

Correlated and Coupled Within-Person Change in Emotional and Behavioral Disturbance in Individuals With Intellectual Disability

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Abstract

Individual change and variation in emotional/behavioral disturbance in children and adolescents with intellectual disability has received little empirical investigation. Based on 11 years of longitudinal data from the Australian Child to Adult Development Study, we report associations among individual differences in level, rate of change, and occasion-specific variation across subscales of the Developmental Behavior Checklist (DBC) with 506 participants who had intellectual disability and were ages 5 to 19 years at study entry. Correlations among the five DBC subscales ranged from .43 to .66 for level, .43 to .88 for rate of change, and .31 to .61 for occasion-specific variation, with the highest correlations observed consistently between disruptive, self-absorbed, and communication disturbance behaviors. These interdependencies among dimensions of emotional/behavioral disturbance provide insight into the developmental dynamics of psychopathology from childhood through young adulthood.

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Using checklists of behavior and emotional problems, researchers have found substantially elevated levels of psychopathology in children and adolescents with intellectual disability (e.g., Dekker, Koot, van der Ende, & Verhulst, 2002; Einfeld et al., 2006; Linna et al., 1999; Richardson & Koller, 1996; Wallander, Dekker, & Koot, 2003). In studies of psychiatric diagnoses in children and adolescents with intellectual disability, investigators have also reported elevated rates of *Diagnostic and Statistical Manual of Mental Disorders–DSM-IV* (American Psychiatric Association, 1994) and *International Classification of Diseases–ICD-10* (World Health Organization, 1992) diagnoses, and significant comorbidity (Dekker & Koot, 2003; Emerson, 2003). Although few researchers have addressed the issue of comorbidity, still fewer have considered associations among longitudinal change in psychopathology symptoms in children and adolescents with intellectual disability.

In a British population study of 5 to 15 year olds with intellectual disability ($N = 264$), Emerson (2003) found that 16% of the sample had two or more *ICD-10* diagnoses and 3% had three or more diagnoses. The children with intellectual disability were 14 times more likely than their typically developing peers to have two diagnoses and 9.4 times more likely to have three diagnoses. Investigating a sample of 474 individuals with intellectual disability (7 to 20 years of age), Dekker and Koot (2003) reported multiple *DSM-IV* disorders in 14.2% of the sample. Of those who had a diagnosis, 36.8% met diagnostic criteria for more than one disorder. Seven percent of the total sample had coexisting anxiety and disruptive diagnoses; 0.4%, comorbid anxiety and mood disorders; 0.8%, comorbid mood and disruptive disorder; and 2.3% met criteria for a diagnosis in each of these groupings. Almost 4% of the total sample had a comorbid diagnosis within the same major diagnostic grouping. Of the children with attention deficit hyperactivity disorder (ADHD), 44% also met criteria for oppositional defiant disorder, and 79.5% of the sample with a mood disorder also met criteria for another *DSM-IV* disorder, most commonly a disruptive disorder.

More information is available on community populations of typically developing children and adolescents. For children and adolescents who do not have intellectual disability but who do have one psychiatric diagnosis, rates of comorbidity

(two or more psychiatric diagnoses) of around 30% to 40% have been reported (Costello et al., 1996; Fergusson, Horwood, & Lynskey, 1993; Newman et al., 1996). Significant relationships have been reported among conduct/oppositional disorders and ADHD, anxiety/mood disorders, and substance use disorders (Costello et al., 1996; Fergusson et al., 1993). Other associations include anxiety and mood disorders, anxiety/mood disorders and substance use, and mood disorders and ADHD (Costello et al., 1996; Fergusson et al., 1993). In terms of gender differences, results have been mixed, with reports of either higher rates of comorbidity in boys (Costello et al., 1996; Fergusson et al., 1993) or no gender differences (Newman et al., 1996). Gender differences in patterns of comorbidity have been observed, with depression and conduct disorder co-occurring in girls but not boys and comorbid depression and substance use disorder in boys but not girls (Costello, Mustillo, Erkanli, Keeler, & Angold, 2003).

Very little research has been undertaken on longitudinal change in patterns of symptomatology. Achenbach (1991) examined psychopathology as described in the Child Behavior Checklist (CBCL) for a population of typically developing Dutch children ages 4 to 11 years. He examined patterns and relationships between types of psychopathology over a 6-year period. Cross-sectionally, results indicated positive associations between all syndromes of the CBCL, with the strongest correlations between those that covered related behaviors (e.g., Aggressive and Delinquent Behavior). There were also significant associations between the Anxious/Depressed syndrome and Aggressive Behavior, Aggressive Behavior and Attention Problems, Aggressive Behavior and Anxious/Depressed, and between both Attention Problems and Aggressive Behavior and Social Problems.

In terms of change over time, Aggressive Behavior at the initiation of a study was significantly associated with Anxious/Depressed, Social Problems, Attention Problems, and Delinquent Behavior at follow-up 6 years later (Verhulst & Vanderende, 1993). Significant correlations of scores across time (.27 and above) were also found between Withdrawn and Anxious/Depressed, Anxious/Depressed and Somatic Complaints, Anxious/Depressed and Aggressive Behavior, Social Problems and Anxious/Depressed, Social Problems and Aggressive Behavior, Social Prob-

lems and Attention Problems, Attention Problems and Aggressive Behavior/Delinquent Behavior, and Delinquent Behavior and Aggressive Behavior. Prinzie, Onghena, and Hellinckx (2006) reported positive associations between individual trajectories of CBCL Aggressive and Delinquent Problem Behavior over a 3-year period in boys and girls 4 to 7 years of age, with a stronger association found for boys.

