Introduction to Bayesian Statistics and Markov Chain Monte Carlo Estimation

PSYC 943 (930): Fundamentals of Multivariate Modeling Lecture 17: October 25, 2012

Today's Class

- An introduction to Bayesian statistics:
 - > What it is
 - > What it does
 - > Why people use it
- An introduction to Markov Chain Monte Carlo (MCMC estimation)
 - How it works
 - Features to look for when using MCMC
 - > Why people use it

AN INTRODUCTION TO BAYESIAN STATISTICS

Bayesian Statistics: The Basics

- Bayesian statistical analysis refers to the use of models where some or all of the parameters are treated as random components
 - Each parameter comes from some type of distribution
- The likelihood function of the data is then augmented with an additional term that represents the likelihood of the prior distribution for each parameter
 - Think of this as saying each parameter has a certain likelihood the height of the prior distribution
- The final estimates are then considered summaries of the posterior distribution of the parameter, conditional on the data
 - In practice, we use these estimates to make inferences, just as we have when using the non-Bayesian approaches we have used throughout this class (e.g., maximum likelihood/least squares)

Bayesian Statistics: Why It Is Used

- Bayesian methods get used because the <u>relative</u> accessibility of one method of estimation (MCMC – to be discussed shortly)
- There are three main reasons why people use MCMC:
- 1. Missing data
 - Multiple imputation: MCMC is used to estimate model parameters then "impute" data
 - > More complicated models for certain types of missing data
- 2. Lack of software capable of handling large sized analyses
 - Have a zero-inflated negative binomial with 21 multivariate outcomes per 18 time points?
- 3. New models/generalizations of models not available in software
 - > Have a new model?
 - > Need a certain link function not in software?

Bayesian Statistics: Perceptions and Issues

- Historically, the use of Bayesian statistics has been controversial
 - The use of certain prior distributions can produce results that are biased or reflect subjective judgment rather than objective science
- Most MCMC estimation methods are computationally intensive
 - Until recently, very few methods available for those who aren't into programming in FORTRAN or C++
- Understanding of what Bayesian methods are and how they work is limited outside the field of mathematical statistics
 - Especially the case in the social sciences
- Over the past 15 years, Bayesian methods have become widespread

 making new models estimable and becoming standard in some social science fields (quantitative psychology and educational measurement)

HOW BAYESIAN METHODS WORK

How Bayesian Statistics Work

- The term Bayesian refers to Thomas Bayes (1701-1761)
 - Formulated Bayes' Theorem
- Bayesian methods rely on Bayes' Theorem:

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

- > P(A) is the **prior distribution (pdf) of A** \rightarrow WHY THINGS ARE BAYESIAN
- > *P*(*B*) is the marginal distribution (pdf) of B
- > P(B|A) is the conditional distribution (pdf) of B, given A
- > P(A|B) is the **posterior distribution (pdf) of A, given B**
- Bayes' Theorem Example...

Imagine a patient takes a test for a rare disease (present 1% of the population) that has a 95% accuracy rate...what is the probability the patient actually has the disease?

Bayes' Theorem Example

Imagine a patient takes a test for a rare disease (present 1% of the population) that has a 95% accuracy rate...what is the probability the patient actually has the disease?

- D = the case where the person actually has the disease
- ND = the case where the person does not have the disease
- + = the test for the disease is positive

The question is asking for: P(D|+) From Bayes' Theorem:

$$P(D|+) = \frac{P(+|D)P(D)}{P(+)}$$

What we know:

$$P(D) = .01$$

 $P(+|D) = .95$

Back to Distributions

- We don't know P(+) directly from the problem, but we can figure it out if we recall how distributions work:
- P(+) is a marginal distribution
- P(+|D) is a conditional distribution
- We can get to the marginal by summing across the conditional: P(+) = P(+|D)P(D) + P(+|ND)P(ND) = .95 * .01 + .05 * .99 = .059
- So, to figure out the answer, if a person tests positive for the disease, the **posterior probability** they actually have the disease is: $P(D|+) = \frac{P(+|D)P(D)}{P(+)} = \frac{.01 * .99}{.059} = .17$

A (Perhaps) More Relevant Example

- The old-fashioned Bayes' Theorem example I've found to be difficult to generalize to your actual data, so...
- Imagine you administer an IQ test to a sample of 50 people

>
$$y_p$$
 = person p's IQ test score

• To put this into a linear-models context, the empty model for Y: $y_p = \beta_0 + e_p$

Where $e_p \sim N(0, \sigma_e^2)$

- From this empty model, we know that:
 - > β_0 is the mean of the Y (the mean IQ)
 - > σ_e^2 is the sample variance of Y
 - > The conditional distribution of Y is then: $f(y_p | \beta_0, \sigma_e^2) \sim N(\beta_0, \sigma_e^2)$

Non-Bayesian Analysis (i.e., Frequentist Approach)

- Up to this point in the class, we have analyzed these data using ML
- For ML, we maximized the joint likelihood of the sample with respect to the two unknown parameters β_0 and σ_e^2 $L(\beta_0, \sigma_e^2) = \prod_{p=1}^N f(y_p | \beta_0, \sigma_e^2) = \prod_{p=1}^N \frac{1}{\sqrt{2\pi\sigma_e^2}} \exp\left(-\frac{(y_p - \beta_0)^2}{2\sigma_e^2}\right)$
- Here, using PROC MIXED, I found:

$$\beta_0 = 102.769$$

 $\sigma_e^2 = 239.490$

• Also, I found:

$$-2LogL = 415.8$$

Setting up a Bayesian Approach

- The (fully) Bayesian approach would treat each parameter as a random instance from some prior distribution
- Let's say you know that this version of the IQ test is supposed to have a mean of 100 and a standard deviation of 15
 - > So β_0 should be 100 and σ_e^2 should be 225
- Going a step further, let's say you have seen results for administrations of this test that led you to believe that the mean came from a normal distribution with a SD of 2.13

