! Residual covariances estimated for same item across time
v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3*;
v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3*;
v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3*;
v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3*;
v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3*;
v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3*;
    
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MODEL FIT INFORMATION

Number of Free Parameters	75
Loglikelihood	
H0 Value	-4430.302
H0 Scaling Correction Factor for MLR	1.4617
H1 Value	-4284.045
H1 Scaling Correction Factor for MLR	1.2029
Information Criteria	
Akaike (AIC)	9010.604
Bayesian (BIC)	9236.900
Sample-Size Adjusted BIC	8999.533
(n* = (n + 2) / 24)	
Chi-Square Test of Model Fit	
Value	283.247*
Degrees of Freedom	114
P-Value	0.0000
Scaling Correction Factor for MLR	1.0327
RMSEA (Root Mean Square Error Of Approximation)	
Estimate	0.099
90 Percent C.I.	0.085 0.114
Probability RMSEA <= .05	0.000
CFI/TLI	
CFI	0.903
TLI	0.870
Chi-Square Test of Model Fit for the Baseline Model	
Value	1896.788
Degrees of Freedom	153
P-Value	0.0000
SRMR (Standardized Root Mean Square Residual)	
Value	0.089



Although the fit is not great, attempts to improve it logically were unsuccessful, so we proceed from here with this as the configural invariance model. The plot of factor loadings on the left foreshadows what will happen when testing metric invariance next...

UNSTANDARDIZED MODEL RESULTS - NOTE ALL MEASUREMENT PARAMETERS DIFFER ACROSS TIME

Factor Loadings and Variances					Means and Intercepts					
	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value	
FACTOR LOADINGS PER OCCASION					Means (FACTOR MEANS FIXED=0 FOR IDENTIFICATION)					
TIME1	BY				TIME1	0.000	0.000	999.000	999.000	
V1T1		3.222	0.267	12.063	0.000	TIME2	0.000	0.000	999.000	
V2T1		1.915	0.274	6.997	0.000	TIME3	0.000	0.000	999.000	
V3T1		2.080	0.209	9.956	0.000					
V4T1		1.975	0.271	7.298	0.000	Intercepts (ARE EXPECTED OUTCOME WHEN FACTOR IS AT 0)				
V5T1		0.931	0.148	6.281	0.000	V1T1	16.077	0.276	58.220	0.000
V6T1		1.441	0.119	12.101	0.000	V1T2	17.226	0.245	70.294	0.000
						V1T3	17.756	0.220	80.620	0.000
TIME2	BY					V2T1	8.672	0.298	29.132	0.000
V1T2		2.863	0.305	9.372	0.000	V2T2	9.981	0.263	37.921	0.000
V2T2		2.072	0.197	10.490	0.000	V2T3	10.442	0.281	37.204	0.000
V3T2		2.133	0.185	11.509	0.000	V3T1	11.970	0.225	53.108	0.000
V4T2		2.098	0.322	6.514	0.000	V3T2	12.467	0.218	57.264	0.000
V5T2		1.175	0.239	4.921	0.000	V3T3	13.029	0.213	61.157	0.000
V6T2		1.512	0.129	11.749	0.000	V4T1	-3.037	0.271	-11.216	0.000
						V4T2	-3.211	0.260	-12.349	0.000
TIME3	BY					V4T3	-2.738	0.249	-11.014	0.000
V1T3		2.550	0.288	8.865	0.000	V5T1	-1.283	0.138	-9.293	0.000
V2T3		1.961	0.230	8.539	0.000	V5T2	-1.664	0.200	-8.338	0.000
V3T3		1.751	0.210	8.323	0.000	V5T3	-1.247	0.166	-7.511	0.000
V4T3		1.678	0.260	6.448	0.000	V6T1	-2.871	0.164	-17.508	0.000
V5T3		1.021	0.170	6.012	0.000	V6T2	-2.413	0.158	-15.316	0.000
V6T3		1.523	0.159	9.574	0.000	V6T3	-2.075	0.152	-13.618	0.000
TIME1	WITH (ESTIMATED FACTOR COVARIANCES)					Residual Variances (VARIANCE PER ITEM THAT IS NOT THE FACTOR)				
TIME2		0.786	0.042	18.827	0.000	V1T1	0.241	0.395	0.610	0.542
TIME3		0.707	0.084	8.456	0.000	V1T2	0.511	0.268	1.907	0.056
						V1T3	0.523	0.349	1.497	0.134
TIME2	WITH					V2T1	8.672	1.022	8.484	0.000
TIME3		0.671	0.089	7.532	0.000	V2T2	5.913	0.617	9.581	0.000
						V2T3	5.142	0.806	6.379	0.000
*** Residual covariances among same item across time ***						V3T1	2.413	0.398	6.067	0.000
V1T1	WITH					V3T2	2.202	0.369	5.972	0.000
V1T2		-0.214	0.250	-0.855	0.393	V3T3	2.381	0.430	5.542	0.000
V1T3		-0.004	0.247	-0.016	0.987	V4T1	7.199	1.036	6.950	0.000
						V4T2	6.765	0.990	6.834	0.000
V1T2	WITH					V4T3	6.456	1.078	5.988	0.000
V1T3		0.113	0.231	0.488	0.626	V5T1	1.824	0.446	4.093	0.000
.....						V5T2	4.676	1.439	3.251	0.001
						V5T3	2.944	0.752	3.913	0.000
Variances (FACTOR VARIANCES FIXED=1 FOR IDENTIFICATION)						V6T1	1.694	0.243	6.974	0.000
TIME1		1.000	0.000	999.000	999.000	V6T2	1.103	0.166	6.643	0.000
TIME2		1.000	0.000	999.000	999.000	V6T3	0.751	0.162	4.630	0.000
TIME3		1.000	0.000	999.000	999.000					

Model 2a. Metric Invariance Model (ALL loadings held equal across time – identified model using Time1 Factor Variance = 1)

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MODEL:
!!!! Model 2a: Metric Longitudinal Invariance

! Factor loadings NOW CONSTRAINED EQUAL ACROSS TIME
Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1* (L1-L6);
Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2* (L1-L6);
Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3* (L1-L6);
! Item intercepts all freely estimated, not labeled
[v1T1* v1T2* v1T3*]; [v2T1* v2T2* v2T3*];
[v3T1* v3T2* v3T3*]; [v4T1* v4T2* v4T3*];
[v5T1* v5T2* v5T3*]; [v6T1* v6T2* v6T3*];
! Residual variances all freely estimated, not labeled
v1T1* v1T2* v1T3* ; v2T1* v2T2* v2T3* ;
v3T1* v3T2* v3T3* ; v4T1* v4T2* v4T3* ;
v5T1* v5T2* v5T3* ; v6T1* v6T2* v6T3* ;
! Factor variance AT TIME 1 fixed=1 for identification
Time1@1 Time2* Time3* ;
! Factor means all fixed=0 for identification
[Time1@0 Time2@0 Time3@0];
! Factor covariances all freely estimated
Time1 Time2 Time3 WITH Time1* Time2* Time3* ;
! Residual covariances estimated for same item across time
v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3* ;
v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3* ;
v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3* ;
v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3* ;
v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3* ;
v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3* ;
    
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Does the metric model (2a) fit worse than the configural model (1)?
 Yes, $-2\Delta LL(df=10) = 19.14, p=.04$

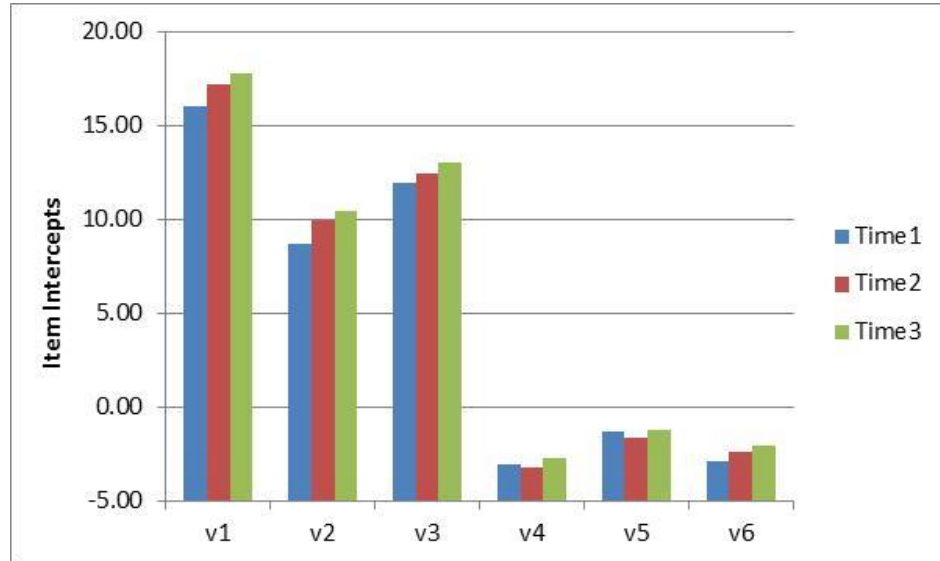
MODEL FIT INFORMATION	
Number of Free Parameters	65
Loglikelihood	
H0 Value	-4442.401
H0 Scaling Correction Factor for MLR	1.4921
H1 Value	-4284.045
H1 Scaling Correction Factor for MLR	1.2029
Information Criteria	
Akaike (AIC)	9014.803
Bayesian (BIC)	9210.926
Sample-Size Adjusted BIC	9005.208
(n* = (n + 2) / 24)	
Chi-Square Test of Model Fit	
Value	301.234*
Degrees of Freedom	124
P-Value	0.0000
Scaling Correction Factor for MLR	1.0514
RMSEA (Root Mean Square Error Of Approximation)	
Estimate	0.097
90 Percent C.I.	0.083 0.111
Probability RMSEA <= .05	0.000
CFI/TLI	
CFI	0.898
TLI	0.875
SRMR (Standardized Root Mean Square Residual)	
Value	0.094
MODEL MODIFICATION INDICES (relevant for testing invariance)	
BY Statements	
	M.I. E.P.C. Std E.P.C. StdYX E.P.C.
BY Statements	
TIME1 BY V1T1	10.377 0.182 0.182 0.058
TIME1 BY V5T1	6.062 -0.054 -0.054 -0.033
TIME3 BY V6T3	7.603 0.201 0.175 0.105
Modification indices suggest that freeing the loading for v1 at Time1 would help, and that matches our observations, so let's try that.	