Einfeld et al. (2006) found that a population of children and adolescents with intellectual disability showed slight improvement in four of the five Developmental Behavior Checklist (DBC) subscales, on average, but Social Relating Problems were found to increase, on average. However, these investigators did not examine the extent to which these reductions in psychopathology tend to occur in the same individuals or whether different individuals are improving in different areas. Knowledge of the relationships between individual trajectories on the five DBC subscales would provide further insight into the developmental dynamics of psychopathology from childhood through young adulthood. For example, are individual differences in the rate of change of social relating behaviors correlated positively with change in self-absorbed behaviors? That is, do children who develop more social relating problems than average also tend to improve less than average in self-absorbed behaviors? Comparing mean trajectories is uninformative and potentially misleading with respect to inference about individual change because means can go up (or down) together over time without this same pattern applying to particular individuals. The resolution of this issue has important theoretical and practical ramifications because it would clarify the nature of psychopathology in this population—whether improvements in individual children typically involve isolated or general behavioral issues.

In this paper, our goal was to extend the univariate and group average findings of Einfeld et al. (2006) to a multivariate individual differences emphasis to address the question of whether individual-level changes across dimensions of psychopathology are correlated. The Australian Child to Adult Development Study, a 14-year (four occasions) Australian epidemiological longitudinal study of young people with intellectual disability, provides rigorous evaluation of the course and pattern of emotional and behavioral problems. The central outcomes

include performance on the subscales of the DBC (Einfeld & Tonge, 1992, 1995, 2002), indicating disruptive/antisocial, self-absorbed, communication disturbance, anxiety, and social relating behaviors, shown with example items in Table 1.

Multivariate growth curve models provide estimates of covariation among individual differences in initial status (i.e., level), rates of linear change, and systematic occasion-specific deviations (i.e., within-person correlation). The growth curve model essentially summarizes each person's data in terms of a regression line, with estimates of each individual's intercept, slope, and occasion-specific residual used as the outcomes in simultaneous analyses. In this case, we have an intercept, slope, and residual for each person on each of the DBC subscales and have evaluated the degree to which these characteristics (i.e., means, rates of change, trajectory deviations) are related. Correlations among the initial levels (i.e., intercepts) indicate similarity in the relative ordering of individuals at their initial time point across outcomes (i.e., relations among individual differences in initial status). Correlations among the slopes indicate the extent to which individual differences in linear change in one outcome are related to individual differences in linear change in another (i.e., correlated change). These correlations are in the time frame of the span of four waves of data collection and, thus, are analogous to slow change in individual traits. Within-person correlations among occasion-specific residuals, often neglected in the modeling of associations between trajectories, provide information regarding state-like, occasion-specific fluctuation in emotional and behavior disturbance after controlling for an individual's trait-like growth trajectories. Whereas the correlation between slopes taps the association of long-term trends, the correlation of residuals captures the extent to which short-term departures from the trend occur together. Within-person variation could result from factors such as transient changes in the child (e.g., illness) or exogenous causes, such as life events or external stressors (e.g., Sliwinski, Smyth, Hofer, & Stawski, 2006). We also evaluated whether a common factor model provides a sufficiently good fit to the structure of covariation among levels, slopes, and residuals. In this evaluation we go further than evaluating the degree of interdependence of the five subscales of the DBC by providing a formal test of whether

Table 1. Developmental Behavior Checklist (DBC) Subscales and Sample Items

Subscale	Sample item
Anxiety	Cries easily or for no reason, separation anxiety Distressed about being alone, fears things and/or situations Nightmares, loss of appetite
Communication Disturbance	Confuses use of pronouns, echolalia, talks to self Speaks in whispers/high pitched voice/other unusual tone or rhythm, overaffectionate, perseveration Unrealistically happy/elated, stands too close to others, doesn't mix with own age group
Disruptive	Abusive, runs away, temper tantrums, hides things Easily led by others, irritable, impulsive, jealous Kicks/hits others, noisy, lights fires, overactive, attention-seeking, stubborn, steals, lies
Self-Absorbed	Aloof, bangs head, poor attention span, mouths objects Pica, grinds teeth, hits self, wanders aimlessly Hums/whines/makes nonspeech noises, repetitive body/hand movements, screams a lot Plays with unusual objects, repetitive activities Stares at lights/spinning objects, laughs for no apparent reason
Social Relating	Aloof, does not show affection, tends to be a loner Does not respond to other's feelings, underactive Somatic symptoms, resists being cuddled/touched/held Sleeps too much

these patterns of covariation are consistent with common factor models of psychopathology.

Method

Sample

We recruited individuals from the epidemiological cohort ($N = 578$) of the Australian Child to Adult Development Study, which began in 1991 and whose participants came from every health, education, and family agency that provided services to children with intellectual disability of all levels. The first wave was composed of individuals ages 4 to 19 years whose families lived in six census districts of the states of New South Wales and Victoria, Australia. These areas were selected as representative of the Australian population in terms of social class, ethnicity, and urban/rural distribution (Einfeld & Tonge, 1996a).