> This indicates the prior distribution for the **mean**...or

 $f(\beta_0) \sim N(100, 2.13^2)$

 Let's also say that you don't really have an idea as for the distribution of the variance, but you have seen it range from 200 to 400, so we can come up with a prior distribution for the variance of:

 $f(\sigma_e^2) \sim U(200,\!400)$

 Here the prior is a uniform distribution meaning all values from 200 to 400 are equally likely

More on the Bayesian Approach

 The Bayesian approach is now to seek to find the **posterior** distribution of the parameters given the data:

 $f(\beta_0, \sigma_e^2 | \boldsymbol{y}_p)$

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

- We can again use Bayes' Theorem (but for continuous parameters): $f(\beta_0, \sigma_e^2 | \mathbf{y}_p) = \frac{f(\mathbf{y}_p | \beta_0, \sigma_e^2) f(\beta_0, \sigma_e^2)}{f(\mathbf{y}_p)} = \frac{f(\mathbf{y}_p | \beta_0, \sigma_e^2) f(\beta_0) f(\sigma_e^2)}{f(\mathbf{y}_p)}$
- Because $f(\mathbf{y}_p)$ essentially is a constant (which involves integrating across β_0 and σ_e^2 to find its value), this term is often referred to as: $f(\beta_0, \sigma_e^2 | \mathbf{y}_p) \propto f(\mathbf{y}_p | \beta_0, \sigma_e^2) f(\beta_0) f(\sigma_e^2)$
- The symbol ∝ is read as "is proportional to" meaning it is the same as when multiplied by a constant

> So it is the same for all values of β_0 and σ_e^2

Unpacking the Posterior Distribution

- $f(\mathbf{y}_p | \beta_0, \sigma_e^2)$ is the **conditional distribution** of the data given the parameters we know this already from our linear model (slide 12) $f(\mathbf{y}_p | \beta_0, \sigma_e^2) = \prod_{p=1}^N f(y_p | \beta_0, \sigma_e^2) = \prod_{p=1}^N \frac{1}{\sqrt{2\pi\sigma_e^2}} \exp\left(-\frac{(y_p - \beta_0)^2}{2\sigma_e^2}\right)$
- $f(\beta_0)$ is the **prior distribution** of β_0 , which we decided would be $N(100,2.13^2)$, giving the height of any β_0 :

$$f(\beta_0) = \frac{1}{\sqrt{2\pi\sigma_{\beta_0}^2}} \exp\left(-\frac{\left(\beta_0 - \mu_{\beta_0}\right)^2}{2\sigma_{\beta_0}^2}\right)$$
$$= \frac{1}{\sqrt{2\pi * 2.13^2}} \exp\left(-\frac{\left(\beta_0 - 100\right)^2}{2 * 2.13^2}\right)$$

Unpacking the Posterior Distribution

• $f(\sigma_e^2)$ is the **prior distribution** of σ_e^2 , which we decided would be U(200,400), giving the height of any value of σ_e^2 as:

$$f(\sigma_e^2) = \frac{1}{b_{\sigma_e^2} - a_{\sigma_e^2}} = \frac{1}{400 - 200} = \frac{1}{200} = .005$$

- Some useful terminology:
 - > The parameters of the model (for the data) get prior distributions
 - The prior distributions each have parameters these parameters are called hyper-parameters
 - The hyper-parameters are not estimated in our example, but could be giving us a case where we would call our priors empirical priors
 - AKA random intercept variance

Up Next: Estimation (first using non-MCMC)

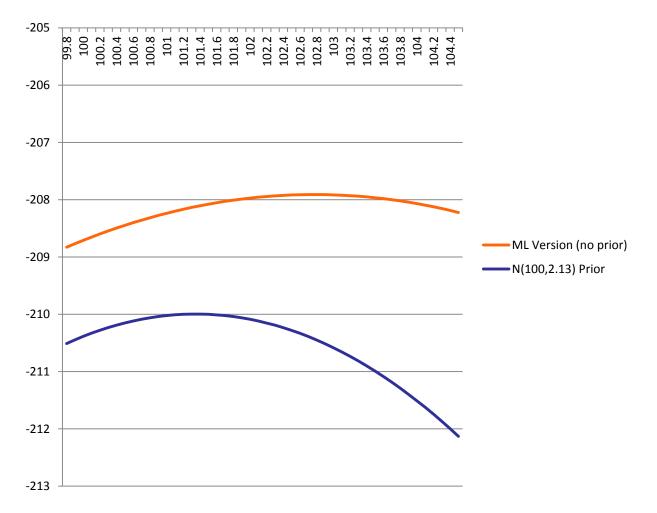
- Although MCMC is commonly thought of as the only method for Bayesian estimation, there are several other forms
- The form analogous to ML (where the value of the parameters that maximize the likelihood or log-likelihood) is called Maximum a Posteriori estimation (MAP)
 - The term modal comes from the maximum point coming at the peak (the mode) of the posterior distribution
- In practice, this functions similar to ML, only instead of maximizing the joint likelihood of the data, we now have to worry about the prior: $f(\beta_0, \sigma_e^2 | \mathbf{y}_p) = \frac{f(\mathbf{y}_p | \beta_0, \sigma_e^2) f(\beta_0) f(\sigma_e^2)}{f(\mathbf{y}_p)} \propto f(\mathbf{y}_p | \beta_0, \sigma_e^2) f(\beta_0) f(\sigma_e^2)$
- Because it is often more easy to work with, the log of this is often used: $\log \left(f(\beta_0, \sigma_e^2 | \mathbf{y}_p) \right) \propto \log f(\mathbf{y}_p | \beta_0, \sigma_e^2) + \log f(\beta_0) + \log f(\sigma_e^2)$

