Reviewing Main Effects in General Linear Models (as estimated using restricted maximum likelihood in SAS PROC MIXED)

The models for this example come from Hoffman (2015) chapter 2. We will be examining the extent to which cognition (as measured by an information test outcome) can be predicted from age (centered at 85 years) grip strength (centered at 9 pounds), sex (with men as the reference group) and dementia status (none = 1, future = 2, and current = 3) in a sample of 550 older adults.

SAS Syntax and Output for Data Manipulation and Data Description:

```
* Defining global variable for file location to be replaced in code below;
 %LET filesave= C:\Dropbox\17_CLDP944\CLDP944_Example02;
* Location for SAS files for these models (uses macro variable filesave);
 LIBNAME filesave "&filesave.";
* Import chapter 2 example data into work library;
DATA work.Chapter2; SET filesave.SAS_Chapter2;
* Center continuous predictors;
age85 = age - 85;
grip9 = grip - 9;
* Creating manual contrasts for dementia groups (to be treated as continuous);
    IF demgroup=1 THEN DO; demNF=0; demNC=0; END; * None group is reference;
ELSE IF demgroup=2 THEN DO; demNF=1; demNC=0; END; * Future group difference;
ELSE IF demgroup=3 THEN DO; demNF=0; demNC=1; END; * Current group difference;
* Labeling new variables - note semi-colon is only at the end of all labels;
LABEL
age85= "age85: Age in Years (0=85)"
grip9= "grip9: Grip Strength in Pounds (0=9)"
       "sexMW: Sex (0=Men, 1=Women)"
sexMW=
demNF=
       "demNF: Dementia Contrast for None=0 vs Future=1"
demNC= "demNC: Dementia Contrast for None=0 vs Current=1";
RUN;
TITLE1 "Chapter 2: Descriptive Statistics for Example Variables";
PROC MEANS NDEC=2 DATA=work.Chapter2; VAR age grip cognition; RUN;
PROC FREQ DATA=work.Chapter2; TABLE sexMW demgroup; RUN;
PROC CORR DATA=work.Chapter2; VAR age grip sexMW cognition; RUN;
TITLE1;
```

New-school default SAS HTML output:

Variable	Label	N	Mean	Std Dev	Minimum	Maximum
age grip cognition	age: Age in Years grip: Grip Strength in Pounds cognition: Information Test Cognitive	550 550 550	84.93 9.11 24.82	3.43 2.98 10.99	80.02 0 0	96.97 19.00 44.00
	Outcome					

		sexMW:	Sex (0)=Me	en, 1=W	/omen)		
sexMW	Fre	equency	Perc	ent		lative lency	Cumulative Percent		
0		227	4	1.27	_	227	41.27		
1		323	58	3.73		550 100.00			
		demgro	up: D	eme	ntia Di	agnos	is (1=None, 2	2=Future, 3=	:Current)
demgro	up	Frequ	ency	P	ercent		Cumulative	Frequency	Cumulative Percent
	1		399		72.55			399	72.55
	2		109		19.82			508	92.36
	3		42		7.64			550	100.00

SAS MIXED Syntax and Output for Empty Model in Equation 2.3

Cognition_i = $\beta_0 + e_i$

TITLE1 'Eq 2.3: Empty Means Model';

PROC MIXED DATA=work.Chapter2 COVTEST NOCLPRINT NAMELEN=100 IC METHOD=REML;

MODEL cognition = / SOLUTION DDFM=BW;

ODS OUTPUT CovParms=CovEmpty;

RUN; TITLE1;

MODEL y = fixed effects of predictors

ODS OUTPUT saves CovParms table to dataset called "CovEmpty" for use in R² macro.

Old-school default SAS listing output (much easier to paste into handouts):

Number of Observations

Number of Observations Read 550
Number of Observations Used 550
Number of Observations Not Used 0

This table tells you how many cases were removed due to incomplete data—make sure you pay attention to this if you are doing any model comparisons (which will need to be based on the same cases to be valid).

Covariance Parameter Estimates → CovParms

This table will list all estimated parameters within the model for the variance. Right now all we have is residual variance, the variance of the e_i residuals. Because this is an empty model with no predictors, this is ALL the variance to be predicted in the cognition outcome.