The response rate was 80.2% for individuals with IQ less than 50, and 78.5% for individuals with IQ above 50. For participants with moderate, severe, or profound intellectual disability, the

ascertainment process is likely to be virtually complete. The population of 4 to 18 year olds in the census area was 172,000, which equates to a prevalence rate in this cohort of 3.04 per 1,000 comparable to a prevalence of 2.94 per 1,000 in 5 to 9 year olds without intellectual disability that was suggested by Quinn (1986). The recruitment procedure was likely to locate almost all children with moderate and more severe intellectual disability but only captures those with mild intellectual disability who receive services. As in other studies, some young people with the mildest forms of intellectual disability blend in to the general population and were not identified because they may not have had impairments in adaptive behavior that require services. Individuals in the cohort with mild intellectual disability may, therefore, be biased towards higher levels of disturbance.

A selection of parents of nonparticipants were contacted by telephone and asked the last question in the DBC regarding whether their child had any major or minor problems with his or her emotions or behavior (Einfeld & Tonge,

1996b). There was no difference between the participants and nonparticipants on this question. The major reason for not participating was an inability to contact or locate the caregivers, presumably because they had moved. Full details about recruitment and participant demographics are provided in Tonge and Einfeld (2003).

The mean age of the entire epidemiological cohort at Wave 1 was 12.1 years ($SD = 4.4$); at Wave 2, 16.5 ($SD = 4.5$); Wave 3, 19.5 ($SD = 4.5$); and Wave 4, 23.5 ($SD = 4.5$). Participation has been consistently high throughout the study. The response rate (excluding the 31 participants who died since Wave 1) was 82.5% at Wave 2 ($N = 477$), 78.5% ($N = 448$) at Wave 3, and 84% ($N = 438$) at Wave 4. Analyses were limited to individuals ages 5 to 19.5 years at the first wave ($N = 506$) because of the few individuals in the extreme range of sample ages. This analysis sample was comprised of 288 males and 218 females. In terms of intellectual disability, 165 were classified with mild intellectual disability (95 males, 70 females), 206 with moderate intellectual disability (112 males, 94 females), and 135 with severe/profound intellectual disability (81 males, 54 females).

Measures

Developmental Behavior Checklist-Primary Care Version (DBC-P). The DBC-P (Einfeld & Tonge, 1992, 1995, 2002), which is the primary measure of psychopathology for young people with intellectual disability ages 4 to 18 years, is a 96-item instrument completed by parents or other primary caregivers reporting problems with emotions or behavior for the target individual over the previous 6-month period (scoring for each item: 0 = not true as far as you know, 1 = somewhat or sometimes true, 2 = very true or often true). For the purposes of this longitudinal analysis, the investigators used the five factor analytically derived subscales: Disruptive/Antisocial, Self-Absorbed, Communication Disturbance Anxiety, and Social Relating. Evidence for content, criterion, construct, and concurrent validity has been demonstrated for the subscales (Einfeld & Tonge, 1992, 1995). The original psychometric validation of the DBC-P was established using the Wave 1 cohort (Einfeld & Tonge, 1995). Subsequent studies and reviews have confirmed that this measure has sound psychometric properties (Dekker, Nunn, & Koot, 2002; Hastings et al., 2001;

Lecavalier & Aman, 2005). All analyses were based on raw total scores for each DBC subscale.

Developmental Behavior Checklist for Adults (DBC-A). This measure, developed by Mohr (2003) and Mohr, Tonge, and Einfeld (2005), is a 107-item caregiver-completed checklist adapted from the DBC-P that includes 12 new items as well as a few other minor modifications. The DBC-A, designed for individuals over age 19, was developed after Wave 3 and implemented in Wave 4. Therefore, the DBC was used to assess all individuals over 19 years of age through Wave 3. Mohr et al. (2005) reported test-retest reliability of .75 and .85 for paid and family caregivers, respectively, and interrater reliability (family caregivers) of .72. Additionally, concurrent validity coefficients of the DBC-A with the Aberrant Behavior Checklist (Aman, Singh, Steward, & Field, 1985) and the PAS-ADD (Moss et al., 1998) are .63 and .61, respectively. For purposes of analysis across all four waves of assessment, the DBC-A was scored in the same manner and, for consistency, using the same items as the DBC-P.

Levels of intellectual disability. Children were categorized as having a mild, moderate, or severe/profound degree of intellectual disability. Categorization was based upon the results of IQ assessments (typically one of the Wechsler measures, as determined by the child's chronological age, or the Stanford-Binet) according to the ranges of intellectual disability specified in the *DSM-IV*. Specifically, the IQ cutoffs used were 50–55 to 70 (mild intellectual disability), 35–40 to 50–55 (moderate intellectual disability), below 35 (severe to profound intellectual disability), dependent upon the SD of the measure used. Assignment to categories was based upon the results of existing assessments as provided by parents/caregivers. In the absence of a current cognitive or developmental assessment, one of the psychologists involved in this study conducted an assessment.

Procedure

In the Australian Child to Adult Development Study, investigators have gathered data on a broad range of potential biopsychosocial risk and protective variables, including whether the person received mental health services (Einfeld & Tonge, 1995). Data collection has taken place at four time points: Wave 1 (1991–1992), Wave 2 (1995–1996), Wave 3 (1999), and Wave 4 (2002–2003)

through use of a questionnaire booklet mailed to the parents and caregivers of the young people with intellectual disability. Of the 506 participants, 60% lived at home and were rated by their mothers (67% if counting any family respondent). Only 14% of children lived in settings outside the home. Of these, 74% were rated by a professional caregiver. The remaining children were either rated by family (16%) or by a professional caregiver (2%) if living arrangement information was missing (< 2%), or had missing data on both rater and living arrangement variables. Few individuals who entered the study while living at home transitioned subsequently to a care ($n = 24$) or independent ($n = 16$) situation. In 71% of transitions to care, rating was subsequently done by a professional caregiver.