Grid Searching for the MAP Estimate of β_0

- To demonstrate, let's imagine we know σ_e² = 239.490
 Later we won't know this...when we use MCMC
- We will use Excel to search over a grid of possible values for β_0
- In each, we will use $\log f(\mathbf{y}_p | \beta_0) + \log f(\beta_0)$
- As a comparison, we will also search over the ML log likelihood function $\log f(\pmb{y}_p | \pmb{\beta}_0)$

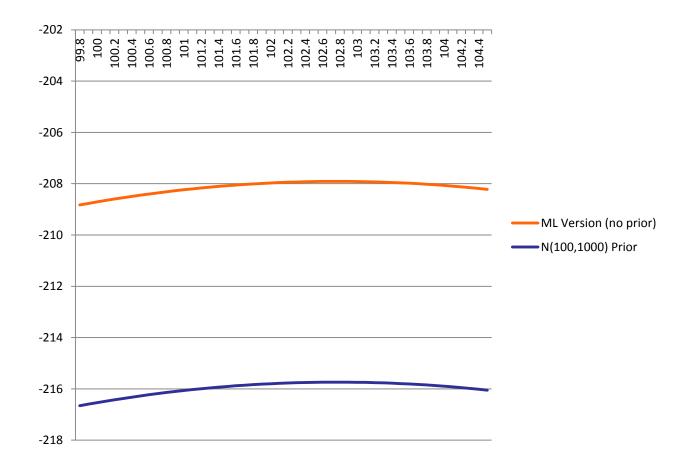
ML v. Prior for β_0 of N(100, 2.13²)

- Maximum for ML: 102.8
- Maximum for Bayes: 101.4 (estimate is closer to mean of prior)



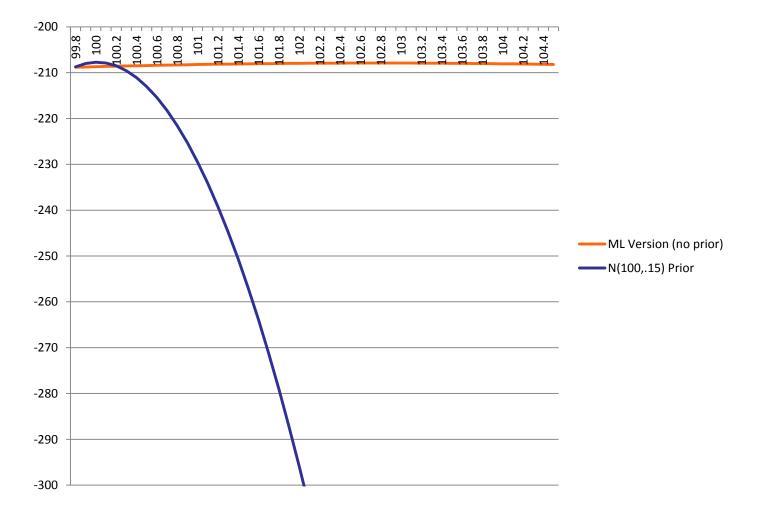
ML vs. Prior for β_0 of N(100, 1000²)

- Maximum for ML: 102.8
- Maximum for Bayes: 102.8



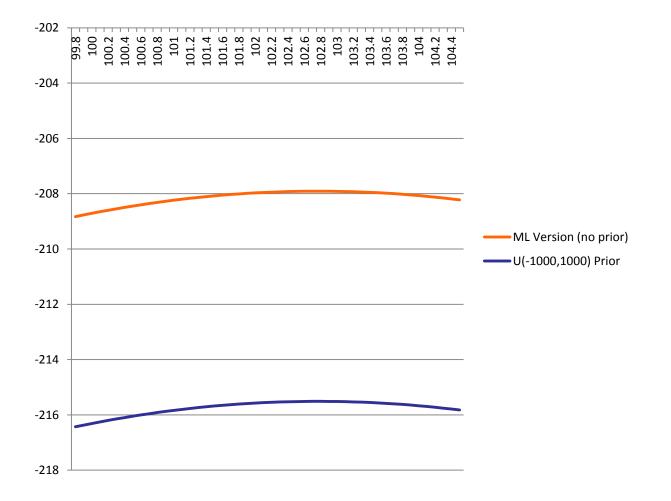
ML vs. Prior for β_0 of N(100, 0.15²)

- Maximum for ML: 102.8
- Maximum for Bayes: 100



ML vs. Prior for β_0 of U(-1000,1000)

- Maximum for ML: 102.8
- Maximum for Bayes: 102.8



- Bayesian \rightarrow parameters have prior distributions
- Estimation in Bayesian → MAP estimation is much like estimation in ML, only instead of likelihood of data, now have to add in likelihood for prior of all parameters
 - But...MAP estimation may be difficult as figuring out derivatives for gradient function (for Newton Raphson) are not always easy
 - > Where they are easy: Conjugate priors → prior distributions that are the same as the posterior distribution (think multilevel with normal outcomes)
- Priors can be informative (highly peaked) or uninformative (not peaked)
 - > Some uninformative priors will give MAP estimates that are equal to ML
- Up next: estimation by brute force: Markov Chain Monte Carlo

MARKOV CHAIN MONTE CARLO ESTIMATION: THE BASICS

How Estimation Works (More or Less)

- Most estimation routines do one of three things:
- Minimize Something: Typically found with names that have "least" in the title. Forms of least squares include "Generalized", "Ordinary", "Weighted", "Diagonally Weighted", "WLSMV", and "Iteratively Reweighted." Typically the estimator of last resort...
- <u>Maximize Something:</u> Typically found with names that have "maximum" in the title. Forms include "Maximum likelihood", "ML", "Residual Maximum Likelihood" (REML), "Robust ML". Typically the gold standard of estimators (and we now know why).
- 3. <u>Use Simulation to Sample from Something:</u> more recent advances in simulation use resampling techniques. Names include "Bayesian Markov Chain Monte Carlo", "Gibbs Sampling", "Metropolis Hastings", "Metropolis Algorithm", and "Monte Carlo". Used for complex models where ML is not available or for methods where prior values are needed.