Model 2b. Partial Metric Invariance Model with loading for v1 at Time 1 free

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MODEL:
! Model 2b: Partial Metric Invariance without v1T1

! Factor loadings NOW CONSTRAINED EQUAL ACROSS TIME EXCEPT v1T1
Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1* (L1a L2-L6);
Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2* (L1-L6);
Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3* (L1-L6);
! Item intercepts all freely estimated, not labeled
[v1T1* v1T2* v1T3*]; [v2T1* v2T2* v2T3*];
[v3T1* v3T2* v3T3*]; [v4T1* v4T2* v4T3*];
[v5T1* v5T2* v5T3*]; [v6T1* v6T2* v6T3*];
! Residual variances all freely estimated, not labeled
v1T1* v1T2* v1T3* ; v2T1* v2T2* v2T3* ;
v3T1* v3T2* v3T3* ; v4T1* v4T2* v4T3* ;
v5T1* v5T2* v5T3* ; v6T1* v6T2* v6T3* ;
! Factor variance AT TIME 1 fixed=1 for identification
Time1@1 Time2* Time3* ;
! Factor means all fixed=0 for identification
[Time1@0 Time2@0 Time3@0];
! Factor covariances all freely estimated
Time1 Time2 Time3 WITH Time1* Time2* Time3* ;
! Residual covariances estimated for same item across time
v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3* ;
v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3* ;
v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3* ;
v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3* ;
v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3* ;
v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3* ;
    
```



MODEL FIT INFORMATION

Number of Free Parameters	66
Loglikelihood	
H0 Value	-4435.669
H0 Scaling Correction Factor for MLR	1.4980
H1 Value	-4284.045
H1 Scaling Correction Factor for MLR	1.2029
Information Criteria	
Akaike (AIC)	9003.337
Bayesian (BIC)	9202.478
Sample-Size Adjusted BIC	8993.595
(n* = (n + 2) / 24)	
Chi-Square Test of Model Fit	
Value	290.301*
Degrees of Freedom	123
P-Value	0.0000
Scaling Correction Factor for MLR	1.0446
RMSEA (Root Mean Square Error Of Approximation)	
Estimate	0.095
90 Percent C.I.	0.081 0.109
Probability RMSEA <= .05	0.000
CFI/TLI	
CFI	0.904
TLI	0.881
SRMR (Standardized Root Mean Square Residual)	
Value	0.091

Does the partial metric model (2b) fit better than the full metric model (2a)? Yes, $-2\Delta LL(df=1) = 7.16, p < .01$

Does the partial metric model (2b) fit worse than the configural model (1)? No, $-2\Delta LL(df=9) = 8.98, p = .44$

No large invariance-related modification indices were found, so we'll call it good! Onto the next model! The plot of intercepts on the left foreshadow what we will find with testing scalar invariance...

2b UNSTANDARDIZED PARTIAL METRIC MODEL RESULTS - ALL FACTOR LOADINGS ARE HELD EQUAL EXCEPT v1T1

		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value					
		Estimate	S.E.	Est./S.E.	P-Value					
TIME1	BY					Means (FACTOR MEANS FIXED=0 FOR IDENTIFICATION)				
	v1T1	3.233	0.261	12.362	0.000	TIME1	0.000	0.000	999.000	999.000
	V2T1	1.950	0.201	9.706	0.000	TIME2	0.000	0.000	999.000	999.000
	V3T1	1.967	0.198	9.910	0.000	TIME3	0.000	0.000	999.000	999.000
	V4T1	1.899	0.224	8.481	0.000	Intercepts - SCALED SO SHOULD BE EQUAL ACROSS TIME				
	V5T1	0.968	0.137	7.055	0.000	V1T1	16.078	0.276	58.267	0.000
	V6T1	1.476	0.131	11.247	0.000	V1T2	17.225	0.245	70.282	0.000
TIME2	BY					V1T3	17.756	0.222	80.036	0.000
	v1T2	2.644	0.234	11.315	0.000	V2T1	8.672	0.298	29.071	0.000
	V2T2	1.950	0.201	9.706	0.000	V2T2	9.980	0.264	37.872	0.000
	V3T2	1.967	0.198	9.910	0.000	V2T3	10.434	0.280	37.245	0.000
	V4T2	1.899	0.224	8.481	0.000	V3T1	11.978	0.225	53.192	0.000
	V5T2	0.968	0.137	7.055	0.000	V3T2	12.468	0.217	57.325	0.000
	V6T2	1.476	0.131	11.247	0.000	V3T3	13.041	0.212	61.441	0.000
TIME3	BY					V4T1	-3.034	0.267	-11.343	0.000
	v1T3	2.644	0.234	11.315	0.000	V4T2	-3.210	0.260	-12.365	0.000
	V2T3	1.950	0.201	9.706	0.000	V4T3	-2.720	0.254	-10.720	0.000
	V3T3	1.967	0.198	9.910	0.000	V5T1	-1.288	0.137	-9.377	0.000
	V4T3	1.899	0.224	8.481	0.000	V5T2	-1.663	0.199	-8.340	0.000
	V5T3	0.968	0.137	7.055	0.000	V5T3	-1.246	0.169	-7.373	0.000
	V6T3	1.476	0.131	11.247	0.000	V6T1	-2.871	0.164	-17.506	0.000
TIME1	WITH					V6T2	-2.414	0.158	-15.319	0.000
	TIME2	0.847	0.078	10.837	0.000	V6T3	-2.087	0.154	-13.571	0.000
	TIME3	0.682	0.124	5.508	0.000	Residual Variances - ITEM VARIANCE THAT IS NOT THE FACTOR				
TIME2	WITH					V1T1	0.170	0.374	0.454	0.650
	TIME3	0.699	0.128	5.473	0.000	V1T2	0.548	0.265	2.070	0.038
*** Residual covariances among same item across time ***										
V1T1	WITH					V1T3	0.509	0.314	1.618	0.106
	V1T2	-0.225	0.249	-0.904	0.366	V2T1	8.702	1.026	8.483	0.000
	V1T3	-0.012	0.236	-0.049	0.961	V2T2	5.895	0.605	9.746	0.000
V1T2	WITH					V2T3	5.177	0.795	6.514	0.000
	V1T3	0.132	0.230	0.573	0.566	V3T1	2.502	0.386	6.484	0.000
.....						V3T2	2.178	0.352	6.183	0.000
Variances (FACTOR VARIANCE AT TIME1=1 FOR IDENTIFICATION)										
	TIME1	1.000	0.000	999.000	999.000	V3T3	2.309	0.416	5.548	0.000
	TIME2	1.162	0.185	6.270	0.000	V4T1	7.172	1.021	7.021	0.000
	TIME3	0.941	0.157	5.999	0.000	V4T2	6.759	0.967	6.990	0.000
						V4T3	6.613	1.128	5.860	0.000
						V5T1	1.829	0.443	4.131	0.000
						V5T2	4.678	1.430	3.272	0.001
						V5T3	2.944	0.760	3.872	0.000
						V6T1	1.707	0.242	7.059	0.000
						V6T2	1.090	0.165	6.599	0.000
						V6T3	0.784	0.170	4.618	0.000

Model 3a. Scalar Invariance Model (all intercepts held equal across over time except v1T1); identified by Time1 mean=0

<p>MODEL:</p> <p>! Model 3a: Full Scalar Invariance without v1T1</p> <p>! Factor loadings NOW CONSTRAINED EQUAL ACROSS TIME EXCEPT v1T1 Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1* (L1a L2-L6); Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2* (L1-L6); Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3* (L1-L6);</p> <p>! Item intercepts NOW CONSTRAINED EQUAL ACROSS TIME EXCEPT v1T1 [v1T1*]; [v1T2* v1T3*] (I1); ! 3a: I1 applies only to 2 and 3 [v2T1* v2T2* v2T3*] (I2); [v3T1* v3T2* v3T3*] (I3); [v4T1* v4T2* v4T3*] (I4); [v5T1* v5T2* v5T3*] (I5); [v6T1* v6T2* v6T3*] (I6);</p> <p>! Residual variances all freely estimated, not labeled v1T1* v1T2* v1T3*; v2T1* v2T2* v2T3*; v3T1* v3T2* v3T3*; v4T1* v4T2* v4T3*; v5T1* v5T2* v5T3*; v6T1* v6T2* v6T3*;</p> <p>! Factor variance AT TIME 1 fixed=1 for identification Time1@1 Time2* Time3*;</p> <p>! Factor mean AT TIME 1 fixed=0 for identification [Time1@0 Time2* Time3*];</p> <p>! Factor covariances all freely estimated Time1 Time2 Time3 WITH Time1* Time2* Time3*;</p> <p>! Residual covariances estimated for same item across time v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3*; v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3*; v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3*; v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3*; v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3*; v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3*;</p> <p>Does the full scalar model (3a) fit worse than the partial metric model (2b)? Yes, $-2\Delta LL(df=9) = 55.13, p < .01$</p> <p>Modification indices suggest that freeing these intercepts would help, so let's try v5T1 first (biggest χ^2 change suggested).</p>	<p>MODEL FIT INFORMATION</p> <p>Number of Free Parameters 57</p> <p>Loglikelihood</p> <p>H0 Value -4461.842 H0 Scaling Correction Factor 1.5846 for MLR H1 Value -4284.045 H1 Scaling Correction Factor 1.2029 for MLR</p> <p>Information Criteria</p> <p>Akaike (AIC) 9037.685 Bayesian (BIC) 9209.670 Sample-Size Adjusted BIC 9029.271 (n* = (n + 2) / 24)</p> <p>Chi-Square Test of Model Fit</p> <p>Value 342.530* Degrees of Freedom 132 P-Value 0.0000 Scaling Correction Factor 1.0381 for MLR</p> <p>RMSEA (Root Mean Square Error Of Approximation)</p> <p>Estimate 0.103 90 Percent C.I. 0.089 0.116 Probability RMSEA <= .05 0.000</p> <p>CFI/TLI</p> <p>CFI 0.879 TLI 0.860</p> <p>SRMR (Standardized Root Mean Square Residual)</p> <p>Value 0.093</p> <p>MODEL MODIFICATION INDICES (relevant for invariance testing) Means/Intercepts/Thresholds</p> <table border="1"> <thead> <tr> <th></th> <th></th> <th>M.I.</th> <th>E.P.C.</th> <th>Std E.P.C.</th> <th>StdYX E.P.C.</th> </tr> </thead> <tbody> <tr> <td>[</td> <td>V2T1</td> <td>]</td> <td>14.761</td> <td>-0.696</td> <td>-0.696</td> <td>-0.189</td> </tr> <tr> <td>[</td> <td>V2T2</td> <td>]</td> <td>5.578</td> <td>0.307</td> <td>0.307</td> <td>0.094</td> </tr> <tr> <td>[</td> <td>V4T1</td> <td>]</td> <td>10.400</td> <td>0.366</td> <td>0.366</td> <td>0.113</td> </tr> <tr> <td>[</td> <td>V4T2</td> <td>]</td> <td>5.167</td> <td>-0.271</td> <td>-0.271</td> <td>-0.084</td> </tr> <tr> <td>[</td> <td>V5T1</td> <td>]</td> <td>20.890</td> <td>-0.027</td> <td>-0.027</td> <td>-0.017</td> </tr> <tr> <td>[</td> <td>V5T2</td> <td>]</td> <td>14.191</td> <td>-0.596</td> <td>-0.596</td> <td>-0.241</td> </tr> </tbody> </table>			M.I.	E.P.C.	Std E.P.C.	StdYX E.P.C.	[V2T1]	14.761	-0.696	-0.696	-0.189	[V2T2]	5.578	0.307	0.307	0.094	[V4T1]	10.400	0.366	0.366	0.113	[V4T2]	5.167	-0.271	-0.271	-0.084	[V5T1]	20.890	-0.027	-0.027	-0.017	[V5T2]	14.191	-0.596	-0.596	-0.241
		M.I.	E.P.C.	Std E.P.C.	StdYX E.P.C.																																												
[V2T1]	14.761	-0.696	-0.696	-0.189																																											
[V2T2]	5.578	0.307	0.307	0.094																																											
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[V4T2]	5.167	-0.271	-0.271	-0.084																																											
[V5T1]	20.890	-0.027	-0.027	-0.017																																											
[V5T2]	14.191	-0.596	-0.596	-0.241																																											

Model 3b. Partial Scalar Invariance Model (all intercepts held equal across over time except v1T1 and v5T1)