Statistical Analysis

We used multivariate growth curve modeling to estimate the associations among individual differences in change in distinct features of psychopathology. In longitudinal designs, time is nested within-person, and methods such as growth curve models (a specific case of mixed or random effects models) are a means of properly addressing the corresponding within-person correlations (see Willett, Singer, & Martin, 1998, for application in psychopathology research). Conceptually, these models involve estimating the regression of the outcome of interest (i.e., DBC subscales) on time for each individual (often labeled Level 1) and predicting the regression parameters of these within-person trajectories (i.e., each participant's level and slope) with between-person covariates (Level 2). The model summarizes individual DBC values at each occasion of measurement in terms of true initial level of disturbance (intercept), slope (rate of change), and occasion-specific residual variance parameters. It is useful to think of this residual as a mix of random error and systematic state-like short-term fluctuations around the model-implied trajectory. Within each growth curve, residuals have an expectation of zero at any particular occasion. Correlations would be expected if individual differences in two or more psychopathology scales exhibited related systematic patterns of fluctuation at each occasion.

Models of correlated age-conditional slopes at the between-person level are typically based on smoothed (e.g., linear) individual model-implied

trajectories over time, with the time-to-time dynamics usually considered unmodeled residual error components. The latent growth metric used here is time in study, with age at baseline used as a Level 2 (between-person) predictor to account for the initial age heterogeneity in the sample. The analysis of coupled change is based on the recognition of the state-like component of these residuals, and attention to covariation among the state portions of two or more growth curves. This provides information regarding the systematic occasion-specific fluctuation across different types of psychopathology indicated by the DBC subscales. For parsimony, the state residuals are assumed to have equal variance across occasions (i.e., they are estimated as a single model parameter) and to be uncorrelated over time within each outcome. However, the between-outcome covariances of these occasion-specific residuals are estimated and interpreted as lower bounds of the degree to which state fluctuations are “coupled” within individuals because these residuals are composed of both systematic and error variance.

The associations of level, slope, and occasion-specific residuals for each pair of growth curves form ordinary covariance matrices that may in turn yield additional information about developmental processes. matrices of (a) level covariances, (b) slope covariances, and (c) occasion-specific residuals can be formed. Factor analysis of these covariance matrices can help reveal the extent to which common underlying causes may be responsible for the estimated correlations and couplings (e.g., Does a common cause account for model estimated levels, model estimated rates of change, and/or the state-like time-specific fluctuations?)

Simultaneous multivariate models of all five DBC subscales were fitted using Mplus v5.0 (Muthén & Muthén, 1998–2008) based on a time-in-study data structure, with individually varying intervals between occasions of measurement. Maximum likelihood estimation was used to accommodate incomplete data (missing values; attrition) and provide unbiased population estimates under the assumption that the data are “missing at random” (i.e., missingness is accounted for by covariates and prior values in a longitudinal study; see Little & Rubin, 1987). The intercept (i.e., level) was specified to be at the first occasion of measurement for each individual, with both the level and linear slope conditional

on age at the Wave 1 baseline. The participant's age was centered relative to the mean Wave 1 age ($M = 12.0$ years, $SD = 3.9$), permitting interpretation of this between-person effect to reflect the average age of the sample. Follow-up occurred an average of 4.5, 7.5, and 11.5 years later for Waves 2, 3, and 4, respectively. Time-in-study was used as the metric of change in order to obtain separate estimates of the between-person (cross-sectional) and within-person (longitudinal) effects of age, respectively represented on Level 2 by centered age at Wave 1 and on Level 1 by time since Wave 1. Other Level 2 predictors included gender (with boys as the reference) and intellectual disability (mild, moderate, or severe, with mild intellectual disability as the reference). For each bivariate growth curve model, the occasion-specific residual variances and covariances were constrained to equality, with one estimate for the residual parameter of each outcome and a single covariance parameter between occasion-specific residual variances. As standardized estimates (i.e., correlations) are not provided by the Mplus software when individually varying time intervals are modeled, these correlations were computed based on the estimated variances and covariances for each of the parameters of interest.

We report results from a reduced (minimally conditioned) model that conditions time-in-study change only on age at baseline, and we contrast these results to those from a fully conditioned model that included age, gender, intellectual disability status, and all two-way interactions as predictors of the growth parameters. The degree to which the associations among rates of change are accounted for by these additional Level 2 predictors provides a basis for understanding the influence of gender and intellectual deficit characteristics on correlated and coupled change processes. Subsequently, secondary analyses of the estimated variances and covariances were undertaken to evaluate whether the covariation among initial status, linear rates of change, and occasion-specific residuals could be accounted for by common factor models.

Results

We first examined the shape of the subscale trajectories and the extent to which trajectory parameters varied between individuals. A model including only fixed effects for linear and quadratic components of time was compared to

a mixed model with the same fixed effects of time but in which an additional random-effect allowed the linear fixed-effect of time to vary over individuals. The difference in the deviance statistics ($-2 \log$ likelihood) between these two models was significant for each outcome, indicating that individual differences in the rate of change (in addition to the level) were significant for all of the outcomes. Because the quadratic fixed-effect of time was not significant, it was not retained; change was, therefore, modeled as a straight line with a fixed (group averaged) and random (individual) component. The linear model provided a good fit to each DBC subscale, and polynomial models did not improve model fit.