How MCMC Estimation Works

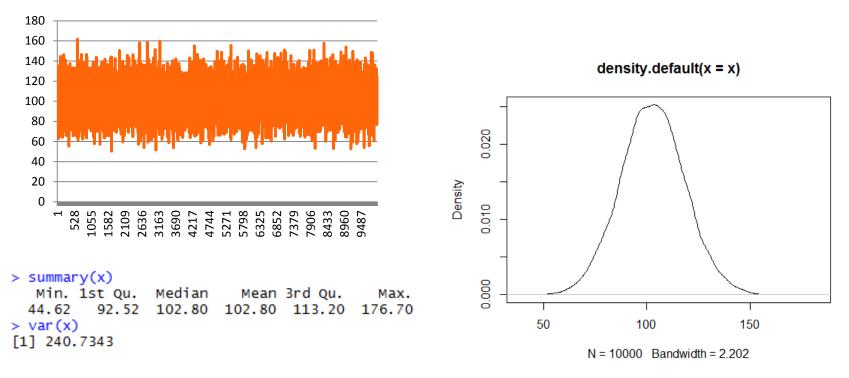
 MCMC estimation works by taking samples from the posterior distribution of the data given the parameters:

$$f(\beta_0, \sigma_e^2 | \mathbf{y}_p) = \frac{f(\mathbf{y}_p | \beta_0, \sigma_e^2) f(\beta_0) f(\sigma_e^2)}{f(\mathbf{y}_p)}$$

- > How is that possible? We don't know $f(y_p)$...but...we'll see...
- After enough values are drawn, a rough shape of the distribution can be formed
 - > From that shape we can take summaries and make them our parameters (i.e., mean)
- How the sampling mechanism happens comes from several different algorithms that you will hear about, the most popular being:
 - > **Gibbs Sampling**: used when $f(\beta_0, \sigma_e^2 | y_p)$ is known
 - Parameter values are drawn and kept throughout the chain
 - > Metropolis-Hastings (within Gibbs): used when $f(\beta_0, \sigma_e^2 | y_p)$ is unknown
 - Parameter values are proposed, then either kept or rejected
 - SAS PROC MCMC uses the latter
 - TRIVIA NOTE: The Metropolis algorithm comes from Chemistry (in 1950)
- In some fields (Physics in particular), MCMC estimation is referred to as Monte Carlo estimation

Sampling Example

- Imagine I wanted to get the shape of a distribution similar to our IQ example (with a mean of 102.8 and a variance of 239.5)
 - > This is essentially Gibbs Sampling
- I will open Excel and draw 10,000 random values from N(102.8, 239.5)
 - You can do this by typing "=norminv(rand(),102.8,SQRT(239.5))"



MCMC Estimation with MHG

- The Metropolis-Hastings algorithm works a bit differently than Gibbs sampling:
- 1. Each parameter (here β_0 and σ_e^2) is given an initial value
- 2. In order, a new value is proposed for each model parameter from some distribution:

$$\beta_0^* \sim Q(\beta_0^*|\beta_0); \sigma_e^{2^*} \sim Q(\sigma_e^{2^*}|\sigma_e^2)$$

3. The proposed value is then accepted as the current value with probability $\max(r_{MHG}, 1)$:

$$r_{MHG} = \frac{f(\mathbf{y}_p | \beta_0^*, \sigma_e^{2^*}) f(\beta_0^*) f(\sigma_e^{2^*}) Q(\beta_0 | \beta_0^*) Q(\sigma_e^{2^*} | \sigma_e^{2^*})}{f(\mathbf{y}_p | \beta_0, \sigma_e^{2^*}) f(\beta_0) f(\sigma_e^{2^*}) Q(\beta_0^* | \beta_0) Q(\sigma_e^{2^*} | \sigma_e^{2^*})}$$

 The process continues for a pre-specified number of iterations (more is better)

Notes About MHG

• The constant in the denominator of the posterior distribution: $f(\beta_0, \sigma_e^2 | \mathbf{y}_p) = \frac{f(\mathbf{y}_p | \beta_0, \sigma_e^2) f(\beta_0) f(\sigma_e^2)}{f(\mathbf{y}_p)}$

...cancels when the ratio is formed

- The proposal distributions $Q(\beta_0^*|\beta_0)$ and $Q(\sigma_e^{2^*}|\sigma_e^2)$ can literally be any statistical distribution
 - > The trick is picking ones that make the chain "converge" quickly
 - Want to find values that lead to moderate number of accepted parameters
 - SAS PROC MCMC/WINBUGS don't make you pick these
- Given a long enough chain, the final values of the chain will come from the posterior distribution
 - > From that you can get your parameter estimates

```
*INITIAL PROC MCMC RUN: PRIORS FROM SLIDE 13 - but without burnin period;

= PROC MCMC DATA=work.normalgen OUTPOST=work.outpost SEED=10252012

        NBI=0 THIN=1 NMC=10000 DIC;

    PARMS b0, errorvar;

    PRIOR b0 ~ N(100,SD=2.13);

    PRIOR errorvar ~ UNIFORM(200,400);

    condmean_y = b0;

    MODEL y ~ N(condmean_y,VAR=errorvar);

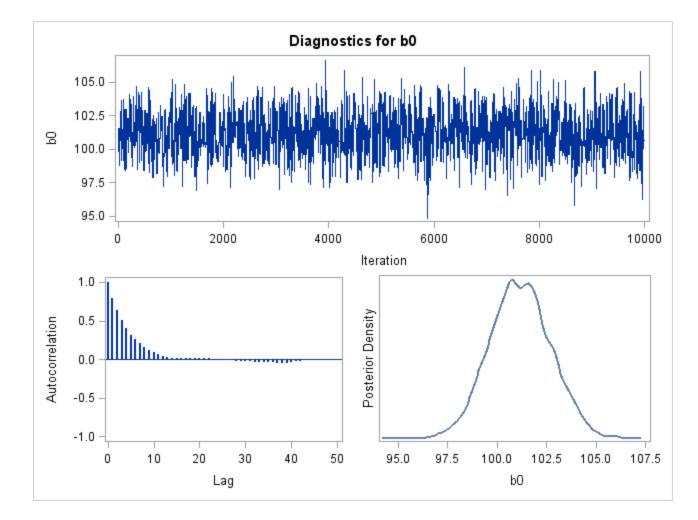
    RUN;
```