```
MODEL: ! Model 3b: Partial Scalar Invariance, no v1T1 v5T1

! Factor loadings NOW CONSTRAINED EQUAL ACROSS TIME EXCEPT v1T1
Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1* (L1a L2-L6);
Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2* (L1-L6);
Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3* (L1-L6);
! Item intercepts NOW CONSTRAINED EQUAL ACROSS TIME, no v1T1 v5T1
[v1T1*]; [v1T2* v1T3*] (I1); ! 3a: I1 applies only to 2 and 3
[v2T1* v2T2* v2T3*] (I2);
[v3T1* v3T2* v3T3*] (I3);
[v4T1* v4T2* v4T3*] (I4);
[v5T1*]; [v5T2* v5T3*] (I5); ! 3b: I5 applies only to 2 and 3
[v6T1* v6T2* v6T3*] (I6);
! Residual variances all freely estimated, not labeled
v1T1* v1T2* v1T3*; v2T1* v2T2* v2T3*;
v3T1* v3T2* v3T3*; v4T1* v4T2* v4T3*;
v5T1* v5T2* v5T3*; v6T1* v6T2* v6T3*;
! Factor variance AT TIME 1 fixed=1 for identification
Time1@1 Time2* Time3*;
! Factor mean AT TIME 1 fixed=0 for identification
[Time1@0 Time2* Time3*];
! Factor covariances all freely estimated
Time1 Time2 Time3 WITH Time1* Time2* Time3*;
! Residual covariances estimated for same item across time
v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3*;
v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3*;
v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3*;
v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3*;
v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3*;
v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3*;
```

Does the partial scalar model (3b) fit *better* than the full scalar model (3a)?
 Yes, $-2\Delta LL(df=1) = 15.16, p < .01$

Does the partial scalar model (3b) fit *worse* than the partial metric model (2b)?
 Yes, $-2\Delta LL(df=8) = 27.84, p < .01$

Modification indices still suggest that freeing these intercepts would help, so let's try v4T1 next (biggest χ^2 change suggested).

MODEL FIT INFORMATION					
Number of Free Parameters				58	
Loglikelihood					
H0 Value				-4450.001	
H0 Scaling Correction Factor				1.5626	
for MLR					
H1 Value				-4284.045	
H1 Scaling Correction Factor				1.2029	
for MLR					
Information Criteria					
Akaike (AIC)				9016.001	
Bayesian (BIC)				9191.004	
Sample-Size Adjusted BIC				9007.440	
				(n* = (n + 2) / 24)	
Chi-Square Test of Model Fit					
Value				318.018*	
Degrees of Freedom				131	
P-Value				0.0000	
Scaling Correction Factor				1.0437	
for MLR					
RMSEA (Root Mean Square Error Of Approximation)					
Estimate				0.097	
90 Percent C.I.				0.084 0.111	
Probability RMSEA <= .05				0.000	
CFI/TLI					
CFI				0.893	
TLI				0.875	
SRMR (Standardized Root Mean Square Residual)					
Value				0.086	
MODEL MODIFICATION INDICES (relevant for invariance testing)					
Means/Intercepts/Thresholds					
		M.I.	E.P.C.	Std E.P.C.	StdYX
E.P.C.					
[V2T1]		11.529	-0.599	-0.599	-0.164
[V2T2]		4.390	0.278	0.278	0.085
[V4T1]		13.795	0.425	0.425	0.132
[V4T2]		6.398	-0.306	-0.306	-0.096

Model 3c. Partial Scalar Invariance Model (all intercepts held equal across over time except v1T1, v5T1, v4T1)

```

MODEL: ! Model 3c: Partial Scalar Invariance, no v1T1 v5T1 v4T1

! Factor loadings NOW CONSTRAINED EQUAL ACROSS TIME EXCEPT v1T1
Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1* (L1a L2-L6);
Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2* (L1-L6);
Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3* (L1-L6);
! Item intercepts NOW CONSTRAINED EQUAL ACROSS TIME,
! no v1T1 v5T1 v4T1
[v1T1*]; [v1T2* v1T3*] (I1); ! 3a: I1 applies only to 2 and 3
[v2T1* v2T2* v2T3*] (I2);
[v3T1* v3T2* v3T3*] (I3);
[v4T1*]; [v4T2* v4T3*] (I4); ! 3c: I4 applies only to 2 and 3
[v5T1*]; [v5T2* v5T3*] (I5); ! 3b: I5 applies only to 2 and 3
[v6T1* v6T2* v6T3*] (I6);
! Residual variances all freely estimated, not labeled
v1T1* v1T2* v1T3*; v2T1* v2T2* v2T3*;
v3T1* v3T2* v3T3*; v4T1* v4T2* v4T3*;
v5T1* v5T2* v5T3*; v6T1* v6T2* v6T3*;
! Factor variance AT TIME 1 fixed=1 for identification
Time1@1 Time2* Time3*;
! Factor mean AT TIME 1 fixed=0 for identification
[Time1@0 Time2* Time3*];
! Factor covariances all freely estimated
Time1 Time2 Time3 WITH Time1* Time2* Time3*;
! Residual covariances estimated for same item across time
v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3*;
v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3*;
v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3*;
v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3*;
v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3*;
v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3*;
    