Table 2 provides the coefficients and standard errors for each DBC subscale based on the multivariate model, with level and slope conditional on baseline age. Between-person differences in age predicted behavior problem level for Anxiety, Self-Absorbed, and Social Relating, with older individuals showing worse problems on Social Relating, but fewer problems on Anxiety or Self-Absorbed. A statistically significant decrease over time (linear slope growth factor) was observed for all subscales, except Social Relating, which indicated increasing problem behaviors over time. Age also predicted rate of change for all subscales except Disruptive, with older individuals showing more rapid decline in Communication Disturbance, slower decline in Anxiety and Self-Absorbed behaviors, and slower increases in Social Relating problems.

Correlations between subscale levels, slopes, and residuals derived from the models including the predictors age, gender, and intellectual disability status were very similar to those from the age-conditional model (correlations differed by $< .02$). Accordingly, results for the age-only models are presented. Table 3 shows the correlations among levels, slopes, and occasion-specific residuals for the age conditional growth model estimates. Correlations ranged from .43 to .66 for initial levels, from .43 to .88 for slopes, and .31 to .61 for within-person residuals. The highest correlations were consistently between Disruptive, Self-Absorbed, and Communication Disturbance behaviors. All covariance estimates from which the correlations were derived were significant at the .05 level.

Correlations among the levels (i.e., expected intercepts at age 12) are similar to what might be

Table 2. Estimates and Standard Errors for Developmental Behavior Checklist (DBC) Subscales

Parameter	Communication			Self-	Social
	Anxiety	Disturbance	Disruptive	Absorbed	Relating
Fixed effects					
Level					
Estimate	4.09*	5.77*	13.49*	14.20*	4.49*
	(0.14)	(0.19)	(0.43)	(0.45)	(0.15)
Age (12)	−0.08*	0.00	−0.02	−0.40*	0.12*
	(0.04)	(0.05)	(0.11)	(0.11)	(0.04)
Rate of change					
Estimate	−0.06*	−0.05*	−0.27*	−0.32*	0.06*
	(0.02)	(0.02)	(0.04)	(0.04)	(0.02)
Age (12)	0.01*	−0.01*	0.01	0.02*	−0.01*
	(0.00)	(0.01)	(0.01)	(0.01)	(0.00)
Variance components					
Level					
	6.27*	13.49*	74.43*	86.61*	7.28*
	(.69)	(1.25)	(6.15)	(7.11)	(0.80)
Slope					
	0.04*	0.07*	0.31*	0.27*	0.03*
	(0.01)	(0.02)	(0.05)	(0.05)	(0.01)
Residual					
	3.80*	6.22*	21.52*	20.73*	4.66*
	(0.28)	(0.44)	(1.67)	(1.67)	(0.34)

Note. Standard errors are shown in parentheses. For details of full model in univariate analysis, see Einfeld et al. (2006). * $p < .01$.

obtained by age-adjusted correlation between scales in a cross-sectional study. In a longitudinal model, however, these levels are based on the linear model for the individual repeated measures and provide more reliable estimates given the correction for occasion-specific variability. The strongest correlations were between the DBC subscales Self-Absorbed and Social Relating, and between Communication Disorder and Disruptive, followed by Anxiety and Communication Disorder, Anxiety and Disruptive, and Anxiety and Social Relating. Scatterplots of the expected level–level associations (based on factor scores output from the longitudinal model) are shown in the upper triangle of Figure 1. Each scatterplot shown in the upper triangle graphically illustrates one of the correlations of levels reported in Table 3. For example, the highest correlation, level of Self-Absorbed with level of Social Relating, $r = .67$, show the tightest elongated scatter cloud. It can also be seen that the level–level scatterplots show a wider scatter at higher levels, indicating that a substantial proportion of individuals can be high on one subscale, but not on the other. For example, it is not uncommon

for a child to have very high levels of disruptive behavior, but only moderate levels of self-absorbed behavior (Row 3, Column 4). In contrast, improvements in disruptive behavior over time are almost universally paired with improvements in self-absorbed behavior (Row 4, Column 3).

The slope–slope correlations represent the extent to which within-person trajectories of different types of problem behaviors are related between-persons. The strongest correlations were among Disruptive, Self-Absorbed, and Communication Disturbance. Figure 1 (lower triangle) is a graphical representation of these correlations, with most individuals who change substantially on one subscale also exhibiting change on the other. For example, the highest correlation, slope of Disruptive with slope of Self-Absorbed, $r = .87$, shows the tightest elongated scatter cloud; whereas while the weakest subscale slope correlation, Anxiety with Social Relating, $r = .39$, shows the most diffuse scatter cloud.