- SEED: random number seed (same number = same output)
- NBI: number of burn in iterations (more on this soon)
- THIN: thinning interval (more on this soon)
- NMC: number of total iterations
- PARMS: list all model parameters
- PRIOR: specify priors for each parameter
- MODEL: specify model for the data (note: MODEL is different from previous SAS PROCs in that you must specify distribution)

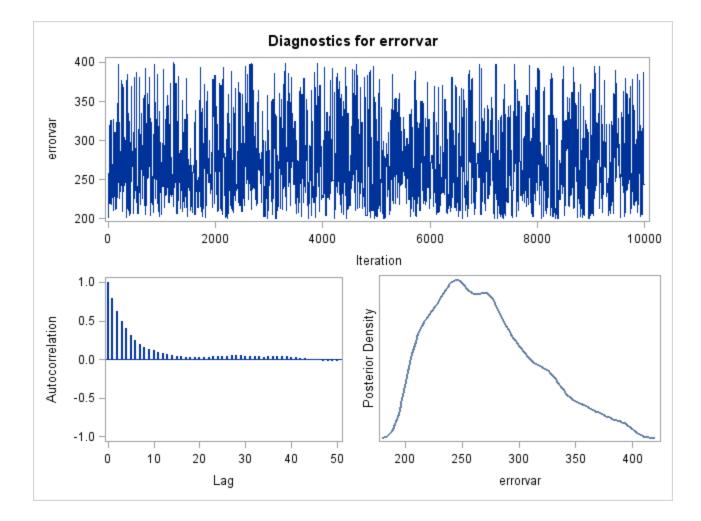
Iteration History from SAS

| | Iteration | ь0 | errorvar | Log Prior Density | Log-Likelihood Value | Log Posterior |
|----|-----------|---------|----------|----------------------|-------------------------|------------------|
| | | | | - | | Density |
| 1 | 1 | 101.2 | 205.7 | -7.1297 | -208.5 | -215.6 |
| 2 | 2 | 100.3 | 258.5 | -6.9864 | -208.5 | -215.5 |
| 3 | 3 | 101.6 | 201.0 | -7.2473 | -208.5 | -215.7 |
| 4 | 4 | 100.2 | 256.3 | -6.9781 | -208.6 | -215.6 |
| 5 | 5 | 100.2 | 256.3 | -6.9781 | -208.6 | -215.6 |
| 6 | 6 | 99.8318 | 228.9 | -6.9765 | -208.9 | -215.9 |
| 7 | 7 | 99.8318 | 228.9 | -6.9765 | -208.9 | -215.9 |
| 8 | 8 | 100.5 | 224.4 | -7.0026 | -208.5 | -215.5 |
| 9 | 9 | 100.5 | 224.4 | -7.0026 | -208.5 | -215.5 |
| 10 | 10 | 100.5 | 224.4 | -7.0026 | -208.5 | -215.5 |
| 11 | 11 | 100.5 | 224.4 | -7.0026 | -208.5 | -215.5 |
| 12 | 12 | 100.5 | 224.4 | -7.0026 | -208.5 | -215.5 |
| 13 | 13 | 100.5 | 224.4 | -7.0026 | -208.5 | -215.5 |
| 14 | 14 | 100.5 | 224.4 | -7.0026 | -208.5 | -215.5 |
| 15 | 15 | 98.8215 | 275.7 | -7.1264 | -209.6 | -216.7 |
| 16 | 16 | 98.8215 | 275.7 | -7.1264 | -209.6 | -216.7 |
| 17 | 17 | 101.4 | 318.8 | -7.1909 | -209.0 | -216.2 |
| 18 | 18 | 101.4 | 318.8 | -7.1909 | -209.0 | -216.2 |
| 19 | 19 | 100.4 | 228.1 | -6.9899 | -208.6 | -215.6 |
| 20 | 20 | 100.4 | 228.1 | -6.9899 | -208.6 | -215.6 |
| 21 | 21 | 101.4 | 306.1 | -7.1903 | -208.8 | -215.9 |
| 22 | 22 | 99.9909 | 280.8 | -6.9734 | -208.9 | -215.9 |
| 23 | 23 | 99.9909 | 280.8 | -6.9734 | -208.9 | -215.9 |
| 24 | 24 | 99.9909 | 280.8 | -6.9734 | -208.9 | -215.9 |
| 25 | 25 | 101.6 | 219.7 | -7.2484 | -208.2 | -215.4 |
| 26 | 26 | 101.6 | 219.7 | -7.2484 | -208.2 | -215.4 |
| 27 | 27 | 101.6 | 219.7 | -7.2484 | -208.2 | -215.4 |
| 28 | 28 | 101.6 | 219.7 | -7.2484 | -208.2 | -215.4 |
| 29 | 29 | 101.6 | 219.7 | -7.2484 | -208.2 | -215.4 |
| 30 | 30 | 101.6 | 219.7 | -7.2484 | -208.2 | -215.4 |
| 31 | 31 | 101.6 | 219.7 | -7.2484 | -208.2 | -215.4 |
| 32 | 32 | 101.6 | 219.7 | -7.2484 | -208.2 | -215.4 |
| 33 | 33 | 101.6 | 219.7 | -7.2484 | -208.2 | -215.4 |
| 34 | 34 | 101.7 | 245.9 | -7.3065 | -208.0 | -215.3 |
| 35 | 35 | 101.7 | 245.9 | -7.3065 | -208.0 | -215.3 |
| 36 | 36 | 101.7 | 245.9 | -7.3065 | -208.0 | -215.3 |
| 37 | 37 | 101.7 | 245.9 | -7.3065 | -208.0 | -215.3 |
| 38 | 38 | 101.7 | 245.9 | -7.3065 | -208.0 | -215.3 |
| 39 | 39 | 101.7 | 245.9 | -7.3065 | -208.0 | -215.3 |