```

Does the partial scalar model (3c) fit *better* than the partial scalar model (3b)? Yes, $-2\Delta LL(df=1) = 9.24, p < .01$

Does the partial scalar model (3c) fit *worse* than the partial metric model (2b)? Eh, $-2\Delta LL(df=7) = 13.99, p = .05$

Although fit is close to not worse than the partial metric model, there is a significant modification index for v2T1, suggesting localized strain. So let's see what happens if we free that one, too.

MODEL FIT INFORMATION					
Number of Free Parameters				59	
Loglikelihood					
H0 Value				-4442.214	
H0 Scaling Correction Factor				1.5647	
for MLR					
H1 Value				-4284.045	
H1 Scaling Correction Factor				1.2029	
for MLR					
Information Criteria					
Akaike (AIC)				9002.427	
Bayesian (BIC)				9180.447	
Sample-Size Adjusted BIC				8993.718	
				(n* = (n + 2) / 24)	
Chi-Square Test of Model Fit					
Value				304.537*	
Degrees of Freedom				130	
P-Value				0.0000	
Scaling Correction Factor				1.0387	
for MLR					
RMSEA (Root Mean Square Error Of Approximation)					
Estimate				0.094	
90 Percent C.I.				0.081 0.108	
Probability RMSEA <= .05				0.000	
CFI/TLI					
CFI				0.900	
TLI				0.882	
SRMR (Standardized Root Mean Square Residual)					
Value				0.092	
MODEL MODIFICATION INDICES (relevant for invariance testing)					
Means/Intercepts/Thresholds					
		M.I.	E.P.C.	Std E.P.C.	StdYX E.P.C.
[v2T1]		8.560	-0.497	-0.497	-0.137

Model 3d. Partial Scalar Invariance Model (all intercepts held equal across over time except v1T1, v5T1, v4T1, v2T1)

```

MODEL: ! Model 3d: Partial Scalar Invariance,
      ! no v1T1 v5T1 v4T1 v2T1

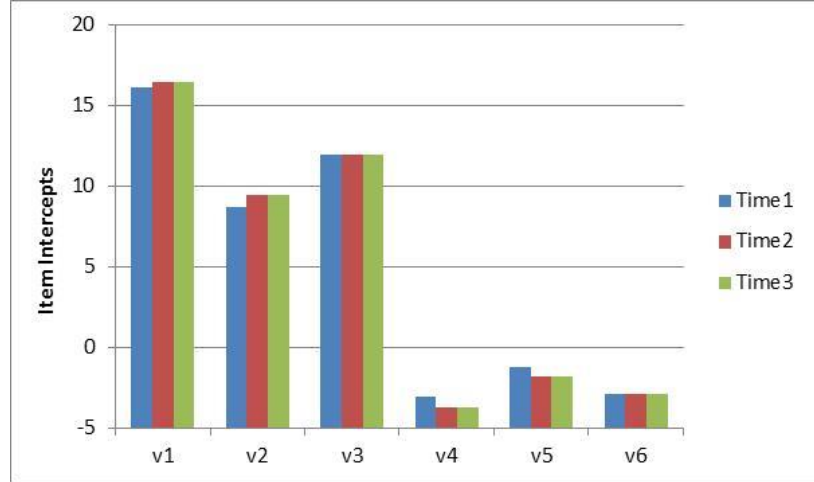
! Factor loadings NOW CONSTRAINED EQUAL ACROSS TIME EXCEPT v1T1
Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1* (L1a L2-L6);
Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2* (L1-L6);
Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3* (L1-L6);
! Item intercepts NOW CONSTRAINED EQUAL ACROSS TIME,
! no v1T1 v5T1 v4T1 v2T1
[v1T1*]; [v1T2* v1T3*] (I1); ! 3a: I1 applies only to 2 and 3
[v2T1*]; [v2T2* v2T3*] (I2); ! 3d: I2 applies only to 2 and 3
[v3T1* v3T2* v3T3*] (I3);
[v4T1*]; [v4T2* v4T3*] (I4); ! 3c: I4 applies only to 2 and 3
[v5T1*]; [v5T2* v5T3*] (I5); ! 3b: I5 applies only to 2 and 3
[v6T1* v6T2* v6T3*] (I6);
! Residual variances all freely estimated, not labeled
v1T1* v1T2* v1T3*; v2T1* v2T2* v2T3*;
v3T1* v3T2* v3T3*; v4T1* v4T2* v4T3*;
v5T1* v5T2* v5T3*; v6T1* v6T2* v6T3*;
! Factor variance AT TIME 1 fixed=1 for identification
Time1@1 Time2* Time3*;
! Factor mean AT TIME 1 fixed=0 for identification
[Time1@0 Time2* Time3*];
! Factor covariances all freely estimated
Time1 Time2 Time3 WITH Time1* Time2* Time3*;
! Residual covariances estimated for same item across time
v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3*;
v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3*;
v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3*;
v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3*;
v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3*;
v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3*;
    
```

MODEL FIT INFORMATION		
Number of Free Parameters		60
Loglikelihood		
H0 Value		-4437.665
H0 Scaling Correction Factor for MLR		1.5560
H1 Value		-4284.045
H1 Scaling Correction Factor for MLR		1.2029
Information Criteria		
Akaike (AIC)		8995.330
Bayesian (BIC)		9176.366
Sample-Size Adjusted BIC		8986.473
(n* = (n + 2) / 24)		
Chi-Square Test of Model Fit		
Value		295.789*
Degrees of Freedom		129
P-Value		0.0000
Scaling Correction Factor for MLR		1.0387
RMSEA (Root Mean Square Error Of Approximation)		
Estimate		0.093
90 Percent C.I.		0.079 0.106
Probability RMSEA <= .05		0.000
CFI/TLI		
CFI		0.904
TLI		0.887
Chi-Square Test of Model Fit for the Baseline Model		
Value		1896.788
Degrees of Freedom		153
P-Value		0.0000
SRMR (Standardized Root Mean Square Residual)		
Value		0.091

Does the partial scalar model (3d) fit better than the partial scalar model (3c)? Yes, $-2\Delta LL(df=1) = 8.73, p < .01$

Does the partial scalar model (3d) fit worse than the partial metric model (2b)? No, $-2\Delta LL(df=6) = 4.35, p = .63$

No invariance-related modification indices remain, so we are done!
The intercepts at the end of Model 3d are shown on the left.



3d UNSTANDARDIZED PARTIAL SCALAR MODEL RESULTS

Two-Tailed					Two-Tailed						
		Estimate	S.E.	Est./S.E.	P-Value			Estimate	S.E.	Est./S.E.	P-Value
TIME1	BY					Means (FACTOR MEAN AT TIME1 FIXED=0 FOR IDENTIFICATION)					
	V1T1	3.231	0.262	12.330	0.000	TIME1	0.000	0.000	999.000	999.000	
	V2T1	1.953	0.201	9.739	0.000	TIME2	0.293	0.081	3.625	0.000	
	V3T1	1.974	0.198	9.989	0.000	TIME3	0.521	0.093	5.612	0.000	
	V4T1	1.904	0.220	8.656	0.000						
	V5T1	0.983	0.138	7.097	0.000	Intercepts					
	V6T1	1.477	0.130	11.353	0.000	V1T1	16.090	0.274	58.684	0.000	
TIME2	BY					V1T2	16.425	0.281	58.364	0.000	
	V1T2	2.629	0.232	11.317	0.000	V1T3	16.425	0.281	58.364	0.000	
	V2T2	1.953	0.201	9.739	0.000	V2T1	8.674	0.294	29.540	0.000	
	V3T2	1.974	0.198	9.989	0.000	V2T2	9.413	0.261	36.036	0.000	
	V4T2	1.904	0.220	8.656	0.000	V2T3	9.413	0.261	36.036	0.000	
	V5T2	0.983	0.138	7.097	0.000	V3T1	11.950	0.225	53.099	0.000	
	V6T2	1.477	0.130	11.353	0.000	V3T2	11.950	0.225	53.099	0.000	
TIME3	BY					V3T3	11.950	0.225	53.099	0.000	
	V1T3	2.629	0.232	11.317	0.000	V4T1	-3.024	0.267	-11.334	0.000	
	V2T3	1.953	0.201	9.739	0.000	V4T2	-3.744	0.299	-12.535	0.000	
	V3T3	1.974	0.198	9.989	0.000	V4T3	-3.744	0.299	-12.535	0.000	
	V4T3	1.904	0.220	8.656	0.000	V5T1	-1.215	0.131	-9.277	0.000	
	V5T3	0.983	0.138	7.097	0.000	V5T2	-1.802	0.207	-8.706	0.000	
	V6T3	1.477	0.130	11.353	0.000	V5T3	-1.802	0.207	-8.706	0.000	
TIME1	WITH					V6T1	-2.854	0.161	-17.688	0.000	
	TIME2	0.850	0.079	10.809	0.000	V6T2	-2.854	0.161	-17.688	0.000	
	TIME3	0.686	0.124	5.543	0.000	V6T3	-2.854	0.161	-17.688	0.000	
TIME2	WITH					Residual Variances (ITEM VARIANCE THAT IS NOT THE FACTOR)					
	TIME3	0.706	0.128	5.519	0.000	V1T1	0.170	0.374	0.454	0.650	
*** Residual covariances among same item across time ***											
V1T1	WITH					V1T2	0.548	0.265	2.070	0.038	
	V1T2	-0.206	0.246	-0.838	0.402	V1T3	0.509	0.314	1.618	0.106	
	V1T3	-0.010	0.233	-0.043	0.966	V2T1	8.702	1.026	8.483	0.000	
V1T2	WITH					V2T2	5.895	0.605	9.746	0.000	
	V1T3	0.130	0.231	0.561	0.575	V2T3	5.177	0.795	6.514	0.000	
.....						V3T1	2.502	0.386	6.484	0.000	
Variances (FACTOR VARIANCE AT TIME1=1 FOR IDENTIFICATION)											
	TIME1	1.000	0.000	999.000	999.000	V3T2	2.178	0.352	6.183	0.000	
	TIME2	1.167	0.187	6.252	0.000	V3T3	2.309	0.416	5.548	0.000	
	TIME3	0.947	0.156	6.054	0.000	V4T1	7.172	1.021	7.021	0.000	

Model 4a. Residual Variance Invariance Model (error variances held equal for all except non-invariant items)