Fit Statistics

-2 Res Log Likelihood 4196.1 AIC (Smaller is Better) 4198.1 AICC (Smaller is Better) 4198.1 BIC (Smaller is Better) 4202.4

Information Criteria

This first "Fit Statistics" table will index relative model fit (stay tuned). The second "Information Criteria" table reports the same info plus other indices. The "parms" in REML refers to the number of estimated parameters in the model for the variance (just 1 residual variance now).

Neg2LogLike Parms AIC AICC HQIC BIC CAIC 4196.1 1 4198.1 4198.1 4199.8 4202.4 4203.4

Solution for Fixed Effects

This "Solution for Fixed Effects" table will list all estimated parameters in the model for the means. It is not printed by default in PROC MIXED.

Intercept $\beta_0 =$

SAS MIXED Syntax and Output for Age, Grip, and Sex (0=M, 1=W) Model in Equation 2.7

Cognition_i = $\beta_0 + \beta_1 (Age_i - 85) + \beta_2 (Grip_i - 9) + \beta_3 (SexMW_i) + e_i$

TITLE1 'Eq 2.7: Age + Grip + Adding Sex (0=M 1=W, as continuous predictor)';

PROC MIXED DATA=work.Chapter2 COVTEST NOCLPRINT NAMELEN=100 IC METHOD=REML;

MODEL cognition = age85 grip9 sexMW / CHISQ SOLUTION DDFM=BW;

CONTRAST "Model R2 F-Test with df=3" age85 1, grip9 1, sexMW 1 / CHISQ;

ODS OUTPUT CovParms=CovAgeGripSex;

RUN; TITLE1;

Covariance Parameter Estimates

		Standard	Z	
Cov Parm	Estimate	Error	Value	Pr > Z
Residual	109.38	6.6200	16.52	<.0001

Information Criteria

Neg2LogLike	Parms	AIC	AICC	HQIC
4141.1	1	4143.1	4143.1	4144.8

CONTRAST provides a multivariate Wald test for the significance of the model (reduction in error variance from adding three fixed effects). Here I have also requested a chi-square value for illustration purposes (F*df = chi-square).

ODS OUTPUT saves CovParms table to dataset called "CovAgeGripSex" for use in R² macro. Thus, you can name each saved dataset whatever you want (but max 32 characters)

Solution for Fixed Effects

		Standard				
Effect	Estimate	Error	DF	t Value	Pr > t	
Intercept	26.9594	0.7389	546	36.49	<.0001	is BO
age85	-0.4338	0.1325	546	-3.27	0.0011	is B1
grip9	0.5460	0.1663	546	3.28	0.0011	is B2
sexMW	-3.7988	0.9904	546	-3.84	0.0001	is B3

Interpret each fixed effect:

Intercept $\beta_0 =$

Main effect of Age β_1 =

Main effect of Grip Strength β_2 =

Main effect of Sex β_3 =

Contrasts

Den

Label	DF	DF	Chi-Square	F Value	Pr > ChiSq	Pr > F
Model R2 F-Test with df=3	3	546	60.11	20.04	<.0001	<.0001

* Call macro to calculate R2 for overall model and change in R2 between models; %ModelR2(CovBase=CovEmpty, CovFewer=CovEmpty, CovMore=CovAgeGripSex);

R2 (% Reduction) Overall and for CovEmpty vs. CovAgeGripSex

Num

This table is created by my macro program to calculate R^2 and change in R^2 .

R2 from

Name	CovParm	Estimate	StdErr	ZValue	ProbZ	Base	R2_Increment
CovEmpty	Residual	120.76	7.2887	16.57	<.0001	0.000000	
CovAgeGripSex	Residual	109.38	6.6200	16.52	<.0001	0.094221	0.094221

Calculate model R^2 = (empty σ_e^2 – current σ_e^2) / (empty σ_e^2) = (120.76 – 109.38) / (120.76) = .09 The df=3 CONTRAST above says that this R^2 is significantly > 0, F(3,546) = 22.04, p < .0001.