Examining the Chain and Posteriors



Examining the Chain and Posteriors

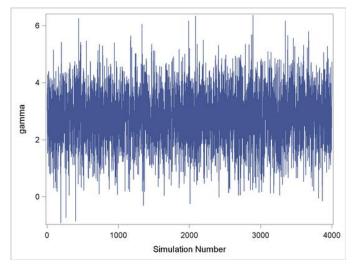


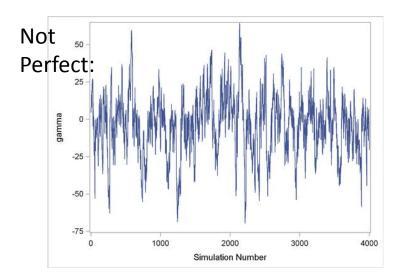
Practical Specifics in MCMC Estimation

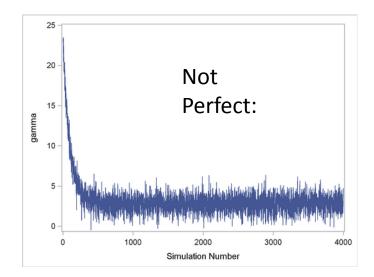
- A burn-in period is used where a chain is run for a set number of iterations before the sampled parameter values are used in the posterior distribution
- Because of the rejection/acceptance process, any two iterations are likely to have a high correlation (called **autocorrelation**) → posterior chains use a **thinning interval** to take every Xth sample to reduce the autocorrelation
 - A high autocorrelation may indicate the standard error of the posterior distribution will be smaller than it should be
- The chain length (and sometimes number of chains) must also be long enough so the rejection/acceptance process can reasonably approximate the posterior distribution
- How does one what values to pick for these? Output diagnostics
 > Trial. And. Error.

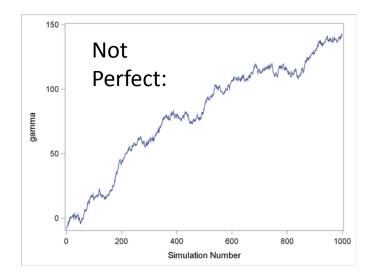
Best Output Diagnostics: the Eye Ball Test

Perfect:









| Geweke Diagnostics | | | | | | |
|--------------------|---------|-------------|--|--|--|--|
| Parameter | Z | $\Pr > z $ | | | | |
| b0 | -0.4489 | 0.6535 | | | | |
| errorvar | 1.0657 | 0.2865 | | | | |

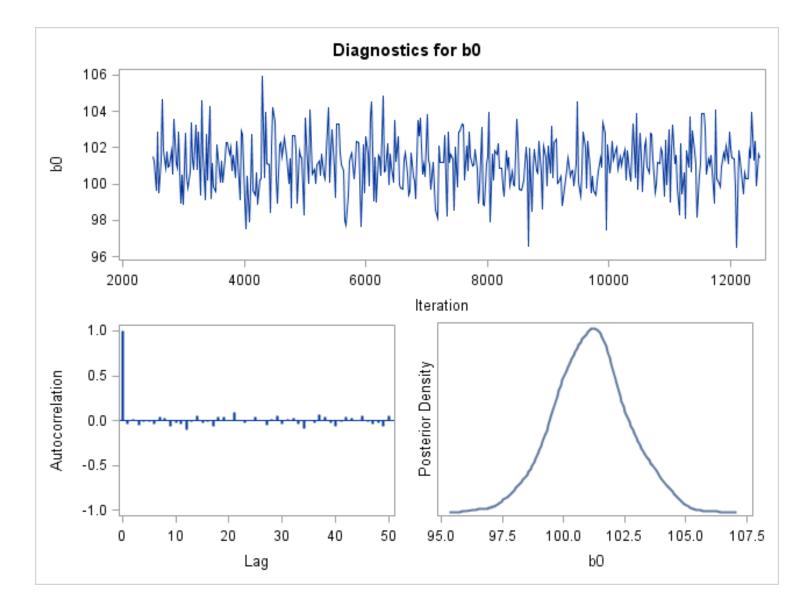
| Posterior Autocorrelations | | | | | | | | |
|----------------------------|--------|--------|--------|---------|--|--|--|--|
| Parameter | Lag 1 | Lag 5 | Lag 10 | Lag 50 | | | | |
| b0 | 0.7942 | 0.3163 | 0.0929 | -0.0074 | | | | |
| errorvar | 0.7914 | 0.3189 | 0.1158 | -0.0280 | | | | |

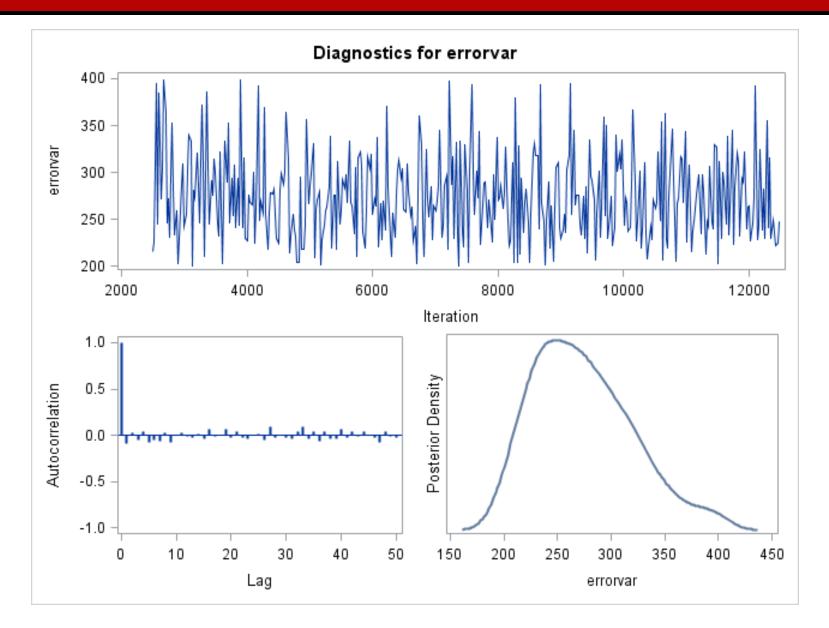
Rerunning Our Analysis: Burn-in and Thinning Interval