```

MODEL: ! Model 4a: Residual Variances
      ! except for non-invariant items

! Factor loadings NOW CONSTRAINED EQUAL ACROSS TIME EXCEPT v1T1
Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1* (L1a L2-L6);
Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2* (L1-L6);
Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3* (L1-L6);
! Item intercepts NOW CONSTRAINED EQUAL ACROSS TIME,
! no v1T1 v5T1 v4T1 v2T1
[v1T1*]; [v1T2* v1T3*] (I1); ! 3a: I1 applies only to 2 and 3
[v2T1*]; [v2T2* v2T3*] (I2); ! 3d: I2 applies only to 2 and 3
[v3T1* v3T2* v3T3*] (I3);
[v4T1*]; [v4T2* v4T3*] (I4); ! 3c: I4 applies only to 2 and 3
[v5T1*]; [v5T2* v5T3*] (I5); ! 3b: I5 applies only to 2 and 3
[v6T1* v6T2* v6T3*] (I6);
! Residual variances NOW CONSTRAINED EQUAL ACROSS TIME
(WHEN POSSIBLE)
v1T1* v1T2* v1T3* (E1); !4a: E1 applies only to 2 and 3
v2T1* v2T2* v2T3* (E2); !4a: E2 applies only to 2 and 3
v3T1* v3T2* v3T3* (E3);
v4T1* v4T2* v4T3* (E4); !4a: E4 applies only to 2 and 3
v5T1* v5T2* v5T3* (E5); !4a: E5 applies only to 2 and 3
v6T1* v6T2* v6T3* (E6);
! Factor variance AT TIME 1 fixed=1 for identification
Time1@1 Time2* Time3*;
! Factor mean AT TIME 1 fixed=0 for identification
[Time1@0 Time2* Time3*];
! Factor covariances all freely estimated
Time1 Time2 Time3 WITH Time1* Time2* Time3*;
! Residual covariances estimated for same item across time
v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3*;
v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3*;
v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3*;
v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3*;
v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3*;
v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3*;
    
```

Does the full residual model (4a) fit worse than the partial scalar model (3d)? Yes, $-2\Delta LL(df=8) = 24.72, p < .01$

Modification indices suggest freeing v5 across Time2 and Time3, so let's try that next.

MODEL FIT INFORMATION				
Number of Free Parameters				52
Loglikelihood				
H0 Value				-4454.592
H0 Scaling Correction Factor				1.5487
for MLR				
H1 Value				-4284.045
H1 Scaling Correction Factor				1.2029
for MLR				
Information Criteria				
Akaike (AIC)				9013.185
Bayesian (BIC)				9170.083
Sample-Size Adjusted BIC				9005.509
				(n* = (n + 2) / 24)
Chi-Square Test of Model Fit				
Value				318.280*
Degrees of Freedom				137
P-Value				0.0000
Scaling Correction Factor				1.0717
for MLR				
RMSEA (Root Mean Square Error Of Approximation)				
Estimate				0.094
90 Percent C.I.				0.080 0.107
Probability RMSEA <= .05				0.000
CFI/TLI				
CFI				0.896
TLI				0.884
SRMR (Standardized Root Mean Square Residual)				
Value				0.095
MODEL MODIFICATION INDICES (relevant for invariance testing)				
Means/Intercepts/Thresholds				
	M.I.	E.P.C.	Std E.P.C.	StdYX E.P.C.
Variances/Residual Variances				
V5T2	12.739	0.755	0.755	0.153
V5T3	12.740	-1.125	-1.125	-0.238
V6T1	13.740	0.421	0.421	0.124
V6T3	7.815	-0.393	-0.393	-0.124

Model 4b. Partial Residual Variance Invariance Model (error variances held equal for all except non-invariant items and v5T2/T3)

```

MODEL: ! Model 4b: Residual Variances
      ! except for non-invariant items, v5T2-v5T3

! Factor loadings NOW CONSTRAINED EQUAL ACROSS TIME EXCEPT v1T1
Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1* (L1a L2-L6);
Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2* (L1-L6);
Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3* (L1-L6);
! Item intercepts NOW CONSTRAINED EQUAL ACROSS TIME,
! no v1T1 v5T1 v4T1 v2T1
[v1T1*]; [v1T2* v1T3*] (I1); ! 3a: I1 applies only to 2 and 3
[v2T1*]; [v2T2* v2T3*] (I2); ! 3d: I2 applies only to 2 and 3
[v3T1* v3T2* v3T3*] (I3);
[v4T1*]; [v4T2* v4T3*] (I4); ! 3c: I4 applies only to 2 and 3
[v5T1*]; [v5T2* v5T3*] (I5); ! 3b: I5 applies only to 2 and 3
[v6T1* v6T2* v6T3*] (I6);
! Residual variances NOW CONSTRAINED EQUAL ACROSS TIME
(WHEN POSSIBLE) except v5T2-v5T3
v1T1* v1T2* v1T3* (E1); !4a: E1 applies only to 2 and 3
v2T1* v2T2* v2T3* (E2); !4a: E2 applies only to 2 and 3
v3T1* v3T2* v3T3* (E3);
v4T1* v4T2* v4T3* (E4); !4a: E4 applies only to 2 and 3
v5T1* v5T2* v5T3*; !4b: 2 and 3 now also separate
v6T1* v6T2* v6T3* (E6);
! Factor variance AT TIME 1 fixed=1 for identification
Time1@1 Time2* Time3*;
! Factor mean AT TIME 1 fixed=0 for identification
[Time1@0 Time2* Time3*];
! Factor covariances all freely estimated
Time1 Time2 Time3 WITH Time1* Time2* Time3*;
! Residual covariances estimated for same item across time
v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3*;
v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3*;
v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3*;
v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3*;
v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3*;
v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3*;
    
```

Does the partial residual model (4b) fit better than the full residual model (4a)? Yes, $-2\Delta LL(df=1) = 10.06, p < .01$

Does the partial residual model (4b) fit worse than the partial scalar model (3d)? Eh, $-2\Delta LL(df=7) = 14.14, p = .05$

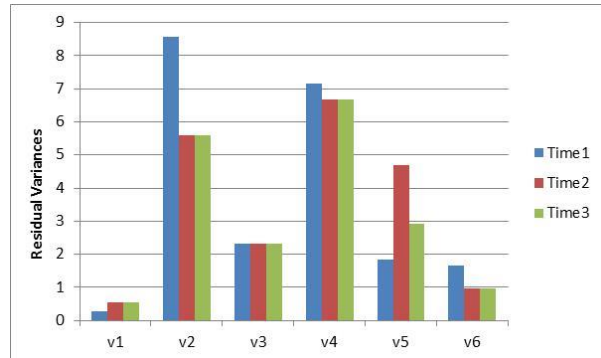
Modification indices suggest freeing v6 from Time1, so let's try that next.

MODEL FIT INFORMATION				
Number of Free Parameters				53
Loglikelihood				
H0 Value				-4447.259
H0 Scaling Correction Factor				1.5823
for MLR				
H1 Value				-4284.045
H1 Scaling Correction Factor				1.2029
for MLR				
Information Criteria				
Akaike (AIC)				9000.518
Bayesian (BIC)				9160.434
Sample-Size Adjusted BIC				8992.694
(n* = (n + 2) / 24)				
Chi-Square Test of Model Fit				
Value				309.384*
Degrees of Freedom				136
P-Value				0.0000
Scaling Correction Factor				1.0551
for MLR				
RMSEA (Root Mean Square Error Of Approximation)				
Estimate				0.092
90 Percent C.I.				0.078 0.105
Probability RMSEA <= .05				0.000
CFI/TLI				
CFI				0.901
TLI				0.888
SRMR (Standardized Root Mean Square Residual)				
Value				0.093
MODEL MODIFICATION INDICES (relevant for invariance testing)				
Means/Intercepts/Thresholds				
	M.I.	E.P.C.	Std E.P.C.	StdYX E.P.C.
Variances/Residual Variances				
v6T1	13.772	0.419	0.419	0.125
v6T3	7.149	-0.373	-0.373	-0.118

Model 4c. Partial Residual Variance Invariance Model (error variances held equal for all except non-invariant items, v5T2/T3, v6T1)

```

MODEL: ! Model 4c: Residual Variances
      ! except for non-invariant items, v5T2-v5T3, v6T1
! Factor loadings NOW CONSTRAINED EQUAL ACROSS TIME EXCEPT v1T1
Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1* (L1a L2-L6);
Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2* (L1-L6);
Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3* (L1-L6);
! Item intercepts NOW CONSTRAINED EQUAL ACROSS TIME,
! no v1T1 v5T1 v4T1 v2T1
[v1T1*]; [v1T2* v1T3*] (I1); ! 3a: I1 applies only to 2 and 3
[v2T1*]; [v2T2* v2T3*] (I2); ! 3d: I2 applies only to 2 and 3
[v3T1* v3T2* v3T3*] (I3);
[v4T1*]; [v4T2* v4T3*] (I4); ! 3c: I4 applies only to 2 and 3
[v5T1*]; [v5T2* v5T3*] (I5); ! 3b: I5 applies only to 2 and 3
[v6T1* v6T2* v6T3*] (I6);
! Residual variances NOW CONSTRAINED EQUAL ACROSS TIME
(WHEN POSSIBLE) except v5T2-v5T3, v6T1
v1T1*; v1T2* v1T3* (E1); !4a: E1 applies only to 2 and 3
v2T1*; v2T2* v2T3* (E2); !4a: E2 applies only to 2 and 3
v3T1* v3T2* v3T3* (E3);
v4T1*; v4T2* v4T3* (E4); !4a: E4 applies only to 2 and 3
v5T1*; v5T2*; v5T3*; !4b: 2 and 3 now also separate
v6T1*; v6T2* v6T3* (E6); !4c: E6 applies only to 2 and 3
! Factor variance AT TIME 1 fixed=1 for identification
Time1@1 Time2* Time3*;
! Factor mean AT TIME 1 fixed=0 for identification
[Time1@0 Time2* Time3*];
! Factor covariances all freely estimated
Time1 Time2 Time3 WITH Time1* Time2* Time3*;
! Residual covariances estimated for same item across time
v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3*;
v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3*;
v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3*;
v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3*;
v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3*;
v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3*;
    