SAS MIXED Syntax and Output for Dementia Group Model in Equation 2.8

$$Cognition_i = \beta_0 + \beta_1 (Age_i - 85) + \beta_2 (Grip_i - 9) + \beta_3 (SexMW_i) + \beta_4 (DemNF_i) + \beta_5 (DemNC_i) + e_i$$

We can use the model equation to calculate the **dementia group means** for predicted cognition:

```
Cognition for None = \beta_0

Cognition for Future = \beta_0 + \beta_4

Cognition for Current = \beta_0 + \beta_5
```

We can determine the **differences between the dementia group means** as follows:

```
None vs. Future = Future – None = (\beta_0 + \beta_4) - (\beta_0) = \beta_4

None vs. Current = Current – None = (\beta_0 + \beta_5) - (\beta_0) = \beta_5

Future vs. Current = Current – Future = (\beta_0 + \beta_5) - (\beta_0 + \beta_4) = \beta_5 - \beta_4 = -\beta_4 + \beta_5
```

These values are then requested via the SAS ESTIMATE statements below...

```
TITLE1 'Eq 2.8: Adding Dementia Group';
TITLE2 'Using Manual Group Contrasts so Reference=None';
TITLE3 'sexMW, demNF, and demNC are all treated as continuous predictors';
PROC MIXED DATA=work.Chapter2 COVTEST NOCLPRINT NAMELEN=100 IC METHOD=REML;
    MODEL cognition = age85 grip9 sexMW demNF demNC / CHISQ SOLUTION DDFM=BW;
    ODS OUTPUT CovParms=CovDem;
```

The first CONTRAST below includes all fixed effects, and thus tests the model R². The second CONTRAST below includes only the new fixed effects, and thus tests the increment to the model R2 from adding dementia group. In this case this is also an "omnibus" ANOVA test for group.

```
CONTRAST "Model R2 F-Test df=3" age85 1, grip9 1, sexmw 1, demNF 1, demNC 1; CONTRAST "Omnibus F-Test for Dementia Group df=2" demNF 1, demNC 1;
```

The first 3 ESTIMATEs request predicted outcomes, so they include the intercept. The last 3 ESTIMATES request slopes for group differences, so they do NOT include the intercept.

```
* Request conditional (adjusted) group means (hold age=85, grip=9, men);
 ESTIMATE "Intercept for None Group" intercept 1 demNF 0 demNC 0; * Given as B0;
 ESTIMATE "Intercept for Future Group" intercept 1 demNF 1 demNC 0; * Not given (B0+B4);
 ESTIMATE "Intercept for Current Group" intercept 1 demNF 0 demNC 1; * Not given (B0+B5);
* Request group differences (unconditional because there are no interactions);
 ESTIMATE "None
                  vs. Future Group" demNF 1 demNC 0; * Given as B4;
                  vs. Current Group" demNF
 ESTIMATE "None
                                              0 demNC 1; * Given as B5;
 ESTIMATE "Future vs. Current Group" demNF -1 demNC 1; * Not given (B5-B4);
RUN; TITLE1; TITLE2; TITLE3;
          Covariance Parameter Estimates
                      Standard
                                      Ζ
Cov Parm
           Estimate
                         Error
                                  Value
                                            Pr > Z
Residual
            88.0709
                        5.3401
                                  16.49
                                            <.0001
                         Information Criteria
Neg2LogLike
             Parms
                         AIC
                                  AICC
                                            HQIC
                                                       BIC
                                                                CAIC
    4016.3
                1
                      4018.3
                                4018.3
                                          4019.9
                                                    4022.6
                                                               4023.6
```

Interpret each fixed effect below:

Intercept β_0 =

Main effect of Age β_1 =

Main effect of Grip Strength β_2 =

Main effect of Sex β_3 =

Main effect of DemNF β_4 =

Main effect of DemNC β_5 =

Solution for Fixed Effects

		Standard			
Effect	Estimate	Error	DF	t Value	Pr > t
Intercept	29.2643	0.6985	544	41.90	<.0001 is B0
age85	-0.4057	0.1189	544	-3.41	0.0007 is B1
grip9	0.6042	0.1498	544	4.03	<.0001 is B2
sexMW	-3.6574	0.8914	544	-4.10	<.0001 is B3
demNF	-5.7220	1.0191	544	-5.61	<.0001 is B4
demNC	-16.4798	1.5228	544	-10.82	<.0001 is B5

Estimates

		Standard			
Label	Estimate	Error	DF	t Value	Pr > t
Intercept for None Group	29.2643	0.6985	544	41.90	<.0001 B0
Intercept for Future Group	23.5424	1.0785	544	21.83	<.0001 B0+B4
Intercept for Current Group	12.7845	1.5302	544	8.35	<.0001 B0+B5
None vs. Future Group	-5.7220	1.0191	544	-5.61	<.0001 B4
None vs. Current Group	-16.4798	1.5228	544	-10.82	<.0001 B5
Future vs. Current Group	-10.7578	1.7080	544	-6.30	<.0001 B5-B4
	Contras	:†9			