• Burn-in = 2500; thinning interval = 25; chain length = 10000

Posterior Autocorrelations Lag 10 Lag 50 Parameter Lag 5 Lag 1 -0.0402 -0.0128 -0.0198 0.0493 b0 -0.0832 -0.0751 -0.0033 -0.0188 errorvar

| Geweke Diagnostics | | | | |
|--------------------|--------|---------|--|--|
| Parameter | Z | Pr > z | | |
| b0 | 1.7722 | 0.0764 | | |
| errorvar | 1.7334 | 0.0830 | | |





Converged? If Yes...now onto the parameters

Posterior Summaries

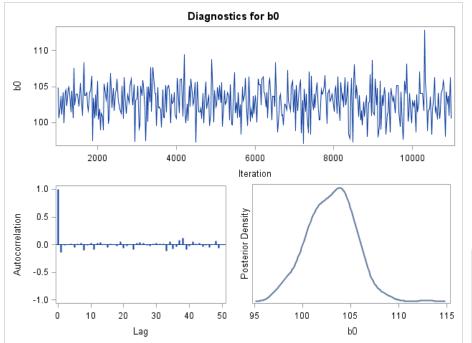
| Parameter | Parameter N | | Mean | | Standard Per | | Perc | centiles | | | | | |
|----------------------|-------------|--|-------------------------|---------------------|--------------|---------------|--------|--------------|-------|-------|-----|------|--|
| | | | | | | Deviation 25% | | , D | | 50% | 7 | 5% | |
| b0 | 40 | 0 | 101 | .1 | 1.4841 | | 100.1 | | 1 | 101.1 | 1(| 02.0 | |
| errorvar | 40 | 0 | 273 | .9 45.256 | | 61 | 238.7 | | 2 | 268.5 | 30 | 02.3 | |
| Posterior In | ter | vals | | | | | | | | | | | |
| Parameter | | Alpha | | Equal-Tail Interval | | | | HPD Interval | | | | | |
| b0 | | 0.050 | | 98.137 | 98.1372 | | 1 98 | | 98.2 | 694 | 104 | 4.2 | |
| errorvar | | 0.050 | | 204.1 | | 384.7 | .7 20 | | 200. | 2 | 359 | 9.7 | |
| | | Deviance Information Criterion | | | | | | | | | | | |
| 2 0700 | | Dbar (posterior mean of deviance) | | | | 417.636 | | | | | | | |
| Have to use DIC | | Dmean (deviance evaluated at posterior mean) | | | | 416 | 16.748 | | | | | | |
| | | pD (effective number of parameters) | | | f | | 0.8 | 0.888 | | | | | |
| PSYC 943: Lecture 17 | | | DIC (smaller is better) | | | | | 418 | 3.524 | | | | |

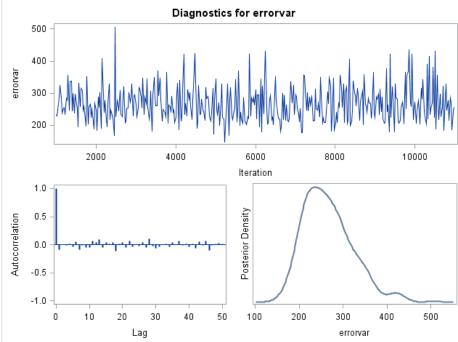
Changing Up the Prior

- To demonstrate how changing the prior affects the analysis, we will now try a few prior distributions for our parameters
- Prior: $\beta_0 \sim U(-10000, 10000); \sigma_e^2 \sim U(0, 5000)$

| Posterior Summaries | | | | | | | |
|---------------------|-----|-------|------------------------------------|-------------|----------|-------|--|
| Parameter | Ν | Mean | Standard DeviationPercer 25% | Percentiles | centiles | | |
| | | | | 25% | 50% | 75% | |
| b0 | 400 | 103.0 | 2.4016 | 101.3 | 103.1 | 104.6 | |
| errorvar | 400 | 266.0 | 55.6614 | 223.5 | 256.8 | 296.5 | |

| Deviance Information Criterion | | | | |
|--|---------|--|--|--|
| Dbar (posterior mean of deviance) | 418.004 | | | |
| Dmean (deviance evaluated at posterior mean) | 416.093 | | | |
| pD (effective number of parameters) | 1.911 | | | |
| DIC (smaller is better) | 419.915 | | | |
| | | | | |