```



MODEL FIT INFORMATION		
Number of Free Parameters		54
Loglikelihood		
H0 Value		-4439.971
H0 Scaling Correction Factor		1.5771
for MLR		
H1 Value		-4284.045
H1 Scaling Correction Factor		1.2029
for MLR		
Information Criteria		
Akaike (AIC)		8987.942
Bayesian (BIC)		9150.876
Sample-Size Adjusted BIC		8979.971
	(n* = (n + 2) / 24)	
Chi-Square Test of Model Fit		
Value		296.084*
Degrees of Freedom		135
P-Value		0.0000
Scaling Correction Factor		1.0533
for MLR		
RMSEA (Root Mean Square Error Of Approximation)		
Estimate		0.089
90 Percent C.I.		0.075 0.103
Probability RMSEA <= .05		0.000
CFI/TLI		
CFI		0.908
TLI		0.895
SRMR (Standardized Root Mean Square Residual)		
Value		0.092

Does the partial residual model (4c) fit better than the partial residual model (4b)? Yes, $-2\Delta LL(df=1) = 11.20, p < .01$

Does the partial residual model (4c) fit worse than the partial scalar model (3d)? No, $-2\Delta LL(df=6) = 3.38, p = .76$

No invariance-related modification indices remain, so we are done!
 The residual variances at the end of Model 4c are shown on the left.
 Next is structural invariance.

4c UNSTANDARDIZED FINAL MEASUREMENT INVARIANCE SOLUTION

Two-Tailed					Two-Tailed					
	Estimate	S.E.	Est./S.E.	P-Value		Estimate	S.E.	Est./S.E.	P-Value	
TIME1	BY				Means (FACTOR MEAN AT TIME1 FIXED=0 FOR IDENTIFICATION)					
V1T1		3.214	0.259	12.409	0.000	TIME1	0.000	0.000	999.000	999.000
V2T1		1.945	0.200	9.735	0.000	TIME2	0.295	0.081	3.654	0.000
V3T1		1.983	0.196	10.094	0.000	TIME3	0.520	0.092	5.668	0.000
V4T1		1.913	0.219	8.741	0.000					
V5T1		0.987	0.138	7.154	0.000	Intercepts - V3 AND V6 ARE HOLDING THIS TOGETHER WITH TIME1				
V6T1		1.470	0.123	11.975	0.000	V1T1	16.089	0.275	58.597	0.000
						V1T2	16.418	0.283	58.056	0.000
TIME2	BY					V1T3	16.418	0.283	58.056	0.000
V1T2		2.644	0.230	11.473	0.000	V2T1	8.675	0.294	29.523	0.000
V2T2		1.945	0.200	9.735	0.000	V2T2	9.416	0.262	35.991	0.000
V3T2		1.983	0.196	10.094	0.000	V2T3	9.416	0.262	35.991	0.000
V4T2		1.913	0.219	8.741	0.000	V3T1	11.950	0.225	53.170	0.000
V5T2		0.987	0.138	7.154	0.000	V3T2	11.950	0.225	53.170	0.000
V6T2		1.470	0.123	11.975	0.000	V3T3	11.950	0.225	53.170	0.000
						V4T1	-3.024	0.266	-11.352	0.000
TIME3	BY					V4T2	-3.750	0.298	-12.565	0.000
V1T3		2.644	0.230	11.473	0.000	V4T3	-3.750	0.298	-12.565	0.000
V2T3		1.945	0.200	9.735	0.000	V5T1	-1.213	0.131	-9.275	0.000
V3T3		1.983	0.196	10.094	0.000	V5T2	-1.803	0.207	-8.720	0.000
V4T3		1.913	0.219	8.741	0.000	V5T3	-1.803	0.207	-8.720	0.000
V5T3		0.987	0.138	7.154	0.000	V6T1	-2.851	0.160	-17.815	0.000
V6T3		1.470	0.123	11.975	0.000	V6T2	-2.851	0.160	-17.815	0.000
						V6T3	-2.851	0.160	-17.815	0.000
TIME1	WITH					Residual Variances - ITEM VARIANCE THAT IS NOT THE FACTOR				
TIME2		0.843	0.078	10.745	0.000	V1T1	0.285	0.342	0.831	0.406
TIME3		0.683	0.124	5.505	0.000	V1T2	0.539	0.233	2.316	0.021
						V1T3	0.539	0.233	2.316	0.021
TIME2	WITH					V2T1	8.562	1.004	8.526	0.000
TIME3		0.692	0.126	5.489	0.000	V2T2	5.592	0.502	11.132	0.000
						V2T3	5.592	0.502	11.132	0.000
*** Residual covariances among same item across time ***										
V1T1	WITH					V3T1	2.312	0.271	8.534	0.000
V1T2		-0.165	0.230	-0.716	0.474	V3T2	2.312	0.271	8.534	0.000
V1T3		0.014	0.212	0.066	0.948	V3T3	2.312	0.271	8.534	0.000
						V4T1	7.139	1.043	6.842	0.000
V1T2	WITH					V4T2	6.686	0.870	7.684	0.000
V1T3		0.153	0.230	0.667	0.505	V4T3	6.686	0.870	7.684	0.000
.....						V5T1	1.829	0.448	4.078	0.000
						V5T2	4.705	1.455	3.233	0.001
						V5T3	2.908	0.749	3.881	0.000
Variances (FACTOR VARIANCE AT TIME1=1 FOR IDENTIFICATION)						V6T1	1.664	0.233	7.138	0.000
TIME1		1.000	0.000	999.000	999.000	V6T2	0.957	0.136	7.039	0.000
TIME2		1.159	0.186	6.231	0.000	V6T3	0.957	0.136	7.039	0.000
TIME3		0.934	0.151	6.171	0.000					

STRUCTURAL INVARIANCE TESTS

Model 5a. Factor Variance Invariance Model

Model 6a. Factor Covariance Invariance Model

```

MODEL: ! Model 5a: Factor Variance Invariance
(rest of code before and after is same as 4c)

! Model 5a: Factor Variance Invariance (all fixed to 1 now)
Time1@1 Time2@1 Time3@1;

MODEL FIT INFORMATION
Number of Free Parameters          52
Loglikelihood
  H0 Value                        -4441.238
  H0 Scaling Correction Factor     1.5848
    for MLR
  H1 Value                        -4284.045
  H1 Scaling Correction Factor     1.2029
    for MLR
Information Criteria
  Akaike (AIC)                    8986.475
  Bayesian (BIC)                  9143.374
  Sample-Size Adjusted BIC        8978.799
    (n* = (n + 2) / 24)
Chi-Square Test of Model Fit
  Value                           297.152*
  Degrees of Freedom              137
  P-Value                         0.0000
  Scaling Correction Factor        1.0580
    for MLR
RMSEA (Root Mean Square Error Of Approximation)
  Estimate                        0.088
  90 Percent C.I.                 0.074  0.102
  Probability RMSEA <= .05        0.000
CFI/TLI
  CFI                             0.908
  TLI                             0.897
SRMR (Standardized Root Mean Square Residual)
  Value                           0.100

Does the factor variance model (5a) fit worse than the partial residual model (4c)? No, -2ΔLL(df=2) = 1.84, p=.40

Factor Covariances...
TIME1 WITH
  TIME2          0.778    0.042   18.375    0.000
  TIME3          0.713    0.087    8.214    0.000
TIME2 WITH
  TIME3          0.662    0.095    6.929    0.000
    
```

```

MODEL: ! Model 6a: Factor Covariance Invariance
(rest of code before and after is same as 5a)

! Model 6a: Factor Covariance Invariance (all fixed equal)
Time1 Time2 Time3 WITH Time1* Time2* Time3* (Fcov);

MODEL FIT INFORMATION
Number of Free Parameters          50
Loglikelihood
  H0 Value                        -4443.654
  H0 Scaling Correction Factor     1.5649
    for MLR
  H1 Value                        -4284.045
  H1 Scaling Correction Factor     1.2029
    for MLR
Information Criteria
  Akaike (AIC)                    8987.308
  Bayesian (BIC)                  9138.172
  Sample-Size Adjusted BIC        8979.927
    (n* = (n + 2) / 24)
Chi-Square Test of Model Fit
  Value                           297.568*
  Degrees of Freedom              139
  P-Value                         0.0000
  Scaling Correction Factor        1.0728
    for MLR
RMSEA (Root Mean Square Error Of Approximation)
  Estimate                        0.087
  90 Percent C.I.                 0.073  0.101
  Probability RMSEA <= .05        0.000
CFI/TLI
  CFI                             0.909
  TLI                             0.900
SRMR (Standardized Root Mean Square Residual)
  Value                           0.100

Does the factor covariance model (6a) fit worse than the factor variance model (5a)? No, -2ΔLL(df=2) = 2.32, p=.31

FACTOR COVARIANCES FROM MODEL 6a (REPRESENT CORRELATIONS):
TIME1 WITH TIME2  0.724    0.053   13.748    0.000
TIME1 WITH TIME3  0.724    0.053   13.748    0.000
TIME2 WITH TIME3  0.724    0.053   13.748    0.000
FACTOR MEANS FROM MODEL 6a (REPRESENT MEAN DIFFERENCES):
TIME1             0.000    0.000   999.000   999.000
TIME2             0.284    0.079    3.605    0.000
TIME3             0.520    0.091    5.700    0.000
    