The occasion-specific residual correlations represent the extent to which perturbations in an individual's trajectory at particular occasions

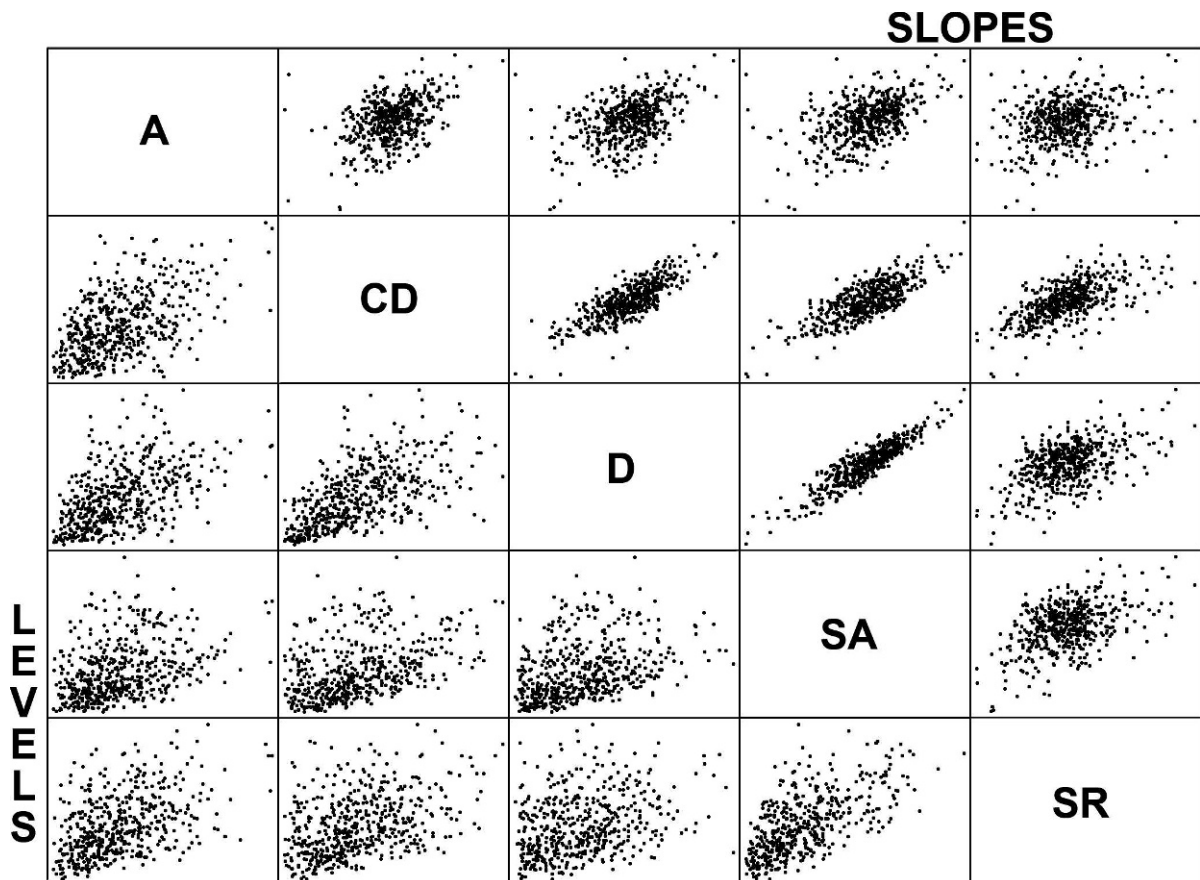


Figure 1. Scatterplots for level–level and slope–slope correlations across Developmental Behavior Checklist (DBC) subscales. A = Anxiety, CD = Communication Disturbances, D = Disruptive, SA = Self-Absorbed, SR = Social Relating.

are related across outcomes after controlling for individual change. The estimates are moderate across most of the scales and provide evidence for systematic occasion-specific fluctuation in emotional and behavioral disturbance. In particular, occasion-specific variation in subscales Disruptive, Communication Disturbance, and Self-Absorbed showed the strongest occasion-level correlations.

An evaluation of whether single common factor models could sufficiently account for the pattern of correlations among levels, slopes, and occasion-specific residuals was undertaken in a second stage analysis of the estimated correlations (reported in Table 3). Factor loadings for separate factor analyses of initial level, linear slope, and occasion-specific residuals are provided in Table 4. These analyses account for 55, 64, and 45% of the variance in the subscale levels, slopes, and occasion-specific residuals, respectively. Conventional standards (Hu & Bentler, 1999) deem the

Tucker-Lewis Index (TLI) $\geq .95$, and the root mean squared error of approximation (RMSEA) $\leq .06$ to be indicative of good model fit, although comparative fit index (CFI) and TLI ≥ 0.90 and RMSEA $\leq .10$ are often considered adequate model fit.

Although loadings were high, and a single factor accounted for a substantial proportion of the variation of initial levels, fit indices lay outside ranges associated with acceptable overall model fit (CFI = .96, TLI = .91, RMSEA = .13). This was also the case for the factor analysis of correlations of rates of change among the DBC subscales (CFI = .97, TLI = .94; RMSEA = .15). A common factor model did, however, provide a very good fit for the occasion-specific residuals (CFI = .99, TLI = .99, RMSEA = .04).