Contrasts

Label	DF	DF	F Value	Pr > F
Model R2 F-Test with df=5	5	544	41.75	<.0001
Omnibus F-Test for Dementia Group with df=2	2	544	67.06	<.0001

* Call macro to calculate R2 for overall model and change in R2; %ModelR2(CovBase=CovEmpty, CovFewer=CovAgeGripSex, CovMore=CovDem);

						R2_from_	
Name	CovParm	Estimate	StdErr	ZValue	ProbZ	Base	R2_Increment
CovEmpty	Residual	120.76	7.2887	16.57	<.0001	-0.00000	
CovAgeGripSex	Residual	109.38	6.6200	16.52	<.0001	0.09422	•
CovDem	Residual	88.0709	5.3401	16.49	<.0001	0.27069	0.17647

Num

Den

Previous model \mathbf{R}^2 = (empty $\sigma_{\mathbf{e}}^2$ – previous $\sigma_{\mathbf{e}}^2$) / (empty $\sigma_{\mathbf{e}}^2$) = (120.76 – 109.38) / (120.76) = .09 Current model \mathbf{R}^2 = (empty $\sigma_{\mathbf{e}}^2$ – current $\sigma_{\mathbf{e}}^2$) / (empty $\sigma_{\mathbf{e}}^2$) = (120.76 – 88.07) / (120.76) = .27 The df=5 CONTRAST above says that current \mathbf{R}^2 is significantly > 0, F(5,544) = 41.75, P(5,544) = .0001.

Change in model R^2 = (current R^2) – (previous R^2) = .27 – .09 = .18 The df=2 CONTRAST above says that change in R^2 is significantly > 0, F(2.544) = 67.06, p < .0001.

SAS MIXED Syntax and Output for Dementia Group Model in Equation 2.8 Using CLASS statement (SAS-coded contrasts instead of manually created contrasts)

Because the default reference group is the HIGHEST group numerically or last alphabetically, I have changed the model to reflect "Current" (group=3) as the reference:

Cognition_i = $\beta_0 + \beta_1 (Age_i - 85) + \beta_2 (Grip_i - 9) + \beta_3 (SexMW_i) + \beta_4 (DemCN_i) + \beta_5 (DemCF_i) + e_i$

```
TITLE1 'Eq 2.8: Adding Dementia Group';
TITLE2 'Categorical Predictor for Dementia Group on CLASS statement';
PROC MIXED DATA=work.Chapter2 COVTEST NOCLPRINT NAMELEN=100 IC METHOD=REML;
    CLASS demgroup; * CLASS statement demgroup replaces previous dem contrasts;
    MODEL cognition = age85 grip9 sexMW demgroup / SOLUTION CHISQ DDFM=BW;
    CONTRAST "Model R2 F-Test with df=5" age85 1, grip9 1, sexmw 1,
                                         demgroup -1 1 0, demgroup -1 0 1;
* Request conditional (adjusted) group means (hold age=85, grip=9, men) and all diffs;
    LSMEANS demgroup / DIFF=ALL AT(age85 grip9 sexMW) = (0 0 0);
* Request conditional (adjusted) group means and all differences;
    LSMEANS demgroup / DIFF=ALL;
*** All of the code below is redundant with LSMEANS, but here is how you get all the info;
 CONTRAST "Omnibus F-Test for Dementia Group with df=2" demgroup -1 1 0, demgroup -1 0 1;
* Request conditional (adjusted) group means (hold age=85, grip=9, men);
 ESTIMATE "Intercept for None Group" intercept 1 demgroup 1 0 0; * Not given (B0+B4);
 ESTIMATE "Intercept for Future Group" intercept 1 demgroup 0 1 0; * Not given (B0+B5);
 ESTIMATE "Intercept for Current Group" intercept 1 demgroup 0 0 1; * Given as B0;
* Request group differences (unconditional because there are no interactions);
 ESTIMATE "None vs. Future Group" demgroup -1 1 0; * Not given (B5-B4);
 ESTIMATE "None vs. Current Group" demgroup -1 0 1; * Given as B5;
 ESTIMATE "Future vs. Current Group" demgroup 0 -1 1; * Given as B4;
```

New output after using the CLASS statement for demgroup:

Omnibus F-Test for Dementia Group with df=2

RUN; TITLE1; TITLE2;

	demgroup (1=None,			Fixed Effects Standard		referer	nce. The o	ther rows	then in	ndic	n group is the ate group rence group.
Effect	3=Curren	t)	Estimate	Error	DI	= t	Value	Pr > t			
Intercept			12.7845	1.5302	54	4	8.35	<.0001	new E	30	
age85			-0.4057	0.1189	54	4	-3.41	0.0007	same E	31	
grip9			0.6042	0.1498	54	4	4.03	<.0001	same E	32	
sexMW			-3.6574	0.8914	54	4	-4.10	<.0001	same E	33	
demgroup	1		16.4798	1.5228	54	4	10.82	<.0001	new E	34	
demgroup	2		10.7578	1.7080	54	4	6.30	<.0001	new E	35	
demgroup	3		0						new re	ef g	roup
	Num	Ty Den	•	Fixed Effects							
Effect	DF	DF	Chi-Squar	e F Value		Pr >	ChiSq	Pr > F			
age85	1	544	•				0.0006	0.0007			
grip9	1	544	16.2	7 16.27		<	<.0001	<.0001			
sexMW	1	544	16.8	3 16.83		<	<.0001	<.0001			
demgroup	2	544	134.1	1 67.06		•	<.0001	<.0001			
			Con	trasts							
				Num		Den					
Label				DF		DF	F Valu	e Pr >	F		
Model R2 F-T		5		544	41.7	5 <.000	01				

67.06

<.0001 now given by default

544

	Estimates					These group means and mean differences, which we requested, are given through LSMEANS below (with less typing).					
Label			Estimate	Standa Err		F t Value	Pr	> t			
Intercept	for None Grou	ıp	29.2643	0.69	985 54	4 41.90	<	.0001			
Intercept for Future Group			23.5424	1.07	785 54	4 21.83	<	.0001			
Intercept for Current Group			12.7845	1.53	302 54	4 8.35	<	.0001			
None vs. Future Group			-5.7220	1.01	91 54	4 -5.61	<	.0001			
None vs. Current Group			-16.4798	1.52	228 54	4 -10.82	<	.0001			
Future vs.	Current Grou	ab	-10.7578	1.70	080 54	4 -6.30	<	.0001			
				Least Squ	ıares Mean	S					
	demgroup: (1=None,										
	2=Future,					Standard					
Effect	3=Current)	age85	grip9	sexMW	Estimate	Error	DF t	Value	Pr > t		
demgroup	1	0.00	0.00	0.00	29.2643	0.6985	544	41.90	<.0001		
demgroup	2	0.00	0.00	0.00	23.5424	1.0785	544	21.83	<.0001		
demgroup	3	0.00	0.00	0.00	12.7845	1.5302	544	8.35	<.0001		
1	three rows are and three rows a						•				
demgroup	2	-0.07	0.11	0.59	21.4923	0.9024	544	23.82	<.0001		
demgroup	3	-0.07	0.11	0.59	10.7345	1.4486	544	7.41	<.0001		
		Differences of Least Squares Means									
	demgroup: (1=None,	demgrou (1=None	,			_					
F.6.6	2=Future,	2=Futur	•	\ _			tandard	D.F.	+ 1/-1	Do a la	
Effect	3=Current)	3=Curre	nt) age8	35 grip9	9 sexMW	Estimate	Error	DF	t Value	Pr > t	
demgroup	1	2	0.0	0.00	0.00	5.7220	1.0191	544	5.61	<.000	
demgroup	1	3	0.0	0.00	0.00	16.4798	1.5228	544	10.82	<.000	
demgroup	2	3	0.0	0.00	0.00	10.7578	1.7080	544	6.30	<.000	
The secon	three rows are and three rows a lemgroup does	are group	mean diff	ferences a	t the samp	le mean valu	es of the	other p	predictors	instead.	
demaroup	1	2	-0.0	0.11	0.59	5.7220	1.0191	544	5.61	<.000	
demgroup	1	2									
demgroup	1	3	-0.0	0.11	0.59	16.4798	1.5228	544	10.82	<.000	

So to CLASS or not to CLASS? Either can work in every circumstance. The use of CLASS for categorical predictors can be more convenient in models with more than one categorical predictor (e.g., to get marginal and cell means for factorial designs), whereas manual group contrasts can be more convenient when most other predictors are continuous, or when some of your effects pertain to only some levels of the grouping variable (i.e., nested effects; stay tuned).

0.59

10.7578

1.7080

544

6.30

<.0001

0.11

-0.07

demgroup

3