```

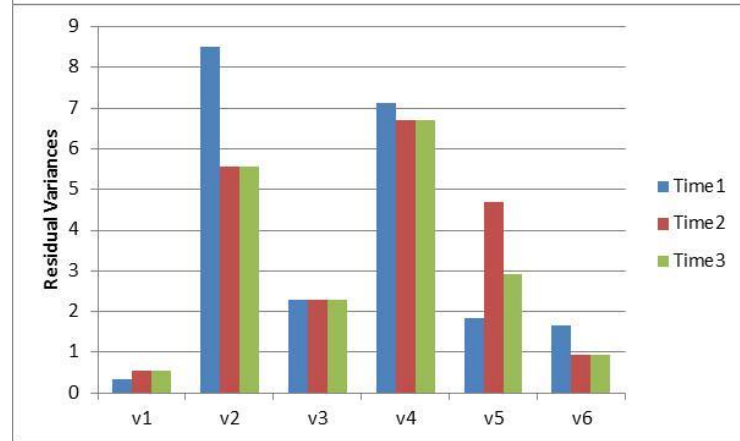
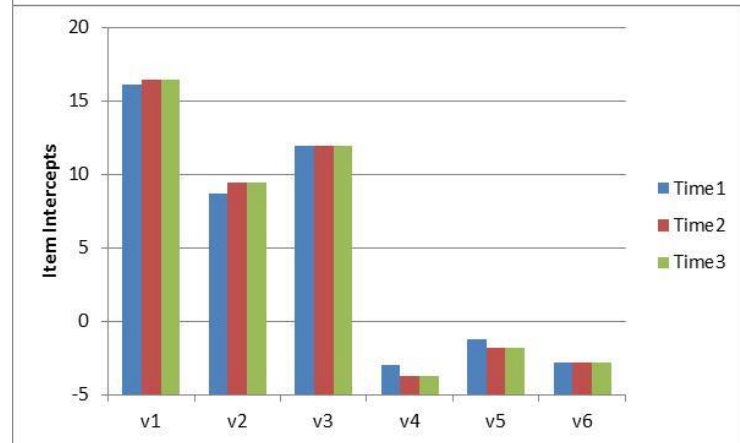
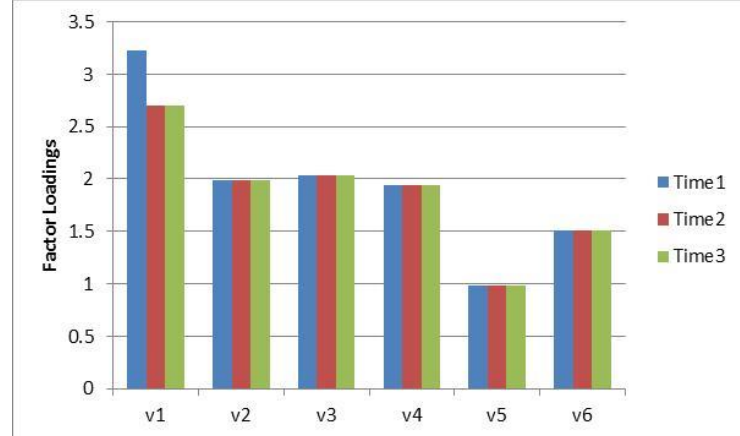

Model 7a. Factor Mean Invariance Model

```

MODEL: ! Model 7a: Factor Mean Invariance
      ! Testing Diff between Time2 and Time3
! Factor loadings NOW CONSTRAINED EQUAL ACROSS TIME EXCEPT v1T1
Time1 BY v1T1* v2T1* v3T1* v4T1* v5T1* v6T1* (L1a L2-L6);
Time2 BY v1T2* v2T2* v3T2* v4T2* v5T2* v6T2* (L1-L6);
Time3 BY v1T3* v2T3* v3T3* v4T3* v5T3* v6T3* (L1-L6);
! Item intercepts NOW CONSTRAINED EQUAL ACROSS TIME,
! no v1T1 v5T1 v4T1
[v1T1*]; [v1T2* v1T3*] (I1); ! 3a: I1 applies only to 2 and 3
[v2T1*]; [v2T2* v2T3*] (I2); ! 3d: I2 applies only to 2 and 3
[v3T1* v3T2* v3T3*] (I3);
[v4T1*]; [v4T2* v4T3*] (I4); ! 3c: I4 applies only to 2 and 3
[v5T1*]; [v5T2* v5T3*] (I5); ! 3b: I5 applies only to 2 and 3
[v6T1* v6T2* v6T3*] (I6);
! Residual variances NOW CONSTRAINED EQUAL ACROSS TIME
(WHEN POSSIBLE) except v5T2-v5T3, v6T1
v1T1*; v1T2* v1T3* (E1); !4a: E1 applies only to 2 and 3
v2T1*; v2T2* v2T3* (E2); !4a: E2 applies only to 2 and 3
v3T1* v3T2* v3T3* (E3);
v4T1*; v4T2* v4T3* (E4); !4a: E4 applies only to 2 and 3
v5T1*; v5T2*; v5T3*; !4b: 2 and 3 now also separate
v6T1*; v6T2* v6T3* (E6); !4c: E6 applies only to 2 and 3
! Factor variance fixed=1 for structural invariance
Time1@1 Time2@1 Time3@1;
! Testing factor mean difference between Time2 and Time3
[Time1@0]; [Time2* Time3*] (Fmean); ! NEW CONSTRAINT
! Factor covariances held equal for structural invariance
Time1 Time2 Time3 WITH Time1* Time2* Time3* (Fcov);
! Residual covariances estimated for same item across time
v1T1 v1T2 v1T3 WITH v1T1* v1T2* v1T3*;
v2T1 v2T2 v2T3 WITH v2T1* v2T2* v2T3*;
v3T1 v3T2 v3T3 WITH v3T1* v3T2* v3T3*;
v4T1 v4T2 v4T3 WITH v4T1* v4T2* v4T3*;
v5T1 v5T2 v5T3 WITH v5T1* v5T2* v5T3*;
v6T1 v6T2 v6T3 WITH v6T1* v6T2* v6T3*;
    