In the case of the occasion-specific residuals (i.e., within-person variation), there were no sources of significant misfit among the observed correlations. Although there is certainly evidence

Table 3. Random Effects Correlation Estimates for DBC Subscales

DBC subscale	Subscale correlation: Conditional on age and age/gender/IQ					
	Level		Slope		Occasion-specific residual	
	Age only	Age/gender/IQ	Age only	Age/gender/IQ	Age only	Age/gender/IQ
Anxiety with						
Communication						
Disturbance	.50	.49	.53	.52	.39	.39
Disruptive	.50	.48	.47	.44	.42	.42
Self-Absorbed	.37	.47	.46	.43	.38	.38
Social Relating	.46	.50	.39	.44	.31	.31
Communication						
Disturbance with						
Disruptive	.60	.57	.80	.81	.50	.50
Self-Absorbed	.46	.64	.85	.83	.52	.52
Social Relating	.42	.52	.56	.57	.40	.40
Disruptive with						
Self-Absorbed	.42	.62	.87	.88	.61	.61
Social Relating	.34	.43	.57	.62	.39	.39
Self-Absorbed with						
Social Relating	.67	.66	.56	.57	.45	.46

Note. All associations were statistically significant at the $p < .01$ level. DBC = Developmental Behavior Checklist.

for substantial shared covariation among levels and rates of change across the DBC subscales, the fit of these models would not generally be regarded as acceptable and indicates that pairs of subscales correlate more or less strongly than would be implied by a model positing a single latent variable “driving” each of initial status, change, and time-specific responses. More specifically, the common factor model did not sufficiently account for the correlations among Self-Absorbed and Anxiety, Social Relating and

Disruptive, and Social Relating and Self-Absorbed subscales. The sources of model misfit for the common factor model of slopes were Communication Disturbance and Anxiety, Disruptive and Communication Disturbance, Self-Absorbed and Anxiety, Self-Absorbed and Disruptive, Social Relating and Anxiety, and Social Relating and Self-Absorbed. This is an indication of distinct individual differences in manifestations of psychopathology across individuals in both a cross-sectional and longitudinal change context. In

Table 4. Confirmatory Factor Analysis of the Intercepts, Slopes, and Occasion-Specific Residuals

DBC ^a subscale	Level		Slope		Occasion-specific residual	
	Factor loading	R ²	Factor loading	R ²	Factor loading	R ²
Anxiety	.62	.38	.50	.25	.53	.28
Communication						
Disturbance	.76	.57	.88	.78	.67	.46
Disruptive	.71	.51	.93	.87	.76	.57
Self-Absorbed	.86	.74	.94	.88	.79	.62
Social Relating	.73	.53	.64	.41	.56	.32

Note. R² is the proportion of total variance explained in the indicator variable by the common factor.

^aDevelopmental Behavior Checklist.

summary, although there are significant associations among levels, slopes, and occasion-specific variation, a single underlying factor may not adequately account for the entire pattern of correlations among levels and slopes but does satisfactorily account for the pattern of within-person variation at each occasion.

Discussion

The present study modeled the interdependence of developmental change and variation in psychopathology as assessed by the DBC (Einfeld & Tonge, 1995) within a population of individuals with intellectual disability. These multivariate analyses extend the univariate analysis of change reported by Einfeld et al. (2006) by considering the structure of change in psychopathology in terms of correlated levels and rates of change between persons and coupling of temporal dynamics within persons. The central finding of this study is evidence for moderate and systematic interdependencies across distinct aspects of emotional and behavioral disturbance. A better understanding of the dynamics (change and variation) in psychopathology from childhood to young adulthood is one of the major strengths of applying multivariate growth curve methodology to longitudinal data.

In this study we adopted a dimensional approach to behavioral and emotional disturbance. An alternative would be a categorical one, either in respect of clinical “caseness” or “non-caseness,” or in respect of a categorical approach to specific psychiatric diagnoses. The disadvantage of the former would be that if many individuals start in the study close to the cutoff, one may see lots of apparent “change” that does not involve notable changes in behavior. In fact, this is the case for the Australian Child to Adult Development sample and the DBC. The modal DBC score was 42, close to the clinical cutoff of 46. However, if many individuals start far from the cutoff (much higher or lower), then larger real changes may not be noticed at all from the cutoff perspective. Stated another way, using cutoffs focuses the analysis on change in a very specific range of behavior rather than across the entire range. With respect to the measurement of specific diagnoses, this presumes that the reliability and validity of a broad range of potential diagnoses has been well-established in this population. With the possible exception of Autistic Disorder and Stereotypic Self-Injury, this is still

not the case (see Einfeld and Aman, 1995, for a discussion of this).

On average, the severity of psychopathology in four of the five DBC subscales declined over the course of the study, with the exception of the Social Relating scale. One possible explanation for the increase in Social Relating problems is that these children and youths appear more physically mature over time and are placed into more demanding social settings (e.g., school, day-care, occupational settings), where the same level of behavioral disturbance may be rated more negatively. Alternatively, the behavior related to most of these items might be self-intensifying and, thus, truly increase over time. The actual items and pattern of loadings for the Social Relating construct shows a broad/mixed item set ranging from nonsocial (e.g., moves slowly, underactive, prefers to do things on his/her own) to behaviors that create social barriers (e.g., avoids eye contact, doesn't respond to other's feelings, doesn't show affection). It is important to consider the meaning of this construct in the developmental context when conceptualizing change in the Social Relating construct (Bontempo et al., 2008). Additionally, the increase in scores on the Social Relating subscale, when considering the population as a whole, may reflect the emergence of depression as the cohort reaches young adulthood because a number of the items of this subscale suggest depressive symptoms (e.g., appears depressed, downcast, or unhappy). This possibility is the subject of ongoing investigation in our research group.