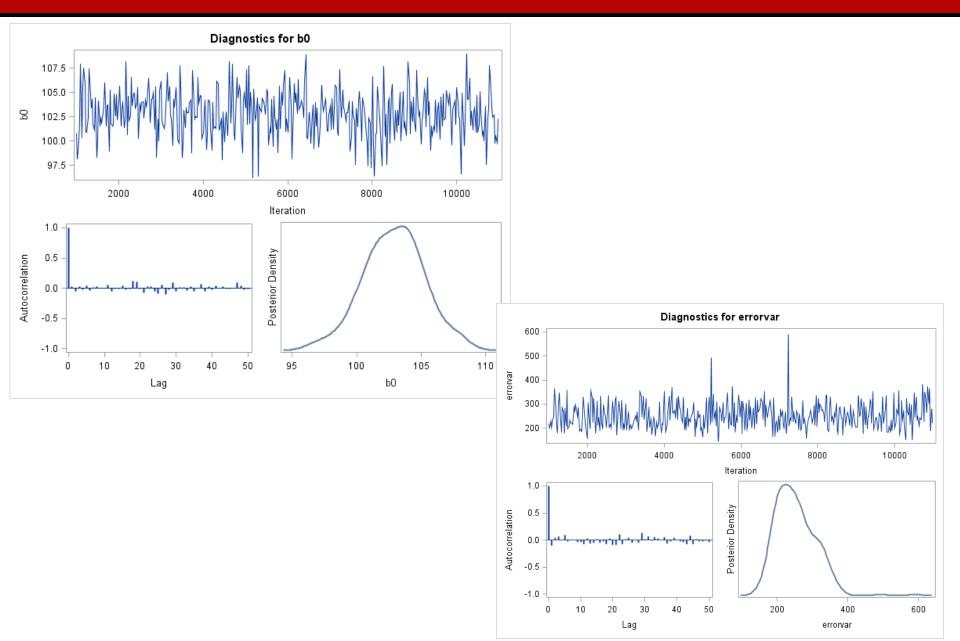


Changing Up the Prior

• Prior: $\beta_0 \sim N(0, 100, 000); \sigma_e^2 \sim \gamma^{-1}(shape = \frac{3}{10}, scale = \frac{10}{3})$

| Posterior Summaries | | | | | | | |
|---------------------|-----|-------|-----------|-------------|-------|-------|--|
| Parameter | N | Mean | Doviation | Percentiles | | | |
| | | | | 25% | 50% | 75% | |
| b0 | 400 | 102.9 | 2.3533 | 101.3 | 102.9 | 104.3 | |
| errorvar | 400 | 252.0 | 54.6698 | 210.1 | 244.9 | 286.7 | |

| Deviance Information Criterion | | | | |
|--|---------|--|--|--|
| Dbar (posterior mean of deviance) | 417.973 | | | |
| Dmean (deviance evaluated at posterior mean) | 415.885 | | | |
| pD (effective number of parameters) | 2.088 | | | |
| DIC (smaller is better) | 420.061 | | | |

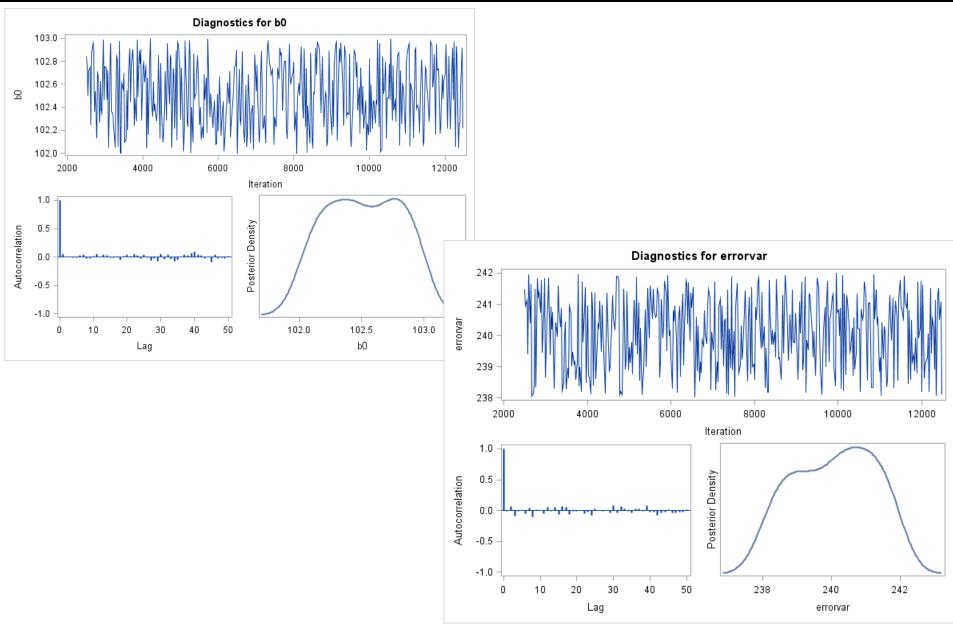


What About an Informative Prior?

• Prior: $\beta_0 \sim U(102,103); \sigma_e^2 \sim U(238,242)$

| Posterior Summaries | | | | | | | |
|---------------------|-----|-------|-----------|-------------|-------|-------|--|
| Parameter | Ν | Mean | Deviation | Percentiles | | | |
| | | | | 25% | 50% | 75% | |
| b0 | 400 | 102.5 | 0.2886 | 102.3 | 102.5 | 102.8 | |
| errorvar | 400 | 240.1 | 1.1440 | 239.1 | 240.2 | 241 | |

| Deviance Information Criterion | | | | |
|--|---------|--|--|--|
| Dbar (posterior mean of deviance) | 415.851 | | | |
| Dmean (deviance evaluated at posterior mean) | 415.833 | | | |
| pD (effective number of parameters) | 0.018 | | | |
| DIC (smaller is better) | 415.869 | | | |



WRAPPING UP

Wrapping Up

- Today was an introduction to Bayesian statistics
 - Bayes = use of prior distributions on parameters
- We used two methods for estimation:
 - MAP estimation far less common
 - MCMC estimation
 - Commonly, people will say Bayesian and mean MCMC but Bayesian is just the addition of priors. MCMC is one way of estimating Bayesian models!
- MCMC is effective for most Bayesian models:
 - > Model likelihood and prior likelihood are all that are needed
- MCMC is estimation by brute force:
 - > Can be very slow, computationally intensive, and disk-space intensive
- But...MCMC runs multiple imputation...which is our next topic