```

Does the factor mean model (7a) fit worse than the factor covariance model (6a)? Yes, $-2\Delta LL(df=1) = 11.15, p < .01$, so we keep Model 6a instead.

MODEL FIT INFORMATION				
Number of Free Parameters			49	
Loglikelihood				
H0 Value			-4448.472	
H0 Scaling Correction Factor			1.5792	
			for MLR	
Means				
TIME1	0.000	0.000	999.000	999.000
TIME2	0.378	0.075	5.014	0.000
TIME3	0.378	0.075	5.014	0.000



6a UNSTANDARDIZED FINAL STRUCTURAL INVARIANCE SOLUTION

Two-Tailed					Two-Tailed				
Estimate					Estimate				
S.E.					S.E.				
Est./S.E.					Est./S.E.				
P-Value					P-Value				
TIME1 BY					Means (FACTOR MEAN AT TIME1 FIXED=0 FOR IDENTIFICATION)				
V1T1	3.229	0.243	13.272	0.000	TIME1	0.000	0.000	999.000	999.000
V2T1	1.993	0.170	11.754	0.000	TIME2	0.284	0.079	3.605	0.000
V3T1	2.029	0.169	12.022	0.000	TIME3	0.520	0.091	5.700	0.000
V4T1	1.939	0.214	9.077	0.000	Intercepts - V3 AND V6 ARE HOLDING THIS TOGETHER WITH TIME1				
V5T1	0.986	0.147	6.701	0.000	V1T1	16.099	0.271	59.420	0.000
V6T1	1.508	0.109	13.821	0.000	V1T2	16.428	0.281	58.488	0.000
TIME2 BY					V1T3	16.428	0.281	58.488	0.000
V1T2	2.704	0.232	11.677	0.000	V2T1	8.681	0.292	29.694	0.000
V2T2	1.993	0.170	11.754	0.000	V2T2	9.423	0.259	36.368	0.000
V3T2	2.029	0.169	12.022	0.000	V2T3	9.423	0.259	36.368	0.000
V4T2	1.939	0.214	9.077	0.000	V3T1	11.956	0.223	53.706	0.000
V5T2	0.986	0.147	6.701	0.000	V3T2	11.956	0.223	53.706	0.000
V6T2	1.508	0.109	13.821	0.000	V3T3	11.956	0.223	53.706	0.000
TIME3 BY					V4T1	-3.018	0.263	-11.463	0.000
V1T3	2.704	0.232	11.677	0.000	V4T2	-3.737	0.292	-12.784	0.000
V2T3	1.993	0.170	11.754	0.000	V4T3	-3.737	0.292	-12.784	0.000
V3T3	2.029	0.169	12.022	0.000	V5T1	-1.210	0.131	-9.269	0.000
V4T3	1.939	0.214	9.077	0.000	V5T2	-1.791	0.203	-8.807	0.000
V5T3	0.986	0.147	6.701	0.000	V5T3	-1.791	0.203	-8.807	0.000
V6T3	1.508	0.109	13.821	0.000	V6T1	-2.847	0.159	-17.889	0.000
TIME1 WITH					V6T2	-2.847	0.159	-17.889	0.000
TIME2	0.724	0.053	13.748	0.000	V6T3	-2.847	0.159	-17.889	0.000
TIME3	0.724	0.053	13.748	0.000	Residual Variances - ITEM VARIANCE THAT IS NOT THE FACTOR				
TIME2 WITH					V1T1	0.351	0.331	1.060	0.289
TIME3	0.724	0.053	13.748	0.000	V1T2	0.562	0.231	2.432	0.015
*** Residual covariances among same item across time ***					V1T3	0.562	0.231	2.432	0.015
V1T1 WITH					V2T1	8.506	0.999	8.511	0.000
V1T2	-0.106	0.225	-0.471	0.638	V2T2	5.563	0.494	11.261	0.000
V1T3	0.038	0.215	0.175	0.861	V2T3	5.563	0.494	11.261	0.000
V1T2 WITH					V3T1	2.288	0.269	8.507	0.000
V1T3	0.130	0.243	0.534		V3T2	2.288	0.269	8.507	0.000
0.593.....					V3T3	2.288	0.269	8.507	0.000
Variances (FACTOR VARIANCES CONSTRAINED EQUAL)					V4T1	7.134	1.041	6.853	0.000
TIME1	1.000	0.000	999.000	999.000	V4T2	6.694	0.873	7.666	0.000
TIME2	1.000	0.000	999.000	999.000	V4T3	6.694	0.873	7.666	0.000
TIME3	1.000	0.000	999.000	999.000	V5T1	1.825	0.446	4.092	0.000
					V5T2	4.705	1.454	3.235	0.001
					V5T3	2.921	0.752	3.887	0.000
					V6T1	1.656	0.235	7.054	0.000
					V6T2	0.942	0.131	7.188	0.000
					V6T3	0.942	0.131	7.188	0.000

Example results section for these analyses:

The extent to which a confirmatory factor model measuring social functioning (with six observed indicators) exhibited measurement invariance and structural invariance over time (i.e., across three occasions taken at six-month intervals) was examined using *Mplus* v. 8.4 (Muthén & Muthén, 1998–2017). Robust maximum likelihood (MLR) estimation was used for all analyses; accordingly, nested model comparisons were conducted using the rescaled difference in the model $-2LL$ values as a function of the difference in model degrees of freedom. A configural invariance model was initially specified in which three correlated factors (i.e., the factor at the three occasions) were estimated simultaneously; all factor means were fixed to 0 and all factor variances were fixed to 1 for identification. Residual covariances between the same indicators across occasions were estimated as well. As shown in Table 1, although the configural invariance model had marginal fit, theoretically reasonable attempts to improve the fit were unsuccessful. Thus, the analysis proceeded by applying parameter constraints in successive models to examine potential decreases in fit resulting from measurement or structural non-invariance over the three occasions.

Equality of the unstandardized indicator factor loadings across occasions was then examined in a metric invariance model. The factor variance was fixed to 1 at time 1 but was freely estimated at times 2 and 3. All factor loadings were constrained equal across time; all intercepts and residual variances were still permitted to vary across time. Factor covariances and residual covariances were estimated as described previously. The metric invariance model fit significantly worse than the configural invariance model $-2\Delta LL(10) = 19.14$, $p = .04$. The modification indices suggested that the loading of indicator 1 at time 1 was a source of misfit and should be freed. After doing so, the partial metric invariance model fit significantly better than the full metric invariance model, $-2\Delta LL(1) = 7.16$, $p < .001$, and the partial metric invariance model did not fit worse than the configural invariance model $-2\Delta LL(9) = 8.98$, $p = .44$. The fact that partial metric invariance (i.e., “weak invariance”) held indicates that the indicators were related to the latent factor equivalently across time, or more simply, that the same latent factor was being measured at each of occasion (with the exception of indicator 1, which was more related to the factor at time 1 than at times 2 or 3).

Equality of the unstandardized indicator intercepts across time was then examined in a scalar invariance model. The factor mean and variance were fixed to 0 and 1, respectively, at time 1 for identification, but the factor mean and variance were then estimated at times 2 and 3. All factor loadings and indicator intercepts were constrained equal across time (except for indicator 1 at time 1); all residual variances were still permitted to differ across time. Factor covariances and residual covariances were estimated as described previously. The scalar invariance model fit significantly worse than the partial metric invariance model, $-2\Delta LL(9) = 55.13$, $p < .01$. The modification indices suggested that the intercept of indicator 5 at time 1 was the largest source of the misfit and should be freed. After doing so, although the partial scalar invariance model had significantly better fit than the full scalar invariance model, $-2\Delta LL(1) = 15.16$, $p < .01$, it still fit worse than the partial metric invariance model, $-2\Delta LL(8) = 27.84$, $p < .001$. The modification indices suggested that the intercept of indicator 4 at time 1 was the largest remaining source of the misfit and should be freed. After doing so, although the new partial scalar invariance model (with the intercepts for indicators 1, 4, and 5 freed at time 1) fit significantly better than the previous partial scalar invariance model (without the intercept for indicator 4 freed at time 1), $-2\Delta LL(1) = 9.24$, $p < .01$, it still fit marginally worse than the partial metric invariance model, $-2\Delta LL(7) = 13.99$, $p = .05$. The modification indices suggested that the intercept of indicator 2 at time 1 was the largest remaining source of the misfit and should be freed. After doing so, the new partial scalar invariance model (with the intercepts for indicators 1, 2, 4 and 5 freed at time 1) fit significantly better than the previous partial scalar invariance model (without the intercept for indicator 2 freed at time 1), $-2\Delta LL(1) = 8.73$, $p < .01$, and it did not fit significantly worse than the partial metric invariance model, $-2\Delta LL(6) = 4.35$, $p = .63$. The fact that partial scalar invariance (i.e., “strong invariance”) held indicates that times 2 and 3 have the same expected response for each indicator at the same absolute level of the trait, or more simply, that the observed differences in the indicator means between times 2 and 3 is due to factor mean differences only. However, indicators 1 and 2 had a lower expected indicator response at the same absolute level of social functioning at time 1 than at time 2 or 3, while indicators 4 and 5 had a higher expected response.

Equality of the unstandardized residual variances across time was then examined in a residual variance invariance model. As in the partial scalar invariance model, the factor mean and variance were fixed to 0 and 1, respectively, for identification at time 1, but the factor mean and variance were still estimated at times 2 and 3. All factor loadings (except for indicator 1 at time 1), item intercepts (except for indicators 1, 2, 4, and 5 at time 1), and all residual variances (except for indicators 1, 2, 4, and 5 at time 1) were constrained to be equal across groups. Factor covariances and residual covariances were estimated as described previously. The residual variance invariance model fit significantly worse than the last partial scalar invariance model, $-2\Delta LL(8) = 24.72$, $p < .01$. The modification indices suggested that the residual variance of indicator 5 at time 2 versus time 3 was the largest remaining source of the misfit and should be

freed. After doing so, the partial residual variance invariance model fit significantly better than the residual invariance model, $-2\Delta LL(1) = 10.06$, $p < .01$, but still fit marginally worse than the last partial scalar invariance model, $-2\Delta LL(7) = 14.14$, $p = .05$. The modification indices suggested that the residual variance of indicator 6 at time 1 was the largest remaining source of the misfit and should be freed. After doing so, the new partial residual variance invariance model (with residual variances for indicators 1, 2, 4, 5, and 6 free at time 1; indicator 5 free at times 2 and 3 also) fit significantly better than the partial residual invariance model (without the residual variance for indicator 6 at time 1 freed), $-2\Delta LL(1) = 11.20$, $p < .01$, and did not fit worse than the last partial scalar invariance model, $-2\Delta LL(6) = 3.38$, $p = .76$. The fact that partial residual variance invariance (i.e., “strict invariance”) held indicates that the amount of indicator variance not accounted for by the factor was the same across times 2 and 3 (except for indicator 5, for which there was more residual variance at time 2). However, 5 out of 6 indicators did not have residual variance invariance at time 1 (although this was required because of a lack of metric or scalar invariance for indicators 1, 2, 4, and 5).

After achieving partial measurement invariance as was just described, structural invariance was then tested with three additional models. First, the factor variance at times 2 and 3 (which had been estimated freely) was constrained to 1 (i.e., to be equal to the factor variance at time 1), resulting in a nonsignificant decrease in fit relative to the last partial residual invariance model, $-2\Delta LL(2) = 1.84$, $p = .40$. Thus, equivalent amounts of individual differences in social functioning were found across time. Second, the factor covariances across time were constrained to be equal (which become factor correlations given a variance of 1 for each factor across time), resulting in a nonsignificant decrease in fit relative to the factor variance invariance model, $-2\Delta LL(2) = 2.32$, $p = .31$. Third, the factor means at times 2 and 3 (which had been estimated freely) was constrained to be equal to each other, resulting in a significant decrease in fit relative to the factor covariance invariance model $-2\Delta LL(1) = 11.15$, $p < .01$, indicating that the factor mean at time 3 was significantly higher than at time 2. The factor mean at time 2 was already significantly different from 0 (the factor mean at time 1), and thus, the three factor means were significantly different, increasing over time.

In conclusion, these analyses showed that partial measurement invariance was obtained over time—that is, the relationships of the indicators to the latent factor of social functioning were equivalent at times 2 and 3, although primarily not equivalent at time 1, as previous described. These analyses also showed that partial structural invariance was obtained over time, such that the same amount of individual differences variance in social functioning was observed with equal covariance over time across occasions (i.e., compound symmetry of the latent factor), although the amount of social functioning on average increased significantly over time. Model parameters from the final model are given in Table 2.

Reference: Muthén, L. K., & Muthén, B.O. (1998–2017). *Mplus user's guide* (8th ed.). Los Angeles, CA: Muthén & Muthén.

(see excel worksheet for Table 1; Table 2 would have unstandardized and standardized estimates and their SEs)

You might also replace all the nested model comparisons tests in the text with a table that provides them instead.