In the current study we found significant correlations between the levels of all of the subscales of the DBC, a finding comparable to earlier work. Correlations between the levels of the DBC subscales can be seen as the underlying dimensional foundation of previous researchers reporting higher rates of diagnostic comorbidity (Dekker & Koot, 2003; Emerson, 2003). The central focus of this report, however, is on the substantial heterogeneity in developmental change as well as in systematic occasion-specific variation. The multivariate analysis of the dimensions of emotional and behavioral disturbance tapped by the DBC subscales indicate moderate to substantial interdependence in long-term change patterns and moderate associations among occasion-specific fluctuations of psychopathology.

Beyond previous research, the multivariate growth model demonstrates that children who show increases or decreases over time on one

subscale tend to exhibit similar changes (relative to other children) on another. Put simply, changes for an individual on one subscale of the DBC tend to be mirrored by change in the other subscales. The strongest relationships were consistently between the Disruptive, Self-Absorbed, and Communication Disturbance subscales. The DBC Disruptive subscale is similar to the Aggressive-Delinquent cross-informant syndrome of the CBCL, being primarily concerned with externalizing behavior problems. The Self-Absorbed subscale reflects withdrawn and nonsocial behaviors, and the Communication Disturbance subscale contains a mix of abnormal communication items and social difficulties. The Self-Absorbed and Communication Disturbance subscales both contain elements of nonsocial behavior and social difficulties. The behaviors described in these three subscales are interrelated in a substantial and statistically significant way. Changes in Anxiety were the least correlated with the other subscales, although even changes on this subscale were significantly correlated with the others.

Common factor models were subsequently evaluated to provide evidence for common versus specific patterns of level, change, and occasion-specific residual variation in psychopathology. Whether or not a common factor model provides a fit to the data has more to do with the general pattern (i.e., consistency) of covariation than to the magnitude of correlation among DBC subscales and should not be taken as direct evidence for a common or unitary cause. Only in the case of factor analysis of the occasion-specific residuals was the pattern of covariation sufficiently consistent with a common factor model. The fit was marginal for level and rate of change in DBC subscales (i.e., some correlations among subscales that were not sufficiently accounted for by the factor model). The common factor model of residuals provides indirect evidence for consistent “state-like” transient behavioral and emotional disturbances across different features of psychopathology. This common covariation may be also related to situational changes in a child’s circumstances, such as contextual stressors associated with changes in living conditions or life events, which are not cumulative. This finding is not an artifact of model misspecification because the linear model provided a satisfactory fit to each DBC subscale and polynomial models did not improve fit. Further examination of individual

differences in within-person variation, in addition to systematic change over long periods of time, is certainly warranted and would benefit from short-term intensive measurement studies of within-person change and variation in psychopathology.

In previous studies our research group identified high levels of behavior disturbance even in the youngest members of the Australian Child to Adult Development cohort (Einfeld et al., 2006; Tonge & Einfeld, 2003). One model that may explain these findings is that the vulnerability to psychopathology is brought on by the same process that causes developmental delay. We hypothesize that the process is brain impairment, sufficiently widespread to cause both intellectual impairment and wide-ranging behavioral and emotional disturbance. That brain impairment is caused by a multiplicity of factors, chiefly genetic, has been described previously for this cohort (Partington, Turner, Mowat, Einfeld, & Tonge, 2000). The early high levels of emotional and behavioral problems, the persistence of these problems even given the relatively small steady decline in degree across all types of disturbance except for social relating behaviors, and the lack of influence of gender on the profile of disturbance, all point to the salience of the underlying organic deficits in brain functioning as the predominant cause of the broad range of types of problem behaviors and emotions. The progressive maturation of the brain, despite persistence of congenital or early acquired impairment, might account for an increase in resilience and a steady, though slow, decline in symptoms of anxiety, disruptive and self-absorbed behaviors, and communication disturbance. Of course, the changes in psychopathology may reflect the benefits of education, family and community support, or acquisition of social and daily living skills. However, the Australian Child to Adult study does not include measures of these possible influences, so no comment can be offered in this regard.

The clinical implications of the continuing effects of impaired brain function on cognition, affect regulation, impulse control, and social and adaptive behavior are that education, family support, and general disability services are needed to address the burden of psychopathology in children with intellectual disability. Unless early intervention can be shown to be effective in altering the course of psychopathology as found in this study, then behavioral and mental health

interventions will need to extend from early intervention through young adulthood.

Although we have examined changes in the severity of psychopathology over time, we are not currently able to evaluate whether the average decline or individual differences in change in psychopathology is related to particular individual interventions or treatments received. The sample was recruited from every health, education, and family agency that provided services to children with intellectual disability of all levels, though especially with respect to moderate, severe, and profound levels of intellectual deficit. In this study we did not limit or encourage individual treatments (e.g., use of psychotropic medications or behavioral interventions). Less than 10% of the individuals in our sample who had clinically significant levels of psychopathology received mental health services (Einfeld & Tonge, 1996a, 1996b). These findings of moderate interdependency of change in different aspects of behavioral and emotional disturbance are, therefore, likely to reflect the natural history of psychopathology that is relatively independent of any specific mental health intervention. The results of this study highlight the links between disruptive behavior problems and withdrawn and nonsocial behaviors in children and adolescents with intellectual disability. When planning interventions, professionals should consider this relationship, ensuring that treatment programs include elements aimed at assisting young people to develop their social interaction and communication skills and providing opportunities for supported social interaction. Research is needed to investigate this relationship further, with high priorities being examination of the broad impact of targeted treatments for disruptive behavior and determination of whether subsequent improvements in social and communication behaviors are evident.

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